

# 1 Review

**Simple vs Simple:** Reject for large  $\frac{p_1(X)}{p_0(X)}$

**One-sided:**  $\Theta \subset \mathbb{R}$ .

$H_0 : \theta \leq \theta_0$  vs  $H_1 : \theta > \theta_0$

If  $\mathcal{P}$  MLR in  $T(X)$ , UMP test rej for large  $T(X)$ .

$= \{P_\theta : \theta \in \Theta \subset \mathbb{R}\}$ .

Test  $\theta \leq \theta_0$  vs  $\theta > \theta_0$ .

$T(X) \in \mathbb{R}$  stochastically increasing in  $\theta$ .

$P_\theta((X) \leq t)$  non-increasing in  $\theta$ .

$\beta(\theta) = P_\theta(T(X) > c)$  non-decreasing in  $\theta$ .

**Example 1.1.**  $X_i \stackrel{\text{iid}}{\sim} p_\theta(x) = p_0(x - \theta), i \leq n$ .

$T(X)$  = sample mean/median.

**Example 1.2.**  $X_i \stackrel{\text{iid}}{\sim} p_\theta(x) = \frac{1}{\theta} p\left(\frac{x}{\theta}\right)$

$T(X) = \sum_i X_i^2, \sum |X_i|$

**Example 1.3.**  $X_i \stackrel{\text{iid}}{\sim} N(\theta, 1). T(X) = \bar{X}$ .

$$\tilde{T}(X) = \mathbb{E}[\tilde{T}(X) \mid \bar{X}] + \text{random noise} \quad (1.1)$$

$$= \bar{X} + \varepsilon \approx N\left(\theta, \frac{1 + \delta}{n}\right) \quad (1.2)$$

# 2 Two-sided tests, unbiasedness

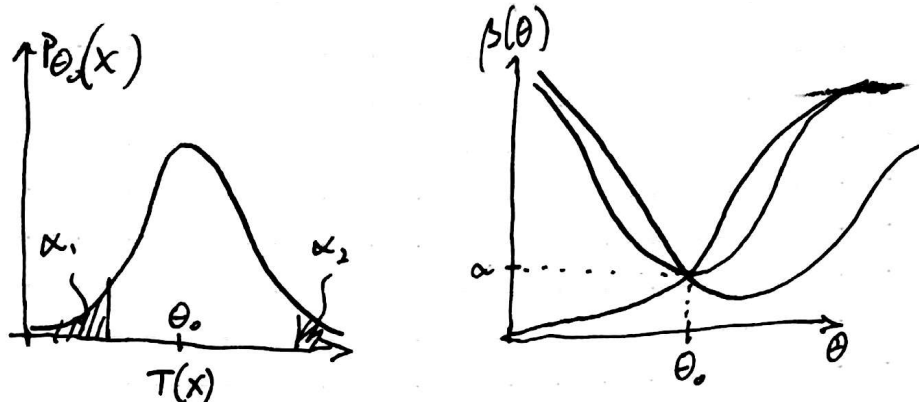
$H_0 : \theta = \theta_0$  vs  $H_1 : \theta \neq \theta_0$

**Example 2.1.**  $X \sim N(\theta, 1), H_0 : \theta = 0$

(a) Reject when  $|X| > z_{n/2}$

(b) Reject when  $X > z_\alpha$

(c) Reject when  $X < z_{1-\alpha_1}$  or  $X > z_{\alpha_2}$  such that  $\alpha_1 + \alpha_2 = \alpha$ .



## 2.1 Equal-tailed test

Reject for extreme  $T(X)$

$$(*)\phi(X) = \begin{cases} 1, & \text{if } T > c_1, T < c_1 \\ 0, & \text{if } T \in (c_1, c_2) \\ \gamma, & \text{if } T = c_i \end{cases} \quad (2.1)$$

Equal-tailed: choose  $c_i, \gamma_i$  so  $\alpha_1 = \alpha_2 = \frac{\alpha}{2}$ .

## 2.2 Unbiased test

A test is *unbiased* if

$$\inf_{\theta \in \Theta_1} \mathbb{E}_{\theta}[\phi(X)] \geq \alpha \quad (2.2)$$

(for exp fam  $\theta_0 \in \Theta^{\circ}, \beta'(\theta_0) = 0$ )

Suppose  $\theta_0 \in \Theta^{\circ}$ . There is a two-sided UMP unbiased level- $\alpha$  test of the form (\*) with  $c_i, \gamma_i$ , chosen to satisfy

$$\mathbb{E}_{\theta_0} \phi(X) = \alpha \quad (2.3)$$

$$\beta'(\theta_0) = \mathbb{E}_{\theta_0}(T(X)(\phi(X) - \alpha)) = 0 \quad (2.4)$$

$$p_0(x) = \begin{cases} \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, & \text{if } x > 0 \\ \frac{1}{2} e^{-|x|}, & \text{if } x < 0 \end{cases} \quad (2.5)$$

## 3 p-Values

**Example 3.1.**  $X \sim N(\theta, 1)$ .  $H_0 : \theta = \theta_0$  vs  $H_1 : \theta \neq \theta_0$ .

$$p(X) = P_{\theta_0}(|X - \theta_0| > |x - \theta_0|) \quad (3.1)$$

$$p(x) = P_{\theta_0}(X > x) \quad (3.2)$$

**Setup:** Testing problem  $\mathcal{P}, H_0, H_1$

Non-random Test  $\phi_\alpha(x)$  for each significance level  $\alpha \in (0, 1)$ .

$$\phi_\alpha(x) = 1_{x \in R_\alpha} \quad (3.3)$$

**Example 3.2.** test  $X \sim N(\theta, 1), H_0 : \theta = 0. R_\alpha = (-\infty, z_{\alpha/2}] \cup [z_{\alpha/2}, \infty)$

Assume rejection regions are *nested*

$$\alpha_1 \leq \alpha_2 \implies R_{\alpha_1} \subset R_{\alpha_2} \quad (3.4)$$

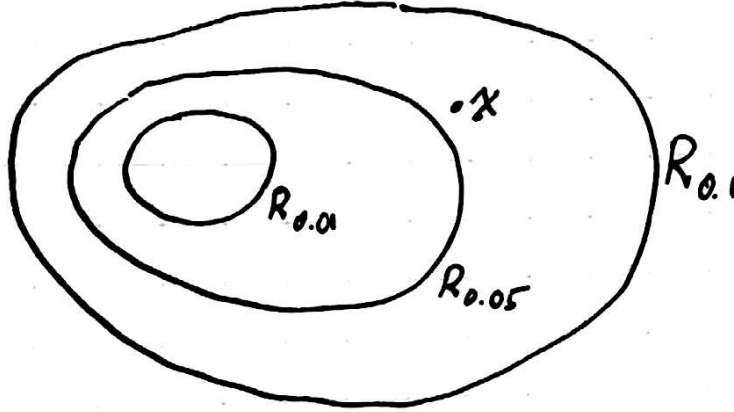
$$(\phi_{\alpha_1}(X) \leq \phi_{\alpha_2}(X)) \quad (3.5)$$

Define *p-value*

$$p(x) = \inf\{\alpha : \phi_\alpha(x) = 1\} \quad (3.6)$$

$$= \inf\{\alpha : x \in R_\alpha\} \quad (3.7)$$

**Picture:**

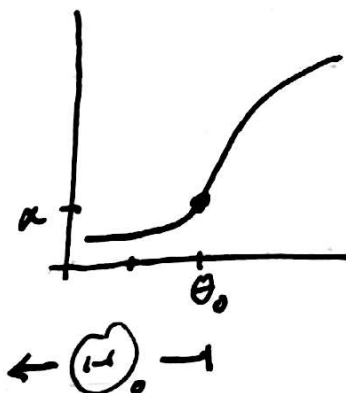


$$\forall \theta \in \Theta_0$$

$$P_\theta(p(X) \leq \alpha) = \sup_{\tilde{\alpha} < \alpha} P_\theta(x \in R_{\tilde{\alpha}}) \leq \sup_{\tilde{\alpha} < \alpha} \tilde{\alpha} = \alpha \quad (3.8)$$

If  $\phi_\alpha(x)$  reject for large  $T(X)$

$$p(x) = \sup_{\theta \in \Theta_0} P_\theta(T(X) > T(x)) \quad (3.9)$$



## 4 Confidence intervals/regions

Accept/reject decision only so informative

**Definition 4.1.** Model  $\mathcal{P}_{1-\alpha} = \{P_\theta : \theta \in \Theta\}$ .

$C(X)$  is a *confidence set* for  $g(\theta)$  if  $P_\theta(g(\theta) \in C(X)) \geq 1 - \alpha$  for all  $\theta$ .