## 1 Plug-in estimator

Suppose we observe

$$X_1, \dots, X_n \stackrel{\text{iid}}{\sim} P \quad X_i \in \mathbb{R}^d$$
 (1.1)

**Example 1.1.** Say  $X_{ij}$  = score of student i on test in subject j.

Combine into 
$$X = \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}$$
 and do PCA to estimate

$$\theta(p) = \underset{\|v\|=1}{\operatorname{arg\,max}} \operatorname{Var}_{p}(v'X_{i})$$
(1.2)

(assume unique).

One natural estimator is

$$\hat{\theta}(X_1, \dots, X_n) = \arg\max_{\|v\|=1} \frac{1}{n} \sum_{i=1}^n \left[ v'(X_i - \bar{X}) \right]^2$$
 (1.3)

where  $\bar{X} = \frac{1}{n} \sum_{i} X_{i}$  is the sample average.

$$\hat{\theta} = \underset{\|v\|=1}{\arg\max} \operatorname{Var}_{\hat{p}_n}(v'X) \tag{1.4}$$

$$\hat{p}_n = \frac{1}{n} \sum_{i=1}^n \delta_{X_i} \tag{1.5}$$

For 
$$A \subset \mathbb{R}^d$$
:  $\hat{p}_n(A) = \frac{\#\{i \le n : X_i \in A\}}{n} \stackrel{\text{a.s.}}{\to} P(A)$  (1.6)

 $\hat{p}$  is called the *empirical distribution*.

The plug-in estimator is

$$\hat{\theta} = \theta(\hat{p}_n) \tag{1.7}$$

#### **Questions:**

- (a) What are bias & variance?
- (b) Can we get a CI?

## 2 Bootstrap

We want  $\operatorname{Var}_{p}(\hat{\theta}_{n})$ .

Bootstrap estimates via plug-in

$$\hat{\text{Var}}(\hat{\theta}_n) = \underbrace{\text{Var}_{\hat{p}_n}(\hat{\theta}_n)}_{\text{can compute directly}}$$
(2.1)

 $\operatorname{Var}_{\hat{p}_n}(\hat{\theta}_n)$  integrates over possible samples

$$X_1^*, \dots, X_n^* \stackrel{\text{iid}}{\sim} \hat{p}_n \tag{2.2}$$

**MC Algorithm**: For  $b = 1, \dots, B$ :

- 1) Sample  $X_1^{*b}, \ldots, X_n^{*b} \stackrel{\text{iid}}{\sim} \hat{p}_n$  with replacement
- 2) Compute  $\hat{\theta}^{*b} = \hat{\theta}(x_1^{*b}, ..., x_n^{*b})$

$$\bar{\theta}^* = \frac{1}{B} \sum_{b=1}^{B} \hat{\theta}^{*b} \tag{2.3}$$

$$\hat{\text{Var}} = \frac{1}{B} \sum_{h=1}^{B} (\hat{\theta}^{*b} - \bar{\theta}^{*})^2 \xrightarrow[B \to \infty]{\text{a.s.}} \text{Var}_{\hat{p}_n}(\hat{\theta}_n)$$
 (2.4)

Notes:

- Not necc for  $\hat{\theta}\theta(\hat{p}_n)$
- $\hat{p_n}$  need not be empirical distribution

#### 2.1 Bias correction

$$\operatorname{Bias}_{p}(\hat{\theta}) = \mathbb{E}_{p}[\hat{\theta}] - \theta(p) \tag{2.5}$$

$$\hat{\text{Bias}}(\hat{\theta}_n) = \text{Bias}_{\hat{\theta}_n}(\hat{\theta}) \tag{2.6}$$

$$= \mathbb{E}_{\hat{p}_n}(\hat{\theta}_n(X_1^*, \dots, X_n^*)) - \theta(\hat{p}_n)$$
(2.7)

If we knew the true bias

$$\bar{\theta}_n = \hat{\theta}_n - \operatorname{Bias}_p(\hat{\theta}_n) \tag{2.8}$$

Unfortunately, we don't know p so we don't know the bias, but we can estimate it with Boostrap

$$\check{\theta_n} = \hat{\theta}_n - \operatorname{Bias}_{\hat{p}_n}(\hat{\theta}_n) \tag{2.9}$$

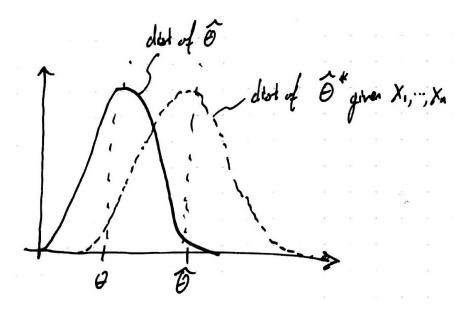


Figure 1: Typical picture showing bias in distribution of estimators

# 3 Bootstrap Confidence Regions

**Definition 3.1.** A *root*  $R_n(X, \theta(p))$  depends on data X and parameter  $\theta$ .

If we know the root's distribution

$$L_n(r; P) = \mathbb{P}_p[R_n(X, \theta(p)) \le r] \tag{3.1}$$

then we can make it into a confidence region C(X; P)

$$C_{\alpha}(X) = \{\theta : L_n(R_n(X;\theta)) \le 1 - \alpha\}$$
(3.2)

$$= \{\theta : R_n(X; \theta) \le L_n^{-1}(1 - \alpha)\}$$
 (3.3)

Then

$$\mathbb{P}_{p}(\theta(p) \in C_{\alpha}(X)) = \mathbb{P}_{p}\left(R_{n}(X, \theta(p)) \le L_{n}^{-1}(1 - \alpha; P)\right) \ge 1 - \alpha \tag{3.4}$$

**Proposal**: Use  $L_n(\cdot, \hat{p}_n)$ 

**Example 3.2.** 
$$R_n = |\hat{\theta}(X) - \theta(p)|.$$
  $\Longrightarrow C(X) = \hat{\theta} = L_n^{-1}(1 - \alpha, \hat{p}_n)$ 

**Example 3.3.** 
$$R_n = \frac{|\hat{\theta} - \theta|}{\hat{\sigma}(X)}$$
 ("studentized root")  $\implies C(X) = \hat{\theta} \pm \hat{\sigma} L_n^{-1} (1 - \alpha, \hat{p}_n)$ 

**Example 3.4.** 
$$R_n = \|\hat{\theta} - \theta(p)\|_{\infty}$$
  $C = [\hat{\theta}_1 = L_n^{-1}(1-\alpha)] \times \cdots \times [\hat{\theta}_d = L_n^{-1}(1-\alpha)]$ 

**Example 3.5.**  $\ell(\hat{\theta}(X)) - \ell(\theta(p))$  [ $\ell$  any loss function]

Operationally (computing  $L_n(r, \hat{p}_n)$ )), for b = 1, ..., B:

$$X_1^{*b}, \dots, X_n^{*b} \stackrel{\text{iid}}{\sim} \hat{p}_n$$

$$R_n^{*b} = R_n(X_1^{*b}, \dots, X_n^{*b}; \theta(\hat{p}_n))$$

Example 3.6. Wald

$$\hat{\theta} \pm \frac{1}{\sqrt{nJ_1(\hat{\theta})}} z_{\alpha/2} \tag{3.5}$$

$$R_n: |\hat{\theta} - \theta| - \sqrt{nJ_1(\hat{\theta})} \Rightarrow |N(0,1)|$$
  
[ $R_n$  is an asymptotic pivot]  $L_n(\cdot; \hat{p}_n) \to L_n(\cdot; p)$