

# 1 Polynomial Approximation Theorems

Main goal is to demonstrate proof techniques, results are not too important.

**Theorem 1.1** (Bernstein). *Given continuous  $f : [0, 1] \rightarrow \mathbb{R}$ , define*

$$f_n(x) := \sum_{m=0}^n \binom{n}{m} x^m (1-x)^{n-m} f(m/n), \quad 0 \leq x \leq 1 \quad (1.1)$$

(polynomial of degree  $n$ ).

*Then  $f_n$  converge uniformly to  $f$  i.e.*

$$\sup_x |f_n(x) - f(x)| \rightarrow 0 \quad (1.2)$$

as  $n \rightarrow \infty$ .

*Proof.* Fix  $x$ . Take i.i.d. Bernoulli( $x$ ) r.v.s  $(X_i, 1 \leq i < \infty)$ . Let  $S_n := \sum_{i=1}^n X_i \sim \text{Binomial}(n, x)$  and note  $f_n(x) = \mathbb{E}f(S_n/n)$ .

Want to bound

$$|f_n(x) - f(x)| = |\mathbb{E}f(S_n/n) - f(x)| \quad (1.3)$$

$$\leq \mathbb{E}|f(S_n/n) - f(x)| \quad (1.4)$$

Want to apply WLLN  $S_n/n \xrightarrow{P} x$ , so split

$$\mathbb{E}|f(S_n/n) - f(x)| = \mathbb{E} \left[ |f(S_n/n) - f(x)| 1_{|S_n/n - x| \leq \delta} + |f(S_n/n) - f(x)| 1_{|S_n/n - x| > \delta} \right] \quad (1.5)$$

From analysis, we have

- $M := \sup |f(X)| \leq \infty$
- (Uniformly Continuous)  $\forall \varepsilon > 0, \exists \delta > 0 : |y_1 - y_2| \leq \delta \implies |f(y_1) - f(y_2)| \leq \varepsilon$

Choosing  $\varepsilon > 0$  and taking  $\delta$  from uniform continuity

$$\mathbb{E} \left[ |f(S_n/n) - f(x)| 1_{|S_n/n - x| \leq \delta} + |f(S_n/n) - f(x)| 1_{|S_n/n - x| > \delta} \right] \quad (1.6)$$

$$\leq \varepsilon + 2MP(|S_n/n - x| > \delta) \quad (1.7)$$

$$\leq \varepsilon \frac{2M}{\delta^2} \text{Var}(S_n/n) \quad (1.8)$$

$$= \varepsilon + \frac{2M}{\delta^2} \frac{x(-x)}{n} \quad (1.9)$$

$$\leq \varepsilon + \frac{M}{2\delta^2} \frac{1}{n} \quad (1.10)$$

$$\leq \varepsilon \quad \text{as } n \rightarrow \infty \quad (1.11)$$

$$= 0 \quad \text{holds } \forall \varepsilon > 0 \quad (1.12)$$

□

## 2 Background for proving a.s. limits

### 2.1 Axioms

For events  $B_n, B$

- $B_n \uparrow B \implies P(B_n) \uparrow P(B)$
- $B_n \downarrow B$  and  $\exists n : P(B_n) < \infty \implies P(B_n) \downarrow P(B)$

**Definition 2.1.** Event  $A_n$  *infinitely often* ( $A_n$  i.o.) means  $\bigcap_{m=1}^{\infty} \bigcup_{n=m}^{\infty} A_n$ .

Event  $A_n$  *ultimately* ( $A_n$  ult.) means  $\bigcup_{m=1}^{\infty} \bigcap_{n=m}^{\infty} A_n$ .

**Proposition 2.2.** *i.o. and ult. are opposites:*  $(A_n \text{ i.o.})^c = (A_n^c \text{ ult.})$

**Lemma 2.3** (Weak).

$$(i) \quad P(A_n \text{ i.o.}) \geq \limsup_n P(A_n)$$

$$(ii) \quad P(A_n \text{ ult.}) \geq \liminf_n P(A_n)$$

*Proof.*

$$P(\bigcup_{n=m}^Q A_n) \geq \max_{m \leq n \leq Q} P(A_n) \quad (2.1)$$

$$P(\bigcup_{n=m}^{\infty} A_n) \geq \sup_{n \geq m} P(A_n) \quad Q \rightarrow \infty \quad (2.2)$$

$$P(A_n \text{ i.o.}) \geq \limsup_n P(A_n) \quad m \rightarrow \infty \quad (2.3)$$

□

**Lemma 2.4** (First Borel-Cantelli). *Events  $(A_n)_{n=1}^\infty$ , if  $\sum_n P(A_n) < \infty$  then  $P(A_n \text{ i.o.}) = 0$ .*

*Proof.*

$$X_n = \sum_{i=1}^n 1_{A_i} = (\text{number of last } n \text{ events that occur}) \quad (2.4)$$

$$X_\infty = \sum_{i=1}^\infty 1_{A_i} = (\text{total number of events that occur}) \quad (2.5)$$

$$\mathbb{E} X_\infty = \sum_{i=1}^\infty P(A_i) \underset{\text{hyp}}{<} \infty \quad (2.6)$$

$$\implies P(X_\infty = \infty) = 0 \quad (2.7)$$

□

**Lemma 2.5** (Second Borel-Cantelli). *For independent events  $(A_i)_{i=1}^\infty$ , if  $\sum_i P(A_i) = \infty$  then  $P(A_n \text{ i.o.}) = 1$ .*

(Many variants under alternative assumptions exist)

*Proof.* Fix  $m$ . We will prove  $P(\cup_{n=m}^\infty A_n) = 1$  or prove  $P(\cap_{n=m}^\infty A_n^c) = 0$ .

Fact:  $0 \leq x \leq 1 \implies 1 - x \leq e^{-x}$ .

By independence

$$P(\cap_{n=m}^q A_n^c) = \prod_{n=m}^q P(A_n^c) \quad (2.8)$$

$$= \prod_{n=m}^q (1 - P(A_n)) \quad (2.9)$$

$$\leq \exp(-\sum_{n=m}^q P(A_n)) \quad (2.10)$$

Let  $q \uparrow \infty$

$$P(\cap_{n=m}^\infty A_n^c) \leq \exp(-\sum_{n=m}^\infty P(A_n)) = 0 \quad (2.11)$$

□

**Lemma 2.6.** *Arbitrary  $\mathbb{R}$ -valued r.v.s  $(Y_n)$  and arbitrary  $-\infty < y < \infty$ .*

*If  $\sum_n P(Y_n \geq y + \varepsilon) < \infty$  for all  $\varepsilon > 0$ , then  $\limsup_n Y_n \leq y$  a.s..*

**Corollary 2.7.** *If  $\sum_n P(|Y_n| \geq \varepsilon) < \infty$ , each  $\varepsilon > 0$ , then  $Y_n \xrightarrow{\text{a.s.}} 0$ .*

**Proposition 2.8** (Deterministic fact for real numbers). *For reals  $(y_n)$  and  $y$ , “ $\limsup_n y_n \leq y$ ” equivalent to “ $y_n \leq y + \varepsilon$ ” ult., each  $\varepsilon > 0$ , equivalent to “ $y_n \leq y + 1/j$ ” ult., each  $j \geq 1$ .*

*Proof.* By Hyp and B-C 1

$$\implies P(Y_n \leq y + 1/j \text{ ult.}) = 1 \quad (2.12)$$

By monotonicity  $P(B_j) = 1 \forall j \implies P(B_j \text{ for all } j) = 1$  hence

$$\implies P(Y_n \leq y + 1/j \text{ ult., each } j \geq 1) = 1 \quad (2.13)$$

$$\implies P(\limsup_n Y_n \leq y) = 1 \quad (2.14)$$

□

**Theorem 2.9** (4th moment SLLN (strong law of large numbers)). *Let  $(X_i, 1 \leq i < \infty)$  be i.i.d.,  $\forall i : \mathbb{E}X_i = 0, \mathbb{E}X^4 < \infty$ . Then*

$$(i) \mathbb{E}S_n^4 \leq 3n^2\mathbb{E}X^4$$

$$(ii) S_n/n \xrightarrow{a.s.} 0 \text{ as } n \rightarrow \infty$$

*Proof.* (i)  $\mathbb{E}S_n^4 = \sum_{i,j,k,l} \mathbb{E}[X_i X_j X_k X_l]$

Note that  $\mathbb{E}[\cdot] = 0$  if some index  $j$  appears only once, e.g.

$$\mathbb{E}[X_4 X_6 X_6 X_6] = \mathbb{E}[X_4] \overset{0}{\mathbb{E}[X_6^3]} = 0 \quad (2.15)$$

Hence

$$\mathbb{E}S_n^4 = n\mathbb{E}X^4 + \binom{4}{2} \binom{n}{2} \mathbb{E}[X_1^2 X_2^2] \quad (2.16)$$

$$= n\mathbb{E}X^4 + 3n(n-1) \underbrace{(\mathbb{E}X^2)^2}_{\leq \mathbb{E}X^4} \quad (2.17)$$

$$\leq 3n^2\mathbb{E}X^4 \quad (2.18)$$

Fix  $\varepsilon > 0$ .

$$P(|S_n/n| \geq \varepsilon) \leq \mathbb{E}|S_n/n|^4 / \varepsilon^4 \quad \text{Markov ineq } \phi(x) = x^4 \quad (2.19)$$

$$\leq \varepsilon^{-4} n^{-4} \times 3n^2 \mathbb{E}X^4 \quad (2.20)$$

$$\leq 3\varepsilon^{-4} \mathbb{E}X^4 n^{-2} \quad (2.21)$$

$$\implies \sum_n P(|S_n/n| \geq \varepsilon) \leq \sum_n 3\varepsilon^{-4} \mathbb{E}X^4 n^{-2} < \infty \quad (2.22)$$

Applying corollary 2.7,  $S_n/n \xrightarrow{a.s.} 0$

□

**Corollary 2.10.** *If  $(A_i)_{i=1}^\infty$  independent Bernoulli( $p$ ),  $S_n = \sum_{i=1}^n 1_{A_i}$ , then  $S_n/n \xrightarrow{a.s.} p$  as  $n \rightarrow \infty$ .*

**Definition 2.11.** Given data real number  $x_1, \dots, x_n$ , define:

**Empirical distribution** Uniform distribution on  $(x_1, x_n)$

**Empirical distribution function**  $G(x) = \frac{1}{n} \sum_{i=1}^n 1_{x_i \leq x}$

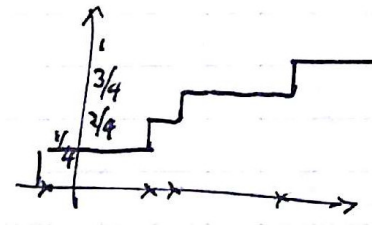


Figure 1: Example empirical distribution function

**Theorem 2.12** (Glivenko-Cantelli).  $(X_i)_{i=1}^\infty$  i.i.d., arbitrary distribution function  $F$ ,  $G_n(\omega, x) = \frac{1}{n} \sum_{i=1}^n 1_{X_i(\omega) \leq x}$  is empirical distribution of  $(X_1(\omega), X_2(\omega), \dots, X_n(\omega))$ .

For fixed  $x$ , events  $\{X_i \leq x\}$  are i.i.d.  $\text{Bern}(f(x))$ .

SLLN for events says

$$G_n(\omega, x) \rightarrow G(x) \text{ a.s. as } n \rightarrow \infty \quad (2.23)$$

for each  $x$ .