

Первое задание.

① $\xi \sim R[0, \Theta]$ $\Theta > 0$ вер. модель
 \vec{x}_n - выборка и $p(x) = \frac{1}{\Theta} \{ (0; \Theta) \}$

$$\tilde{\Theta}_1 = 2\bar{x} = 2 \cdot \frac{1}{n} \sum_{i=1}^n x_i$$

$$\tilde{\Theta}_2 = x_{\min}; \quad \tilde{\Theta}_3 = x_{\max};$$

$$\tilde{\Theta}_4 = x_1 + \frac{1}{n-1} \cdot \sum_{i=2}^n x_i$$

1) $\tilde{\Theta}_1$: несмещенность и состоят-ть.

все одинак

$$M[2 \cdot \frac{1}{n} \sum x_i] = \frac{2}{n} M[\sum x_i] = \frac{2}{n} \cdot \sum Mx_i = 2M\xi = \Theta \Rightarrow \text{несмещ.}$$
$$\left\{ M\xi = \int_0^\Theta x \cdot \frac{1}{\Theta} dx = \frac{\Theta}{2} \right\}$$

$$D[2 \cdot \frac{1}{n} \sum x_i] = \frac{4}{n^2} D[\sum x_i] = \frac{4}{n^2} \sum D x_i =$$

все одинак

$$= \frac{4}{n} D\xi = \frac{1}{3n} \Theta^2 \xrightarrow{n \rightarrow \infty} 0$$

$$\left\{ \begin{aligned} D\xi &= M\xi^2 - M^2\xi \\ M\xi^2 &= \int_0^\Theta x^2 \cdot \frac{1}{\Theta} dx = \frac{\Theta^2}{3} \\ D\xi &= \frac{\Theta^2}{3} - \frac{\Theta^2}{4} = \frac{\Theta^2}{12} \end{aligned} \right\}$$

состоят.
(но дост. усл.)
сост.

$$2) \tilde{\Theta}_2 = x_{\min} = x_{(1)}$$

$$M[\tilde{\Theta}_2] = \int_0^\Theta y q(y) dy \Leftrightarrow$$

$$x_i \sim R(0, \Theta) \quad x_i \sim F(x)$$

$$x_{(1)} \sim 1 - (1 - F(y))^n$$

$$q(y) = n \cdot (1 - F(y))^{n-1} \cdot \frac{1}{\Theta} \{ (0, \Theta) \}$$

$$\Leftrightarrow \int_0^\Theta y \cdot n \left(1 - \frac{y}{\Theta}\right)^{n-1} \cdot \frac{1}{\Theta} dy = \left\{ t = 1 - \frac{y}{\Theta} \right\} =$$

$$= - \int_1^0 \Theta (1-t) n t^{n-1} \cdot \frac{1}{\Theta} \Theta dt = n \Theta \left(\int_0^1 t^{n-1} dt - \int_0^1 t^n dt \right)$$

$$= n \Theta \cdot \frac{1}{n} - n \Theta \cdot \frac{1}{n+1} = \frac{\Theta}{n+1} - \text{среднее значение}$$

$$\tilde{\Theta}_2' = (n+1) \tilde{\Theta}_2 = (n+1) x_{\min} - \text{генератор случайных}$$

$$D[\tilde{\Theta}_2] = M[\tilde{\Theta}_2^2] - M^2[\tilde{\Theta}_2] \Leftrightarrow$$

$$\{ M[\tilde{\Theta}_2^2] = \int_0^\Theta y^2 n \left(1 - \frac{y}{\Theta}\right)^{n-1} \cdot \frac{1}{\Theta} dy = \left\{ t = 1 - \frac{y}{\Theta} \right\} =$$

$$= - \int_1^0 \Theta^2 (1-t)^2 n t^{n-1} \cdot \frac{1}{\Theta} \Theta dt = n \Theta^2 \left(\int_0^1 t^{n-1} dt - \right.$$

$$\left. - 2 \int_0^1 t^n dt + \int_0^1 t^{n+1} dt \right) = n \Theta^2 \left(\frac{1}{n} - 2 \cdot \frac{1}{n+1} + \frac{1}{n+2} \right) =$$

$$= n \Theta^2 \left(\frac{(n+1)(n+2) - 2n(n+2) + n(n+1)}{n(n+1)(n+2)} \right) =$$

$$= \Theta^2 \left(\frac{n^2 + 3n + 2 - 2n^2 - 4n + n^2 + n}{(n+1)(n+2)} \right) = \frac{2 \Theta^2}{(n+1)(n+2)}$$

$$\begin{aligned} & \textcircled{=} \frac{2\theta^2}{(n+1)(n+2)} - \frac{\theta^2}{(n+1)^2} = \frac{2\theta^2 n + 2\theta^2 - \theta^2 n - 2\theta^2}{(n+1)^2(n+2)} = \\ & = \frac{\theta^2 n}{(n+1)^2(n+2)} \xrightarrow{n \rightarrow \infty} 0, \text{ по следств. / не можем } \text{позв. дост. уст.} \end{aligned}$$

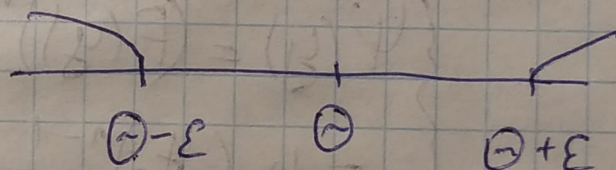
Проверяем по определ. состоят ли.

$$\tilde{\theta}_2 \xrightarrow{P} \theta \quad \forall \theta > 0$$

$$\forall \varepsilon > 0 \quad P(|\tilde{\theta}_2 - \theta| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$\tilde{\theta}_2 \geq \theta + \varepsilon \text{ или } \tilde{\theta}_2 \leq \theta - \varepsilon$$

x_{\min}



$$x_i \sim R(0, \theta) \rightarrow P(x_{\min} \geq \theta + \varepsilon) = 0.$$

$$P(x_{\min} \leq \theta - \varepsilon) = P(x_{\min} < \theta - \varepsilon) = \Phi(\theta - \varepsilon).$$

в стандарт.

$$\begin{aligned} \Phi(y) &= 1 - (1 - F(y))^n = 1 - (1 - F(\theta - \varepsilon))^n = \\ &= 1 - \left(1 - \frac{\theta - \varepsilon}{\theta}\right)^n = 1 - \left(\frac{\varepsilon}{\theta}\right)^n \xrightarrow{n \rightarrow \infty} 1 \quad \begin{array}{l} 0 < \theta - \varepsilon < \theta \\ 0 < \varepsilon < \theta \end{array} \end{aligned}$$

т.е. $\exists \varepsilon$, что $\rightarrow 0 \Rightarrow$ оценка не явл. сост.

Для $\tilde{\theta}_2'$ по определ. состоят ли.

$$\tilde{\theta}_2' \xrightarrow{P} \theta, \quad \forall \theta > 0.$$

$$\forall \varepsilon > 0 \quad P(|\tilde{\theta}_2' - \theta| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$\begin{aligned}
 P(|(n+1)X_{\min} - \Theta| \geq \varepsilon) &\geq P(X_{\min}(n+1) \geq \Theta + \varepsilon) = \\
 &= P\left(X_{\min} \geq \frac{\Theta + \varepsilon}{n+1}\right) = 1 - P\left(X_{\min} < \frac{\Theta + \varepsilon}{n+1}\right) = \\
 &= 1 - \Phi\left(\frac{\Theta + \varepsilon}{n+1}\right) = \left(1 - F\left(\frac{\Theta + \varepsilon}{n+1}\right)\right)^n \stackrel{n \geq n_0}{=} \\
 &= \left(1 - \frac{\Theta + \varepsilon}{\Theta(n+1)}\right)^n \xrightarrow{n \rightarrow \infty} e^{-\frac{\Theta + \varepsilon}{\Theta}} > 0. \\
 &\Rightarrow \underline{\text{не абн. соем.}}
 \end{aligned}$$

$$3) \tilde{\Theta}_3 = x_{\max}$$

$$\mu[\tilde{\Theta}_3] = \int_0^\Theta y q(y) dy \quad \textcircled{=}$$

$$\begin{cases} \psi(y) = (F(y))^n; & q(y) = \psi'(y) = n(F(y))^{n-1} \cdot F'(y) = \\ & = n \left(\frac{y}{\Theta}\right)^{n-1} \cdot \frac{1}{\Theta} \cdot \{ (0; \Theta) \}. \end{cases}$$

$$\textcircled{=} \int_0^\Theta y^n \frac{n}{\Theta^n} dy = \frac{n}{\Theta^n} \cdot \frac{\Theta^{n+1}}{n+1} = \frac{\Theta n}{n+1} - \text{чекнг.}$$

$$\tilde{\Theta}_3^1 = \frac{n+1}{n} x_{\max} - \underline{\text{чекнг.}}$$

$$D[\tilde{\Theta}_3] = \mu[\tilde{\Theta}_3^2] - \mu^2[\tilde{\Theta}_3] \quad \textcircled{=}$$

$$\begin{cases} \mu[\tilde{\Theta}_3^2] = \int_0^\Theta y^2 q(y) dy = \int_0^\Theta \frac{n}{\Theta^n} y^{n+1} dy = \\ & = \frac{n}{\Theta^n} \cdot \frac{\Theta^{n+2}}{n+2} \end{cases}$$

$$\textcircled{=} \frac{\Theta^2 \cdot n}{n+2} - \frac{\Theta^2 n^2}{(n+1)^2} = \frac{\Theta^2 n^2}{(n+2)(n+1)} - \frac{\Theta^2 n^2}{(n+1)^2}$$

$$= \Theta^2 \left(\frac{n^2 + 2n^2 + n - n^2 - 2n^2}{(n+1)^2 (n+2)} \right) = \frac{\Theta^2 n}{(n+1)^2 (n+2)} \xrightarrow{n \rightarrow \infty} 0$$

по дост. ген. невен (сильнее)

Проверяем по опр. сост-ти:

$$\tilde{\Theta}_3 \xrightarrow{P} \Theta \quad \forall \Theta > 0$$

$$\forall \varepsilon > 0 \quad P(|\tilde{\Theta}_3 - \Theta| \geq \varepsilon) \xrightarrow{n \rightarrow \infty} 0$$

$$P(|x_{\max} - \Theta| \geq \varepsilon) = P(x_{\max} \leq \Theta - \varepsilon) = 0.$$

\Rightarrow сост-на

$$D[\tilde{\Theta}_3] = D\left[\frac{n+1}{n} x_{\max}\right] = \frac{(n+1)^2}{n^2} \frac{\Theta^2 n}{(n+1)^2 (n+2)} \xrightarrow{n \rightarrow \infty} 0$$

(и неувен.)

\Rightarrow сост-на (по дост. ген.)

$$4) \quad \tilde{\Theta}_4 = \frac{\xi_n}{x_1} + \frac{1}{(n-1)} \sum_{i=2}^n x_i$$

$$M[\tilde{\Theta}_4] = M\left[x_1 + \frac{1}{(n-1)} \sum_{i=2}^n x_i\right] = M[x_1] + \frac{1}{n-1} \sum_{i=2}^n M[x_i] = M\xi + M\xi = 2M\xi = \Theta$$

$$D[\tilde{\Theta}_4] = D[x_1] + \frac{1}{(n-1)^2} \sum_{i=2}^n D[x_i] = \frac{\Theta^2}{12} +$$

$$+ \frac{1}{n-1} \cdot \frac{\Theta^2}{12} \xrightarrow{n \rightarrow \infty} \frac{\Theta^2}{12} \quad (\text{дост. ген. не работ.})$$

Проверяем по опр.:

$$\tilde{\Theta}_4 \xrightarrow{P} \Theta$$

$$\left. \begin{aligned} \textcircled{1} \quad \xi_n &\xrightarrow{P} \xi \quad \eta_n \xrightarrow{P} \eta \\ \xi_n + \eta_n &\xrightarrow{P} \xi + \eta \end{aligned} \right\} \text{св-ва}$$

② ЗБЧ хвильових: ξ_i незав. і однако.

расср. и $M\xi_i < \infty$, тогда $\frac{1}{n} \sum_{i=1}^n \xi_i \xrightarrow{P} M\xi_i$

$$\xi_n \xrightarrow{P} x_1$$

$$\eta_n \xrightarrow{P} M\xi = \frac{\Theta}{2}$$

$$\text{т.е. } \tilde{\Theta}_4 \xrightarrow{P} \xi + \frac{\Theta}{2}$$

(опр не вблн.) \Rightarrow

не явл. сост.

Сравнение оценок:

$$D\tilde{\Theta}_1 = \frac{\Theta^2}{3n}$$

$$D\tilde{\Theta}_3' = \frac{\Theta^2}{(n+2)/n}$$

$$\frac{1}{3n} \geq \frac{1}{(n+2)/n}$$

$$n+2 \geq 3 \Rightarrow \boxed{n \geq 1}$$

т.е. $\tilde{\Theta}_3'$ эффективнее $\tilde{\Theta}_1$.