

Results

Search results

Preregistered analyses

As preregistered, we ran three random-effects meta-analysis models, one for each outcome variable. Figure 3 shows posterior distributions from these analyses.

Perceived injustice. Across 201,912 participants from 123 samples in 84 studies, we found strong evidence for a weak association ($r = -.07, [-.10, -.04]$) between intergroup contact and perceived injustice, with $> 99.9\%$ of posterior samples for the mean correlation coefficient falling below zero. We found evidence that correlation coefficients varied across studies ($\tau_J = .14, [.12, .17]$) and across samples within studies ($\tau_I = .08, [.05, .12]$). Based on these analyses, we predicted that 80% of studies would result in correlation coefficients between $-.25$ and $.11$ and that researchers would need sample sizes of at least 2,538 participants to find significant associations ($\alpha = .05$, two-sided) in 80% of their studies.¹

Collective action. Across 119,085 participants from 37 samples in 24 studies, we found some evidence for a weak association ($r = -.06, [-.13, .02]$) between intergroup contact and collective action, with 93.6% of posterior samples for the mean correlation coefficient falling below zero. We found evidence that correlation coefficients varied across studies ($\tau_J = .16, [.12, .23]$) and across samples within studies ($\tau_I = .09, [.06, .14]$). Based on these analyses, we predicted that 80% of studies would result in correlation coefficients between $-.27$ and $.16$ and that researchers would need sample sizes of at least 1,882 participants to find significant associations ($\alpha = .05$, two-sided) in 80% of their studies.

Policy support. Across 13,703 participants from 19 samples in 14 studies, we found some evidence for a weak association ($r = -.07, [-.14, -.00]$) between intergroup contact and policy support, with 98.1% of posterior samples for the mean correlation coefficient falling below zero. We found evidence that correlation coefficients varied across studies ($\tau_J = .10, [.06, .18]$) and, to a lesser extent, across samples within studies ($\tau_I = .03, [.00, .12]$). Based on these analyses, we predicted that 80% of studies would result in correlation coefficients between $-.21$ and $.07$ and that researchers would need sample sizes of at least 3,208 participants to find significant associations ($\alpha = .05$, two-sided) in 80% of their studies.

¹Sample sizes are based on posterior predictions from the three models, which implied that, for 80% of studies, the absolute correlation coefficient would be $|r| > .039$ for perceived injustice, $|r| > .045$ for collective action, and $|r| > .035$ for policy support.

As preregistered, we ran another three random-effects meta-analysis models to estimate the relationships between the three outcome variables. As we were not interested in the direction of these relationships, we used cross-sectional correlation coefficients as effect sizes for longitudinal studies. Across 111,753 participants from 25 samples in 14 studies, we found evidence for a moderate association ($r = .31, [.23, .38]$) between perceived injustice and collective action. Across 6,244 participants from 12 samples in 9 studies, we found evidence for a moderate association ($r = .23, [.08, .35]$) between perceived injustice and policy support. Across 8,558 participants from 6 samples in 3 studies, we found evidence for a moderate association ($r = .30, [.13, .42]$) between collective action and policy support.

Robustness checks

First, we assessed to what extent our findings were sensitive to choosing narrower, $\mu \sim \text{Normal}(0, 0.1)$, or wider, $\mu \sim \text{Normal}(0, 1)$, prior distributions. Choosing narrower or wider prior distribution did not affect mean effect size estimates for perceived injustice ($\Delta r = -.00, [-.05, .04]$ and $\Delta r = .00, [-.05, .05]$), collective action ($\Delta r = -.01, [-.11, .09]$ and $\Delta r = -.00, [-.10, .11]$), and policy support ($\Delta r = -.01, [-.10, .08]$ and $\Delta r = .00, [-.09, .10]$). Second, we assessed to what extent our findings were sensitive to including or excluding influential studies by repeating the preregistered analyses J times while leaving out one of J studies each time and by calculating the mean absolute difference (MAD) for the estimated mean effect size across left-out studies. For perceived injustice ($MAD = .02, [.01, .04]$), collective action ($MAD = .04, [.02, .09]$), and policy support ($MAD = .03, [.02, .08]$), the MAD was small. Leaving out the most influential study, for example, did not change estimates of the mean effect size for the three outcomes ($\Delta r = .00, [-.04, .05]$; $\Delta r = .02, [-.09, .12]$; $\Delta r = -.02, [-.11, .07]$). Together, these analyses showed that our findings were robust to choosing different prior distributions and to excluding influential studies.

References

- Pettigrew, T. F., & Tropp, L. R. (2006). A meta-analytic test of intergroup contact theory. *Journal of Personality and Social Psychology*, 90(5), 751–783. <https://doi.org/10.1037/0022-3514.90.5.751>

Figure 1

Flow diagram illustrating the preregistered search strategy, study selection, and data collection

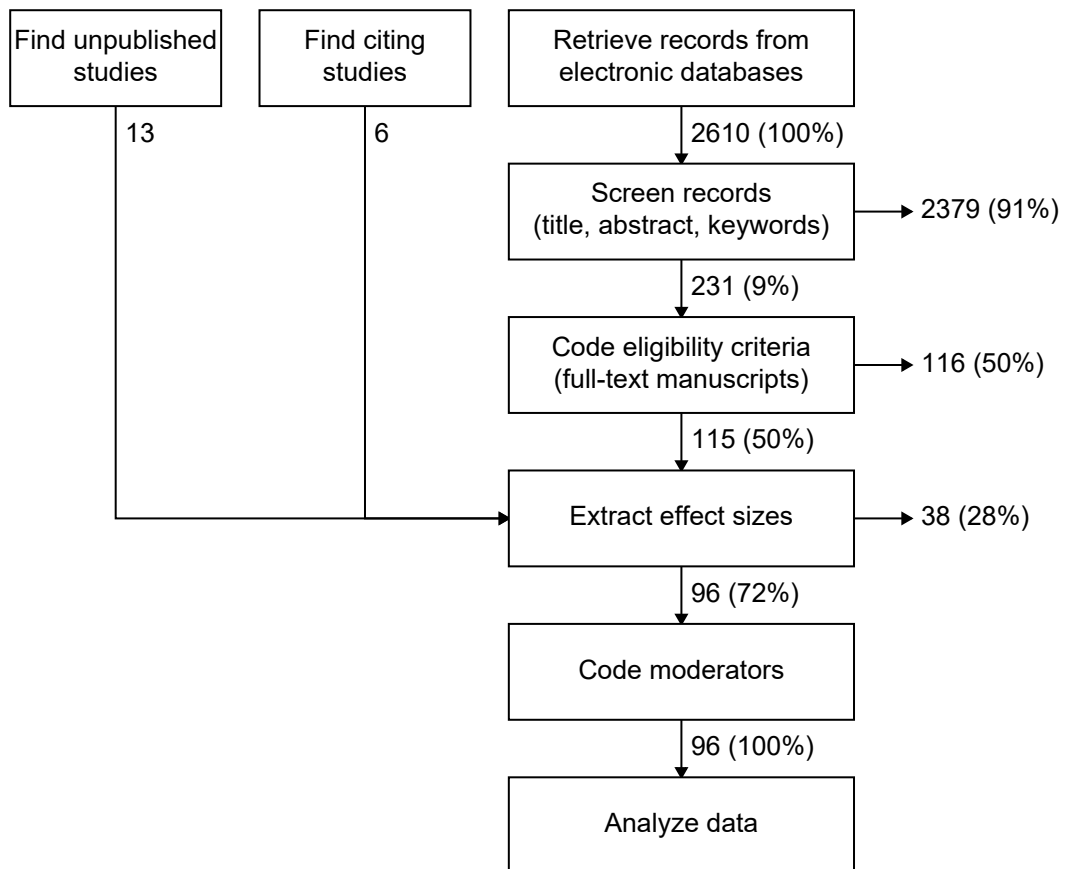
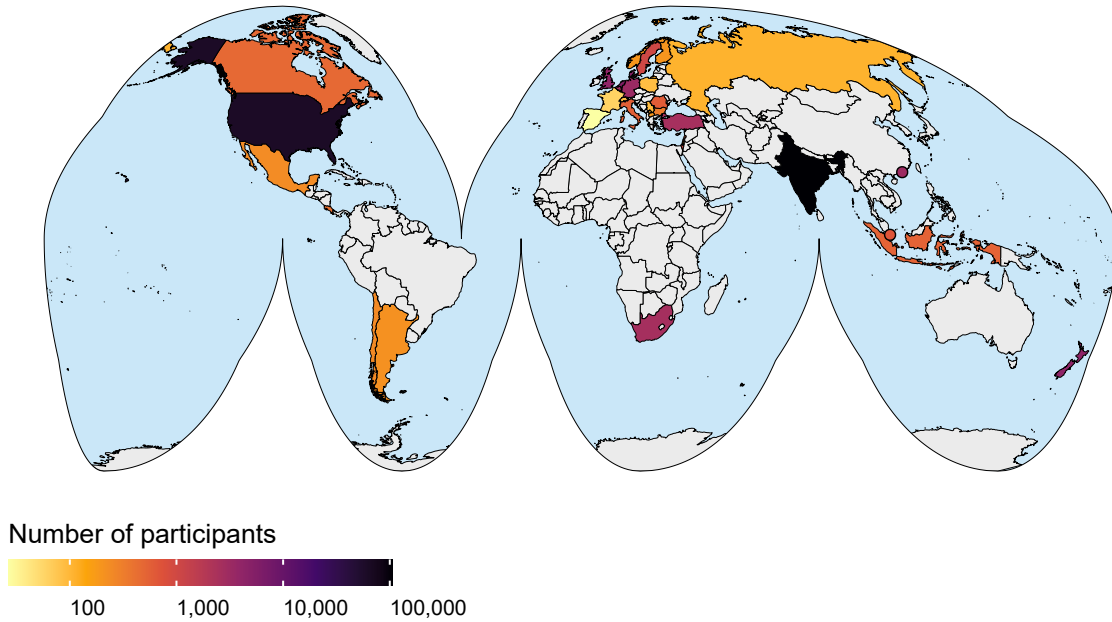
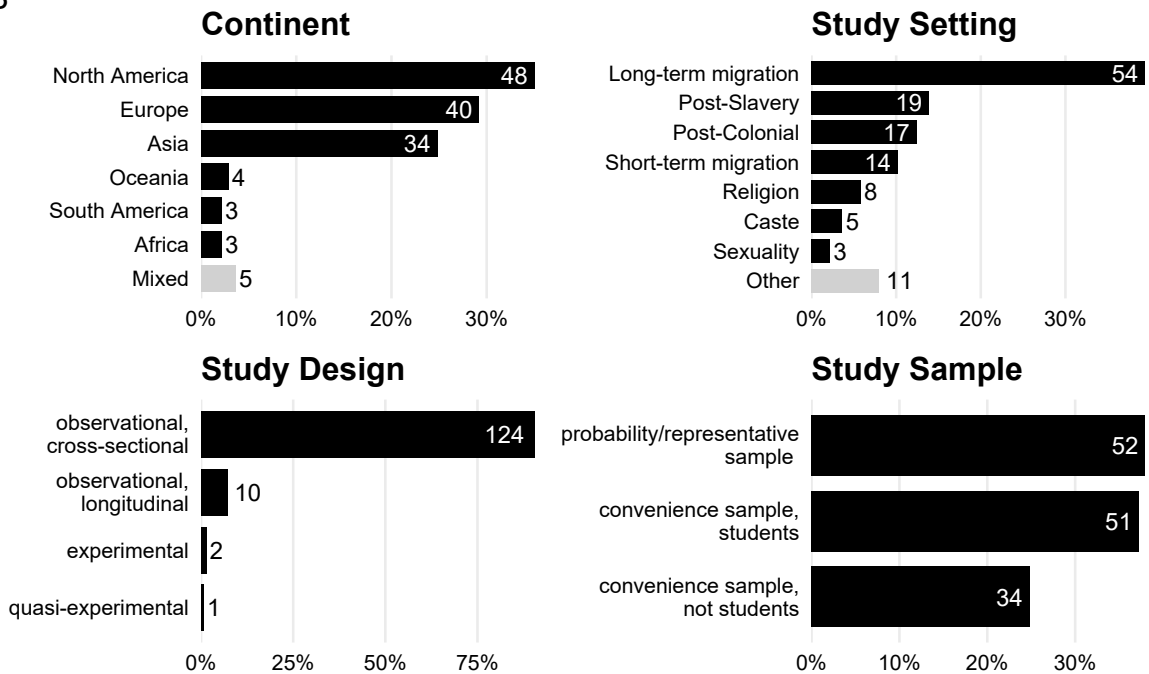


Figure 2
Overview of the relevant literature

A

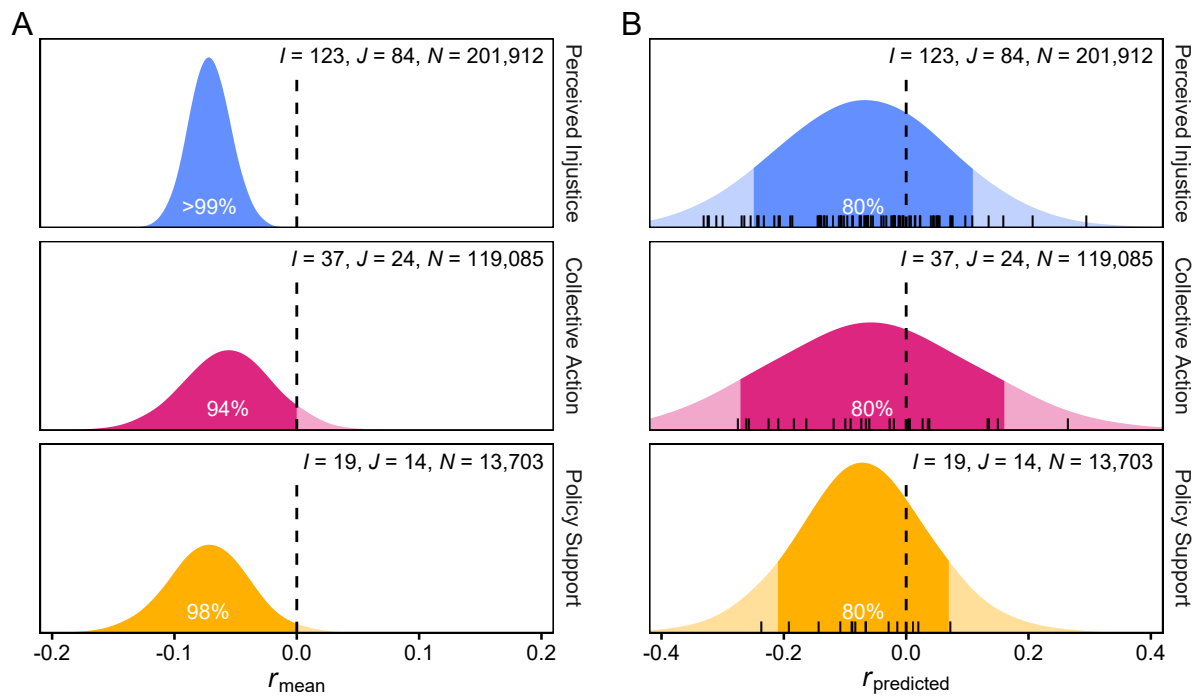


B



Note. **A** Map of all countries included in the meta-analysis with combined sample sizes. **B** Proportion of eligible samples in each category as well as the absolute number of samples in each category.

Figure 3
Posterior distributions from the preregistered random-effect meta-analysis models



Note. **A** Posterior distributions for the estimated mean correlation coefficients, highlighting the proportion of posterior samples for which $r_{\text{mean}} < 0$. **B** Posterior predictive distributions for the estimated study-wise correlation coefficients, based on point estimates of the μ and τ_j parameters, with point estimates for the estimated correlation coefficients for all studies in the sample.