

SHEET - An Introduction to Statistical Learning
Chapter 3 - Linear Regression

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1 Linear Regression

1.1 Courses' Demonstrations

Let $\hat{y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_i$ be the prediction for Y based on the i th value of X_i . Then $e_i = y_i - \hat{y}_i$ represents the i th residual - this is the difference between the i th observed response value and the i th response value that is predicted by our linear model. We define the residual sum of squares (RSS) as

$$\begin{aligned} \text{RSS} &= e_1^2 + e_2^2 + \cdots + e_n^2 \\ &= (y_1 - \hat{\beta}_0 - \hat{\beta}_1 x_1)^2 + (y_2 - \hat{\beta}_0 - \hat{\beta}_1 x_2)^2 + \cdots + (y_n - \hat{\beta}_0 - \hat{\beta}_1 x_n)^2 \end{aligned}$$

The least squares approach chooses $\hat{\beta}_0$ and $\hat{\beta}_1$ to minimize the RSS. Using some calculus, one can show that the minimizers are

$$\begin{aligned} \hat{\beta}_1 &= \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x} \end{aligned} \tag{3.4}$$

where \bar{y} is the sample mean, defined as

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i$$

Demonstration:

We search RSS as

$$f(\hat{\beta}_0, \hat{\beta}_1) = \min_{\hat{\beta}_0, \hat{\beta}_1} \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2$$

We get the minimum where

$$\begin{aligned} &\begin{cases} \frac{\partial f}{\partial \hat{\beta}_0} = 0 \\ \frac{\partial f}{\partial \hat{\beta}_1} = 0 \end{cases} \\ &\begin{cases} -2 \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \\ -2 \sum_{i=1}^n x_i (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i) = 0 \end{cases} \end{aligned}$$

$$\begin{aligned}
& \begin{cases} \sum_{i=1}^n y_i - n\hat{\beta}_0 - \hat{\beta}_1 \sum_{i=1}^n x_i = 0 \\ \sum_{i=1}^n x_i y_i - \hat{\beta}_0 \sum_{i=1}^n x_i - \hat{\beta}_1 \sum_{i=1}^n x_i^2 = 0 \end{cases} \\
& \begin{cases} \bar{y} - \hat{\beta}_0 - \hat{\beta}_1 \bar{x} = 0 \\ \sum_{i=1}^n x_i y_i - n\hat{\beta}_0 \bar{x} - \hat{\beta}_1 \sum_{i=1}^n x_i^2 = 0 \end{cases} \\
& \begin{cases} \hat{\beta}_0 = \bar{y} + \hat{\beta}_1 \bar{x} \\ 0 = \sum_{i=1}^n x_i y_i - n\hat{\beta}_0 \bar{x} - \hat{\beta}_1 \sum_{i=1}^n x_i^2 \end{cases} \\
& \begin{cases} \hat{\beta}_0 = \bar{y} + \hat{\beta}_1 \bar{x} \\ 0 = \sum_{i=1}^n x_i y_i - n(\bar{y} + \hat{\beta}_1 \bar{x}) \bar{x} - \hat{\beta}_1 \sum_{i=1}^n x_i^2 \end{cases} \\
& \begin{cases} \hat{\beta}_0 = \bar{y} + \hat{\beta}_1 \bar{x} \\ 0 = \sum_{i=1}^n x_i y_i - n\bar{y} \bar{x} + \hat{\beta}_1 (n\bar{x}^2 - \sum_{i=1}^n x_i^2) \end{cases} \\
& \begin{cases} \hat{\beta}_0 = \bar{y} + \hat{\beta}_1 \bar{x} \\ 0 = \sum_{i=1}^n x_i y_i - n\bar{y} \bar{x} + \hat{\beta}_1 \sum_{i=1}^n (x_i^2 - \bar{x}^2) \end{cases}
\end{aligned}$$

$$\begin{aligned}
E[(X - E[X])(Y - E[Y])] &= E[XY - X * E[Y] - E[X] * Y + E[X] * E[Y]] \\
&= E[XY] - E[X * E[Y]] - E[E[X] * Y] + E[E[X] * E[Y]] \\
&= E[XY] - E[X] * E[Y] - E[X] * E[Y] + E[X] * E[Y] \\
&= E[XY] - E[X] * E[Y]
\end{aligned}$$

$$\begin{aligned}
E[(X - E[X])^2] &= E[X^2 - 2 * X * E[X] + E[X]^2] \\
&= E[X^2] - E[2 * X * E[X]] + E[E[X]^2] \\
&= E[X^2] - 2 * E[X] * E[X] + E[X]^2 \\
&= E[X^2] - E[X]^2
\end{aligned}$$

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$$\begin{cases} \hat{\beta}_0 = \bar{y} + \hat{\beta}_1 \bar{x} \\ \hat{\beta}_1 = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2} \end{cases}$$

By developping we can have

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n (y_i - \bar{y})(x_i)}{\sum_{i=1}^n (x_i - \bar{x})^2} = \frac{\sum_{i=1}^n (y_i)(x_i - \bar{x})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

We assume that the True relationship between X and Y takes the form

$$Y = f(X) + \varepsilon$$

for some unknown function f , where ε is a mean-zero random error term. If f is to be approximated by a linear function then we can write the relationship as

$$Y = \beta_0 + \beta_1 X + \varepsilon \tag{3.5}$$

How accurate is the sample mean $\hat{\mu}$ as an estimate of μ ? In general, we answer this question by computing the standard error of $\hat{\mu}$, written as the standard error $SE(\hat{\mu})$. A reasonable estimate is $\hat{\mu} = \bar{y}$. We have the well-known formula:

$$Var(\hat{\mu}) = SE(\hat{\mu}^2) = \frac{\sigma^2}{n}$$

where σ is the standard deviation of each of the realizations y_i of Y .

Demonstration:

$$\bar{\mu} = \frac{1}{n} \sum_{i=1}^n \mu_i$$

$$E[\bar{\mu}] = E\left[\frac{1}{n} \sum_{i=1}^n \mu_i\right] = \frac{1}{n} \sum_{i=1}^n E[\mu_i] = \mu$$

$$Var(\hat{\mu}) = Var(\bar{y}) = E[(\bar{y} - E[\bar{y}])^2]$$

$$Var(\hat{\mu}) = E\left[\left(\frac{1}{n} \sum_{i=1}^n y_i - E\left[\frac{1}{n} \sum_{i=1}^n y_i\right]\right)^2\right]$$

$$Var(\hat{\mu}) = \frac{1}{n^2} E\left[\left(\sum_{i=1}^n y_i - E[y_i]\right)^2\right]$$

$$a_i = y_i - E[y_i]$$

$$Var(\hat{\mu}) = \frac{1}{n^2} E\left[\left(\sum_{i=1}^n a_i\right)^2\right] = \frac{1}{n^2} E\left[\left(\sum_{i=1}^n a_i\right) \left(\sum_{j=1}^n a_j\right)\right]$$

$$= \frac{1}{n^2} E\left[\sum_{i=1}^n a_i^2 + 2 \sum_{1 \leq i < j \leq n} a_i a_j\right]$$

$$= \frac{1}{n^2} \left(\sum_{i=1}^n E[a_i^2] + 2 \sum_{1 \leq i < j \leq n} E[a_i a_j] \right)$$

$E[a_i] = E[y_i - E[y_i]] = E[y_i] - E[y_i] = 0$ and a_i, a_j are independent

$$Var(\hat{\mu}) = \frac{1}{n^2} E\left[\sum_{i=1}^n (y_i - E[y_i])^2\right] = \frac{1}{n} E\left[\frac{1}{n} \sum_{i=1}^n (y_i - E[y_i])^2\right] = \frac{1}{n} E[\sigma^2]$$

$$Var(\hat{\mu}) = SE(\hat{\mu}^2) = \frac{\sigma^2}{n}$$

1.2 Answers of Exercises