

Decision and Estimation in Data Processing

Chapter I. Random Signals

I.1 Random variables

Random variables

- ▶ A **random variable** is a variable that holds a value produced by a (partially) random phenomenon
 - ▶ basically it is a *name* attached to an arbitrary value
 - ▶ short notation: r.v.
- ▶ Typically denoted as X , Y etc..
- ▶ Examples:
 - ▶ X = The value of a dice
 - ▶ V_{in} = The value of the voltage in one point of a circuit

Off-topic: Glossary

- ▶ “*i.e.*” = *id est* = “that is” = “adică”
- ▶ “*e.g.*” = *exempli gratia* = “for example” = “de exemplu”

Realizations

- ▶ **A realization** of a random variable = one possible value it can take
 - ▶ e.g. the value 3 of a dice
 - ▶ at different times, one may get different realizations
- ▶ **Sample space** Ω = the set of all values that can be taken by a random variable X
 - ▶ i.e. the set of all possible realizations
- ▶ Example: rolling a dice
 - ▶ The r.v. is denoted as X
 - ▶ We might get a realization $X = 6$
 - ▶ But we could have got any value from the sample space

$$\Omega = \{1, 2, 3, 4, 5, 6\}$$

Discrete and continuous random variables

- ▶ **Discrete** random variable: if Ω is a discrete set
 - ▶ Example: value of a dice
- ▶ **Continuous** random variable: if Ω is a continuous set
 - ▶ Example: a voltage value

Why random variables?

- ▶ Random variables are a great model for **noise**
- ▶ Examples:
 - ▶ Measure a voltage in a circuit
 - ▶ Measure several times, the value is never precisely the same. The values always *varies* a little.
 - ▶ i.e. it is affected by noise

Probability Mass Function

- ▶ Consider a **discrete** r.v. A
- ▶ The **probability mass function (PMF)** = the probability that A has value x

$$w_A(x) = P\{A = x\}$$

- ▶ Also known as the **distribution** of A
- ▶ Example: what is the PMF of a dice? Plot on board.

Computing probability based on PMF

- ▶ Probability that A is equal to some value v

$$P\{A = v\} = w_A(v)$$

- ▶ Probability that A is between a and b (including):

$$P\{a \leq A \leq b\} = \sum_{x=a}^b w_A(x)$$

Cumulative Distribution Function

- ▶ The **cumulative distribution function (CDF)** = the probability that the value of A is smaller or equal than x

$$F_A(x) = P\{A \leq x\}$$

- ▶ In Romanian: “*funcție de repartiție*”
- ▶ Example: what is the CDF of a dice? Plot on board.
- ▶ For discrete r.v., the CDF is “stairwise”

Computing probability based on CDF

- ▶ Probability that A is equal to some value v

$$P\{A = v\} = F_A(v) - F_A(v - 1)$$

- ▶ Probability that A is between a and b (including):

$$P\{a \leq A \leq b\} = F_A(b) - F_A(a - 1)$$

Relation between PMF and CDF

- ▶ CDF is the *cumulative sum* (i.e. the integral) of PMF

$$F_A(x) = \sum_{\text{all } t \leq x} w_A(t)$$

- ▶ Example for dice: easy to notice graphically

Probability Density Function

- ▶ Consider a **continuous** r.v. A
 - ▶ assume it takes values in some interval $[a, b]$
- ▶ The **Probability Density Function (PDF)** of A = probability that the value of A is in a small vicinity *epsilon* around x , divided by *epsilon*
- ▶ Denoted as $w_A(x)$, also known as **the distribution** of A
- ▶ Informally, the PDF gives the probability that the value of A is **close to** x

Probability of an exact value

- ▶ The probability that a continuous r.v. A is **exactly** equal to a value x is **zero**
 - ▶ because there are an infinity of possibilities (continuous)
 - ▶ That's why we can't define a probability mass function like for discrete r.v.
- ▶ That's why the PDF says **in a small vicinity** around some value x , and not precisely equal to x

Computing probability based on PDF

- ▶ Probability that A is equal to some value v is always 0

$$P\{A = v\} = 0$$

- ▶ Probability that A is between a and b = integral of PDF from a to b :

$$P\{a \leq A \leq b\} = \int_a^b w_A(x) dx$$

Cumulative Distribution Function

- ▶ The **cumulative distribution function (CDF)** = the probability that the value of A is smaller or equal than x

$$F_A(x) = P\{A \leq x\}$$

- ▶ In Romanian: “*funcție de repartiție*”
- ▶ Same definition as for discrete r.v.

Computing probability based on CDF

- ▶ Probability that A is between a and b :

$$P\{a \leq A \leq b\} = F_A(b) - F_A(a)$$

- ▶ Doesn't matter if we consider closed or open interval
 - ▶ $[a, b]$ or (a, b)
 - ▶ why?

Relation between PDF and CDF

- ▶ CDF is **the integral** of PMF
- ▶ PDF is **the derivative** of CDF

$$F_A(x) = \int_{-\infty}^x w_A(x) dx$$

$$\begin{aligned} w_A(x) &= \frac{dF_A(x)}{dx} \\ &= \lim_{\epsilon \rightarrow 0} \frac{F_A(x + \epsilon) - F_A(x - \epsilon)}{2\epsilon} \\ &= \lim_{\epsilon \rightarrow 0} \frac{P(A \in [x - \epsilon, x + \epsilon])}{2\epsilon} \end{aligned}$$

Graphical interpretation

- ▶ Probability that a continuous r.v. A is between a and b is **the area below the PDF**
 - ▶ i.e. the integral from a to b
- ▶ Probability that A is exactly equal to a certain value is zero
 - ▶ the area below a single point is zero

Discrete vs continuous r.v.

Comparison of discrete vs continuous random variables:

- ▶ The CDF $F_A(x)$ is defined identically, means same thing
- ▶ The PDF/PMF $w_A(x)$ is the derivative of CDF
 - ▶ for continuous r.v.:
 - ▶ it is a proper derivative
 - ▶ it means probability to be “around” x
 - ▶ for discrete r.v.:
 - ▶ sort of “discrete derivative”
 - ▶ it means probability to be exactly equal to x

Properties of random variables

CDF:

- ▶ The CDF is always ≥ 0
- ▶ The CDF is always monotonously increasing (non-decreasing)
- ▶ The CDF starts from 0 and goes up to 1

$$F_A(-\infty) = 0 \quad F_A(\infty) = 1$$

PDF/PMF:

- ▶ The PDF/PMF are always ≥ 0
- ▶ Integral/sum over all of the PDF/PMF = 1

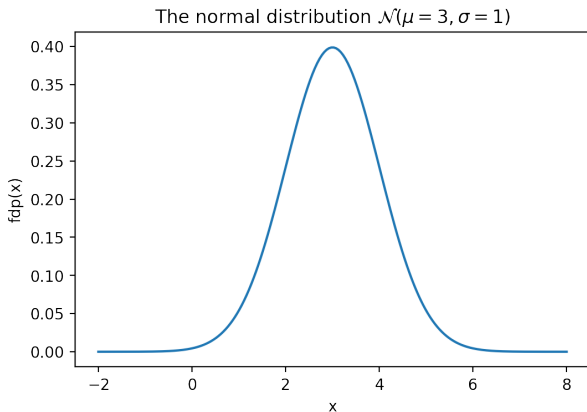
$$\int_{-\infty}^{\infty} w_A(x) dx = 1$$

$$\sum_{x=-\infty}^{\infty} w_A(x) = 1$$

The normal distribution

► Probability density function

$$w_A(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$



The normal distribution

- ▶ Has two parameters:
 - ▶ **Average value** μ = “center” of the function
 - ▶ **Standard deviation** σ = “width” of the function
 - ▶ Small σ = narrow and tall
 - ▶ Big σ = wide and low
- ▶ The front constant is just for normalization (ensures that integral = 1)
- ▶ Extremely often encountered in real life
- ▶ Any real value is possible ($w_A(x) > 0, \forall x \in \mathbb{R}$)
- ▶ Usually denoted as $\mathcal{N}(\mu, \sigma^2)$

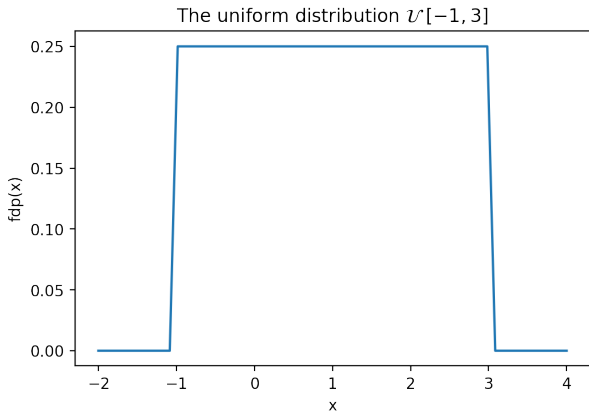
The normal distribution

- ▶ The distribution decreases as x gets farther from μ
 - ▶ Because of the term $-(x - \mu)^2$ at the exponent
 - ▶ Most likely values: around μ ($x - \mu = 0$)
 - ▶ Values closer to μ are more likely, values farther from μ are less likely
- ▶ The function describes a preference for values around μ , with decreasing preference when getting farther from μ

The uniform distribution

- The probability density function = a constant, between two endpoints

$$w_A(x) = \begin{cases} \frac{1}{b-a}, & x \in [a, b] \\ 0, & \text{elsewhere} \end{cases}$$



The uniform distribution

- ▶ Has two parameters: the limits a and b of the interval
- ▶ The “height” of the function is $\frac{1}{b-a}$
 - ▶ in order for the integral to be 1
- ▶ Only values from the interval $[a, b]$ are possible
 - ▶ value cannot be outside interval (probability is 0)
- ▶ Denoted as $\mathcal{U} [a, b]$

Other distributions

- ▶ Many other distributions exist, relevant for particular applications

Computing probabilities for the normal distribution

- ▶ How to compute \int_a^b for a normal distribution?
 - ▶ Can't be done with algebraic formula, non-elementary function
- ▶ Use *the error function*:

$$\operatorname{erf}(z) = \frac{2}{\sqrt{\pi}} \int_0^z e^{-t^2} dt$$

- ▶ The CDF of a normal distribution $\mathcal{N}(\mu, \sigma^2)$

$$F_A(X) = \frac{1}{2} \left(1 + \operatorname{erf} \left(\frac{x - \mu}{\sigma \sqrt{2}} \right) \right)$$

- ▶ The values of $\operatorname{erf}()$ are available / are computed numerically
 - ▶ e.g. on Google, search for $\operatorname{erf}(0.5)$
 - ▶ Other useful values:
 - ▶ $\operatorname{erf}(-\infty) = -1$
 - ▶ $\operatorname{erf}(\infty) = 1$

Exercise

Exercise:

- ▶ Let X be a r.v. with distribution $\mathcal{N}(3, 2)$. Compute the probability that $X \in [2, 4]$

Sum of constant + random variable

- ▶ Consider a random variable A
- ▶ What is $B = 5 + A$?

Answer:

- ▶ B is also a random variable
- ▶ B has same type of distribution, but the function is “shifted” by 5 to the right

Example:

- ▶ A is normal variable with $w_A(x) = \mathcal{N}(\mu = 3, \sigma^2 = 2)$
- ▶ What is the distribution of $B = 5 + A$?
- ▶ Answer: $w_B(x) = \mathcal{N}(\mu = 8, \sigma^2 = 2)$

R.v. as functions of other r.v.

- ▶ A function applied to a r.v. produces another r.v.
- ▶ Examples: if B is a r.v. with distribution $\mathcal{U}[0, 10]$, then
 - ▶ $C = 5 + A$ is another r.v., with distribution $\mathcal{U}[5, 15]$
 - ▶ $D = A^2$ is also another r.v.
 - ▶ $E = \cos(A)$ is also another r.v.
- ▶ Reason: since A is random, the values B , C , D are also random
- ▶ A , B , C , D are *not independent*
 - ▶ A certain value of one of them automatically implies the value of the others

Multiple random variables

- ▶ Consider a system with two continuous r.v. A and B
- ▶ What is the probability that the pair (A, B) has values around (x, y) ?
- ▶ Distribution of the values of (A, B) is described by:
 - ▶ joint probability density function $w_{AB}(x, y)$
 - ▶ joint cumulative density function $F_{AB}(x, y)$

Multiple random variables

- ▶ Joint cumulative distribution function:

$$F_{AB}(x, y) = P \{A \leq x \cap B \leq y\}$$

- ▶ Joint probability density function:

$$w_{AB}(x, y) = \frac{\partial^2 P_{AB}(x, y)}{\partial x \partial y}$$

- ▶ The joint PDF gives the probability that the value of the pair (A, B) is in a vicinity of (x, y)
- ▶ Similar for discrete random variables

$$w_{AB}(x, y) = P \{A = x \cap B = y\}$$

Independent random variables

- ▶ Two v.a. A and B are **independent** if the value of one of them does not influence in any way the value of the other
- ▶ For independent r.v., the probability that A is around x and B is around y is **the product** of the two probabilities

$$w_{AB}(x, y) = w_A(x) \cdot w_B(y)$$

- ▶ Relation holds for CDF / PDF / PMF, continuous or discrete r.v.
- ▶ Same for more than two r.v.

Independent random variables

Exercise:

- ▶ Compute the probability that three r.v. X , Y and Z i.i.d. $\mathcal{N}(-1, 1)$ are all positive simultaneously
 - ▶ ***i.i.d*** = “independent and identically distributed”

Multiple normal variables

- ▶ Consider a set of N normal r.v. (A_1, \dots, A_N) , with different μ_i , but same σ
- ▶ Then probability that (A_1, \dots, A_N) is around (x_1, \dots, x_N) is

$$w_{A_1, \dots, A_N}(x_1, \dots, x_N) = \frac{1}{(\sigma\sqrt{2\pi})^N} e^{-\frac{(x_1 - \mu_1)^2 + \dots + (x_N - \mu_N)^2}{2\sigma^2}}$$

- ▶ The probability depends on the **Euclidean distance** between $\mathbf{x} = (x_1, \dots, x_N)$ and $\boldsymbol{\mu} = (\mu_1, \dots, \mu_N)$

Euclidean distance

- ▶ **Euclidean (geometric) distance** between two N-dimensional vectors:

$$d(\mathbf{u}, \mathbf{v}) = \|\mathbf{u} - \mathbf{v}\| = \sqrt{(u_1 - v_1)^2 + \dots + (u_N - v_N)^2}$$

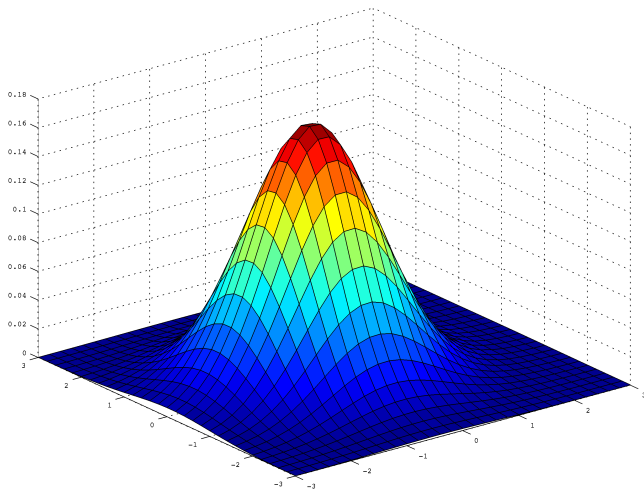
- ▶ One-dimensional: $\|\mathbf{u} - \mathbf{v}\| = |u - v|$
- ▶ 2D: $\|\mathbf{u} - \mathbf{v}\| = \sqrt{(u_1 - v_1)^2 + (u_2 - v_2)^2}$
- ▶ 3D: $\|\mathbf{u} - \mathbf{v}\| = \sqrt{(u_1 - v_1)^2 + (u_2 - v_2)^2 + (u_3 - v_3)^2}$
- ▶ ...
- ▶ N-dimensional: $\|\mathbf{u} - \mathbf{v}\| = \sqrt{\sum_{i=1}^N (u_i - v_i)^2}$
- ▶ ...
- ▶ Continuous signals: $\|\mathbf{u} - \mathbf{v}\| = \sqrt{\int_{-\infty}^{\infty} (u(t) - v(t))^2 dt}$

Multiple normal variables

- ▶ Probability of N normal random variables, independent, with same σ but possibly different μ_i depends on the **squared Euclidean distance to the mean vector** $\mu = (\mu_1, \dots, \mu_N)$
 - ▶ Close to μ : higher probability
 - ▶ Far from μ : lower probability
 - ▶ Two points at same distance from μ have same probability

2D normal distribution

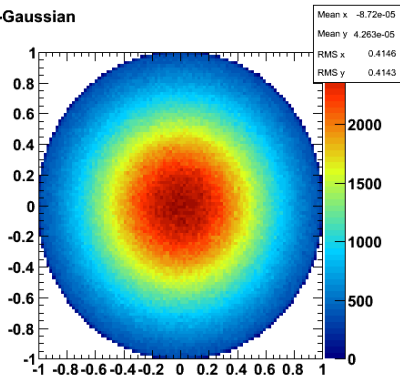
- Distribution of 2 normal random variables (2D normal distribution)



2D normal distribution - top view

- ▶ View from top
- ▶ Here, $\mu = (0, 0)$
- ▶ Probability density decreases as distance from center increases, in circles (symmetrically)

2D-Gaussian



Statistical averages

- ▶ R.v. are described by statistical averages (“*moments*”)
- ▶ **The average value** (moment of order 1)
- ▶ Continuous r.v.:

$$\bar{A} = E\{A\} = \int_{-\infty}^{\infty} x \cdot w_A(x) dx$$

- ▶ Discrete r.v.:

$$\bar{A} = E\{A\} = \sum_{x=-\infty}^{\infty} x \cdot w_A(x)$$

- ▶ (Example: the entropy of $H(X)$ = the average value of the information)
- ▶ Usual notation: μ

Properties of the average value

- ▶ Computing the average value is a **linear** operation
 - ▶ because the underlying integral / sum is a linear operation

- ▶ Linearity

$$E\{c_1A + c_2B\} = c_1E\{A\} + c_2E\{B\}$$

- ▶ Or:

$$E\{cA\} = cE\{A\}, \forall c \in \mathbb{R}$$

$$E\{A + B\} = E\{A\} + E\{B\}$$

- ▶ No proof given here

Average squared value

- ▶ **Average squared value** = average value of the squared values
- ▶ Moment of order 2
- ▶ Continuous r.v.:

$$\overline{A^2} = E\{A^2\} = \int_{-\infty}^{\infty} x^2 \cdot w_A(x) dx$$

- ▶ Discrete r.v.:

$$\overline{A^2} = E\{A^2\} = \sum_{-\infty}^{\infty} x^2 \cdot w_A(x)$$

- ▶ Interpretation: average of squared values = average power of a signal

Variance

- ▶ **Variance**= average squared value of the difference to the average value
- ▶ Continuous r.v.:

$$\sigma^2 = \overline{\{A - \mu\}^2} = \int_{-\infty}^{\infty} (x - \mu)^2 \cdot w_A(x) dx$$

- ▶ Discrete r.v.:

$$\sigma^2 = \overline{\{A - \mu\}^2} = \sum_{-\infty}^{\infty} (x - \mu)^2 \cdot w_A(x)$$

- ▶ Interpretation: how much do the values vary around the average value
 - ▶ $\sigma^2 = \text{large}$: large spread around the average value
 - ▶ $\sigma^2 = \text{small}$: values are concentrated around the average value

Relation between the three values

- Relation between the average value, the average squared value, and the variance:

$$\begin{aligned}\sigma^2 &= \overline{\{A - \mu\}^2} \\ &= \overline{A^2 - 2 \cdot A \cdot \mu + \mu^2} \\ &= \overline{A^2} - 2\mu\overline{A} + \mu^2 \\ &= \overline{A^2} - \mu^2\end{aligned}$$

Sum of random variables

- ▶ Sum of two or more **independent** r.v. is also a r.v.
- ▶ Its distribution = the **convolution** of the distributions of the two r.v.
- ▶ If $C = A + B$

$$w_C(x) = w_A(x) \star w_B(x)$$

- ▶ Particular case: if A and B are normal r.v., with $\mathcal{N}(\mu_A, \sigma_A^2)$ and $\mathcal{N}(\mu_B, \sigma_B^2)$, then:
 - ▶ C is also a normal r.v., with $\mathcal{N}(\mu_C, \sigma_C^2)$, having:
 - ▶ average = sum of the two averages: $\mu_C = \mu_A + \mu_B$
 - ▶ variance = sum of the two variances: $\sigma_C^2 = \sigma_A^2 + \sigma_B^2$

I.2 Random processes

Random process

- ▶ A **random process** = a sequence of random variables indexed in time
- ▶ **Discrete-time** random process $f[n]$ = a sequence of random variables at discrete moments of time
 - ▶ e.g.: a sequence 50 of throws of a dice, the daily price on the stock market
- ▶ **Continuous-time** random process $f(t)$ = a continuous sequence of random variables at every moment
 - ▶ e.g.: a noise voltage signal, a speech signal
- ▶ Every sample from a random process is a (different) random variable!
 - ▶ e.g. $f(t_0)$ = value at time t_0 is a r.v.

Realizations of random processes

- ▶ A **realization** of the random process = a particular sequence of realizations of the underlying r.v.
 - ▶ e.g. we see a given noise signal on the oscilloscope, but *we could have seen any other realization just as well*
- ▶ When we consider a random process = we consider the set of all possible realizations

Distributions of order 1 of random processes

- ▶ Every sample $f(t_1)$ from a random process is a random variable
 - ▶ it is described by a **distribution of order 1**
 - ▶ has a CDF $F_1(x; t_1)$
 - ▶ has a PDF $w_1(x; t_1) = \frac{dF_1(x; t_1)}{dx}$
 - ▶ everything depends on the time moment t_1
- ▶ The sample at time t_2 is a different random variable with **possibly different** functions
 - ▶ has a different CDF $F_1(x; t_2)$
 - ▶ has a different PDF $w_1(x; t_2) = \frac{dF_1(x; t_2)}{dx}$
- ▶ These functions specify how the value of one sample is distributed
- ▶ The index w_1 indicates we consider a single random variable (distribution of order 1)
- ▶ Same for discrete-time random processes

Distributions of order 2

- ▶ A pair of random variables $f(t_1)$ and $f(t_2)$ form a system of 2 r.v.
 - ▶ they are described by a **distribution of order 2**
 - ▶ have a joint CDF $F_2(x_i, x_j; t_1, t_2)$
 - ▶ have a joint PDF $w_2(x_i, x_j; t_1, t_2) = \frac{\partial^2 F_2(x_i, x_j; t_1, t_2)}{\partial x_i \partial x_j}$
 - ▶ depend on time moments t_1 and t_2
- ▶ These functions specify how the pair of values is distributed
- ▶ Same for discrete-time random processes

Distributions of order n

- ▶ Generalize to n samples of the random process
- ▶ A set of n random variables $f(t_1), \dots, f(t_n)$ from the random process $f(t)$
 - ▶ are described by **distribution of order n**
 - ▶ have joint CDF $F_n(x_1, \dots, x_n; t_1, \dots, t_n)$
 - ▶ have joint PDF $w_n(x_1, \dots, x_n; t_1, \dots, t_n) = \frac{\partial^2 F_n(x_1, \dots, x_n; t_1, \dots, t_n)}{\partial x_1 \dots \partial x_n}$
 - ▶ depend on time moments t_1, t_2, \dots, t_n
- ▶ These functions specify how the whole set of n values is distributed
- ▶ Same for discrete-time random processes

Statistical averages

Random processes are characterized using statistical and temporal averages (*moments*)

For continuous random processes:

1. Average value

$$\overline{f(t_1)} = \mu(t_1) = \int_{-\infty}^{\infty} x \cdot w_1(x; t_1) dx$$

2. Average squared value (*valoarea patratica medie*)

$$\overline{f^2(t_1)} = \int_{-\infty}^{\infty} x^2 \cdot w_1(x; t_1) dx$$

Statistical averages - variance

3. Variance (= varianța)

$$\sigma^2(t_1) = \overline{\{f(t_1) - \mu(t_1)\}^2} = \int_{-\infty}^{\infty} (x - \mu(t_1))^2 \cdot w_1(x; t_1) dx$$

- The variance can be computed as:

$$\begin{aligned}\sigma^2(t_1) &= \overline{\{f(t_1) - \mu(t_1)\}^2} \\ &= \overline{f(t_1)^2 - 2f(t_1)\mu(t_1) + \mu(t_1)^2} \\ &= \overline{f^2(t_1)} - \mu(t_1)^2\end{aligned}$$

- Note:

- these three values are calculated across all realizations, at time t_1
- they characterize only the sample at time t_1
- at a different time t_2 , the r.v. $f(t_2)$ is different so all average values might be different

Statistical averages - autocorrelation

4. The autocorrelation function

$$R_{ff}(t_1, t_2) = \overline{f(t_1)f(t_2)} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_1 x_2 w_2(x_1, x_2; t_1, t_2) dx_1 dx_2$$

5. The correlation function (for different random processes $f(t)$ and $g(t)$)

$$R_{fg}(t_1, t_2) = \overline{f(t_1)g(t_2)} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_1 y_2 w_2(x_1, y_2; t_1, t_2) dx_1 dy_2$$

► Note:

- these functions may have different values for a different pair of values (t_1, t_2)

Discrete random processes

For **discrete random processes**, nothing changes (except notation from $f(t)$ to $f[t]$):

Temporal averages

- ▶ What to do when we only have access to a single realization $f^{(k)}(t)$?
- ▶ Compute values **for a single realization** $f^{(k)}(t)$, **across all time moments**
- ▶ For continuous random processes:

1. Temporal average value

$$\overline{f^{(k)}(t)} = \mu^{(k)} = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T f^{(k)}(t) dt$$

2. Temporal average squared value

$$\overline{[f^{(k)}(t)]^2} = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T [f^{(k)}(t)]^2 dt$$

3. Temporal variance

$$\sigma^2 = \overline{\{f^{(k)}(t) - \mu^{(k)}\}^2} = \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T (f^{(k)}(t) - \mu^{(k)})^2 dt$$

- ▶ The variance can be computed as:

$$\sigma^2 = \overline{[f^{(k)}(t)]^2} - [\mu^{(k)}]^2$$

- ▶ Note:

- ▶ these values do not depend anymore on time t (integrated)

Temporal autocorrelation

4. The temporal autocorrelation function

$$\begin{aligned} R_{ff}(t_1, t_2) &= \overline{f^{(k)}(t_1 + t)f^{(k)}(t_2 + t)} \\ &= \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T f^{(k)}(t_1 + t)f^{(k)}(t_2 + t)dt \end{aligned}$$

5. The temporal correlation function (for different random processes $f(t)$ and $g(t)$)

$$\begin{aligned} R_{fg}(t_1, t_2) &= \overline{f^{(k)}(t_1 + t)g^{(k)}(t_2 + t)} \\ &= \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-T}^T f^{(k)}(t_1 + t)g^{(k)}(t_2 + t)dt \end{aligned}$$

Discrete random processes

For **discrete random processes**, replace \int with \sum , T with N , and divide to $2N + 1$ instead of $2T$

$$1. \overline{f^{(k)}[t]} = \mu^{(k)} = \lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{t=-N}^N f^{(k)}[t]$$

$$2. \overline{[f^{(k)}[t]]^2} = \lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{t=-N}^N (f^{(k)}[t])^2$$

$$3. \sigma^2 = \overline{\{f^{(k)}[t] - \mu^{(k)}\}^2} = \lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{t=-N}^N (f^{(k)}[t] - \mu^{(k)})^2$$

Discrete random processes

4. Temporal autocorrelation:

$$\begin{aligned} R_{ff}(t_1, t_2) &= \overline{f^{(k)}[t_1 + t]f^{(k)}[t_2 + t]} \\ &= \lim_{N \rightarrow \infty} \frac{1}{2N + 1} \sum_{t=-N}^N f^{(k)}[t_1 + t]f^{(k)}[t_2 + t] \end{aligned}$$

5. Temporal correlation:

$$\begin{aligned} R_{fg}(t_1, t_2) &= \overline{f^{(k)}[t_1 + t]g^{(k)}[t_2 + t]} \\ &= \lim_{N \rightarrow \infty} \frac{1}{2N + 1} \sum_{t=-N}^N f^{(k)}[t_1 + t]g^{(k)}[t_2 + t] \end{aligned}$$

Finite length realizations

If the realization is not from time $-\infty$ to ∞ , but only from a t_{min} to t_{max} , just use $\int_{t_{min}}^{t_{max}}$ or $\sum_{t_{min}}^{t_{max}}$ for the temporal averages

- Example: Compute the temporal averages for the finite-length realization

$$\{1, -1, 2, -2, 3, -3, 4, -4, 5, -5\}$$

Statistical and temporal averages

- ▶ Statistical averages are usually the relevant values
 - ▶ but they require to know the distributions
- ▶ In real life, with unknown signals, we can only measure one realization
 - ▶ so we can only compute the temporal values for one realization
- ▶ Fortunately, in many cases they are the same (ergodicity, see later)

Stationary random processes

- ▶ All the statistical averages are dependent on the time
 - ▶ i.e. they might be different for a sample at t_2
- ▶ **Stationary** random process = when all statistical averages are **identical if we shift the time origin** (e.g. delay the signal)
- ▶ Equivalent definition: if all the PDF are identical when shifting the time origin

$$w_n(x_1, \dots, x_n; t_1, \dots, t_n) = w_n(x_1, \dots, x_n; t_1 + \tau, \dots, t_n + \tau)$$

- ▶ Basically, nothing should depend on the time t

Strict-sense and wide-sense stationary

- ▶ Strictly stationary / strongly stationary / strict-sense stationary:
 - ▶ relation holds for every n
- ▶ Weakly stationary / wide-sense stationary:
 - ▶ relation holds only for $n = 1$ and $n = 2$ (the most used)

Consequences of stationarity

- ▶ For $n = 1$:

$$w_1(x_i; t_1) = w_1(x_i; t_2) = w_1(x_i)$$

- ▶ The average value, average squared value, variance of a sample are all **identical** for any time t

$$\overline{f(t)} = \text{constant}, \forall t$$

$$\overline{f^2(t)} = \text{constant}, \forall t$$

$$\sigma^2(t) = \text{constant}, \forall t$$

Consequences of stationarity

- ▶ For $n = 2$:

$$w_2(x_i, x_j; t_1, t_2) = w_2(x_i, x_j; 0, t_2 - t_1) = w_2(x_i, x_j; t_2 - t_1)$$

- ▶ The autocorrelation function depends only on the **time difference** $\tau = t_2 - t_1$ between the samples

$$R_{ff}(t_1, t_2) = R_{ff}(0, t_2 - t_1) = R_{ff}(\tau) = \overline{f(t)f(t + \tau)}$$

- ▶ Depends on a single value $\tau =$ time difference of the two samples

Consequences of stationarity

- ▶ Definition of autocorrelation function for **stationary** r.p:
 - ▶ the function now depends on $\tau = t_2 - t_1$, instead of t_1 and t_2
- ▶ Statistical autocorrelation: no change
- ▶ Temporal autocorrelation:
 - ▶ for continuous r.p.

$$\begin{aligned} R_{ff}(\tau) &= \overline{f(t)f(t+\tau)} \\ &= \lim_{T \rightarrow \infty} \frac{1}{T} \int_{-T/2}^{T/2} f^{(k)}(t)f^{(k)}(t+\tau)dt \end{aligned}$$

- ▶ for discrete r.p.

$$\begin{aligned} R_{ff}(\tau) &= \overline{f(t)f(t+\tau)} \\ &= \lim_{N \rightarrow \infty} \frac{1}{2N+1} \sum_{t=-N}^N f^{(k)}[t]f^{(k)}[t+\tau] \end{aligned}$$

- ▶ finite length: limit the integrals / sums to the length of the signal,
 $\int_{t_{min}}^{t_{max}}$ or $\sum_{t_{min}}^{t_{max}}$

Consequences of stationarity

- ▶ Same for correlation function between two different r.p
- ▶ Depends only on the **time difference** $\tau = t_2 - t_1$ between the samples

$$R_{fg}(t_1, t_2) = R_{fg}(0, t_2 - t_1) = R_{fg}(\tau) = \overline{f(t)g(t + \tau)}$$

- ▶ Definition is similar to the autocorrelation definition on the previous slide

Interpretation of autocorrelation

- ▶ $R_{ff}(\tau)$ = the average value of the product of two samples which are time τ apart
 - ▶ e.g. tells us if the two samples vary in same direction or not
- ▶ Same for correlation, but the samples are taken from different r.p f and g
- ▶ Example:
 - ▶ $R_{ff}(0.5) > 0$ means two samples separated by 0.5 seconds tend to vary in same direction (both positive, both negative \Rightarrow their product is mostly positive)
 - ▶ $R_{ff}(1) < 0$ means two samples separated by 1 second tend to vary in opposite directions (when one is positive, the other is negative \Rightarrow their product is mostly negative)
 - ▶ $R_{ff}(2) = 0$ means two samples separated by 2 seconds are uncorrelated (their product is 0 on average, so equally positive and negative)

Ergodic random processes

- ▶ In practice, we have access to a single realization
- ▶ **Ergodic** random process = the temporal averages on any realization are equal to the statistical averages
- ▶ Ergodicity means:
 - ▶ We can compute / estimate all averages from a single realization (any)
 - ▶ but the realization must be very long ($\text{length} \rightarrow \infty$) for precise results
 - ▶ Realizations are all similar to the others, statistically
 - ▶ so a single realization is characteristic of the whole process

Ergodic random processes

- ▶ Most random processes we care about are ergodic and stationary
 - ▶ e.g. voltage noises
- ▶ Example of non-ergodic process:
 - ▶ throw a dice, then the next 50 values are identical to the first
 - ▶ a single realization is not characteristic

I.3 More on autocorrelation

The Power Spectral Density of a random process

- ▶ The Power Spectral Density (PSD) $S_{ff}(\omega)$ is the power of the random process at every frequency f ($\omega = 2\pi f$)
- ▶ The PSD describes how the power of a signal is distributed in frequency
 - ▶ e.g. some random processes have more power at low frequency, others at high frequency etc.
- ▶ The power in the frequency band $[f_1, f_2]$ is equal to $\int_{f_1}^{f_2} S_{ff}(\omega) d\omega$
- ▶ The whole power of the signal is $P = \int_{-\infty}^{\infty} S_{ff}(\omega) d\omega$
- ▶ The PSD is a measurable quantity
 - ▶ it can be determined experimentally
 - ▶ it is important in practical (engineering) applications

The Wiener-Khinchin theorem

- *Rom: teorema Wiener-Hincin*

Theorem:

- **The Power Spectral Density = the Fourier transform of the autocorrelation function**

$$S_{ff}(\omega) = \int_{-\infty}^{\infty} R_{ff}(\tau) e^{-j\omega\tau} d\tau$$

$$R_{ff}(\tau) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{ff}(\omega) e^{j\omega\tau} d\omega$$

- No proof
 - Makes a relation between two rather different domains
 - autocorrelation function: a *statistical* property
 - PSD function: a *physical* property (relevant for engineering purposes)

White noise

- ▶ **White noise** = a random process with autocorrelation function equal to a Dirac function

$$R_{ff}(\tau) = \delta(\tau)$$

- ▶ is a random process: every sample of white noise is a random variable
 - ▶ autocorrelation is a Dirac: autocorrelation is 0 for any $\tau \neq 0$
 - ▶ any two different samples ($\tau \neq 0$) have zero correlation (are uncorrelated)
 - ▶ values of any two different samples are not related
- ▶ Power spectral density of white noise = Fourier transform of a Dirac = a constant $\forall \omega$
 - ▶ equal distribution of power at all frequencies up to ∞
- ▶ White noise can have any distribution (normal, uniform etc.)
 - ▶ the term “white noise” doesn’t refer to the distribution of sample values, but to the fact that all samples are unrelated to each other

Band-limited white noise

- ▶ In real life, power spectral density goes to 0 at very high frequencies
 - ▶ because total power $P = \int_{-\infty}^{\infty} S_{ff}\omega$ cannot be infinite
 - ▶ known as “*band-limited white noise*”
- ▶ In this case, autocorrelation = approximately a Dirac, but not infinitely thin
 - ▶ samples which are very close are necessarily a bit correlated
 - ▶ e.g. due to small parasitic capacities

- ▶ **AWGN** = Additive White Gaussian Noise
 - ▶ is the usual type of noise considered in applications
- ▶ It means:
 - ▶ additive: the noise is added to the original signal (e.g. not multiplied with it)
 - ▶ gaussian: the samples have normal distribution
 - ▶ white: the samples are uncorrelated (unrelated) with each other

Properties of the autocorrelation function

1. Is even

$$R_{ff}(\tau) = R_{ff}(-\tau)$$

- ▶ Proof: change variable in definition

2. At infinite it goes to a constant

$$R_{ff}(\infty) = \overline{f(t)}^2 = \text{const}$$

- ▶ Proof: two samples separated by ∞ are independent

3. Is maximum in 0

$$R_{ff}(0) \geq R_{ff}(\tau)$$

- ▶ Proof: start from $\overline{(f(t) - f(t + \tau))^2} \geq 0$
- ▶ Interpretation: different samples might vary differently, but a sample always varies identically with itself

Properties of the autocorrelation function

4. Value in 0 = the power of the random process

$$R_{ff}(0) = \frac{1}{2\pi} \int_{-\infty}^{\infty} S_{ff}(\omega) d\omega$$

- Proof: Put $\tau = 0$ in inverse Fourier transform of Wiener-Khinchin theorem

5. Variance = difference between values at 0 and ∞

$$\sigma^2 = R_{ff}(0) - R_{ff}(\infty)$$

- Proof: $R_{ff}(0) = \overline{f(t)^2}$, $R_{ff}(\infty) = \overline{f(t)}^2$

Autocorrelation of filtered random processes

- ▶ Consider a stationary random process applied as input to a LTI system
 - ▶ either continuous-time: input $x(t)$, system $H(s)$, output $y(t)$
 - ▶ or discrete-time: input $x[n]$, system $H(z)$, output $y[n]$
- ▶ How does the autocorrelation of y depend on that of the input x ?
 - ▶ y is the convolution between x and the impulse response h

Computations

- For discrete-time processes

$$\begin{aligned}R_{yy}(\tau) &= \overline{y[n]y[n+\tau]} \\&= \overline{\sum_{k_1=-\infty}^{\infty} h[k_1]x[n-k_1] \sum_{k_2=-\infty}^{\infty} h[k_2]x[n+\tau-k_2]} \\&= \sum_{k_1=-\infty}^{\infty} \sum_{k_2=-\infty}^{\infty} h[k_1]h[k_2]\overline{x[n-k_1]x[n+\tau-k_2]} \\&= \sum_{k_1=-\infty}^{\infty} \sum_{k_2=-\infty}^{\infty} h[k_1]h[k_2]R_{xx}[\tau-k_1+k_2]\end{aligned}$$

- From Wiener-Hincin theorem:

$$S_{ff}(\omega) = \sum_{\tau=-\infty}^{\infty} R_{ff}(\tau)e^{-j\omega\tau}$$

Computations

- Therefore

$$S_{yy}(\omega) = \sum_{\tau=-\infty}^{\infty} \sum_{k_1=-\infty}^{\infty} \sum_{k_2=-\infty}^{\infty} h[k_1]h[k_2]R_{xx}[\tau - k_1 + k_2]e^{-j\omega\tau}$$

- Change of variable: $\tau - k_1 + k_2 = u$

- then $\tau = u + k_1 - k_2$

$$\begin{aligned} S_{yy}(\omega) &= \sum_{u=-\infty}^{\infty} \sum_{k_1=-\infty}^{\infty} \sum_{k_2=-\infty}^{\infty} h[k_1]h[k_2]R_{xx}[u]e^{-j\omega(u+k_1+k_2)} \\ &= \sum_{u=-\infty}^{\infty} R_{xx}[u]e^{-j\omega u} \sum_{k_1=-\infty}^{\infty} h[k_1]e^{-j\omega k_1} \sum_{k_2=-\infty}^{\infty} h[k_2]e^{j\omega k_2} \\ &= S_{xx}(\omega) \cdot H(\omega) \cdot H^*(\omega) \\ &= S_{xx}(\omega) \cdot |H(\omega)|^2 \end{aligned}$$

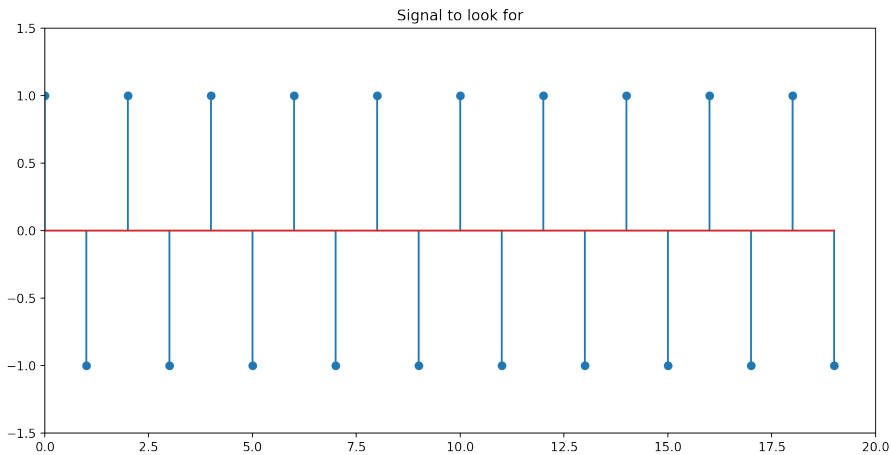
$$S_{yy}(\omega) = S_{xx}(\omega) \cdot |H(\omega)|^2$$

- ▶ The PSD of y = the PSD of x multiplied with the squared amplitude response of the filter
- ▶ Same relation is valid for continuous processes as well

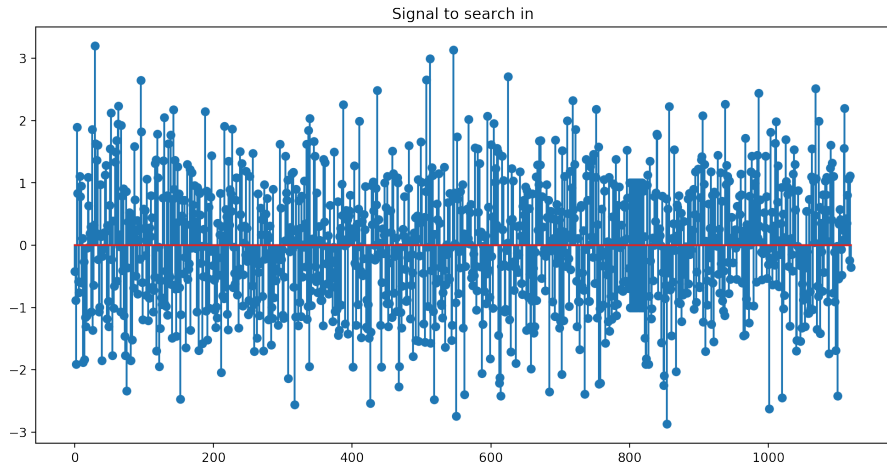
Applications of (auto)correlation

- ▶ Searching for a certain part in a large signal
- ▶ Correlation of two signals = measure of **similarity** of the two signals
 - ▶ The correlation function measures the similarity of a signal with all the shifted versions of the other
 - ▶ Example at blackboard
- ▶ Correlation can be used to locate data
 - ▶ The (auto)correlation function has large values when the two signals match
 - ▶ Large value when both positive and negative areas match,
 - ▶ Small values when they don't match

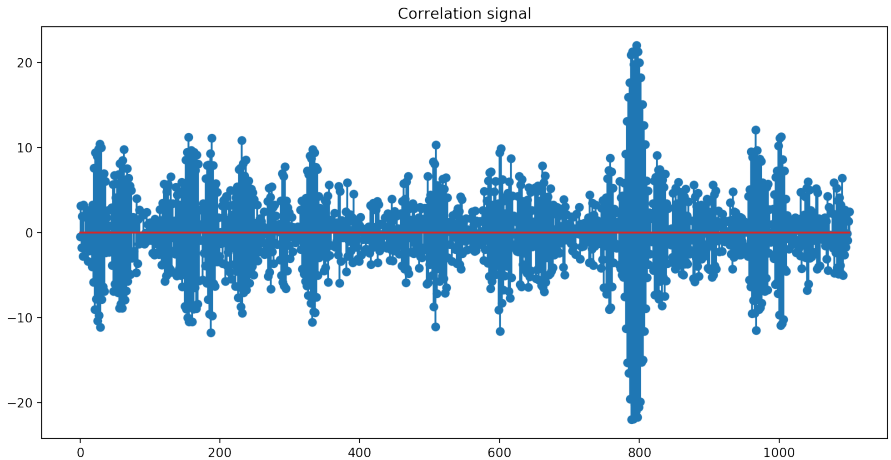
The signal to look for



The complete signal



Correlation result



System identification

- ▶ Determining the impulse response of an unknown LTI system
- ▶ Based on correlation between input and output of the system



Figure 2: System identification setup

System identification

$$\begin{aligned} R_{fg}(\tau) &= \overline{f[n]g[n+\tau]} \\ &= \overline{f[n] \sum_{k=-\infty}^{\infty} h[k]f[n+\tau-k]} \\ &= \sum_{k=-\infty}^{\infty} h[k] \overline{f[n]f[n+\tau-k]} \\ &= \sum_{k=-\infty}^{\infty} h[k] R_{ff}[\tau-k] \\ &= h[\tau] \star R_{ff}[\tau] \end{aligned}$$

- ▶ If the input f is **white noise** with power A , $R_{ff}[n] = A \cdot \delta[n]$, and

$$R_{fg}(\tau) = h[\tau] \star R_{ff}[\tau] = A \cdot h[\tau] \star \delta[\tau] = A \cdot h[\tau]$$

- ▶ Then the correlation is proportional with the impulse response of the unknown system