**Title**: Neurodevelopmental features among infants with congenital hypotonia: An observational cross-sectional trial.

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## Abstract

**Objective**: […]. **Material and methods**: […]. **Results**: […]. **Conclusion**: […].

**Keywords**: […].

# Introduction

Both physical and psychological signs of early childhood development have been shown to be representative and relevant markers for the identification and monitoring of overall growth in early life ([Di Rosa et al., 2016](#ref-di2016predictive)) and can, therefore, be used in the screening of children at risk of developmental delay to support early referral and need further assessment to determine if they are eligible for early intervention services ([Bruder, 2010](#ref-bruder2010early); [Guralnick, 2017](#ref-guralnick2017early)).

Currently, a plethora of tools have been proposed to assess the developmental continuum of infants. In this sense, the Ages and Stages Questionnaire, Third Edition (ASQ-3) has been proposed as a global screening tool, parent/caregivers-oriented, that assesses five domains of development in children aged 0 to 5.5 years ([Singh, Yeh, & Blanchard, 2017](#ref-singh2017ages)). Current evidence suggests that the ASQ-3 is an accurate, cost-effective yet parent-friendly instrument for screening, monitoring children up to pre-school age, and can help identify and exclude neurodevelopmental impairments in very preterm-born children ([Ballantyne, Benzies, McDonald, Magill-Evans, & Tough, 2016](#ref-ballantyne2016risk); [Kerstjens et al., 2015](#ref-kerstjens2015ages); [Singh et al., 2017](#ref-singh2017ages)).

Hypotonia has been defined both as decreased muscle tone or floppiness, involving a wide range and levels of progression ([Gabis et al., 2021](#ref-gabis2021weak); [Harris, 2008](#ref-harris2008congenital)). There are multiple forms of neuromuscular, metabolic and genetic conditions associated with hypotonia and it may be a sign of neurodevelopmental delay, that may predispose to cognitive impairment ([Riou, Ghosh, Francoeur, & Shevell, 2009](#ref-riou2009global)). Given that hypotonia, and the hyperlaxity and motor delay associated with it, may impair the infant’s capacity to engage with its surroundings, critical visual cues may be ignored, potentially leading both to an impairment of learning and cognitive development ([Harris, 2008](#ref-harris2008congenital)), hence the need to explore the neurodevelopmental attributes of infants with hypotonia.

However, and to the best of our knowledge, there is no robust evidence characterising the observed variation in developmental traits in infants with known hypotonia across age. Therefore, our main objective in this study was to describe and model the relationship between sociodemographics, prematurity and neurodevelopmental levels based on ASQ-3 scores in infants with diagnosed hypotonia.

# Material y methods

## Participants

[…].

## Instruments

### Instrument 1

[…].

### Instrument 2

[…].

## Procedures

[…].

[…].

## Statistical analysis

Data is presented as median (*Mdn*) and interquartile range (*IQR*) for continuous variables; for categorical/discrete variables, the absolute and relative sample size was reported.

A non-parametric approach was used since the underlying distribution of continuous measured outcomes, assessed through analytical and graphical methods, did not follow a Gaussian distribution.

In order assess the differences in developmental scores between males and females, the *Wilcoxon* rank-sum test was used, meanwhile the chi-square test () was used to evaluate goodness-of-fit () and independence of factors ().

Generalized additive models (GAM) were used to describe linear and non-linear relationships in the form of smooth terms between developmental characteristics, represented through penalized regression splines ([Wood, 2011](#ref-wood2011fast)). Restricted maximum likelihood method was used for the estimation of the smoothing parameters, and thin-plate regression splines as the smoothing basis, as they are the optimal smoother of any given basis dimension/rank ([Wood, 2003](#ref-wood2003thin)). To describe the smooth terms by means of quasi-linear segments, we used approximative effect derivatives with 95% confidence intervals (CI95%).

To account for any source of variability coming from subject’s sex, evaluators and the type of relationship between the infants with the respondents, we incorporate them as random effects in the fitted models in the form of penalized parametric terms ([Wood, N., Pya, & S"afken, 2016](#ref-wood2016smoothing)).

A probability of committing a type I () error of less than 5% (*p* < 0.05), was considered sufficient evidence for statistical significance in hypothesis testing. All the statistical analyses were computed and implemented in the R programming language ([R Core Team, 2021](#ref-rlanguage)). GAMs and the corresponding model estimates were computed using the *mgcv* and *modelbased* packages, each with well documented functions and methods ([Makowski, Ben-Shachar, Patil, & Lüdecke, 2020](#ref-dominique2020estimation); [Wood, 2017](#ref-wood2017generalized)). Complementary R packages were used for visualization purposes ([Lüdecke et al., 2021](#ref-daniel2021see); [Wickham, 2016](#ref-hadley2016ggplot2)).

# Results

From a total of 234 subjects with congenital hypotonia, 94 (40.2%) were females and 140 (59.8%) males ( (1) = 9.04, *p* = 0.003, = 0.19, CI0.95%[0.09, 1]). The developmental characteristics of the sample can be seen in [Table 1](#tab1).

When modelling the effect of chronological age on developmental domains, corrected for prematurity, we observed a significant non-linear relationship on communication scores ( (5.2, 224.04) = 13.43, *p* < 0.001), that reflect an overall negative marginal effect ( = -2.36, CI95%[-3.47, -1.25], (224.04) = -4.2, *p* < 0.001), however, this was not true when assessing the direction of the effect in the age range between 0 to 6.8 ( = 0.49, CI95%[-0.89, 1.86], (224.04) = 0.45, *p* = 0.319), and neither in the 18.4 to 48 months old group ( = 0.45, CI95%[-1.32, 2.23], (224.04) = 0.42, *p* = 0.593), whereas the effect tend to be positive but non-significant. The relationship between developmental domains, corrected age and their effect derivatives can be seen in [Figure 1](#fig1).

When analysing the motor skills domain, we found a significant non-linear effect of corrected age on gross motor scores, (5.24, 226.75) = 6.19, *p* < 0.001, which had an overall positive effect ( = 1.95, CI95%[0.66, 3.25], (226.75) = 2.97, *p* = 0.003), however, the slope varied as a function of age, with a negative effect in the 0 to 6.8 age range ( = -2.94, CI95%[-4.55, -1.34], (226.75) = -3.7, *p* = 0.004), but in the 9.7 to 15.5 interval, this relationship was inverted ( = 1.86, CI95%[0.61, 3.11], (226.75) = 2.93, *p* = 0.009), although, in the 7.3 to 9.2 ( = 0.02, CI95%[-1.12, 1.17]) and 16 to 48 age range ( = -0.02, CI95%[-2.05, 2.01]) the slope was non-significant and virtually zero ( (226.75) = 0.06, *p* = 0.45, and (226.75) = 0.07, *p* = 0.646 respectively).

Despite the fact that a similar non-linear effect was observed when inspecting the influence of corrected age in the fine motor domain scores ( (2.59, 226.77) = 4.2, *p* = 0.005), it was not possible to estimate a significant overall effect different from zero ( = 0.04, CI95%[-0.45, 0.52], (226.77) = 0.14, *p* = 0.886), nevertheless, it was only in the 22.3 to 38.3 age range where a significant negative effect was observed ( = -0.79, CI95%[-1.45, -0.12], (226.77) = -2.34, *p* = 0.022).

Problem solving abilities were significantly influenced by corrected age ( (5.66, 227.01) = 3.65, *p* = 0.001), with an overall negative effect ( = -1.87, CI95%[-3.17, -0.57], (227.01) = -2.83, *p* = 0.005), and just like the other domains, this relationship was modified across corrected age. In this sense, from the 0 to 5.8 age interval, we found that for every increase in one month in corrected age, we can expect a proportional increase in 2.81 points (CI95%[1.18, 4.44], (227.01) = 3.49, *p* = 0.002) in the problem solving domain, while in the 9.2 to 14.1 age range this relationship is modified inversely, mainly because in this range we observed that for every increase in one month in corrected age, we could expect a decrease in 1.59 points (CI95%[-2.82, -0.37], (227.01) = -2.55, *p* = 0.015) in the same domain. Other age intervals had no significant slope different from zero (age range [6.3, 8.7], = 0.05, CI95%[-1.1, 1.2], (227.01) = 0.06, *p* = 0.395; age range [14.5, 48.0], = 0.03, CI95%[-1.99, 2.04], (227.01) = -0.06, *p* = 0.55).

Unlike the others, socio-individual domain was not influenced by corrected age ( (1, 231.58) = 1.16, *p* = 0.282).

# Discussion

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[…].

# Conclusion

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# Acknowledgment

[…].

# Conflicts of interest

[…].

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**Table 1**. Overall baseline and developmental characteristics of the sample and grouped by sex. 1 Data is presented as sample size, and *Mdn* (*IQR*); 2 p-values are computed from the *Wilcoxon* rank-sum test.

**Figure 1**. Relationship between corrected age (in months) and developmental domains. Left panel: regression lines and shaded area represent predicted values estimated from GAM models and their CI95%, points and error bars represent the mean and standard error at 5-month age intervals. Right panel: effect derivatives representing how the effect of corrected age (in months) in developmental domains changes across corrected age. Significant areas consider CI95% that did not cross zero.