Review 1

Commit to memory: *Sub-gaussian* mean $\mathbb{E}e^{\lambda X} \leq e^{\frac{\lambda^2\sigma^2}{2}}$

- λ is *tilting* the distribution, picking out a moment
- σ is a *scale parameter*, like a variance

$$\implies P(|X - \mu| > t) \le e^{-\frac{t^2}{2\sigma^2}}$$

Sub-exponential example: For χ^2

$$\mathbb{E}e^{\lambda X} = \frac{e^{-\lambda}}{\sqrt{1-2}} \le e^{2\lambda^2} \le e^{2\lambda^2}, \quad |\lambda| < \frac{1}{2}$$
 (1.1)

Equivalent definitions of sub-Gaussian:

(a)
$$\exists \sigma \text{ st } \mathbb{E} e^{\lambda X} \leq e^{\frac{\lambda^2 \sigma^2}{2}}$$

(b)
$$\exists c_0 > 0 \text{ st } \mathbb{E}e^{\lambda X} < \infty$$
, $|\lambda| \le c_0$

(c)
$$\exists c_1, c_2 > 0 \text{ st } P(|X| > t) \le c_1 e^{-c_2 t}, \quad t > 0$$

(d)
$$\sup_{k\geq 0} \left(\frac{\mathbb{E}|X|_k}{k!}\right)^{1/k}$$
 is finite

Tail bounds on sub-Exponential

Let X be zero-mean (subtract $\mathbb{E}X$ off otherwise).

$$P(X \ge t) \le \exp\left(-\lambda t + \frac{\lambda^2 \nu}{2}\right), \quad \lambda < \frac{1}{\alpha}$$
 (2.1)

Find infimum over $0 \le \lambda < \frac{1}{\alpha}$.

Unconstrianed optimum

$$\lambda^* = \frac{t}{\nu^2} \stackrel{?}{=} \frac{1}{\alpha} \tag{2.2}$$

(2.3)

If $t < \frac{\nu^2}{\alpha}$ (small deviations), then

$$\implies P(X \ge t) \le \exp\left(-\frac{t^2}{2\nu^2}\right)$$
 (2.4)

"Gaussian-like."

Otherwise $t > \frac{v^2}{\alpha}$ (large deviations),

$$\implies \phi(\lambda) = -\lambda t + \frac{\lambda^2 \nu}{2} \tag{2.5}$$

$$\frac{d\phi}{d\lambda} = -t + \frac{\nu^2}{2} \le -\frac{\nu^2}{\alpha} + \lambda \nu^2 \tag{2.6}$$

$$= \nu^2 \left(\nu - \frac{1}{\alpha} \right) \le 0 \tag{2.7}$$

$$\implies \lambda^* = \frac{1}{\alpha} \tag{2.8}$$

$$\implies P(X \ge t) \le \exp\left(-\frac{t}{2\alpha}\right) \tag{2.9}$$

"Exponential tail."

In summary, there are two regimes depending on how far out (t)

$$P(X - \mu \ge t) \le \begin{cases} e^{-\frac{t^2}{2\nu^2}}, & 0 \le t \le \frac{\nu^2}{\alpha} \\ e^{-\frac{t}{2\alpha}}, & t > \frac{\nu}{\alpha} \end{cases}$$
 (2.10)

3 Bernstein's Condition

TODO: Fig 4.1 TODO: Fig 4.2

Definition 3.1. A RV X satisfies *Bernstein's condition* if

- $\sigma^2 = \mathbb{E}X^2 \mu^2$ exists
- $\exists b > 0$ such that $\mathbb{E}(X \mu)^k \leq \frac{1}{2}k!\sigma^2b^{k-2}$ for all $k = 3, 4, \cdots$

Remark 3.2. This holds, e.g., for bounded RVs, improves on Hoeffding

Remark 3.3. Bounded \implies Sub-Exponential

Theorem 3.4. Bernstein's condition implies sub-exponential *Proof.*

$$\mathbb{E}e^{\lambda(X-\mu)} = 1 + \frac{\lambda^2 \sigma^2}{2} + \sum_{k=3}^{\infty} \lambda^k \frac{E(X-\mu)^k}{k!}$$
(3.1)

$$\leq 1 + \frac{\lambda^2 \sigma^2}{2} + \frac{\lambda^2 \sigma^2}{2} \sum_{k=3}^{\infty} \lambda^{k-2} b^{k-2}$$
 (3.2)

$$=1+\frac{\lambda^2\sigma^2}{2}\left(\frac{1}{1-|\lambda|b}\right), \quad |\lambda|<\frac{1}{b} \tag{3.3}$$

$$\leq e^{\frac{\lambda^2 \sigma^2/2}{1-b|\lambda|}} \leq e^{\frac{\lambda^2(\sqrt{2}\sigma)^2}{2}} \implies SE(\sqrt{2}\sigma, 2b) \tag{3.4}$$

Remark 3.5. $\mathbb{E}e^{\lambda(X-\mu)} \leq e^{\frac{\lambda^2\sigma^2/2}{1-b|\lambda|}}$ can be used to give more tighter bounds than the sub-exponential (special case when $\lambda = \frac{1}{2h}$)

Aside: This is a convexity bound, follows from the general framework of conjugate duality (*Fenchenl transform*)

$$\frac{1}{1-x} = \sum_{k=0}^{\infty} x^k f(t) \tag{3.5}$$

$$f^*(\lambda) := \sup_{\lambda} (\lambda t - f(t))$$
 (3.6)

$$|x| < 1 \tag{3.7}$$

$$1 + t \le e^t \tag{3.8}$$

TODO: Fig 4.3

The following is worth memorizing:

Proposition 3.6 (Bernstein's inequality). From

$$\mathbb{E}e^{\lambda(X-\mu)} \le e^{\frac{\lambda^2\sigma^2/2}{1-b|\lambda|}} \tag{3.9}$$

we get

$$P(|X - \mu| \ge t) \le 2 \exp\left(\lambda t - \frac{\lambda^2 \sigma^2}{2(1 - b|\lambda|)}\right)$$
(3.10)

$$=2e^{-\frac{t^2}{2(bt+\sigma^2)}}\tag{3.11}$$

where the last line is obtained by setting $\lambda = \frac{t}{bt + \sigma^2}$.

Whenever you have bounded RVs, instead of reaching for Hoeffding, consider Bernstein.

Proposition 3.7 (Sum of sub-exponentials). *If* $X_i \sim SE(\nu_i, \alpha_i)$, then $\sum_i X_i \sim SE(\sqrt{\sum_i \nu_i^2}, \max_i \alpha_i)$

Proof. Let X_i be independent with mean μ_i and sub-exponential parameters ν_i , α_i .

$$\mathbb{E}e^{\lambda \sum_{i=1}^{n} (X_i - \mu_i)} = \prod_{i=1}^{n} \mathbb{E}e^{\lambda (X_i - \mu_i)}$$
(3.12)

$$\leq \prod_{i=1}^{n} e^{\lambda^2 \nu_i^2 / 2} \tag{3.13}$$

$$=e^{\lambda^2 \sum_i \nu_i^2/2} \tag{3.14}$$

This holds for
$$\lambda \leq \frac{1}{\alpha_i}$$
 for all i , so $\lambda \leq \frac{1}{\max_i \alpha_i}$.

Example 3.8. Let $Y = \sum_{k=1}^{n} Z_k^2$ where $Z_k \sim N(0,1)$. $Y \sim \chi_n^2$ and since it's a sum of sub-exponentials, $Y \sim SE(2\sqrt{n},4)$.

Know $Z_K^2 \sim SE(2,4)$, so

$$P\left(\left|\frac{1}{n}\left(\sum_{k=1}^{n} Z_k^2 - 1\right)\right| \ge t\right) \le 2e^{-\frac{nt^2}{8}}$$
 (3.15)

Example 3.9 (Johnson-Lindenstrauss). N > 2 points in \mathbb{R}^d , say u^1, \dots, u^N . Project onto $m \ll d$ dimensions, $F : \mathbb{R}^d \to \mathbb{R}^m$.

Want to preserve metric approximately (i.e. with "fudge-factor" δ)

$$1 - \delta \le \frac{\|F(u^i) - F(u^j)\|_2^2}{\|u^i - u^j\|_2^2} \le 1 + \delta \tag{3.16}$$

Can do this with high probability if

$$m \gtrsim \frac{1}{\delta^2} \log N \tag{3.17}$$