

# Surrogate-based Bayesian Inversion for the Model Calibration of Fire Insulation Panels

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# Surrogate-based Bayesian Inversion for the Model Calibration of Fire Insulation Panels

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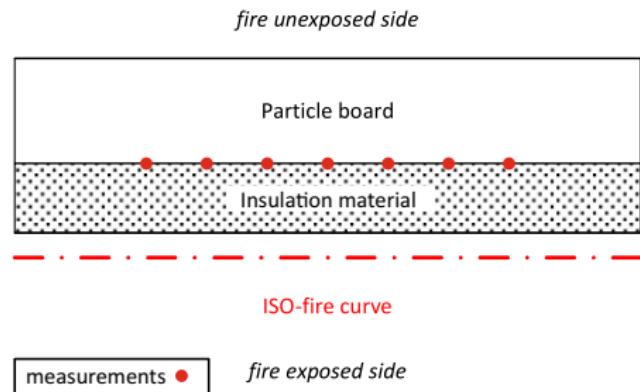
# Motivation

## Efficient inversion

Timber engineering relies on accurate model calibration for:

- simulation of fire hazards (component additive method)
- code based design

Experiments are carried out under standardized conditions for different insulation materials.



Breu (2016), Isofloc Insulation

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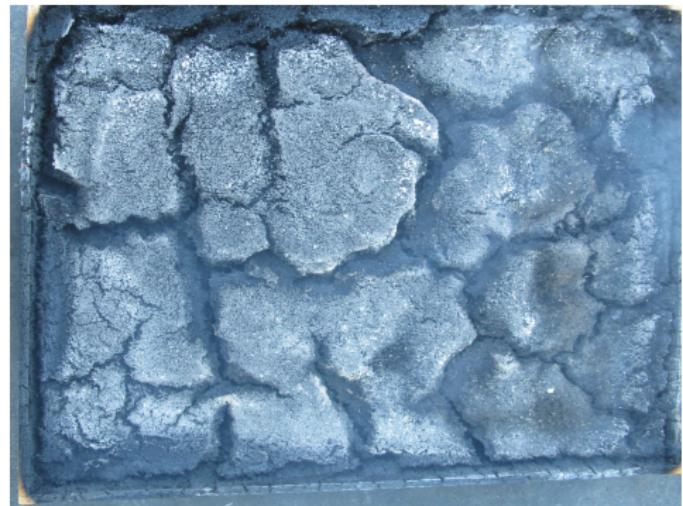
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# Outline

## 1 Introduction

- Problem formulation
- Bayesian inversion

## 2 Inversion with surrogate forward modelling

## 3 Inversion with surrogate likelihood

## 4 Conclusion

# Experimental data and forward model

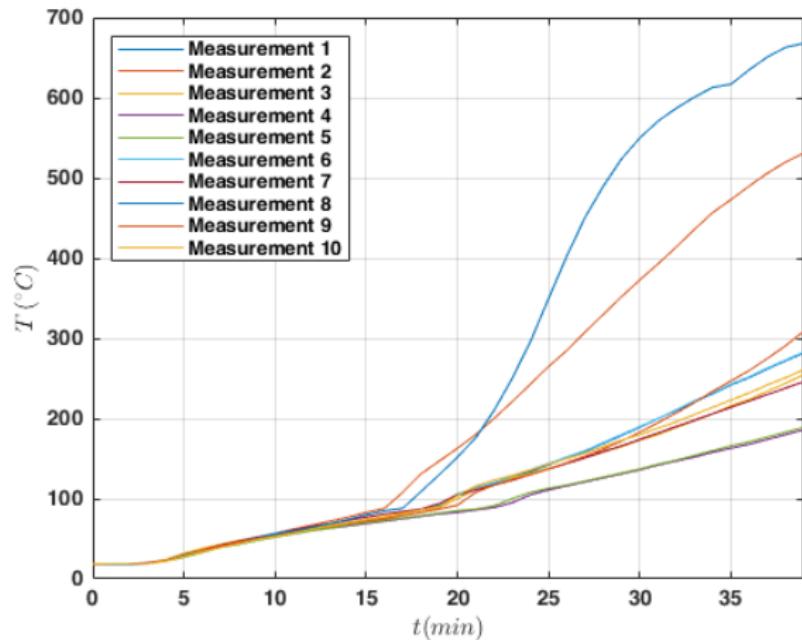
## Experimental data $y$

Huge variations in measurements due to:

- cracking
- inhomogeneous insulation material

## Forward model $\mathcal{M}(X)$

1D heat transfer problem with temperature dependent effective material parameters solved with FEM.



# Experimental data and forward model

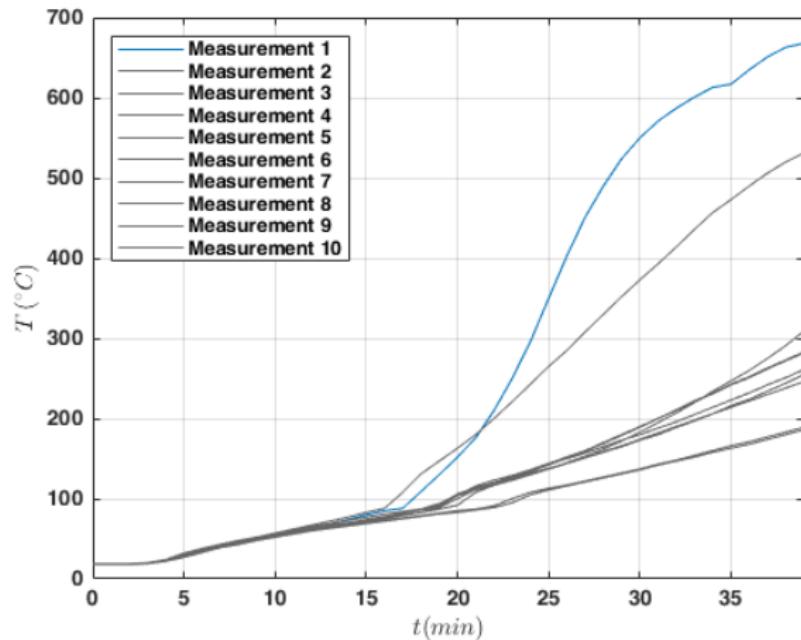
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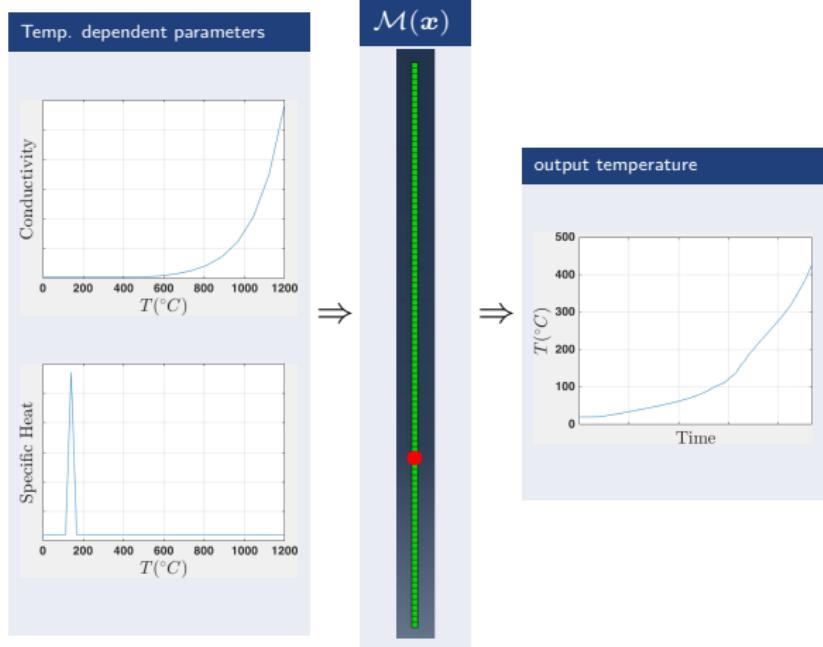
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# Inverse problem

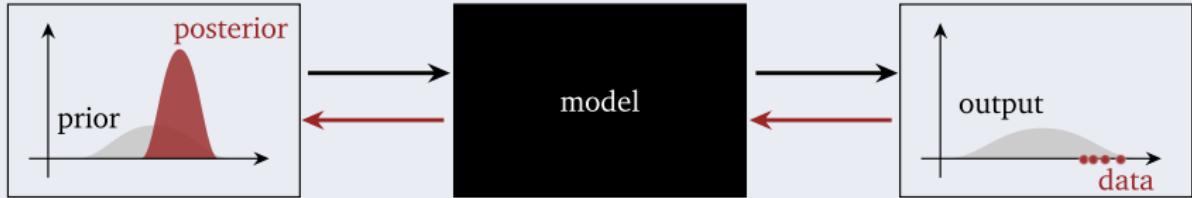
## Bayes' theorem

With observed data regarded as  $\mathbf{Y}|\mathbf{x} \sim \pi(\mathbf{y}|\mathbf{x})$

$$\pi(\mathbf{x}|\mathbf{y}) \propto \pi(\mathbf{y}|\mathbf{x})\pi(\mathbf{x})$$

with the **Posterior**  $\pi(\mathbf{x}|\mathbf{y})$ , the **Prior**  $\pi(\mathbf{x})$  and the **Likelihood**  $\pi(\mathbf{y}|\mathbf{x})$

## Black-Box model

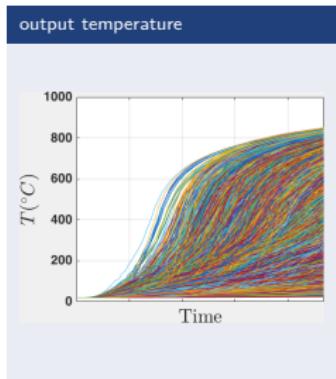
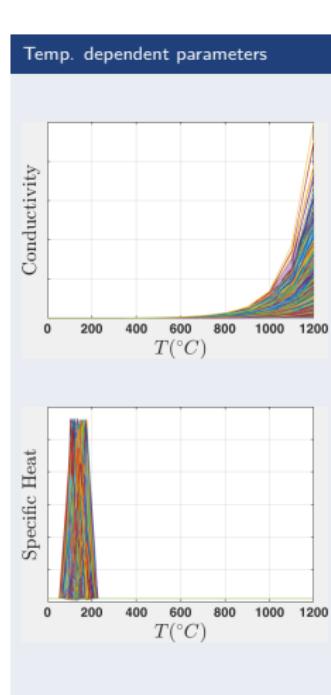


# Prior in insulation problem

## Parametrization

Define prior  $\pi(\mathbf{x})$  on temp. dependent parameter curve parameters:

$$\mathbf{x} \sim \mathcal{U}(\mathbf{x}_l, \mathbf{x}_u)$$



# Likelihood in insulation problem

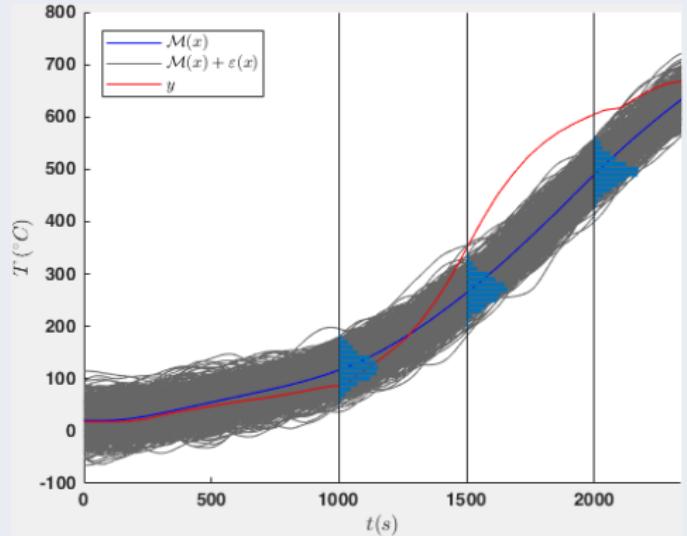
Likelihood function  $\pi(\mathbf{y}|\mathbf{x})$

Pdf of the data  $\mathbf{y}$  conditioned on the parameters  $\mathbf{x}$  as a function of these parameters:

$$\mathbf{Y}|\mathbf{x} \sim \mathcal{N}(\mathcal{M}(\mathbf{x}) - \mathbf{y}, \Sigma(\mathbf{x}))$$

Normal distribution describes the discrepancy between the model predictions  $\mathcal{M}(\mathbf{x})$  and the observed data  $\mathbf{y}$ .

Experimental curve vs. predictions



# Outline

① Introduction

② Inversion with surrogate forward modelling

Surrogate forward modelling

MCMC

Results

③ Inversion with surrogate likelihood

④ Conclusion

# Inversion with surrogate forward modelling

## Method

Two-stage procedure:

- build a sufficiently accurate surrogate model of  $\mathcal{M}(\mathbf{X})$ 
  - ① run  $N$  forward runs of  $\mathbf{y} = \mathcal{M}(\mathbf{x})$
  - ② construct  $\mathcal{M}^{PC}(\mathbf{X}) \approx \mathcal{M}(\mathbf{X})$
- perform inference
  - ① set up a standard method (MAP, MLE, MCMC, etc.)
  - ② use  $\mathcal{M}^{PC}(\mathbf{X})$  instead of  $\mathcal{M}(\mathbf{X})$

# PCE + PCA

## Principal Component Analysis

The  $N = 235$  output dimensions of  $\mathbf{Y} = \mathcal{M}(\mathbf{X})$  are compressed to  $N' = 7$  dimensions of a set of principal components  $\mathbf{Z}$  through *PCA*:

$$\log(\mathbf{Y}) \approx \bar{\mu}_{\mathbf{Y}} + \sum_{p=1}^{N'} \tilde{z}_p(\mathbf{X}) \bar{\phi}_p$$

with the empirical mean  $\bar{\mu}_{\mathbf{Y}}$  and a set of  $N'$  eigenvectors of the empirical covariance matrix  $\bar{\phi}_p$  of  $\log(\mathbf{Y})$ .

## Polynomial Chaos Expansion

The  $N'$  retained principal components are expanded using an orthogonal polynomial basis  $\Psi_{\alpha}(\mathbf{X})$ :

$$\tilde{z}_p(\mathbf{X}) = \sum_{\alpha \in \mathbb{N}^M} a_{p,\alpha} \Psi_{\alpha}(\mathbf{X}) \approx \sum_{\alpha \in \mathcal{A}_{\gamma}} a_{p,\alpha} \Psi_{\alpha}(\mathbf{X})$$

# Surrogate model $\mathcal{M}^{PC}(\boldsymbol{x})$

$\mathcal{M}^{PC}(\boldsymbol{x})$

- After computing the coefficients the response random vector can be represented as:

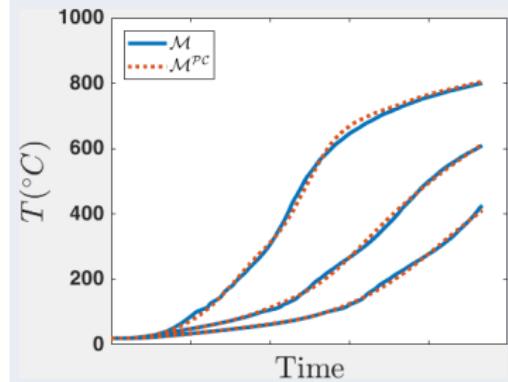
$$\mathbf{Y} \approx \exp \left[ \bar{\mu}_{\mathbf{Y}} + \sum_{p=1}^{N'} \left( \sum_{\alpha \in \mathcal{A}_\gamma} a_{p,\alpha} \Psi_\alpha(\mathbf{X}) \right) \bar{\phi}_p \right]$$

which can be used as a metamodel for  $\mathcal{M}(\boldsymbol{x})$ .

- The transformation to the log-space assures:

$$\mathcal{M}^{PC}(\boldsymbol{x}) \in \mathbb{R}_+^N$$

Comparison of  $\mathcal{M}$  and  $\mathcal{M}^{PC}$



The surrogate model is accurate enough!

# MCMC

## Solution of the inverse problem

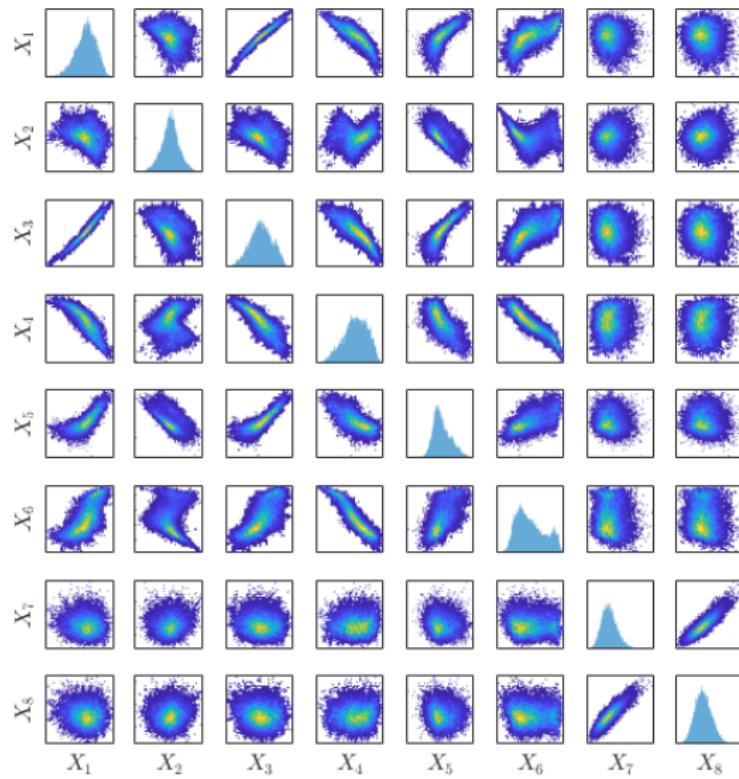
We are interested in sampling from the posterior distribution  $\pi(\mathbf{x}|\mathbf{y})$ :

$$\mathbf{X}|\mathbf{y} \sim \pi(\mathbf{x}|\mathbf{y})$$

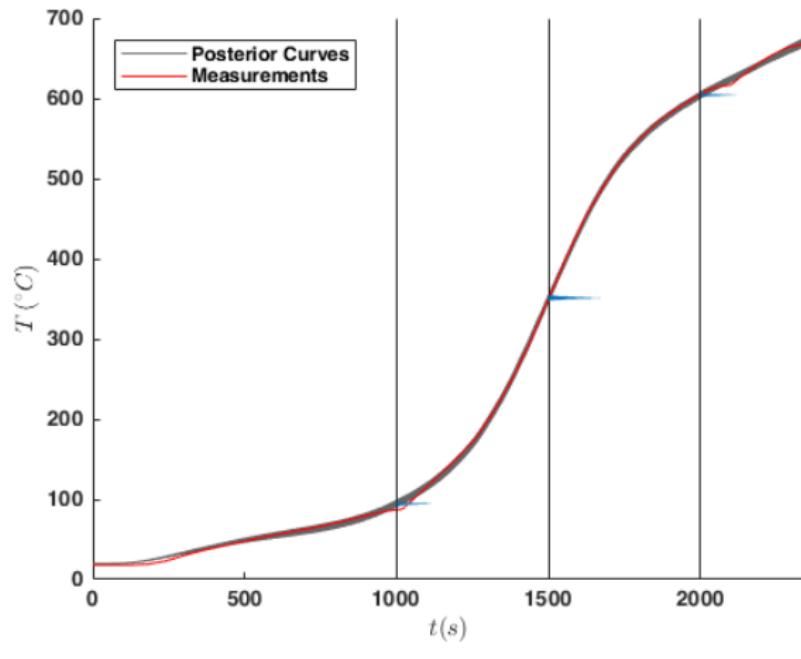
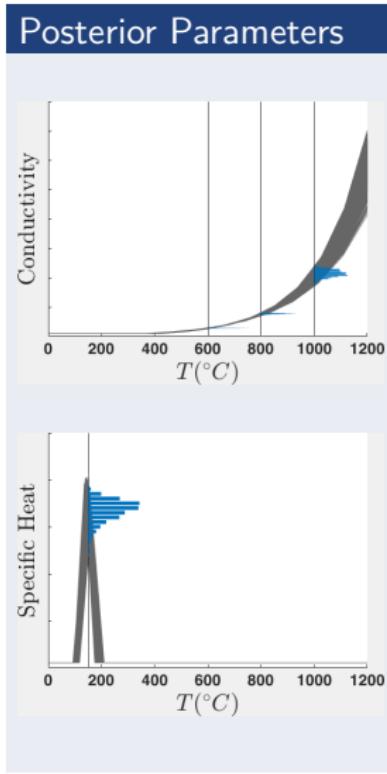
## AIES

We use the **affine-invariant ensemble sampler** (AIES) presented in Goodman, Weare (2010)

- + invariant under affine transformations of the target distribution
- + multiple parallel chains simultaneously explore the target distribution
- + little tuning necessary
- slow (no parallelization)

Posterior distribution  $\pi(x|y)$ 

# Posterior curves



# Outline

- ① Introduction
- ② Inversion with surrogate forward modelling
- ③ Inversion with surrogate likelihood
  - Surrogate likelihood (SLE)
  - Results
- ④ Conclusion

# Inversion with surrogate likelihood

## Method

### One-stage procedure:

- build sufficiently accurate surrogate model of  $\pi(\mathbf{y}|\mathbf{x})$ 
  - ① run  $N$  forward runs of  $\mathcal{L}(\mathbf{x}) = \pi(\mathbf{y}|\mathbf{x})$
  - ② construct  $\mathcal{L}^{PC}(\mathbf{X}) \approx \mathcal{L}(\mathbf{X})$
  - ③ inference is merely post-processing of  $\mathcal{L}^{PC}(\mathbf{X})$

## Spectral likelihood expansion

directly build a surrogate model for the likelihood:

$$\mathcal{L}(\mathbf{X}) \approx \mathcal{L}^{PC}(\mathbf{X}) = \sum_{\alpha \in \mathcal{A}_\gamma} a_\alpha \Psi_\alpha(\mathbf{X})$$

Nagel, J. and Sudret, B., Spectral likelihood expansions for Bayesian inference, J. Comp. Phys. (2016)

# Spectral likelihood expansion

## Inference

- By post-processing the coefficients  $a_\alpha$ , an approximation to the posterior distribution can be derived easily:

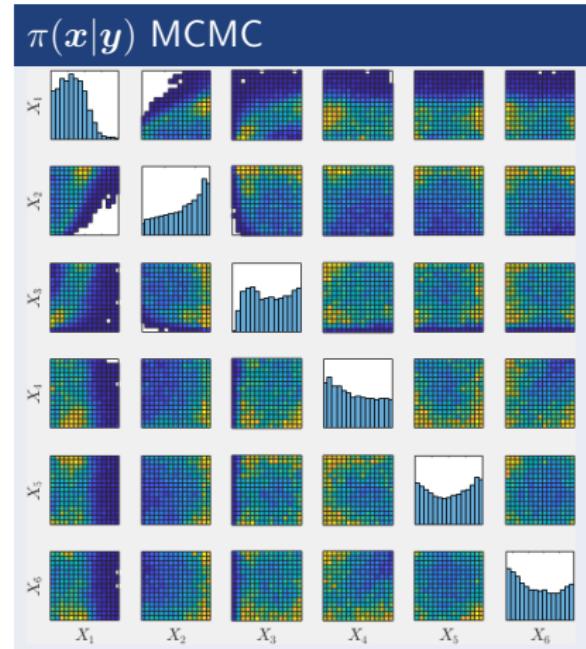
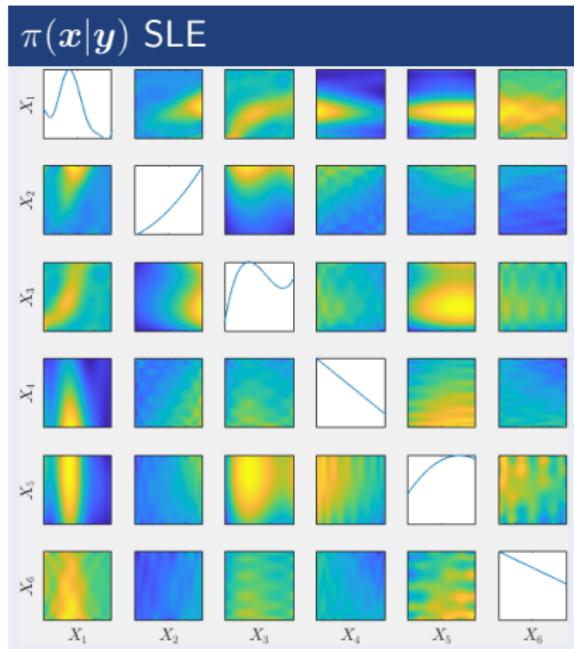
$$\pi(\boldsymbol{x}|\boldsymbol{y}) \approx \frac{1}{a_0} \mathcal{L}^{PC}(\boldsymbol{x}) \pi(\boldsymbol{x})$$

- The posterior moments of the marginals can be derived analytically:

$$\text{E}[X_j|\boldsymbol{y}] \approx \frac{1}{a_0} \langle x_j, \mathcal{L}_j^{PC} \rangle_{L^2_{\pi_j}}$$

$$\text{Var}[X_j|\boldsymbol{y}] \approx \frac{1}{a_0} \langle (x_j - \text{E}[X_j|\boldsymbol{y}])^2, \mathcal{L}_j^{PC} \rangle_{L^2_{\pi_j}}$$

# Posterior distribution $\pi(\mathbf{x}|\mathbf{y})$ SLE vs. MCMC

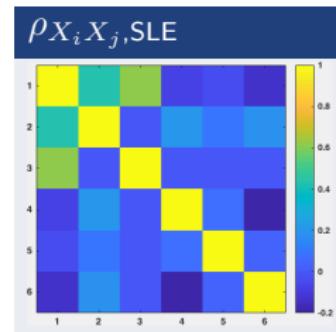
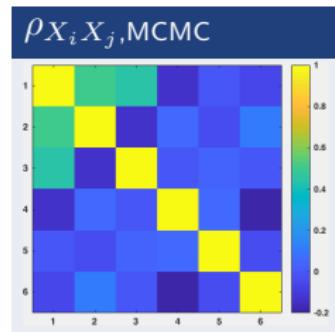


Large model error assumed!

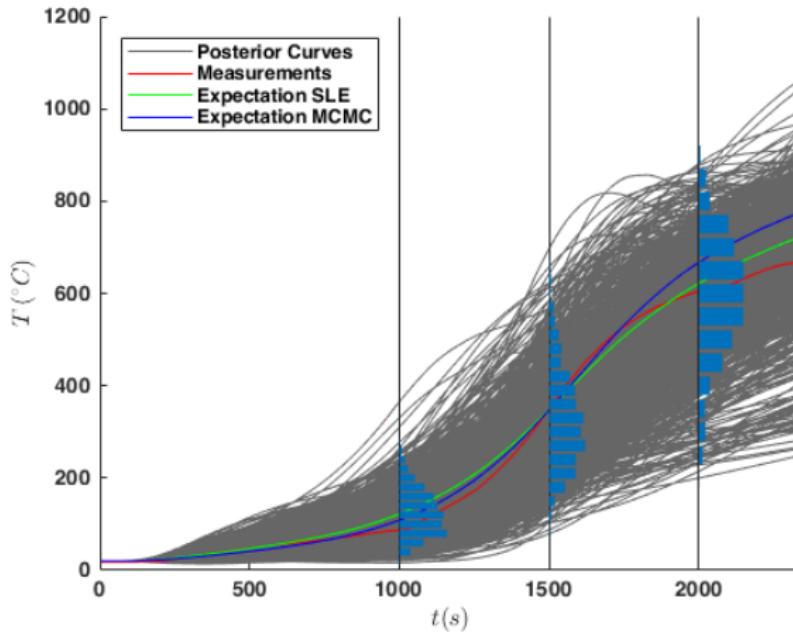
# Posterior moments of $\pi(\mathbf{x}|\mathbf{y})$ SLE vs. MCMC

## Comparison

	$X_1$	$X_2$	$X_3$	$X_4$	$X_5$	$X_6$
$E[X_j \mathbf{y}]_{\text{MCMC}}$	427.4	$3.308 \cdot 10^{-4}$	0.016	$5.12 \cdot 10^4$	142	51.11
$E[X_j \mathbf{y}]_{\text{SLE}}$	446.5	$3.708 \cdot 10^{-4}$	0.015	$4.431 \cdot 10^4$	144	50.3
$\varepsilon$ (%)	4.47	12.09	6.25	13.46	1.41	1.58
$\text{Var}[X_j \mathbf{y}]_{\text{MCMC}}$	$5.667 \cdot 10^3$	$1.669 \cdot 10^{-9}$	$6.616 \cdot 10^{-5}$	$1.118 \cdot 10^9$	628.1	758.3
$\text{Var}[X_j \mathbf{y}]_{\text{SLE}}$	$4.707 \cdot 10^3$	$9.242 \cdot 10^{-9}$	$6.265 \cdot 10^{-5}$	$8.855 \cdot 10^8$	475.3	652.9
$\varepsilon$ (%)	20.4	81.94	5.6	26.26	32.15	16.14



# Posterior curves



## SLE pros/cons

- + sampling free
- + deterministic approach for inversion
- + only one source of error (Quality of PCE)
- does not work with spiked likelihood functions (yet!)
- high dimensions (PCE based)

# Outline

- ① Introduction
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- ④ Conclusion

# Conclusion & outlook

- Bayesian inversion is suitable for calibrating transient heat conduction models of fire experiments
- Classical MCMC can be used together with advanced surrogate models (principal component analysis + polynomial chaos expansion)
- Spectral likelihood expansion is a promising, **sampling free** approach that requires some further tuning
- Multiple experiments shall be combined using hierarchical Bayesian models



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# Thank You!

# Questions?