

1 Patterns in coin-tossing

Say in words (exercise X 's general pattern):

Fix pattern $HTTHT$. Toss fair coin until see this pattern — requires W tosses where W is random, $t \leq W < \infty$ a.s.

What is $\mathbb{E}W$?

Consider the strategy

- bet \$1 that toss i is H
- if win, bet \$2 that toss $i + 1$ is T
- if win, bet \$4 that toss $i + 2$ is T
- if win, bet \$8 that toss $i + 3$ is H
- if win, bet \$16 that toss $i + 4$ is T

Do strategy for each $1 \leq i \leq W$, stop after toss W .

W is a stopping time, so the optional sampling theorem implies $\mathbb{E}[\text{profit}] = 0$.

$$\text{cost} = W \tag{1.1}$$

$$\text{return} = 32 + 4 = 36 \tag{1.2}$$

$$\text{profit} = \text{return} - \text{cost} \tag{1.3}$$

$$0 = \mathbb{E}[\text{profit}] = \mathbb{E}[36 - W] = 36 - \mathbb{E}W \tag{1.4}$$

$$\mathbb{E}W = 36 \tag{1.5}$$

2 MG proof of Radon-Nikodym

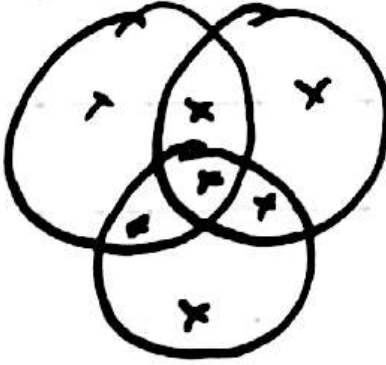
Theorem 2.1 (Radon-Nikodym). *Let (S, \mathcal{S}, μ) be a probability space, $\mathcal{S} = \sigma(A_i, i \geq 0)$ countable events.*

If $\nu \ll \mu$, $\nu(S) < \infty$, then \exists measurable $h : S \rightarrow [0, \infty]$ such that $\nu(A) = \int_A h d\mu$ for all $A \in \mathcal{S}$.

$h = \frac{d\nu}{d\mu}$ is the Radon-Nikodym density of ν wrt μ .

Heuristics: $h(s) = \frac{d\nu}{d\mu}(s) = \lim_{A \downarrow \{s\}} \frac{\nu(A)}{\mu(A)}$.

Proof. Define $\mathcal{F}_n = \sigma(A_i, 1 \leq i \leq n)$ finite field with 2^n atoms.



Define $X_n(s) = \frac{\nu(F)}{\mu(F)}$ for atom $F \ni s$.

$= \nu(F)$ for each $F \in \mathcal{F}_n$

Claim: (X_n, \mathcal{F}_n) is a MG.

Justification: Take $G \in \mathcal{F}_{n-1}$.

$$G = \underbrace{(G \cap A_n)}_{G_1} \cup \underbrace{(G \cap A_n^c)}_{G_2} \quad (2.1)$$

$$\mathbb{E}X_n 1_G = \mathbb{E}X_n 1_{G_1} + \mathbb{E}X_n 1_{G_2} \quad (2.2)$$

$$\text{Equation (2.3)} \implies = \nu(G_1) + \nu(G_2) = \nu(G) = \mathbb{E}X_{n-1} 1_G \quad (2.3)$$

$$X_{n-1} = \mathbb{E}[X_n \mid \mathcal{F}_{n-1}] \quad (2.4)$$

By MG convergence theorem, $X_n \rightarrow X_\infty$ for some $X_\infty \geq 0$ a.s. If we prove $(X_n, n \geq 1)$ is UI, then Theorem from Lecture 20 implies $X_n = \mathbb{E}[X_\infty \mid \mathcal{F}_n]$ hence

$$\mathbb{E}X_\infty 1_F = \mathbb{E}X_n 1_F = \nu(F) \quad \text{eq. (2.3)} \quad (2.5)$$

$$\mathbb{E}X_\infty 1_F = \nu(F) \quad \forall F \in \cup_n \mathcal{F}_n \quad (2.6)$$

$$\mathbb{E}X_\infty 1_F = \nu(F) \quad \forall F \in \sigma(\cup_n \mathcal{F}_n) = \mathcal{S} \quad (2.7)$$

$$\nu(F) = \mathbb{E}_\mu X_\infty 1_F = \int_F X_\infty d\mu \quad (2.8)$$

$$(2.9)$$

So X_∞ is R-N density $\frac{d\nu}{d\mu}$ and we are done.

Lemma 2.2. Suppose $\nu \ll \mu$. $\forall \varepsilon > 0, \exists \delta(\varepsilon) > 0$ such that

$$\mu(A) \leq \delta(\varepsilon) \implies \nu(A) \leq \varepsilon \quad (2.10)$$

Proof. If false for ε , $\exists A_n$ such that $\mu(A_n) \leq 2^{-n}$ and $\nu(A_n) > \varepsilon$. Consider $\Lambda = \{A_n \text{ i.o.}\}$. By Borel-Cantelli, $\mu(\Lambda) = 0$ but $\nu(\Lambda) \geq \varepsilon$, contradicting $\nu \ll \mu$. \square

Claim: (X_n) is UI.

Justification: By eq. (2.3), $\mathbb{E}X_n 1_{X_n \geq b} = \nu(X_n \geq b)$. Given $\varepsilon > 0$, take b such that $\frac{\nu(S)}{b} \leq \delta(\varepsilon)$. Then

$$\nu(X_n \geq b) \underset{\text{Markov}}{\leq} \frac{\mathbb{E}X_n}{b} = \frac{\nu(S)}{b} \leq \delta(\varepsilon) \quad (2.11)$$

$$\text{lemma 2.2} \nu(X_N \geq b) \leq \varepsilon \quad (2.12)$$

$$\implies \sup_n \mathbb{E}X_n 1_{X_n \geq b} \leq \varepsilon \quad (2.13)$$

So UI. \square

The above proof relies on martingale convergence theorem for existence of R-N density X_∞ . It also only holds for countable events.

3 Azuma's inequality

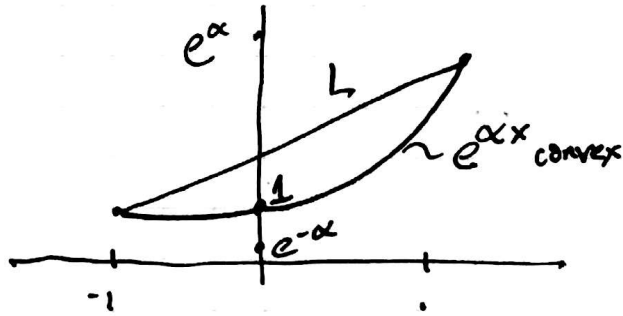
Theorem 3.1 (Azuma's inequality). Let $S_n = \sum_{i=1}^n X_i$ be a MG with $|X_i| \leq 1$ a.s. Then for $\lambda > 0$,

$$P(S_n \geq \lambda\sqrt{n}) \leq e^{-\lambda^2/2} \quad (3.1)$$

so

$$P(|S_n| \geq \lambda\sqrt{n}) \leq 2e^{-\lambda^2/2} \quad (3.2)$$

Lemma 3.2. If $\mathbb{E}Y = 0$ and $|Y| \leq 1$, then $\mathbb{E}e^{\alpha Y} \leq e^{\alpha^2/2}$ for all α .



Proof of Lemma 3.2.

$$\mathbb{E}e^{\alpha Y} \underset{\text{convexity}}{\leq} \mathbb{E}L(Y) \underset{\text{linear}}{=} L(\mathbb{E}Y) = L(0) = (e^{\alpha} + e^{-\alpha})/2 \underset{\text{calculus}}{\leq} e^{\alpha^2/2} \quad (3.3)$$

Calculus: coefficient of α^{2n} in series expansion

$$\frac{1}{(2n)!} \leq \frac{1}{2^n n!} \quad (3.4)$$

□

Proof of Azuma (Theorem 3.1). Apply lemma 3.2 to conditional distribution of X_i given \mathcal{F}_{i-1}

$$\mathbb{E}[e^{\alpha X_i} \mid \mathcal{F}_{i-1}] \leq e^{\alpha^2/2} \quad (3.5)$$

$$\mathbb{E}[e^{\alpha S_n} \mid \mathcal{F}_{n-1}] = e^{\alpha S_{n-1}} \mathbb{E}[e^{\alpha X_n} \mid \mathcal{F}_{n-1}] \leq e^{\alpha^2/2} e^{\alpha S_{n-1}} \quad (3.6)$$

$$(3.7)$$

Take \mathbb{E} and apply tower property

$$\mathbb{E}[e^{\alpha S_n}] \leq e^{\alpha^2/2} \mathbb{E}[e^{\alpha S_{n-1}}] \quad (3.8)$$

$$\mathbb{E}[e^{\alpha S_n}] \leq \left(e^{\alpha^2/2}\right)^n = e^{\alpha^2 n/2} \quad (3.9)$$

Applying Markov inequality with $\phi = \exp$

$$P(S_n \geq \lambda \sqrt{n}) \leq \frac{\mathbb{E} e^{\alpha S_n}}{e^{\alpha \lambda \sqrt{n}}} \leq e^{\alpha^2 n/2 - \alpha \lambda \sqrt{n}} \quad (3.10)$$

Minimize over α by taking $\alpha = \lambda / \sqrt{n}$

$$P(S_n \geq \lambda \sqrt{n}) \leq e^{\alpha^2 n/2 - \alpha \lambda \sqrt{n}} = e^{-\lambda^2/2} \quad (3.11)$$

□

4 Method of bounded differences

Corollary 4.1. Take $(\xi_i, 1 \leq i \leq n)$ independent, arbitrary state spaces.

Take \mathbb{R} -valued $Z = f(\xi_1, \xi_2, \dots, \xi_n)$ such that if $\tilde{x} = (x_1, \dots, x_n)$ and $\tilde{y} = (y_1, \dots, y_n)$ differ in one coordinate only (i.e. $|\{i : y_i \neq x_i\}| = 1$), then $|f(\tilde{x}) - f(\tilde{y})| \leq 1$.

Then $P(|Z - \mathbb{E}Z| \geq \lambda \sqrt{n}) \leq 2e^{-\lambda^2/2}$ for $\lambda > 0$.

This is useful for analysis of random algorithms: consider randomized traveling salesman where the tour \tilde{x} is changed at a single location $y_i \neq x_i$.

Proof. WLOG assume $\mathbb{E}Z = 0$. Write $S_m = \mathbb{E}[Z \mid \mathcal{F}_m]$ where $\mathcal{F}_m = \sigma(\xi_i, 1 \leq i \leq m)$, so $(S_m, 1 \leq m \leq n)$ is a MG.

If we can show “ S_m has bounded differences”

$$|S_m - S_{m-1}| \leq 1 \quad (4.1)$$

then Azuma’s inequality (theorem 3.1) yields the desired conclusion.

Lemma 4.2. *If Y is such that any 2 possible values within 1, then $|Y - \mathbb{E}Y| \leq 1$.*

Proof. $\min \text{supp } Y \leq Y \leq \max \text{supp } Y$ and $\min \text{supp } Y \leq \mathbb{E}Y \leq \max \text{supp } Y$ so

$$|Y - \mathbb{E}Y| \leq \max \text{supp } Y - \min \text{supp } Y \leq 1 \quad (4.2)$$

□

If we know all $(\xi_i, i \neq m)$ then apply lemma 4.2 conditionally

$$|Z - \underbrace{\mathbb{E}[Z \mid \xi_i, i \neq m]}_{Z^*}| \leq 1 \quad (4.3)$$

$$(4.4)$$

Lemma 4.3. *If W is independent of (Y, \mathcal{G}) , then $\mathbb{E}[Y \mid \mathcal{G}, W] = \mathbb{E}[Y \mid \mathcal{G}]$.*

Proof. Exercise.

□

By lemma 4.3

$$\mathbb{E}[Z^* \mid \mathcal{F}_m] = \mathbb{E}[Z^* \mid \mathcal{F}_{m-1}, \xi_m] = \mathbb{E}[Z^* \mid \mathcal{F}_{m-1}] \quad (4.5)$$

$$\text{tower property} \implies = \mathbb{E}[Z \mid \mathcal{F}_{m-1}] \quad (4.6)$$

$$(4.7)$$

This is nice: Z and Z^* are both conditioned on the same σ -field \mathcal{F}_m so

$$|S_m - S_{m-1}| = |\mathbb{E}[Z \mid \mathcal{F}_m] - \mathbb{E}[Z \mid \mathcal{F}_{m-1}]| \quad (4.8)$$

$$= |\mathbb{E}[Z \mid \mathcal{F}_m] - \mathbb{E}[Z^* \mid \mathcal{F}_m]| \quad (4.9)$$

$$\leq \mathbb{E}[|Z - Z^*| \mid \mathcal{F}_m] \quad (4.10)$$

$$\text{Equation (4.3)} \implies \leq 1 \quad (4.11)$$

□