

Random variables

- ► A **random variable** is a variable that holds a value produced by a (partially) random phenomenon
 - basically it is a name attached to an arbitrary value
 - short notation: r.v.
- ► Typically denoted as X, Y etc..
- Examples:
 - $\triangleright X =$ The value of a dice
 - $ightharpoonup V_{in} =$ The value of the voltage in one point of a circuit

Off-topic: Glossary

- "i.e." = id est = "that is" = "adică"
- "e.g." = exampli gratia = "for example" = "de exemplu"

Realizations

- ▶ A realization of a random variable = one possible value it can take
 - e.g. the value 3 of a dice
 - at different times, one may get different realizations
- ▶ Sample space Ω = the set of all values that can be taken by a random variable X
 - i.e. the set of all possible realizations
- ► Example: rolling a dice
 - ► The r.v. is denoted as X
 - We might get a realization X = 6
 - But we could have got any value from the sample space

$$\Omega = \{1,2,3,4,5,6\}$$

Discrete and continuous random variables

- **Discrete** random variable: if Ω is a discrete set
 - Example: value of a dice
- **Continuous** random variable: if Ω is a continuous set
 - Example: a voltage value

Why random variables?

- ► Random variables are a great model for **noise**
- Examples:
 - Measure a voltage in a circuit
 - Measure several times, the value is never precisely the same. The values always varies a little.
 - ▶ i.e. it is affected by noise

Probability Mass Function

- Consider a discrete r.v. A
- ► The **probability mass function (PMF)** = the probability that *A* has value *x*

$$w_A(x) = P\{A = x\}$$

- ▶ Also known as the **distribution** of *A*
- Example: what is the PMF of a dice? Plot on board.

Computing probability based on PMF

ightharpoonup Probability that A is equal to some value v

$$P\{A=v\}=w_A(v)$$

Probability that A is between a and b (including):

$$P\left\{a \leq A \leq b\right\} = \sum_{x=a}^{b} w_{A}(x)$$

Cumulative Distribution Function

▶ The **cumulative distribution function (CDF)** = the probability that the value of A is smaller or equal than x

$$F_A(x) = P\{A \le x\}$$

- ► In Romanian: "funcție de repartitie"
- Example: what is the CDF of a dice? Plot on board.
- ► For discrete r.v., the CDF is "stairwise"

Computing probability based on CDF

Probability that A is equal to some value v

$$P\{A = v\} = F_A(v) - F_A(v - 1)$$

▶ Probability that *A* is between *a* and *b* (including):

$$P\{a \le A \le b\} = F_A(b) - F_A(a-1)$$

Relation between PMF and CDF

► CDF is the *cumulative sum* (i.e. the integral) of PMF

$$F_A(x) = \sum_{A|I \ t \leq x} w_A(t)$$

Example for dice: easy to notice graphically

Probability Density Function

- Consider a continuous r.v. A
 - ightharpoonup assume it takes values in some interval [a, b]
- ▶ The **Probability Density Function (PDF)** of A = probability that the value of A is in a small vicinity *epsilon* around x, divided by *epsilon*
- ▶ Denoted as $w_A(x)$, also known as **the distribution** of A
- Informally, the PDF gives the probability that the value of A is close to x

Probability of an exact value

- ► The probability that a continuous r.v. A is exactly equal to a value x is zero
 - because there are an infinity of possibilities (continuous)
 - ► That's why we can't define a probability mass function like for discrete r.v.
- ► That's why the PDF says **in a small vicinity** around some value *x*, and not precisely equal to *x*

Computing probability based on PDF

Probability that A is equal to some value v is always 0

$$P\{A = v\} = 0$$

▶ Probability that A is between a and b = integral of PDF from a to b:

$$P\left\{a \leq A \leq b\right\} = \int_a^b w_A(x) dx$$

Cumulative Distribution Function

▶ The cumulative distribution function (CDF) = the probability that the value of A is smaller or equal than x

$$F_A(x) = P\{A \le x\}$$

- ► In Romanian: "funcție de repartiție"
- Same definition as for discrete r.v.

Computing probability based on CDF

Probability that A is between a and b:

$$P\{a \le A \le b\} = F_A(b) - F_A(a)$$

- ▶ Doesn't matter if we consider closed or open interval
 - [a,b] or (a,b)
 - ► why?

Relation between PDF and CDF

- CDF is the integral of PMF
- ▶ PDF is **the derivative** of CDF

$$F_A(x) = \int_{-\infty}^x w_A(x) \mathrm{d}x$$

$$w_A(x) = \frac{\mathrm{d}F_A(x)}{\mathrm{d}x}$$

$$= \lim_{\epsilon \to 0} \frac{F_A(x+\epsilon) - F_A(x-\epsilon)}{2\epsilon}$$

$$= \lim_{\epsilon \to 0} \frac{P(A \in [x-\epsilon, x+\epsilon])}{2\epsilon}$$

Graphical interpretation

- Probability that a continuous r.v. A is between a and b is the area below the PDF
 - ▶ i.e. the integral from a to b
- ▶ Probability that A is exactly equal to a certain value is zero
 - the area below a single point is zero

Discrete vs continuous r.v.

Comparison of discrete vs continous random variables:

- ▶ The CDF $F_A(x)$ is defined identically, means same thing
- ▶ The PDF/PMF $w_A(x)$ is the derivative of CDF
 - for continuous r.v.:
 - it is a proper derivative
 - it means probability to be "around" x
 - for discrete r.v:
 - sort of "discrete derivative"
 - it means probability to be exactly equal to x

Properties of random variables

CDF:

- ▶ The CDF is always ≥ 0
- ► The CDF is always monotonously increasing (non-decreasing)
- ▶ The CDF starts from 0 and goes up to 1

$$F_A(-\infty) = 0$$
 $F_A(\infty) = 1$

PDF/PFM:

- ► The PDF/PMF are always > 0
- ▶ Integral/sum over all of the PDF/PMF = 1

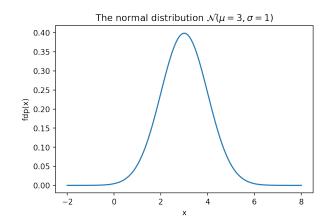
$$\int_{-\infty}^{\infty} w_A(x) \mathrm{d}x = 1$$

$$\sum_{x=-\infty}^{\infty} w_{A}(x) = 1$$

The normal distribution

Probability density function

$$w_{\mathcal{A}}(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$



The normal distribution

- Has two parameters:
 - **Average value** $\mu =$ "center" of the function
 - **Standard deviation** σ = "width" of the function
 - ightharpoonup Small $\sigma = \text{narrow and tall}$
- lacktriangle The front constant is just for normalization (ensures that integral =1)
- Extremely often encountered in real life
- ▶ Any real value is possible $(w_A(x) > 0, \forall x \in \mathbb{R})$
- ▶ Usually denoted as $\mathcal{N}(\mu, \sigma^2)$

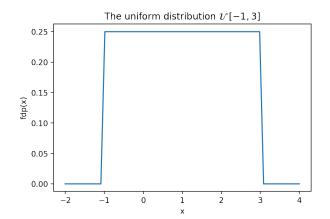
The normal distribution

- lacktriangle The distribution decreases as x gets farther from μ
 - ▶ Because of the term $-(x \mu)^2$ at the exponent
 - ▶ Most likely values: around μ ($x \mu = 0$)
 - lacktriangle Values closer to μ are more likely, values farther from μ are less likely
- The function describes a preference for values around μ , with decreasing preference when getting farther from μ

The uniform distribution

▶ The probability density function = a constant, between two endpoints

$$w_A(x) = \begin{cases} \frac{1}{b-a}, & x \in [a,b] \\ 0, & elsewhere \end{cases}$$



The uniform distribution

- ▶ Has two parameters: the limits a and b of the interval
- ▶ The "height" of the function is $\frac{1}{b-a}$
 - ▶ in order for the integral to be 1
- ▶ Only values from the interval [a, b] are possible
 - value cannot be outside interval (probability is 0)
- ▶ Denoted as $\mathcal{U}[a, b]$

Other distributions

Many other distributions exist, relevant for particular applications

Computing probabilities for the normal distribution

- ► How to compute \int_a^b for a normal distribution?
 - Can't be done with algebraic formula, non-elementary function
- Use the error function:

$$erf(z) = \frac{2}{\sqrt{\pi}} \int_0^z e^{-t^2} dt$$

▶ The CDF of a normal distribution $\mathcal{N}(\mu, \sigma^2)$

$$F_A(X) = \frac{1}{2}(1 + erf(\frac{x - \mu}{\sigma\sqrt{2}}))$$

- ▶ The values of *erf()* are available / are computed numerically
 - \triangleright e.g. on Google, search for *erf* (0.5)
 - Other useful values:
 - $ightharpoonup erf(-\infty) = -1$
 - $ightharpoonup erf(\infty) = 1$

Exercise

Exercise:

Let X be a r.v. with distribution $\mathcal{N}(3,2)$. Compute the probability that $X \in [2,4]$

Sum of constant + random variable

- Consider a random variable A
- $\blacktriangleright \text{ What is } B = 5 + A?$

Answer:

- B is also a random variable
- ▶ B has same type of distribution, but the function is "shifted" by 5 to the right

Example:

- ▶ A is normal variable with $w_A(x) = \mathcal{N}(\mu = 3, \sigma^2 = 2)$
- ▶ What is the distribution of B = 5 + A?
- Answer: $w_B(x) = \mathcal{N}(\mu = 8, \sigma^2 = 2)$

R.v. as functions of other r.v.

- A function applied to a r.v. produces another r.v.
- lacktriangle Examples: if B is a r.v. with distribution \mathcal{U} [0, 10], then
 - ightharpoonup C = 5 + A is another r.v., with distribution \mathcal{U} [5, 15]
 - \triangleright $D = A^2$ is also another r.v.
 - ightharpoonup E = cos(A) is also another r.v.
- ▶ Reason: since A is random, the values B, C, D are also random
- A, B, C, D are not independent
 - ► A certain value of one of them automatically implies the value of the others

Exercise

Exercise:

Let A be a r.v. with distribution $\mathcal{N}(3,2)$. Compute the probability that $A \in [2,4]$

Multiple random variables

- Consider a system with two continuous r.v. A and B
- ▶ What is the probability that the pair (A, B) has values around (x, y)?
- Distribution of the values of (A, B) is described by:
 - **)** joint probability density function $w_{AB}(x, y)$
 - ▶ joint cumulative density function $F_{AB}(x, y)$

Multiple random variables

Joint cumulative distribution function:

$$F_{AB}(x,y) = P\left\{A \le x \cap B \le y\right\}$$

▶ Joint probability density function:

$$w_{AB}(x,y) = \frac{\partial^2 P_{AB}(x,y)}{\partial x \partial y}$$

- ▶ The joint PDF gives the probability that the value of the pair (A, B) is in a vicinity of (x, y)
- ► Similar for discrete random variables

$$w_{AB}(x,y) = P\{A = x \cap B = y\}$$

Independent random variables

- ► Two v.a. X and Y are **independent** if the value of one of them does not influence in any way the value of the other
- ► For independent r.v., the probability that X is around x and Y is around y is **the product** of the two probabilities

$$w_{AB}(x, y) = w_A(x) \cdot w_B(y)$$

- Relation holds for CDF / PDF / PMF, continuous or discrete r.v.
- ► Same for more than two r.v.

Independent random variables

Exercise:

- Compute the probability that three r.v. X, Y and Z i.i.d. $\mathcal{N}(-1,1)$ are all positive simultaneously
 - *i.i.d* = "independent and identically distributed"

Multiple normal variables

- ► Consider a set of *N* normal r.v. $(A_1,...A_N)$, with different μ_i , but same σ
- ▶ Then probability that $(A_1,...A_N)$ is around $(x_1,...x_N)$ is

$$w_{A_1,...A_N}(x_1,...x_N) = \frac{1}{(\sigma\sqrt{2\pi})^N} e^{\frac{(x_1-\mu_1)^2+...+(x_N-\mu_N)^2}{2\sigma^2}}$$

The probability depends on the **Euclidean distance** between $\mathbf{x} = (x_1, ... x_N)$ and $\mu = (\mu_1, ... \mu_N)$

Euclidean distance

Euclidean (geometric) distance between two N-dimensional vectors:

$$d(\mathbf{u}, \mathbf{v}) = \|\mathbf{u} - \mathbf{v}\| = \sqrt{(u_1 - v_1)^2 + ... + (u_N - v_N)^2}$$

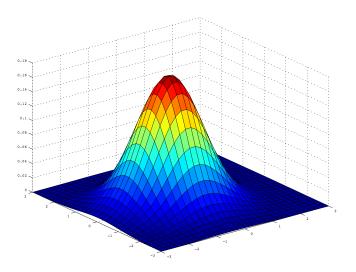
- ▶ One-dimensional: $\|\mathbf{u} \mathbf{v}\| = |u v|$
- ► 2D: $\|\mathbf{u} \mathbf{v}\| = \sqrt{(u_1 v_1)^2 + (u_2 v_2)^2}$
- ▶ 3D: $\|\mathbf{u} \mathbf{v}\| = \sqrt{(u_1 v_1)^2 + (u_2 v_2)^2 + (u_3 v_3)^2}$
- **.**..
- ▶ N-dimensional: $\|\mathbf{u} \mathbf{v}\| = \sqrt{\sum_{i=1}^{N} (u_i v_i)^2}$
- **...**
- ▶ N-dimensional: $\|\mathbf{u} \mathbf{v}\| = \sqrt{\int_{-\infty}^{\infty} (u(t) v(t))^2 dt}$

Multiple normal variables

- Probability of N normal random variables, independent, with same σ but possibly different μ_i depends on the **squared Euclidean** distance to the mean vector $\mu = (\mu_1, ... \mu_N)$
 - ightharpoonup Close to μ : higher probability
 - Far from μ : lower probability
 - ightharpoonup Two points at same distance from μ have same probability

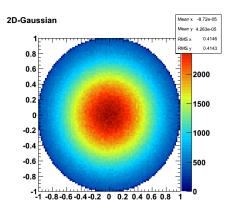
2D normal distribution

▶ Distribution of 2 normal random variables (2D normal distribution)



2D normal distribution - top view

- View from top
- Here, $\mu = (0,0)$
- Probability density decreases as distance from center increases, in circles (symmetrically)



Statistical averages

- R.v. are described by statistical averages ("moments")
- ► The average value (moment of order 1)
- Continuous r.v.:

$$\overline{X} = E\{X\} = \int_{-\infty}^{\infty} x \cdot w_X(x) dx$$

Discrete r.v.:

$$\overline{X} = E\{X\} = \sum_{x=-\infty}^{\infty} x \cdot w_X(x) dx$$

- ightharpoonup (Example: the entropy of H(X) = the average value of the information)
- ▶ Usual notation: μ

Properties of the average value

- ► Computing the average value is a **linear** operation
 - because the underlying integral / sum is a linear operation
- Linearity

$$E\{aX + bY\} = aE\{X\} + bE\{Y\}$$

Or:

$$E\{aX\} = aE\{X\}, \forall a \in \mathbb{R}$$
$$E\{X+Y\} = E\{X\} + E\{Y\}$$

No proof given here

Average squared value

- ► Average squared value = average value of the squared values
- ► Moment of order 2
- Continuous r.v.:

$$\overline{X^2} = E\{X^2\} = \int_{-\infty}^{\infty} x^2 \cdot w_X(x) dx$$

Discrete r.v.:

$$\overline{X^2} = E\{X^2\} = \sum_{x=0}^{\infty} x^2 \cdot w_X(x) dx$$

▶ Interpretation: average of squared values = average energy of a signal

Variance

- ▶ Variance= average squared value of the difference to the average value
- Continuous r.v.:

$$\sigma^2 = \overline{\{X - \mu\}^2} = \int_{-\infty}^{\infty} (x - \mu)^2 \cdot w_X(x) dx$$

Discrete r.v.:

$$\sigma^2 = \overline{\{X - \mu\}^2} = \sum_{-\infty}^{\infty} (x - \mu)^2 \cdot w_X(x) dx$$

- ▶ Interpretation: how much do the values vary around the average value
 - σ^2 = large: large spread around the average value
 - $ightharpoonup \sigma^2 = {\sf small:}$ values are concentrated around the average value

Relation between the three values

▶ Relation between the average value, the average squared value, and the dispersion:

$$\sigma^{2} = \overline{\{X - \mu\}^{2}}$$

$$= \overline{X^{2} - 2 \cdot X \cdot \mu + \mu^{2}}$$

$$= \overline{X^{2}} - 2\mu \overline{X} + \mu^{2}$$

$$= \overline{X^{2}} - \mu^{2}$$

Sum of random variables

- ▶ Sum of two or more **independent** r.v. is also a r.v.
- lts distribution = the **convolution** of the distributions of the two r.v.
- $\blacktriangleright \text{ If } Z = X + Y$

$$w(z) = w(x) \star w(y)$$

- ▶ Particular case: if X and Y are normal r.v., with $\mathcal{N}(\mu_X, \sigma_X^2)$ and $\mathcal{N}(\mu_Y, \sigma_Y^2)$, then:
 - ▶ Z is also a normal r.v., with $\mathcal{N}(\mu_Z, \sigma_Z^2)$, having:
 - average = sum of the two averages: $\mu_Z = \mu_X + \mu_Y$
 - dispersion = sum of the two dispersions: $\sigma_Z^2 = \sigma_X^2 + \sigma_Y^2$

I.2 Random processes

Random process

- ▶ A random process = a sequence of random variables indexed in time
- **Discrete-time** random process f[n] = a sequence of random variables at discrete moments of time
 - e.g.: a sequence 50 of throws of a dice, the daily price on the stock market
- ▶ Continuous-time random process f(t) = a continuous sequence of random variables at every moment
 - e.g.: a noise voltage signal, a speech signal
- ▶ Every sample from a random process is a (different) random variable!
 - e.g. $f(t_0) = \text{value at time } t_0 \text{ is a r.v.}$

Realizations of random processes

- ▶ A **realization** of the random process = a particular sequence of realizations of the underlying r.v.
 - e.g. we see a given noise signal on the oscilloscope, but we could have seen any other realization just as well
- When we consider a random process = we consider the set of all possible realizations

Distributions of order 1 of random processes

- ightharpoonup Every sample $f(t_1)$ from a random process is a random variable
 - \blacktriangleright with CDF $F_1(x; t_1)$
 - with PDF $w_1(x; t_1) = \frac{dF_1(x; t_1)}{dx}$
- ► The sample at time t₂ is a different random variable with **possibly different** functions
 - with CDF $F_1(x; t_2)$
 - with PDF $w_1(x; t_2) = \frac{dF_1(x; t_2)}{dx}$
- ▶ These functions specify how the value of one sample is distributed
- The index w_1 indicates we consider a single random variable from the process (distributions of order 1)
- Same for discrete-time random processes

Distributions of order 2

- A pair of random variables $f(t_1)$ and $f(t_2)$ sampled from the random process f(t) have
 - ightharpoonup joint CDF $F_2(x_i, x_i; t_1, t_2)$
 - ▶ joint PDF $w_2(x_i, x_j; t_1, t_2) = \frac{\partial^2 F_2(x_i, x_j; t_1, t_2)}{\partial x_i \partial x_i}$
- ► These functions specify how the pair of values is distributed (distributions of order 2)
- ► Same for discrete-time random processes

Distributions of order n

- Generalize to n samples of the random process
- A set of n random variables $f(t_1), ... f(t_n)$ sampled from the random process f(t) have
 - ightharpoonup joint CDF $F_n(x_1,...x_n;t_1,...t_n)$
 - ▶ joint PDF $w_n(x_1,...x_n;t_1,...t_n) = \frac{\partial^2 F_n(x_1,...x_n;t_1,...t_n)}{\partial x_1...\partial x_n}$
- ► These functions specify how the whole set of *n* values is distributed (distributions of order *n*)
- ► Same for discrete-time random processes

Statistical averages

Random processes are characterized using statistical / temporal averages (moments)

1. Average value

$$\overline{f(t_1)} = \mu(t_1) = \int_{-\infty}^{\infty} x \cdot w_1(x; t_1) dx$$

2. Average squared value (valoarea patratica medie)

$$\overline{f^2(t_1)} = \int_{-\infty}^{\infty} x^2 \cdot w_1(x; t_1) dx$$

Statistical averages - variance

3. Variance (= dispersia)

$$\sigma^{2}(t_{1}) = \overline{\{f(t_{1}) - \mu(t_{1})\}^{2}} = \int_{-\infty}^{\infty} (x - \mu(t_{1})^{2} \cdot w_{1}(x; t_{1}) dx$$

▶ The variance can be computed as:

$$\sigma^{2}(t_{1}) = \overline{\{f(t_{1}) - \mu(t_{1})\}^{2}}$$

$$= \overline{f(t_{1})^{2} - 2f(t_{1})\mu(t_{1}) + \mu(t_{1})^{2}}$$

$$= \overline{f^{2}(t_{1})} - \mu(t_{1})^{2}$$

- ► Note:
 - lacktriangle these three values are calculated across all realizations, at time t_1
 - they characterize only the sample at time t₁
 - ▶ at a different time t_2 , the r.v. $f(t_2)$ is different so all average values might be different

Statistical averages - autocorrelation

4. The autocorrelation function

$$R_{ff}(t_1, t_2) = \overline{f(t_1)f(t_2)} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_1 x_2 w_2(x_1, x_2; t_1, t_2) dx_1 dx_2$$

5. The correlation function (for different random processes f(t) and g(t))

$$R_{fg}(t_1, t_2) = \overline{f(t_1)g(t_2)} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_1 y_2 w_2(x_1, y_2; t_1, t_2) dx_1 dy_2$$

- ► Note:
 - hese functions may have different values for a different pair of values (t_1,t_2)

Temporal averages

- ▶ What to do when we only have access to a single realization $f^{(k)}(t)$?
- Compute values for a single realization $f^{(k)}(t)$, across all time moments
- 1. Temporal average value

$$\overline{f^{(k)}(t)} = \mu^{(k)} = \lim_{T \to \infty} \frac{1}{T} \int_{T/2}^{T/2} f^{(k)}(t) dt$$

2. Temporal average squared value

$$\overline{[f^{(k)}(t)]^2} = \lim_{T \to \infty} \frac{1}{T} \int_{-T/2}^{T/2} [f^{(k)}(t)]^2 dt$$

Temporal variance

3. Temporal variance

$$\sigma^2 = \overline{\{f^{(k)}(t) - \mu^{(k)}\}^2} = \lim_{T \to \infty} \frac{1}{T} \int_{-T/2}^{T/2} (f^{(k)}(t) - \mu^{(k)})^2 dt$$

▶ The variance can be computed as:

$$\sigma^2 = \overline{[f^{(k)}(t)]^2} - [\mu^{(k)}]^2$$

- ► Note:
 - these values do not depend anymore on time t (integrated)

Temporal autocorrelation

4. The temporal autocorrelation function

$$R_{ff}(t_1, t_2) = \overline{f^{(k)}(t_1 + t)f^{(k)}(t_2 + t)}$$

$$= \lim_{T \to \infty} \frac{1}{T} \int_{-T/2}^{T/2} f^{(k)}(t_1 + t)f^{(k)}(t_2 + t)dt$$

5. The temporal correlation function (for different random processes f(t) and g(t))

$$R_{fg}(t_1, t_2) = \overline{f^{(k)}(t_1 + t)g^{(k)}(t_2 + t)}$$

$$= \lim_{T \to \infty} \frac{1}{T} \int_{-T/2}^{T/2} f^{(k)}(t_1 + t)g^{(k)}(t_2 + t)dt$$

Statistical and temporal averages

- Statistical averages are usually the relevant values
- But in real life, we can only compute the temporal values
- ► Fortunately, in many cases they are the same (ergodicity, see later)

Stationary random processes

- ▶ All the statistical averages are dependent on the time
 - ightharpoonup i.e. they might be different for a sample at t_2
- ► Stationary random process = when all statistical averages are identical if we shift the time origin (e.g. delay the signal)
- ► Equivalent definition: if all the PDF are identical when shifting the time origin

$$w_n(x_1,...x_n;t_1,...t_n) = w_n(x_1,...x_n;t_1+\tau,...t_n+\tau)$$

Basically, nothing should depend on the time t

Strict-sense and wide-sense stationary

- Strictly stationary / strongly stationary / strict-sense stationary:
 - relation holds for every *n*
- Weakly stationary / wide-sense stationary:
 - relation holds only for n = 1 and n = 2 (the most used)

Consequences of stationarity

For n = 1:

$$w_1(x_i; t_1) = w_1(x_i; t_2) = w_1(x_i)$$

► The average value, average squared value, variance of a sample are all **identical** for any time *t*

$$\overline{f(t)} = constant, \forall t$$
 $\overline{f^2(t)} = constant, \forall t$
 $\sigma^2(t) = constant, \forall t$

Consequences of stationarity

▶ For n = 2:

$$w_2(x_i, x_j; t_1, t_2) = w_2(x_i, x_j; 0, t_2 - t_1) = w_2(x_i, x_j; t_2 - t_1)$$

The autocorrelation function depends only on the **time difference** $\tau = t_2 - t_1$ between the samples

$$R_{ff}(t_1, t_2) = R_{ff}(0, t_2 - t_1) = R_{ff}(\tau) = \overline{f(t)f(t + \tau)}$$

- lacktriangle Is the average value of a product of two samples time au apart
- lacktriangle Depends on a single value au= time difference of the two samples

Consequences of stationarity

- ▶ Same for correlation function between two different r.p
- lacktriangle Depends only on the **time difference** $au=t_2-t_1$ between the samples

$$R_{fg}(t_1, t_2) = R_{fg}(0, t_2 - t_1) = R_{fg}(\tau) = \overline{f(t)g(t + \tau)}$$

lacktriangle Is the average value of a product of two samples time au apart

Ergodic random processes

- ▶ In practice, we have access to a single realization
- ► **Ergodic** random process = the temporal averages on any realization are equal to the statistical averages
- ▶ We can compute / estimate all averages from a single realization (any)
 - lacktriangle the realization must be very long (length $ightarrow\infty$) for precise results
- Realizations are all similar to the others, statistically
 - a single realization is characteristic of the whole process

Ergodic random processes

- Most random processes we care about are ergodic and stationary
 - e.g. noises
- Example of non-ergodic process:
 - throw a dice, then the next 50 values are identical to the first
 - a single realization is not characteristic



The Power Spectral Density of a random process

- ► The Power Spectral Density (PSD) $S_{ff}(\omega)$ is the power of the random process at every frequency $f(\omega = 2\pi f)$
- ► The PSD describes how the power of a signal is distributed in frequency
 - e.g. some random processes have more power at low frequency, others at high frequency etc.
- ▶ The power in the frequency band $[f_1,f_2]$ is equal to $\int_{f_1}^{f_2} S_{ff}(\omega) d\omega$
- ▶ The whole power of the signal is $\int_{-\infty}^{\infty} S_{ff}(\omega) d\omega$
- ► The PSD is basically a measurable quantity
 - it can be determined experimentally
 - it is important in practical (engineering) applications

The Wiener-Khinchin theorem

Rom: teorema Wiener-Hincin

Theorem:

► The Power Spectral Density = the Fourier transform of the autocorrelation function

$$S_{ff}(\omega) = \int_{-\infty}^{\infty} R_{ff}(\tau) e^{-j\omega\tau} d\tau$$

$$R_{ff}(au) = rac{1}{2\pi} \int_{-\infty}^{\infty} S_{ff}(\omega) e^{j\omega au} d\omega$$

- No proof
- Makes a relation between two rather different domains
 - autocorrelation function: a statistical property
 - ▶ PSD function: a *physical* property (relevant for engineering purposes)

White noise

► White noise = a random process with autocorrelation function equal to a Dirac function

$$R_{ff}(\tau) = \delta(\tau)$$

- Any two different samples $(\tau \neq 0)$ have zero correlation (are uncorrelated)
 - they do not vary similarly
- Power spectral density = Fourier transform of a Dirac = a constant
 - lacktriangle has equal power at all frequencies up to ∞
- ▶ In real life, power goes to 0 for very high frequencies
 - "band-limited white noise"
 - Samples which are very close are necessarily somewhat correlated
- White noise can have almost any distribution
 - normal, uniform etc.

Properties of the autocorrelation function

1. Is even

$$R_{ff}(au) = R_{ff}(- au)$$

- ▶ Proof: change variable in definition
- 2. At infinite it goes to a constant

$$R_{ff}(\infty) = \overline{f(t)}^2 = const$$

- ightharpoonup Proof: two samples separated by ∞ are independent
- 3. Is maximum in 0

$$R_{ff}(0) \geq R_{ff}(\tau)$$

- ▶ Proof: start from $(f(t) f(t + \tau))^2 \ge 0$
- Interpretation: different samples might vary differently, but a sample always varies identically with itself

Properties of the autocorrelation function

4. Value in 0 = the power of the random process

$$R_{ff}(0) = rac{1}{2\pi} \int_{-\infty}^{\infty} S_{ff}(\omega) d\omega$$

- ightharpoonup Proof: Put au=0 in inverse Fourier transform of Wiener-Khinchin theorem
- 5. Variance = difference between values at 0 and ∞

$$\sigma^2 = R_{ff}(0) - R_{ff}(\infty)$$

▶ Proof: $R_{ff}(0) = \overline{f(t)^2}$, $R_{ff}(\infty) = \overline{f(t)}^2$

Autocorrelation of filtered random processes

- Consider a stationary random process applied as input to a LTI system
 - either continuous-time: input x(t), system H(s), output y(t)
 - or discrete-time: input x[n], system H(z), output y[n]
- \triangleright How does the autocorrelation of y depend on that of the input x?
- y is the convolution of x and the impulse response h

Computations

For discrete-time processes

$$R_{yy}(\tau) = \overline{y[n]y[n+\tau]}$$

$$= \sum_{k_1 = -\infty}^{\infty} h[k_1]x[n-k_1] \sum_{k_2 = -\infty}^{\infty} h[k_2]x[n+\tau-k_2]$$

$$= \sum_{k_1 = -\infty}^{\infty} \sum_{k_2 = -\infty}^{\infty} h[k_1]h[k_2]\overline{x[n-k_1]x[n+\tau-k_2]}$$

$$= \sum_{k_1 = -\infty}^{\infty} \sum_{k_2 = -\infty}^{\infty} h[k_1]h[k_2]R_{xx}[\tau-k_1+k_2]$$

From Wiener-Hincin theorem:

$$S_{ff}(\omega) = \sum_{\tau=-\infty}^{\infty} R_{ff}(\tau) e^{-j\omega\tau}$$

Computations

Therefore

$$S_{yy}(\omega) = \sum_{\tau = -\infty}^{\infty} \sum_{k_1 = -\infty}^{\infty} \sum_{k_2 = -\infty}^{\infty} h[k_1] h[k_2] R_{xx} [\tau - k_1 + k_2] e^{-j\omega\tau}$$

- ▶ Change of variable: $\tau k_1 + k_2 = u$
 - ▶ then $\tau = u + k_1 k_2$

$$S_{yy}(\omega) = \sum_{u=-\infty}^{\infty} \sum_{k_1=-\infty}^{\infty} \sum_{k_2=-\infty}^{\infty} h[k_1] h[k_2] R_{xx}[u] e^{-j\omega(u+k_1+k_2)}$$

$$= \sum_{u=-\infty}^{\infty} R_{xx}[u] e^{-j\omega u} \sum_{k_1=-\infty}^{\infty} h[k_1] e^{-j\omega k_1} \sum_{k_2=-\infty}^{\infty} h[k_2] e^{j\omega k_2}$$

$$= S_{xx}(\omega) \cdot H(\omega) \cdot H *^{(\omega)}$$

$$= S_{xx}(\omega) \cdot |H(\omega)|^2$$

Result

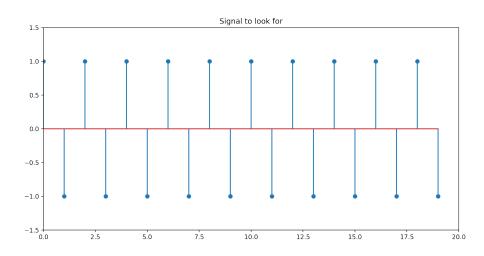
$$S_{yy}(\omega) = S_{xx}(\omega) \cdot |H(\omega)|^2$$

- ▶ The PSD of y =the PSD of x multiplied with the squared amplitude response of the filter
- Same relation is valid for continuous processes as well

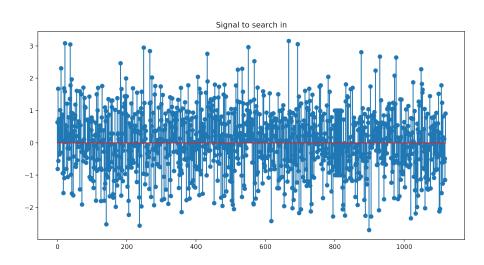
Applications of (auto)correlation

- Searching for a certain part in a large signal
- ► Correlation of two signals = measure of **similarity** of the two signals
 - ► The correlation function measures the similarity of a signal with all the shifted versions of the other
 - Example at blackboard
- Correlation can be used to locate data
 - The (auto)correlation function has large values when the two signals match
 - Large value when both positive and negative areas match,
 - ► Small values when they don't match

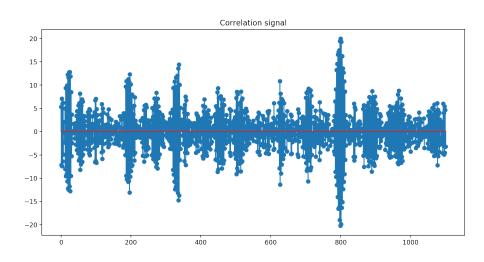
The signal to look for



The complete signal



Correlation result



System identification

- ▶ Determining the impulse response of an unknown LTI system
- Based on correlation between input and output of the system

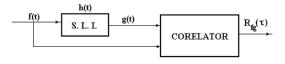


Figure 2: System identification setup

System identification

$$R_{fg}(\tau) = \overline{f[n]g[n+\tau]}$$

$$= \overline{f[n]} \sum_{k=-\infty}^{\infty} h[k]f[n+\tau-k]$$

$$= \sum_{k=-\infty}^{\infty} h[k]\overline{f[n]f[n+\tau-k]}$$

$$= \sum_{k=-\infty}^{\infty} h[k]R_{ff}[\tau-k]$$

$$= h[\tau] \star R_{ff}[\tau]$$

▶ If the input f is **white noise** with power A, $R_{ff}[n] = A \cdot \delta[n]$, and

$$R_{fg}(\tau) = h[\tau] \star R_{ff}[\tau] = A \cdot h[\tau] \star \delta[\tau] = A \cdot h[\tau]$$

► Then the correlation is proportional with the impulse response of the unknown system