1 Johnson-Lindenstrauss wrap up

Setup:

- (a) $N \ge 2$ points in \mathbb{R}^d
- (b) $\{u^1, u^2, \dots, u^N\}$
- (c) $F: \mathbb{R}^d \to \mathbb{R}^m$

Want to reduce \mathbb{R}^d to \mathbb{R}^m , $m \ll d$, such that

$$(1 - \delta) \le \frac{\|F(u^i) - F(u^j)\|_2^2}{\|u^i - u^j\|_2^2} \le (1 + \delta)$$
(1.1)

When *F* is linear, can represent $F(u^i) = Xu$ where $X \in \mathbb{R}^{m \times d}$.

Consider $X \stackrel{\text{ind.}}{\sim} N(0,1)^{m \times d}$. To scale appropriately, $F(u) = \frac{Xu}{\sqrt{m}}$.

Let $x_i \in \mathbb{R}^d$ be the *i*th row of *X*.

$$\Longrightarrow \left\langle x_i, \frac{u}{\|u\|_2} \right\rangle \sim N(0, 1) \tag{1.2}$$

$$\Longrightarrow Y = \frac{\|Xu\|_2^2}{\|u\|_2^2} = \sum_{i=1}^m \left\langle x_i, \frac{u}{\|u\|_2} \right\rangle^2 \sim \chi_m^2 \tag{1.3}$$

From the concentration bound for χ^2 (see lecture on sub-exponentials), for sufficiently small δ ($\delta \in [0,1)$])

$$\Longrightarrow P\left(\left|\frac{1}{m}\frac{\|Xu\|_2^2}{\|u\|_2^2} - 1\right| \ge \delta\right) \le 2e^{-\frac{m\delta^2}{8}} \tag{1.4}$$

$$\Longrightarrow P\left(\frac{\|F(u)\|_2^2}{\|u\|_2^2} \notin [1-\delta, 1+\delta]\right) \le 2e^{-\frac{m\delta^2}{8}} \tag{1.5}$$

Take a union bound across all $\binom{N}{2}$ pairs of u^i , u^j

$$P\left(\frac{\|F(u^{i}) - F(u^{j})\|_{2}^{2}}{\|u^{i} - u^{j}\|_{2}^{2}} \notin [1 - \delta, 1 + \delta] \text{ for some } (i, j)\right) \le {N \choose 2} 2e^{-\frac{m\delta^{2}}{8}}$$
(1.6)

$$\leq N^2 e^{-\frac{m\delta^2}{8}} \tag{1.7}$$

How to choose *m*?

$$\varepsilon = P\left(\frac{\|F(u^{i}) - F(u^{j})\|_{2}^{2}}{\|u^{i} - u^{j}\|_{2}^{2}} \notin [1 - \delta, 1 + \delta] \text{ for some } (i, j)\right)$$
(1.8)

$$\varepsilon \le N^2 e^{-\frac{m\delta^2}{8}} \tag{1.9}$$

$$\implies m \ge \frac{16}{\delta^2} \log \frac{N}{\sqrt{\varepsilon}} \tag{1.10}$$

Memorize: J-L result means that need to choose $m \sim \log N$ to preserve distances.

2 Martingale concentration inequalities

Definition 2.1. Given a martingale Y_k , a martingale difference sequence ("innovation" in econ/EE) is

$$D_k = Y_k - Y_{k-1} \tag{2.1}$$

Telescoping the sum, we have

$$Y_n - Y_0 = \sum_{k=1}^n D_k (2.2)$$

Notice that

$$\mathbb{E}[D_{k+1} \mid \mathcal{F}_k] = \mathbb{E}[Y_{k+1} \mid \mathcal{F}_k] - \mathbb{E}[Y_k \mid \mathcal{F}_k] = Y_k - Y_k = 0$$
(2.3)

Theorem 2.2 (Wainwright 2.3). *Suppose* $\mathbb{E}[e^{\lambda D_k} \mid \mathcal{F}_{k-1}] \leq e^{\frac{\lambda^2 v_k^2}{2}}$ *for* $|\lambda| < \frac{1}{\alpha_k}$. *Then*

$$P\left(\left|\sum_{k=1}^{n} D_{k}\right| \ge t\right) \le \begin{cases} 2e^{-\frac{t^{2}}{2\sum_{k=1}^{n} v_{k}^{2}}}, & if \ 0 \le t \le \frac{\sum_{k=1}^{n} v_{k}^{2}}{\alpha^{*}} \\ 2e^{-\frac{t}{2\alpha^{*}}}, & otherwise \end{cases}$$
(2.4)

Exercise 2.3. Prove this

Similar to Hoeffding bound, for martingales we have the following:

Theorem 2.4 (Azuma-Hoeffding). *Suppose* $D_k \in (a_k, b_k)$. *Then*

$$P\left(\left|\sum_{k=1}^{n}\right| \ge t\right) \le 2e^{-\frac{2t^2}{\sum_{k=1}^{n}(b_k - a_k)^2}} \tag{2.5}$$

This is particularly useful for the *Doob martingale*

$$f(X) = f(X_1, ..., X_n), \{X_i\} \text{ ind}$$
 (2.6)

$$Y_k = \mathbb{E}[f(X) \mid X_1, \dots, X_k] \tag{2.7}$$

 Y_k is a martingale by the tower property.

Example 2.5. Random graph, for $1 \le i < j \le n$ each $X_{ij} \sim \text{Bernoulli}(p)$ represents if edge exists. Let f(X) be the number of cliques in the graph represented by adjacency matrix X. Then Y_k is an *edge-revealing* martingale, and $\mathbb{E}[f(X)]$ is the expected number of cliques in G(n, p).

Lemma 2.6 (Bounded differences). Let $x, x' \in \mathbb{R}^n$. Define $x_j^{\setminus k} = \begin{cases} x_j, & \text{if } j \neq k \\ x_k', & j = k \end{cases}$.

Suppose $|f(x) - f(x^{\setminus k})| \leq L_k$ (Lipschitz property in Hamming distance). Then

$$P(|f(x) - \mathbb{E}f(x)| \ge t) \le 2e^{-\frac{2t^2}{\sum_{k=1}^n L_k^2}}$$
(2.8)

Example 2.7 (*U*-statistics, r = 2).

$$U = \frac{1}{\binom{n}{2}} \sum_{j < k} g(X_i, X_j), \quad \|g\|_{\infty} \le b$$
 (2.9)

$$U(x) - U(x^{\setminus k}) \le \frac{1}{\binom{n}{2}} \sum_{j:j \neq k} |g(x_j, x_k) - g(x_j, x_k')|$$
 (2.10)

$$\leq \frac{(n-1)(2b)}{\binom{n}{2}} = \frac{4b}{n} \tag{2.11}$$

$$\implies P(|U - \mathbb{E}U| \ge t) \le 2e^{-\frac{nt^2}{8b^2}} \tag{2.12}$$

Can invert this to get a confidence interval.

Example 2.8 (Clique number in Erdös-Rényi graphs). G = (V, E), |V| = n.

C(G) = cardinality of largest clique ("clique number").

 $(i,j) \in E$ with probability p.

Let $X_{ij} = 1_{(i,j) \in E}$, so $C(G) = f(\{X_{ij}\})$.

 $\mathcal{C}(G)$ satisfies a bounded difference; if G' is a graph with one edge added/removed, then $|\mathcal{C}(G) - \mathcal{C}(G')| \leq 1$ so

$$P(|\mathcal{C}(G) - \mathcal{C}(G')| \ge \delta) \le 2e^{-\delta^2}$$
(2.13)

3 Lipschitz functions of Gaussians

Definition 3.1. f is *Lipschitz* if \exists Lipschitz constant $L \ge 0$ such that

$$|f(x) - f(y)| \le L||x - y||_2, \quad \forall (x, y) \in \mathbb{R}^n$$
 (3.1)

Theorem 3.2 (Wainwright 2.4). Let $X_1, \ldots, X_n \stackrel{i.i.d.}{\sim} N(0,1)$. Let $f \sim Lipschitz(L)$. Then

$$f(X) - \mathbb{E}f(X) \sim SG(\sigma), \quad \sigma \le L$$
 (3.2)

$$\implies P(|f(X) - \mathbb{E}f(X)| \ge t) \le 2e^{-\frac{t^2}{2L^2}} \tag{3.3}$$

Hence, any Lipschitz function concentrates in a way that depends on the Lipschitz constant.

TODO: Read this proof before next class, there's a Lemma introducing a trick TODO: Read Vershynin about geometric functional analysis