

1 Large Deviation Theorem

If $a_n \sim ce^{\beta n}$ as $n \rightarrow \infty$, then $\frac{1}{n} \log a_n \rightarrow \beta = \text{"asymptotic growth rate."}$

Today $\beta < 0$. $\xrightarrow{\text{decay}}$ $\xrightarrow{\text{growth}}$

IID (X_i) . $S_n = \sum_{i=1}^n X_i$. $\mathbb{E}X = \mu$. Fix $a > \mu$, $P(X \geq a) > 0$.

Consider $P\left(\frac{S_n}{n} \geq a\right)$. This $\rightarrow 0$ as $n \rightarrow \infty$ by WLLN. How fast? We already have

Proposition 1.1 (General large-deviation inequality).

$$P(Y \geq y) \leq \inf_{\theta \geq 0} \frac{\mathbb{E}e^{\theta Y}}{e^{\theta y}} \quad (1.1)$$

Definition 1.2. The transform of X

$$\phi(\theta) = \mathbb{E} \exp(\theta X) \quad (1.2)$$

Assume $\theta^* = \sup\{\theta : \phi(\theta) < \infty\} > 0$. Going back to "how fast?"

By general LD inequality and independence of X_i

$$P\left(\frac{S_n}{n} \geq a\right) = P(S_n \geq an) \quad (1.3)$$

$$\leq \inf_{\theta} \frac{\mathbb{E} \exp(\theta S_n)}{\exp(\theta an)} = \inf_{\theta} \frac{\mathbb{E} \prod_{i=1}^n \exp(\theta X_i)}{\exp(\theta an)} \quad (1.4)$$

$$= \inf_{\theta} \frac{\prod_{i=1}^n \mathbb{E} \exp(\theta X_i)}{\exp(\theta an)} \quad (1.5)$$

$$= \left(\inf_{\theta} \frac{\phi(\theta)}{e^{\theta a}} \right)^n \quad (1.6)$$

$$\frac{1}{n} \log P\left(\frac{S_n}{n} \geq a\right) \leq \inf_{\theta} [\log \phi(\theta) - a\theta] \quad (1.7)$$

So we have

$$\forall n : \frac{1}{n} \log P\left(\frac{S_n}{n} \geq a\right) \leq \inf_{\theta} [\log \phi(\theta) - a\theta] \quad (1.8)$$

Theorem 1.3. As $n \rightarrow \infty$

$$\lim \frac{1}{n} \log P\left(\frac{S_n}{n} \geq a\right) = \inf_{\theta} [\log \phi(\theta) - a\theta] \quad (1.9)$$

1.1 Proof outline

Three steps:

- (a) Analysis of $\phi(\theta)$
- (b) Tilting lemma
- (c) Put together

Lemma 1.4. $\phi'(0+) = \mu$

$$\frac{d}{d\theta}\phi(\theta) = \frac{d}{d\theta}\mathbb{E}e^{\theta X} \stackrel{?}{=} \mathbb{E}\frac{d}{d\theta}e^{\theta X} = \mathbb{E}[Xe^{\theta X}] \quad \forall \theta \quad (1.10)$$

$$\theta = 0 : \phi'(0+) = \mathbb{E}X \quad (1.11)$$

How to justify “?” in detail?

Proof. Know $\frac{e^{\theta X}-1}{\theta} \rightarrow X$ as $\theta \downarrow 0$. **Want:** $\mathbb{E}\left[\frac{e^{\theta X}-1}{\theta}\right] \rightarrow \mathbb{E}X$.

Tool for going from sequence convergence to convergence of an expectation: Dominated Convergence Theorem.

$$x > 0 : e^{\theta x} - 1 = \int_0^{\theta x} e^y dy \leq \theta x e^{\theta x} \quad (1.12)$$

$$x < 0 : |e^{\theta x} - 1| = \int_{\theta x}^0 e^y dy \leq |\theta x| \quad (1.13)$$

$$\implies |e^{\theta x} - 1| \leq \theta |x| \max(1, e^{\theta x}) \quad (1.14)$$

For $0 < \theta \leq \theta_0$

$$\mathbb{E}\left[\frac{e^{\theta X}-1}{\theta}\right] \leq |X| \max(1, e^{\theta_0 X}) \quad (1.15)$$

Hypothesis: $\exists \theta_1$ such that $\mathbb{E}e^{\theta_1 X} < \infty$.

Choose $\theta_0 < \theta_1$

$$\implies \mathbb{E}[|X| \max(1, e^{\theta X})] < \infty$$

Now $\mathbb{E}\left[\frac{e^{\theta X}-1}{\theta}\right] \ll |X| \max(1, e^{\theta X})$ and $|X| \max(1, e^{\theta X})$ is integrable.

Apply DCT. □

Lemma 1.5. $\phi'(0+) = \mu$ and for $0 < \theta < \theta^*$

$$\phi'(\theta) = \mathbb{E}[Xe^{\theta X}] \quad (1.16)$$

$$\phi''(\theta) = \mathbb{E}[X^2 e^{\theta X}] \quad (1.17)$$

Proof. Write out as integrals, apply Fubini-Tonelli theorem. □

Suppose X discrete (so $\phi(\theta) = \sum_x e^{\theta x} P(X = x)$). Fix θ . Define a dist for \hat{X} by

$$P(\hat{X} = x) = \frac{e^{\theta x} P(X = x)}{\phi(\theta)} \quad (1.18)$$

$$\mathbb{E} \hat{X} = \sum_x x P(\hat{X} = x) = \frac{\sum_x x e^{\theta x} P(X = x)}{\phi(\theta)} \quad (1.19)$$

$$= \frac{\mathbb{E} X e^{\theta X}}{\phi(\theta)} = \frac{\phi'(\theta)}{\phi(\theta)} = \frac{d}{d\theta} \log \phi(\theta) \quad (1.20)$$

$$\mathbb{E}(\hat{X}^2) = \frac{\mathbb{E}[X^2 e^{\theta X}]}{\phi(\theta)} = \frac{\phi''(\theta)}{\phi(\theta)} \quad (1.21)$$

$$\text{Var}(\hat{X}) = \mathbb{E}(\hat{X}^2) - (\mathbb{E} \hat{X})^2 \quad (1.22)$$

$$= \frac{\phi''(\theta)}{\phi(\theta)} - \left(\frac{\phi'(\theta)}{\phi(\theta)} \right)^2 \quad (1.23)$$

$$= \frac{d}{d\theta} \left(\frac{\phi'(\theta)}{\phi(\theta)} \right) = \frac{d}{d\theta} \log \phi(\theta) \quad (1.24)$$

The transform of X , $\phi(\theta)$, encodes information about the moments.

For general X , define distribution of \hat{X} by Radon-Nikodym density $(\frac{d\nu}{d\mu})$

$$\frac{dP(\hat{X} \in \cdot)}{dP(X \in \cdot)}(x) = \frac{e^{\theta x}}{\phi(\theta)} \quad (1.25)$$

Lemma 1.6. $\mathbb{E} \hat{X} = \frac{d}{d\theta} \log \phi(\theta)$, $\text{Var} \hat{X} = \frac{d^2}{d\theta^2} \log \theta$

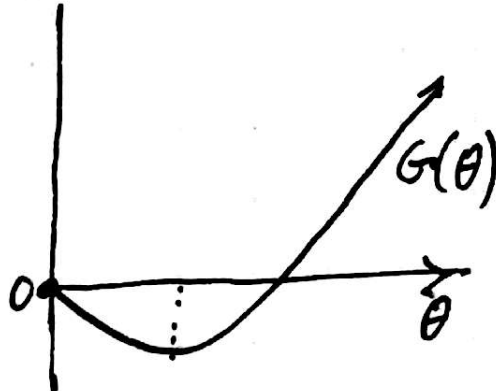
1.1.1 Study $G(\theta) = \log \phi(\theta) - a\theta$

$$G'(0+) = (\log \phi(\theta))' - a = \mu - a \quad (1.26)$$

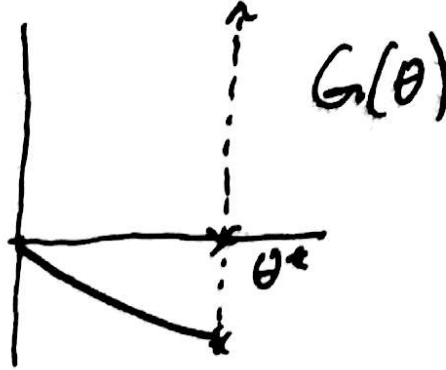
$$G''(\theta) = \text{Var} X_\theta > 0 \text{ on } 0 < \theta < \theta^* \quad (1.27)$$

$$G(0) = 0 \quad (1.28)$$

So G is strictly convex on $(0, \theta^*)$. Easy to show $G(\theta) \rightarrow \infty$ as $\theta \rightarrow \infty$.



Find \inf_{θ} of $G(\theta)$ by solving $G'(\theta) = 0 \implies \frac{\phi'(\theta)}{\phi(\theta)} = a$.



Case 1: \exists solution $\theta_a \in (0, \theta^*)$ of equation $\frac{\phi'(\theta)}{\phi(\theta)} = a$

Assume case 1. Choose $\theta \in (\theta_a, \theta^*)$. Consider tilted distribution $\hat{X} = \hat{X}_{\theta}$.

$$\mathbb{E} \hat{X} = \frac{d}{d\theta} \log \phi(\theta) > \frac{d}{d\theta} \log \phi(\theta) |_{\theta=\theta_a} \quad (1.29)$$

TODO: ??? $\mathbb{E} \hat{X} > a$ and $\mathbb{E} \hat{X}_{\theta} \downarrow a$ as $\theta \downarrow \theta_a$ [Check!].

Fix $b > \mathbb{E} \hat{X}_{\theta}$.

Trick: Apply WLLN to tilted (\hat{X}_i)

$$\frac{P(\hat{X}_1 = x_1, \dots, \hat{X}_n = x_n)}{P(X_1 = x_1, \dots, X_n = x_n)} = \frac{e^{\theta \sum_i x_i}}{\phi^n(\theta)} \implies \underbrace{\frac{P(\hat{S}_n = s)}{P(S_n = s)}}_{\text{interp as Radon-Nikodym density}} = \frac{e^{\theta s}}{\phi^n(\theta)} \quad (1.30)$$

$$\frac{P(y_1 \leq \hat{S}_n \leq y_2)}{P(y_1 \leq S_n \leq y_2)} \leq \frac{e^{\theta y_2}}{\phi^n(\theta)} \quad (1.31)$$

$$P(a \leq \frac{S_n}{n} \leq b) \geq e^{-\theta b_n} \phi^n(\theta) \quad (1.32)$$

$$P(a \leq \frac{S_n}{n} \leq b) \rightarrow 1 \quad \text{as } n \rightarrow \infty \quad (1.33)$$

where we use $y_1 = an, y_2 = bn$.

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log P(\frac{S_n}{n} \geq a) \geq -b\theta + \log \phi(\theta) \quad (1.34)$$

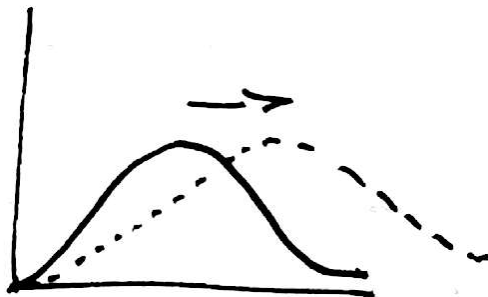
Let $\theta \downarrow \theta_a$.

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log P(\frac{S_n}{n} \geq a) \geq -b\theta_a + \log \phi(\theta_a) \quad (1.35)$$

True $\forall b > a$, let $b \downarrow a$

$$\liminf_{n \rightarrow \infty} \frac{1}{n} \log P\left(\frac{S_n}{n} \geq a\right) \geq -a\theta_a + \log \phi(\theta_a) = G(\theta_a) \quad (1.36)$$

Why is it called tilting?



2 Intro to next lecture

Suppose X, Y continuous.

TODO: Format better

$$P(x, y) = P(X = x, Y = y) \leftrightarrow \quad (2.1)$$

$$\text{marginal dist } p_X(x) = P(X = x) \leftrightarrow \quad (2.2)$$

$$\text{conditional dist of } Y \text{ given } X = x \quad p_{Y|X}(y|x) = P(Y = y|X = x) \leftrightarrow y \mapsto f_{Y|X}(y|x) \text{ conditional density of } Y \text{ given } X = x \quad (2.3)$$

$$\text{relation } p(x, y) = p_X(x)p_{Y|X}(y, x) \leftrightarrow \quad (2.4)$$