

# STAT 330: Lecture 9

2024 Spring

June 11, 2024

Last Lecture Conditional Distribution & Conditional Expectation

$h(x) = E(Y|X=x)$  is a function of  $x$  and  $y$  only.

$h(X) = E(Y|X)$  is a random variable.

function of  $X$  and  $Y$  only

$$\text{var}(Y|X) = E(Y^2|X) - \{E(Y|X)\}^2 \quad E(Y) = E\{E(Y|X)\}$$

This Lecture Conditional Distribution and Joint MGF  
Expectation.

## (Cont') Conditional Expectation

Variance Decomposition Formula (Law of Total Variance)

$$\text{Var}(Y) = E\{\text{Var}(Y|X)\} + \text{Var}\{E(Y|X)\}$$

proof:  $\text{Var}(Y|X) = E(Y^2|X) - \{E(Y|X)\}^2$

$$\begin{aligned} E\{\text{Var}(Y|X)\} &= E\{E(Y^2|X)\} - E\{\{E(Y|X)\}^2\} \\ &= E(Y^2) - E\{\{E(Y|X)\}^2\} \end{aligned}$$

$$\begin{aligned} \text{Var}\{E(Y|X)\} &= E\{\{E(Y|X)\}^2\} - (E\{E(Y|X)\})^2 \\ &= E\{\{E(Y|X)\}^2\} - \{E(Y)\}^2 \end{aligned}$$

$$E\{\text{Var}(Y|X)\} + \text{Var}\{E(Y|X)\} = E(Y^2) - \{E(Y)\}^2 = \text{Var}(Y)$$

①.  $\text{Var}(Y) \geq E\{\text{Var}(Y|X)\}$

e.g. Y is weight. X is height.

$\text{Var}(Y)$ : uncertainty in weight in the population.

$\text{Var}(Y|X)$ : uncertainty in weight after knowing height.

If we know the height, the uncertainty in weight will decrease.

$E\{\text{Var}(Y|X)\}$ : average uncertainty if we know height.

If  $X \perp Y$ ,  $\text{Var}\{E(Y|X)\} = \text{Var}\{E(Y)\} = 0$

$$\text{Var}(Y) = E\{\text{Var}(Y|X)\}$$

②.  $\text{Var}(Y) = E\{\text{Var}(Y|X)\} + \text{Var}\{E(Y|X)\}$

e.g. X = {Canada, US, Mexico}, Y = weight.

$\text{Var}(Y)$ : variance of weight in North America.

$E\{\text{Var}(Y|X)\}$ : average within country variance.

$$\text{Var}(Y|X=CA) \quad \text{Var}(Y|X=US) \quad \text{Var}(Y|X=Mx)$$

average by taking expectation with respect to X

$$\text{Var}\{E(Y|X)\} : E(Y|X=CA) \quad E(Y|X=US) \quad E(Y|X=Mx)$$

variance of mean among different countries.

Example 1  $Y \sim \text{Poi}(\theta)$ .  $X|Y=y \sim \text{Bin}(y, p)$ . Find  $\text{Var}(X)$

$$\text{Solution: } \text{Var}(X) = E\{\text{Var}(X|Y)\} + \text{Var}\{E(X|Y)\}$$

$$\text{Var}(X|Y) = Yp(1-p), E(X|Y) = Yp.$$

$$E\{\text{Var}(X|Y)\} = p(1-p)E(Y), \quad \text{Var}\{E(X|Y)\} = \text{Var}(Yp) \\ = \theta p(1-p) \quad = p^2\theta$$

$$\frac{\text{Var}(X) = \theta p(1-p) + p^2\theta = p\theta}{X \sim \text{Poi}(\theta p) \quad \text{Var}(X) = p\theta} \quad \text{Same result.}$$

Example 2  $X \sim \text{Unif}(0,1)$  &  $Y|X=x \sim \text{Bin}(10, x)$   
Find  $E(Y)$ ,  $\text{Var}(Y)$

$$E(Y) = E\{E(Y|X)\} = E(10X) = 10E(X) = 5.$$

$$\begin{aligned} \text{Var}(Y) &= E\{\text{Var}(Y|X)\} + \text{Var}\{E(Y|X)\} \\ &= E\{10X(1-X)\} + \text{Var}(10X) \end{aligned}$$

$$\begin{aligned} \text{Var}(X) &= \frac{1}{12} \\ E(X^2) &= \frac{1}{12} + \left(\frac{1}{2}\right)^2 = \frac{100}{12} + 10 \cdot \left\{ \frac{1}{2} - \frac{1}{3} \right\} = 10 \\ &= \frac{1}{3} \end{aligned}$$

Example 3  $Y \sim \text{Poi}(\theta)$ .  $X|Y=y \sim \text{Bin}(y, p)$ , Find mgf of  $X$

$$M(t) = E\{e^{tX}\} = E\{E(e^{tX}|Y)\}$$

$$E(e^{tx} | Y=y) = \sum_{x=0}^y \binom{y}{x} \cdot p^x (1-p)^{y-x} e^{tx} = \sum_{x=0}^y \binom{y}{x} (pe^t)^x (1-p)^{y-x}$$

$$\text{Let } P' = \frac{Pe^t}{Pe^t + 1 - P} \quad 1 - P' = \frac{1 - P}{Pe^t + 1 - P} \quad = \sum_{x=0}^{\infty} \binom{y}{x} (P')^x (1 - P')^{y-x} \cdot \frac{(Pe^t + 1 - P)^y}{(Pe^t + 1 - P)^y}$$

$\downarrow$   
Bin(y, P')

$$M(t) = E \{ (pe^t + (1-p))^Y \}$$

$$= \sum_{y=0}^{\infty} (pe^t + 1 - p)^y \cdot \frac{\theta^y e^{-\theta}}{y!} = \sum_{y=0}^{\infty} \frac{\{f_0(pe^t + 1 - p)\}^y}{y!} \cdot e^{-\theta}$$

$$\text{Let } \Theta' = \Theta \cdot (pe^t + 1 - p) = \sum_{y=0}^{\infty} \frac{(\Theta')^y}{y!} e^{-\Theta'} e^{\Theta' - \Theta}$$

$$= e^{\theta^t - \theta} = e^{P\theta(e^t - 1)}$$

$$= e^{\theta_1 - \theta} = e^{p\theta(e^{\theta} - 1)}$$

## Joint Moment Generating Function

Definition:

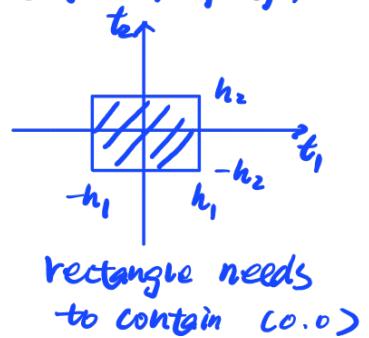
Bivariate:  $X, Y$  rvs.  $M(t_1, t_2) = E[e^{t_1 X + t_2 Y}]$  is called the joint mgf for  $(X, Y)$  if  $M(t_1, t_2)$  exists for  $|t_1| < h_1, |t_2| < h_2$ , for some  $h_1 > 0, h_2 > 0$ .

Multivariate:  $X_1 \dots X_n$  rvs.

$$M(t_1, \dots, t_n) = E[e^{t^T X}] \text{ where}$$

$$\begin{aligned} t^T &= [t_1, \dots, t_n], \quad t^T X = \sum_{i=1}^n t_i X_i \\ X^T &= [X_1, \dots, X_n] \end{aligned}$$

if  $|t_i| < h_i$ , for some  $h_i > 0, i=1, \dots, n$ .



Properties: ① From joint mgf to marginal mgf.

Suppose  $M(t_1, t_2)$  for  $|t_1| < h_1, |t_2| < h_2, h_1 > 0, h_2 > 0$ .

$$M_X(t_1) = M(t_1, t_2=0) = E(e^{t_1 X}), \quad |t_1| < h_1$$

$$M_Y(t_2) = M(t_1=0, t_2) = E(e^{t_2 Y}), \quad |t_2| < h_2.$$

② Independence

$$X \perp Y \iff M(t_1, t_2) = M_X(t_1) \cdot M_Y(t_2)$$

$X_1 \dots X_n$  are independent,  $\iff$

$$M(t_1, \dots, t_n) = \prod_{i=1}^n M_{X_i}(t_i)$$

③ Uniqueness

Joint mgf & joint dist'n are one-to-one.

Examples:Example 1:  $f(x,y) = e^{-y}$ ,  $0 < x < y$ . Find joint and marginal m.g.f.s.

$$\begin{aligned}
 M(t_1, t_2) &= E(e^{t_1 X + t_2 Y}) \\
 &= \int_0^\infty \int_0^y e^{t_1 x + t_2 y} \cdot e^{-y} dx dy \\
 &= \int_0^\infty [e^{-(1-t_1-t_2)y} - e^{-(1-t_2)y}] dy \\
 &= \frac{1}{1-t_1-t_2} \cdot \frac{1}{1-t_2}, \quad t_2 < 1, \quad t_1+t_2 < 1
 \end{aligned}$$

$$M_X(t_1) = M(t_1, t_2=0) = \frac{1}{1-t_1}, \quad t_1 < 1$$

$$M_Y(t_2) = M(t_1=0, t_2) = \frac{1}{(1-t_2)^2}, \quad t_2 < 1$$

$$M_X(t_1) \cdot M_Y(t_2) \neq M(t_1, t_2) \Rightarrow X \not\perp Y$$

Example 2: Suppose  $X \sim \text{Poi}(\theta_1)$ ,  $Y \sim \text{Poi}(\theta_2)$ ,  $X \perp Y$ Find out the distn of  $X+Y$ .

$$\begin{aligned}
 \text{Solution: } M_{X+Y}(t) &= E(e^{tX+tY}) = \{E(e^{tX})\} \{E(e^{tY})\} \\
 &= \exp\{\theta_1(e^t-1)\} \exp\{\theta_2(e^t-1)\} \\
 &= \exp\{(\theta_1+\theta_2)(e^t-1)\} \sim \text{Poi}(\theta_1+\theta_2)
 \end{aligned}$$

