Feature selection and regularised regression

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Goals of modelling

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- 1. **Predictive strength:** How well can we reconstruct the observed data? Has been most important so far.
- 2. Model/variable selection: Which variables are part of the true model? This is about uncovering structure to allow for mechanistic understanding.

Feature Selection

Consider the model

$$y = X\beta + \varepsilon$$

where

▶ $\mathbf{y} \in \mathbb{R}^n$ is the outcome, $\mathbf{X} \in \mathbb{R}^{n \times (p+1)}$ is the design matrix, $\boldsymbol{\beta} \in \mathbb{R}^{p+1}$ are the regression coefficients, and $\boldsymbol{\varepsilon} \in \mathbb{R}^n$ is the additive error

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- ► Centring the outcome $(\frac{1}{n}\sum_{l=1}^{n}y_{l}=0)$ and features removes the need to estimate the intercept

Feature selection as motivation

Analytical solution exists when $\mathbf{X}^\mathsf{T}\mathbf{X}$ is invertible

$$\widehat{\boldsymbol{\beta}}_{\mathrm{OLS}} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{y}$$

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Solutions: Regularisation or **feature selection**

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- Care with cross-validation and multiple testing necessary

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Summary

- ▶ **Pro:** Fast and easy
- ► Con: Filtering mostly operates on single features and is not geared towards a certain method
- ► Care with cross-validation and multiple testing necessary
- ► Filtering is often more of a pre-processing step and less of a proper feature selection step

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- ► As discreet procedures, all of these methods **exhibit high variance** (small changes could lead to different predictors being selected, resulting in a potentially very different model)

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► Softer regularisation methods can help

$$\widehat{\boldsymbol{\beta}} = \operatorname*{arg\,min}_{\boldsymbol{\beta}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} + \lambda \|\boldsymbol{\beta}\|_{q}^{q}$$

where λ is a tuning parameter and $q \ge 1$ or $q = \infty$.

Feature selection

Feature selection can be addressed in multiple ways

- Filtering: Remove variables before the actual model for the data is built
 - ▶ Often crude but fast
 - ► Typically only pays attention to one or two features at a time (e.g. F-Score, MIC) or does not take the outcome variable into consideration (e.g. PCA)
- Wrapping: Consider the selected features as an additional hyper-parameter
 - computationally very heavy
 - most approximations are greedy algorithms
- ► Embedding: Include feature selection into parameter estimation through penalisation of the model coefficients
 - ▶ Naive form is equally computationally heavy as wrapping
 - Soft-constraints create biased but useful approximations

Regularised regression

Constrained and regularised regression

The optimization problem

$$\underset{\boldsymbol{\beta}}{\operatorname{arg\,min}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} \quad \text{subject to} \quad \|\boldsymbol{\beta}\|_{q}^{q} \leq t$$

for q > 0 is equivalent to

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when $q \ge 1$. This is the **Lagrangian** of the constrained problem.

Note: Constraints are convex for all $q \ge 1$ but not differentiable in $\beta = 0$ for q = 1.

Ridge regression

For q=2 the constrained problem is **ridge regression** (Tikhonov regularisation)

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If $\mathbf{X}^{\mathsf{T}}\mathbf{X} = \mathbf{I}_{p}$, then

$$\widehat{oldsymbol{eta}}_{\mathrm{ridge}}(\lambda) = rac{\widehat{oldsymbol{eta}}_{\mathrm{OLS}}}{1+\lambda},$$

i.e. $\widehat{\pmb{\beta}}_{\mathrm{ridge}}(\lambda)$ is **biased** but has **lower variance**.

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The analytical solution for ridge regression becomes $(n \ge p)$

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Ridge regression **acts strongest** on principal components with **lower eigenvalues**, e.g. in presence of correlation between features.

Effective degrees of freedom

Recall the **hat matrix** $\mathbf{H} = \mathbf{X}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}$ in OLS. The trace of \mathbf{H}

$$\operatorname{tr}(H) = \operatorname{tr}(\mathbf{X}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}) = \operatorname{tr}(\mathbf{X}^{\mathsf{T}}\mathbf{X}(\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}) = \operatorname{tr}(\mathbf{I}_p) = p$$

is equal to the trace of $\widehat{\Sigma}$ and the **degrees of freedom** for the regression coefficients.

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In analogy define for ridge regression

$$\mathbf{H}(\lambda) := \mathbf{X}(\mathbf{X}^{\mathsf{T}}\mathbf{X} + \lambda \mathbf{I}_p)^{-1}\mathbf{X}^{\mathsf{T}}$$

and

$$df(\lambda) := tr(\mathbf{H}(\lambda)) = \sum_{j=1}^{p} \frac{d_j^2}{d_j^2 + \lambda},$$

the effective degrees of freedom.

Lasso regression

For q = 1 the constrained problem is known as the lasso

$$\widehat{\pmb{\beta}}_{\mathrm{lasso}}(\lambda) = \mathop{\arg\min}_{\pmb{\beta}} \| \mathbf{y} - \mathbf{X} \pmb{\beta} \|_2^2 + \lambda \| \pmb{\beta} \|_1$$

- ▶ Smallest *q* in penalty such that constraint is still convex
- ► Produces sparse solutions (many coefficients exactly equal to zero) and therefore performs feature selection

Intuition for the penalties (I)

Assume the OLS solution $oldsymbol{eta}_{ ext{OLS}}$ exists and set

$$\mathbf{r} = \mathbf{y} - \mathbf{X}\boldsymbol{\beta}_{\mathrm{OLS}}$$

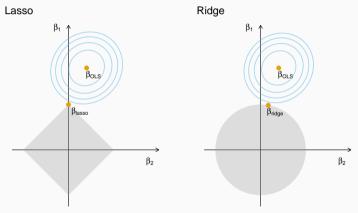
it follows for the residual sum of squares (RSS) that

$$\begin{aligned} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} &= \|(\mathbf{X}\boldsymbol{\beta}_{\mathrm{OLS}} + \mathbf{r}) - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} \\ &= \|(\mathbf{X}(\boldsymbol{\beta} - \boldsymbol{\beta}_{\mathrm{OLS}}) - \mathbf{r}\|_{2}^{2} \\ &= (\boldsymbol{\beta} - \boldsymbol{\beta}_{\mathrm{OLS}})^{\mathsf{T}}\mathbf{X}^{\mathsf{T}}\mathbf{X}(\boldsymbol{\beta} - \boldsymbol{\beta}_{\mathrm{OLS}}) - 2\mathbf{r}^{\mathsf{T}}\mathbf{X}(\boldsymbol{\beta} - \boldsymbol{\beta}_{\mathrm{OLS}}) + \mathbf{r}^{\mathsf{T}}\mathbf{r} \end{aligned}$$

which is an **ellipse** (at least in 2D) centred on $\beta_{\rm OLS}$.

Intuition for the penalties (II)

The least squares RSS is minimized for β_{OLS} . If a constraint is added ($\|\beta\|_q^q \le t$) then the RSS is minimized by the closest β possible that fulfills the constraint.

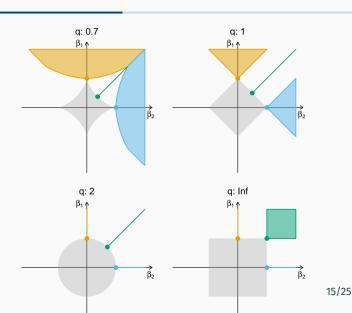


The blue lines are the contour lines for the RSS.

Intuition for the penalties (III)

Depending on q the different constraints lead to different solutions. If $\beta_{\rm OLS}$ is in one of the coloured areas or on a line, the constrained solution will be at the corresponding dot.

Sparsity only for $q \le 1$ **Convexity** only for $q \ge 1$



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Special case: $X^TX = I_p$. Then

$$\frac{1}{2} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|_{2}^{2} + \lambda \|\boldsymbol{\beta}\|_{1} = \frac{1}{2} \mathbf{y}^{\mathsf{T}} \mathbf{y} - \underbrace{\mathbf{y}^{\mathsf{T}} \mathbf{X}}_{=\boldsymbol{\beta}_{\mathrm{OLS}}^{\mathsf{T}}} \boldsymbol{\beta} + \frac{1}{2} \boldsymbol{\beta}^{\mathsf{T}} \boldsymbol{\beta} + \lambda \|\boldsymbol{\beta}\|_{1} = g(\boldsymbol{\beta})$$

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How do we find the solution $\widehat{\beta}$ in presence of the **non-differentiable** penalisation $\|\beta\|_1$?

For $\mathbf{X}^{\mathsf{T}}\mathbf{X} = \mathbf{I}_p$ the target function can be written as

$$\underset{\beta}{\operatorname{arg\,min}} \sum_{j=1}^{p} -\beta_{\mathrm{OLS},j}\beta_{j} + \frac{1}{2}\beta_{j}^{2} + \lambda|\beta_{j}|$$

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Each case results in

$$\widehat{\beta}_{\mathrm{lasso},j} = \mathrm{sign}(\beta_{\mathrm{OLS},j})(|\beta_{\mathrm{OLS},j}| - \lambda)_{+} = \mathrm{ST}(\beta_{\mathrm{OLS},j},\lambda),$$

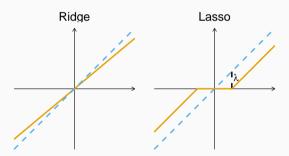
where

- $x_+ = x \text{ if } x > 0 \text{ or } 0 \text{ otherwise,}$
- and ST is called the soft-thresholding operator

Relation to OLS estimates

Both ridge regression and the lasso estimates can be written as functions of $\boldsymbol{\beta}_{\text{OLS}}$ if $\mathbf{X}^{\mathsf{T}}\mathbf{X} = \mathbf{I}_p$.

$$eta_{\mathrm{ridge},j} = rac{eta_{\mathrm{OLS},j}}{1+\lambda} \quad ext{and} \quad \widehat{eta}_{\mathrm{lasso},j} = \mathrm{sign}(eta_{\mathrm{OLS},j})(|eta_{\mathrm{OLS},j}| - \lambda)_{+}$$



Visualisation of the transformations applied to the OLS estimates.

Shrinkage and effective degrees of freedom

When λ is fixed, the **shrinkage** of the lasso estimate $\beta_{\rm lasso}(\lambda)$ compared to the OLS estimate $\beta_{\rm OLS}$ is defined as

$$s(\lambda) = \frac{\|\boldsymbol{\beta}_{\text{lasso}}(\lambda)\|_1}{\|\boldsymbol{\beta}_{\text{OLS}}\|_1}$$

Note: $s(\lambda) \in [0,1]$ with $s(\lambda) \to 0$ for increasing λ and $s(\lambda) = 1$ if $\lambda = 0$

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Prostate cancer dataset

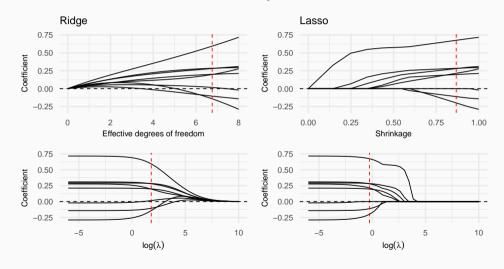
Prostate cancer dataset

Data to examine the correlation between the level of a prostate cancer-specific substance and a number of clinical measurements in men who just before partial or full removal of the prostate in patients.

- ightharpoonup n = 67 samples
- ► A continuous response on the log-scale
- p = 8 features
 - e.g. log cancer volume, log prostate weight or age of patient

Regularisation paths for varying λ

Red dashed lines indicate the λ selected by cross-validation



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- Numerical solution possible, e.g. with **coordinate descent** where each β_j is updated separately with the remaining β_i with $i \neq j$ fixed
- ► As for ridge regression, estimates are biased
- ▶ **Degrees of freedom** are equal to the number of non-zero coefficients

Potential caveats of the lasso (I)

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- ► Correlations: Many non-relevant variables correlated with relevant variables can lead to the selection of the wrong model, even for large *n*
- ► Irrepresentable condition: Split X such that X₁ contains all relevant variables and X₂ contains all irrelevant variables. If

$$|(\mathbf{X}_{\mathbf{2}}^{\mathsf{T}}\mathbf{X}_{1})(\mathbf{X}_{1}^{\mathsf{T}}\mathbf{X}_{1})^{-1}| < 1 - \eta$$

for some $\eta > 0$ then the lasso is (almost) guaranteed to pick the true model

Potential caveats of the lasso (II)

In practice, both the sparsity of the true model and the irrepresentable condition cannot be checked.

Assumptions and domain knowledge have to be used

Take-home message

- ► Filtering and wrapping methods useful for feature selection in practice but can be unprincipled or have high variance
- Regularised regression can help in numerically unstable situations (such as in ridge regression)
- ▶ The lasso can in addition perform variable selection