

1 Last class

(X_n) sub-MG wrt (\mathcal{F}_n) , i.e. $\mathbb{E}[X_n | \mathcal{F}_{n-1}] \geq X_{n-1}$.

(H_n) predictable process, bounded. Interpret $H_n = \#$ shares held on day n .

Define $Y = H \cdot X$ by $Y_0 = 0, \Delta_n^Y = H_n \Delta_n^X$

Then (Y_n) is a sub-MG provided $H_n \geq 0$

Remark 1.1. If $\mathbb{E}[Z1_A] \geq 0$ for all $A \in \mathcal{G}$, then $\mathbb{E}[Z | \mathcal{G}] \geq 0$ a.s.

2 Today

Corollary 2.1. (X_n) a sub-MG. $0 \leq T_1 \leq T_2 \leq t_0$ stopping times. Then $\mathbb{E}[X_{T_2} | \mathcal{F}_{T_1}] \geq X_{T_1}$

Proof. Fix event $A \in \mathcal{F}_{T_1}$. Strategy:

If A happens, buy 1 share at T_1 , sell at T_2 .

If A doesn't happen, do nothing.

In math: $H_n = 1_A 1_{T_1 < n \leq T_2}$.

Want to check H is predictable, that is

$$A \cap \{T_1 < n \leq T_2\} \in \mathcal{F}_{n-1} \quad (2.1)$$

$$= A \cap \underbrace{\{T_1 \leq n-1\}}_{\substack{\in \mathcal{F}_{n-1} \\ \text{def of } A \in \mathcal{F}_{T_1}}} \setminus \underbrace{A \cap \{T_2 \leq n-1\}}_{\in \mathcal{F}_{n-1}} \text{ because } T_2 \geq T_1 \quad (2.2)$$

So H is predictable.

So (Y_n) is a sub-MG, $Y_n = (X_{T_2 \wedge n} - X_{T_1 \wedge n})1_A$

$$\implies \mathbb{E}Y_{t_0} \geq \mathbb{E}Y_0 = 0 \quad (2.3)$$

$$\implies \mathbb{E}[(X_{T_2} - X_{T_1})1_A] \geq 0 \quad \forall A \in \mathcal{F}_{T_1} \quad (2.4)$$

$$\implies \mathbb{E}[X_{T_2} - X_{T_1} | \mathcal{F}_{T_1}] \geq 0 \text{ a.s.} \quad ?? \quad (2.5)$$

□

3 Optional Sampling Theorem (OST)

Theorem 3.1 (Basic version). If (X_n) is a (sub-)MG,

$0 = T_0 \leq T_1 \leq T_2 \leq \dots$ are stopping times,

$T_i \leq t_i$ (constant) for every i ,

then $(X_{T_i}, i = 0, 1, 2, \dots)$ is a (sub-)MG wrt $(\mathcal{F}_{T_i}, i = 0, 1, 2, \dots)$.

In particular, $\mathbb{E}X_T \geq \mathbb{E}X_0$ for sub-MG, $= \mathbb{E}X_0$ for MG.

Many other versions without restriction $T \leq t_0$ exist.

4 Maximal inequalities

$$X_N^* = \max(X_0, X_1, \dots, X_N)$$

always $P(X_N^* \geq x) \leq \sum_{n=0}^N P(X_n \geq x)$

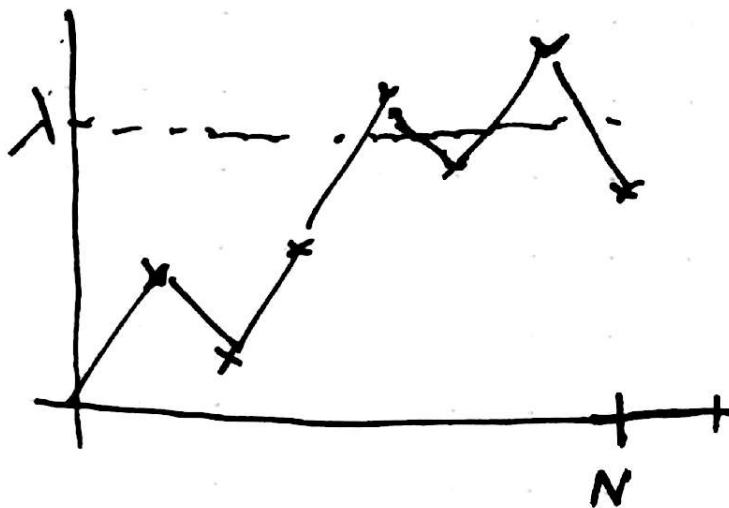
MG get better than above

if independent $P(X_N^* \geq x) = 1 - \prod_{n=0}^N P(X_n \leq x)$

Lemma 4.1. (X_n) a super-MG, $X_n \geq 0$ a.s.

Write $X^* = \sup_n X_n$, so $X_N^* \uparrow X^*$ as $N \rightarrow \infty$.

Then $P(X^* \geq \lambda) \leq \frac{\mathbb{E}X_0}{\lambda}$, all $\lambda > 0$



Proof. Define $T = \min\{n : X_n \geq \lambda\}$. Apply OST to 0 and $T \wedge N$.

$$\implies \mathbb{E}X_0 \geq \mathbb{E}X_{T \wedge N} = \mathbb{E}X_T 1_{T \leq N} + \mathbb{E}X_n 1_{T > N} \quad (4.1)$$

$$\geq \lambda P(T \leq N) + 0 \quad (4.2)$$

$$\implies P(T \leq N) \leq \lambda^{-1} \mathbb{E}X_0 \quad (4.3)$$

$$\implies P(X_N^* \geq \lambda) \leq \lambda^{-1} \mathbb{E}X_0 \quad (4.4)$$

$$N \rightarrow \infty \implies P(X^* > \lambda) \leq \lambda^{-1} \mathbb{E}X_0 \quad (4.5)$$

$$\text{Apply to } \lambda_j \uparrow \lambda \text{ (check)} \implies P(X^* \geq \lambda) \leq \lambda^{-1} \mathbb{E}X_0 \quad (4.6)$$

□

Lemma 4.2 (Doob's L_1 maximal inequality). (X_n) sub-MG. For $\lambda > 0$

$$\lambda P(X_N^* \geq \lambda) \leq \mathbb{E}[X_N 1_{X_N^* \geq \lambda}] \leq \mathbb{E}X_N^+ = \mathbb{E} \max(X, 0) \quad (4.7)$$

Proof. $T = \min\{n : X_n \geq \lambda\}$.

Apply OST to $T \wedge N$ and $N \implies \mathbb{E}X_{T \wedge N} \leq \mathbb{E}X_N$.

$$\implies \mathbb{E}X_T 1_{T \leq N} + \mathbb{E}X_n 1_{T > N} \leq \mathbb{E}X_N 1_{TN} + \mathbb{E}X_N 1_{T > N}$$

$$X_T \geq \lambda \implies \lambda P(T \leq N) \leq \mathbb{E}X_N 1_{T \leq N} = \mathbb{E}X_n 1_{X_N^* \geq \lambda}$$

□

Corollary 4.3. If (X_n) is a MG then (because $Y_n = |X_n|$ is a sub-MG)

$$\lambda P(\max_{0 \leq n \leq N} |X_n| \geq \lambda) \leq \mathbb{E}|X_n|/\lambda \quad (4.8)$$

Also, $Z_n = X_n^2$ is a sub-MG (provided $\mathbb{E}X_n^2 < \infty$).

Apply Lemma to (Z_n)

$$\lambda P(\max_{0 \leq n \leq N} X_n^2 \geq \lambda) \leq \mathbb{E}X_N^2 \quad (4.9)$$

$$\lambda^2 P(\max_{0 \leq n \leq N} X_n^2 \geq \lambda^2) \leq \mathbb{E}X_N^2 \quad (4.10)$$

$$P(\max_{0 \leq n \leq N} |X_n| \geq \lambda) \leq \lambda^{-2} \mathbb{E}X_N^2 \quad (4.11)$$

Different bounds for same quantity (c.f. Markov/Chebyshev)

Lemma 4.4 (Doob's L_2 maximal inequality). (X_n) sub-MG.

$$\mathbb{E}(0 \vee X_N^*)^2 \leq 4\mathbb{E}(X_N^+)^2 \quad (4.12)$$

Proof.

$$\underbrace{\mathbb{E}(0 \vee Z)^2}_a = 2 \int_0^\infty \lambda P(Z \geq \lambda) d\lambda \quad (4.13)$$

$$\implies \mathbb{E}(0 \vee X_N^*)^2 = 2 \int_0^\infty \lambda P(X_N^* \geq \lambda) d\lambda \quad (4.14)$$

$$\leq 2 \int_0^\infty \mathbb{E}[X_N 1_{X_N^* \geq \lambda}] d\lambda \quad (4.15)$$

$$\leq 2 \int_0^\infty \mathbb{E}[X_N^+ 1_{X_N^* \geq \lambda}] d\lambda \quad (4.16)$$

$$= 2 \mathbb{E} \left[X_N^+ \int_0^\infty 1_{X_N^* \geq \lambda} d\lambda \right] \quad (4.17)$$

$$= 2 \mathbb{E}[X_N^+ (0 \vee X_N^*)] \quad (4.18)$$

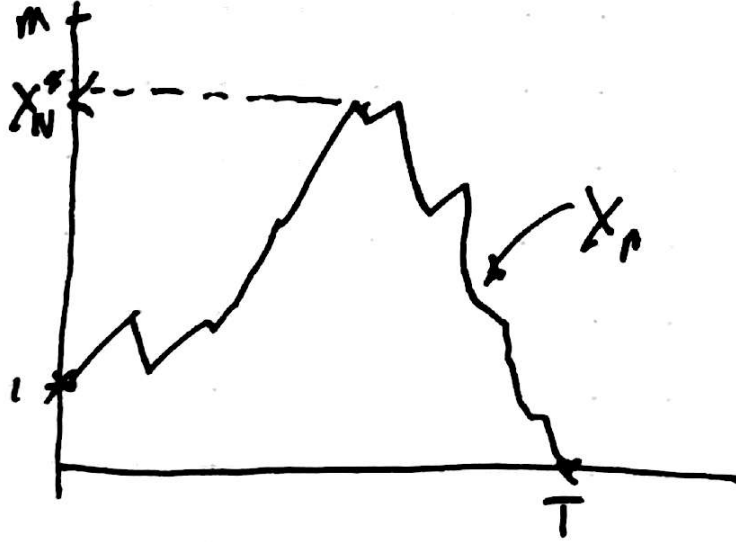
$$\text{Cauchy-Schwarz} \leq 2 \sqrt{\underbrace{\mathbb{E}[(X_N^+)^2]}_b \times \underbrace{\mathbb{E}[(0 \vee X_N^*)^2]}_a} \quad (4.19)$$

$$a \leq 2\sqrt{ba} \implies a \leq 4b. \quad \square$$

If we use Hölder instead of Cauchy-Schwarz

$$\mathbb{E}[(0 \vee X_N^*)^p] \leq \left(\frac{p}{p-1} \right)^p \mathbb{E}[(X_N^+)^p] \quad 1 < p < \infty \quad (4.20)$$

Example 4.5 (not true for $p = 1$). $X_0 = 1$, simple symmetric random walk on \mathbb{Z} , stop at $T = \min\{n \geq 1 : X_n = 0\}$.



(X_n) is a MG. $\mathbb{E}X_n = 1 \forall n$.

$X_N^* \uparrow X^* = \sup_n X_n$. Elementary $P(X^* \geq m) = m^{-1}$.

$\implies \mathbb{E}X^+ = \infty \implies \mathbb{E}X_N^* \uparrow \infty$.

$\therefore \mathbb{E}X_n = 1 \forall n$, so cannot bound the ratio for $p = 1$.