

# 1 Review

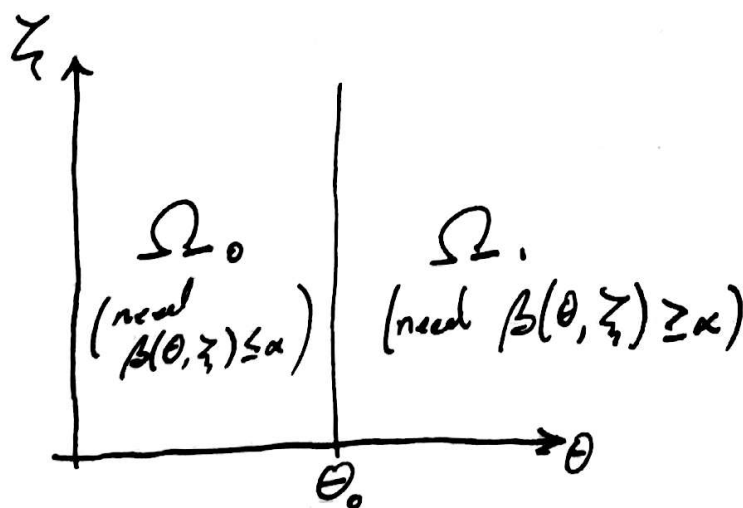
Testing with nuisance parameters  $H_0 : \theta \leq \theta_0$  vs  $H_1 : \theta > \theta_0$ .

$$p_{\theta, \zeta}(x) = e^{\theta T(x) + \zeta' U(x) - A(\theta, \zeta)} h(x) \quad (1.1)$$

$$\phi^*(x) = \psi(T(x), U(x)) \quad (1.2)$$

$$\psi(t, u) = \begin{cases} 1, & \text{if } t > c(u) \\ \gamma(u), & \text{if } t = c(u) \\ 0, & \text{if } t < c(u) \end{cases} \quad (1.3)$$

$$\mathbb{E}_{\theta_0}[\phi \mid U] \stackrel{\text{a.s.}}{=} \alpha \quad (1.4)$$



## 2 Proof UMPU test for multi-param exp fam

(1)  $\phi^*$  works,  $\theta \leq \theta_0$

$$\mathbb{E}_{\theta, \zeta}[\phi^*(X)] = \mathbb{E}_{\theta, \zeta}[\mathbb{E}_{\theta}[\phi^* \mid U]] \leq \alpha \text{ if } \theta \leq \theta_0 \quad (2.1)$$

$$(\geq \alpha \text{ if } \theta \geq \theta_0) \quad (2.2)$$

(2) Recall Keener Theorem 2.4. Since  $\mathbb{E}_{\theta, \zeta}[|\phi(X)|] < \infty$ ,  $\beta(\theta, \zeta)$  is continuous, and all derivatives can be computed by differentiating under the integral.

$$\implies \beta_{\phi}(\theta_0, \zeta) = \alpha \quad \forall \zeta, \forall \phi \text{ level-}\alpha \text{ unbiased}$$

(3) Write  $Q_{\theta_0} = \{q_{\zeta}(x) = p_{\theta_0, \zeta}(x)\}$ . Then  $q_{\zeta}(x) = e^{\zeta'U(x) - A(\theta_0, \zeta)} \underbrace{e^{\theta_0' T(x)} h(x)}_{U(X)}$ .  $U(X)$  complete sufficient.

$$f(u) = \mathbb{E}_{\theta_0}[\phi(X) \mid U(X) = u] \quad (2.3)$$

$$\mathbb{E}_{\theta_0, \zeta}[f(U)] = \alpha \quad \forall \zeta \quad (2.4)$$

$$\implies f(U) \stackrel{\text{a.s.}}{=} \alpha \quad (2.5)$$

(4) Suppose  $\phi(X)$  is level- $\alpha$  given  $U = u$

$$\mathbb{E}_{\theta, \zeta}[\phi(X)] = \mathbb{E}_{\theta, \zeta}[\mathbb{E}_{\theta}[\phi(X) \mid U]] \leq \mathbb{E}_{\theta, \zeta}[\mathbb{E}_{\theta}[\phi^*(X) \mid U]] \leq \mathbb{E}_{\theta, \zeta}[\phi^*(X)] \quad (2.6)$$

Step (2) says power is  $\alpha$  on the boundary of  $\Omega_0 = \{(\theta, \zeta) : \beta(\theta, \zeta) \leq \alpha\}$ .

Step (3) says conditional power is  $\alpha$  on the boundary.  $U$  is complete because of some open set business wrt  $\Omega_0$ .

Step (2)-(3) reduces to a 1-dimensional problem.

Step (4) shows that  $\phi^*$  is uniformly most powerful.

### 3 Examples

**Example 3.1** (Logistic Regression).  $x_1, \dots, x_n \in \mathbb{R}^d$ .

$$y_i \stackrel{\text{indep}}{\sim} \text{Bern}\left(\frac{e^{\beta'x}}{1+e^{\beta'x}}\right)$$

$$p_{\beta}(y \mid x) = \prod_{i=1}^n e^{\beta'x_i y_i - A_i(\beta)} = e^{\beta'(x'y) - A(\beta)} \quad (3.1)$$

Test  $H_0 : \beta_1 \leq 0$  vs  $\beta_1 > 0$

Condition on  $X_2' y = u \iff$  Condition on  $y$

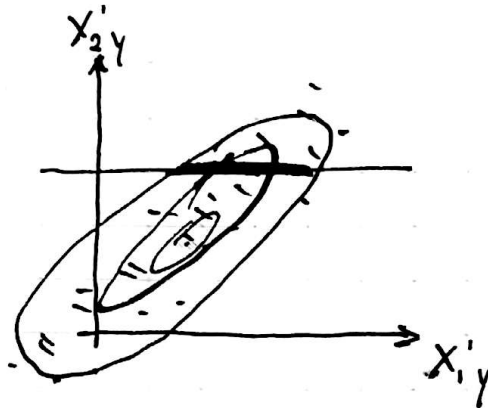


Figure 1: **TODO: Does this belong here?**

**Example 3.2** (One-sample  $t$ -test).  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} \mathcal{N}(\mu, \sigma^2)$ ,  $\sigma^2 > 0$  unknown.

$$p_{\mu, \sigma^2}(x) = \left( \frac{1}{2\pi\sigma^2} \right)^{n/2} e^{\frac{\mu}{\sigma^2} \sum_{i=1}^n X_i - \frac{1}{2\sigma^2} \sum_{i=1}^n X_i^2 - \frac{n\mu^2}{2\sigma^2}} \quad (3.2)$$

$$\theta = \frac{\mu}{\sigma^2} \quad (3.3)$$

$$\zeta = \frac{-1}{2\sigma^2} \quad (3.4)$$

Test  $H_0 : \mu \leq 0 \iff \theta \leq 0$  vs  $H_1 : \mu > 0 \iff \theta > 0$

Condition on  $\|X\|^2 = u$ :

$$\implies p_{0, \sigma^2}(x \mid \|X\|^2 = u) \propto e^{0 \sum_{i=1}^n X_i} 1_{\|X\|^2=u}$$

( $X \sim$  uniform on sphere of radius  $\sqrt{u}$ )

$$\iff \frac{X}{\|X\|} \mid \|X\| \sim \text{Unif}(\text{unit sphere})$$

Optimal test rejects when

$$T(X) = n\bar{X} > c(u) \iff \text{reject when } \frac{\bar{X}}{\|X\|} > 0 \quad (3.5)$$

$$S^2 = \sum_i (X_i - \bar{X})^2 = \|X\|^2 - n\bar{X}^2 \iff \text{reject when } \frac{\bar{X}}{\sqrt{S^2/(n-1)}} \stackrel{\text{large}}{\sim} t_{n-1} \quad (3.6)$$

$$B(X) = \frac{\bar{X}}{\|X\|} \quad D(X) = \frac{\bar{X}}{\sqrt{S}} \quad (3.7)$$

$$B^2 = \frac{X^2}{\|X\|^2} \quad D^2 = \frac{X^2}{\|X\| - X^2} \quad (3.8)$$

$$\frac{1}{D^2} = \frac{1}{B^2} - 1 \quad (3.9)$$

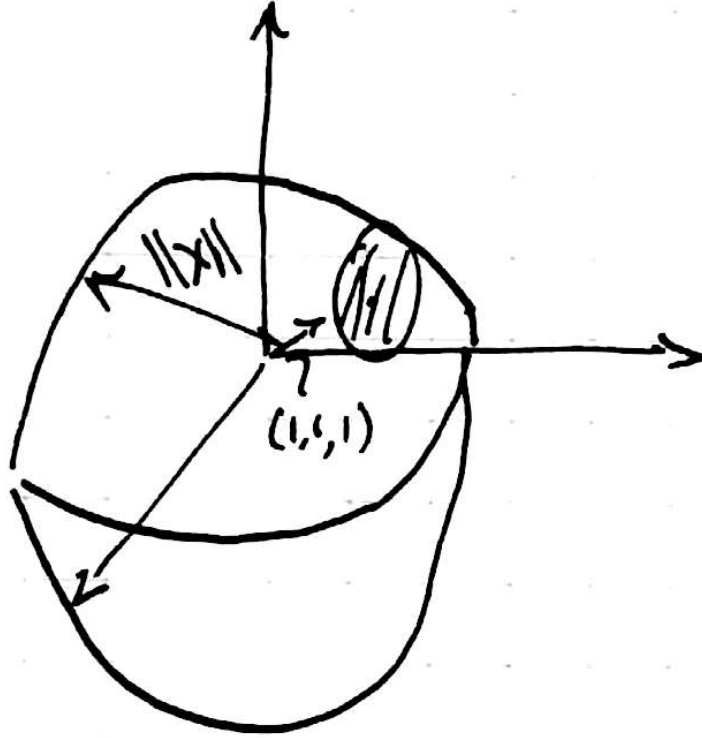


Figure 2: Rejection region is a cap on the sphere  $\{\|X\| = \sqrt{u}\}$

**Example 3.3** (Permutation Tests).  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} P$ .

$Y_1, \dots, Y_m \stackrel{\text{iid}}{\sim} Q$ .

$H_0 : P = Q, H_1 : P \neq Q$ .

Complete sufficient statistics under  $H_0$  are the order statistics

$$Z = (X_1, \dots, X_n, Y_1, \dots, Y_m) \quad (3.10)$$

$$U(X, Y) = (Z_{(1)}, Z_{(2)}, \dots, Z_{(n)}) \quad (3.11)$$

Choose any test stat  $T(X, Y)$ . Under  $H_0$  (not true under  $H_1$ )

$$T(X, Y) = \bar{X} - \bar{Y} \quad (3.12)$$

$$(X, Y) \mid U(X, Y) \sim \text{Unif over all permutations of } (Z_1, \dots, Z_{m+n}) \quad (3.13)$$

$$P_{H_0}(T(X, Y) > 0 \mid U(X, Y)) = \frac{1}{(n+m)!} \sum_{\pi \in S_{n+m}} 1_{T((Z_{\pi(1)}, \dots, Z_{\pi(n)}), (Z_{\pi(n+1)}, \dots, Z_{\pi(n+m)})) > 0} \quad (3.14)$$

Reject when  $T(X, Y)$  large.

Choose  $c$  such that under  $H_0$  but not  $H_a$ : conditional power is  $\alpha \implies$  marginal power is  $\alpha$ .

**TODO: WTF?** This means that we should choose order statistics  $Z$ , not entire data, because the entire data statistic does not differ between null and alternative hypotheses.

## 4 Tests in linear models