

## 1 Independence (undergrad)

Events  $A$  and  $B$  are *independent* if  $P(A \cap B) = P(A)P(B)$ .

R.V.s  $X, Y$  independent if  $P(X \leq x, Y \leq y) = P(X \leq x)P(Y \leq y)$ .

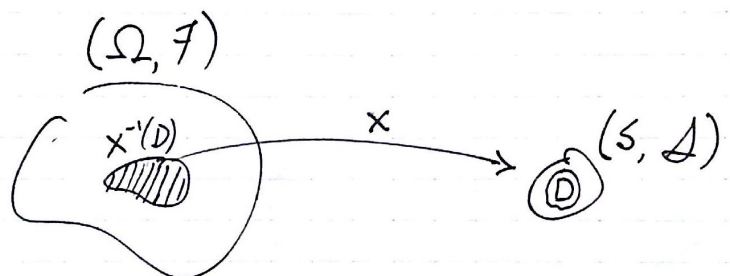
**Idea:** knowing the value of  $X$  doesn't change the probabilities for  $Y$

## 2 Independence (MT version)

$(\Omega, \mathcal{F}, P)$  a probability space. Consider  $\mathcal{B}_1, \mathcal{B}_2$  sub- $\sigma$ -fields of  $\mathcal{F}$ .

**Definition 2.1.** Call  $\sigma$ -fields  $\mathcal{B}_1$  and  $\mathcal{B}_2$  *independent* if

$$P(B_1 \cap B_2) = P(B_1)P(B_2) \quad \forall B_i \in \mathcal{B}_i \quad (2.1)$$



$X$  is measurable  $\implies X^{-1}(D) \in \mathcal{F}$  for all  $D \in \mathcal{S}$ . The collection  $\{X^{-1}(D) : D \in \mathcal{S}\}$  is a sub- $\sigma$ -field of  $\mathcal{F}$ , call it  $\sigma(X)$  "the  $\sigma$ -field generated by  $X$ ."

**Definition 2.2.** Call RVs  $X_1$  and  $X_2$  independent if  $\sigma(X_1)$  and  $\sigma(X_2)$  are independent.

**Theorem 2.3.** For RVs  $X_1, X_2$  with  $X_i$  taking values in  $(S_i, \mathcal{S}_i)$ , the following are equivalent:

- (i)  $X_1$  and  $X_2$  are independent
- (ii)  $P(X_1 \in B_1, X_2 \in B_2) = P(X_1 \in B_1)P(X_2 \in B_2)$  for all  $B_i \in \mathcal{S}_i$
- (iii)  $P(X_1 \in B_1, X_2 \in B_2) = P(X_1 \in B_1)P(X_2 \in B_2)$  for all  $B_i \in \mathcal{A}_i$  where  $\mathcal{A}_i$  is a  $\pi$ -class,  $\sigma(\mathcal{A}_i) = \mathcal{S}_i$
- (iv)  $\mathbb{E}[h_1(X_1)h_2(X_2)] = (\mathbb{E}h_1(X_1))(\mathbb{E}h_2(X_2))$  for all bounded measurable  $h_i : \mathcal{S}_i \rightarrow \mathbb{R}$

**Comments:**

- (iv) extends to integrable  $h_i(X_i)$
- If  $X_i$  are  $\mathbb{R}$ -valued, (iii) can be used to show independence is equivalent to  $P(X_1 \leq x_1, X_2 \leq x_2) = P(X_1 \leq x_1)P(X_2 \leq x_2)$  for all  $x_i \in \mathbb{R}$
- The fact:

if  $X_1, X_2$  are independent, then  $g_1(X_1), g_2(X_2)$  independent (arbitrary measurable  $g_i$ )

is true because  $\sigma(g(X)) \leq \sigma(X)$ .

*Proof outline.* (iv)  $\implies$  (ii)  $\implies$  (iii) special cases

(ii)  $\implies$  (iv) by “monotone class argument”:

(iv) holds for  $h_i = 1_{B_i}$  indicator functions

$\therefore$  holds for  $h_i$  simple functions

$\therefore$  holds for  $h_i$  bounded measurable functions

Want to use Dynkin’s  $\pi - \lambda$  Lemma:

**Step 1** Fix  $B_2 \in \mathcal{A}_2$ . Consider the collection

$$\mathcal{L} = \{A \in \mathcal{S}_1 : P(X_1 \in A, X_2 \in B_2) = P(X_1 \in A)P(X_2 \in B_2)\} \quad (2.2)$$

Check  $\mathcal{L}$  is a  $\lambda$ -class.

By hypothesis,  $\mathcal{L} \supset \mathcal{A}_1$ . Dynkin’s lemma implies  $\mathcal{L} = \mathcal{S}_1$

**Step 2** Consider

$$\mathcal{L}' = \{B_2 \in \mathcal{S}_2 : P(X_1 \in B_2, X_2 \in B_2) = P(X_1 \in B_1, X_2 \in B_2) \quad \forall B_1 \in \mathcal{S}_1\} \quad (2.3)$$

Check  $\mathcal{L}'$  is a  $\lambda$ -class (use linearity property).

By step 1,  $\mathcal{L}' \supset \mathcal{A}_2$ .

By Dynkin’s,  $\mathcal{L}' \supset \sigma(\mathcal{A}_2) = \mathcal{S}_2 \implies$  (ii)

□

**Definition 2.4.**  $\mathcal{B}_1, \mathcal{B}_2, \dots, \mathcal{B}_n$  are (mutually) independent means

$$P(\cap_{i=1}^n B_i) = \prod_{i=1}^n P(B_i) \quad \forall B_i \in \mathcal{B}_i \quad (2.4)$$

This is *stronger* than pairwise independence.

**Example 2.5.**  $X, Y$  for die throws, events  $\{X = 3\}, \{Y = 6\}, \{X = Y\}$ . These events are only pairwise independent, not mutually independent.

**Example 2.6.**  $X_1, X_2$  independent uniform on  $\{0, 1, \dots, n-1\}$ . Define  $X_3 = X_1 + X_2$  modulo  $n$ . Then  $(X_1, X_2, X_3)$  are pairwise independent, not mutually independent.

**Claim.** If  $X_1, \dots, X_k, \dots, X_n$  independent, then  $f(X_1, \dots, X_k)$  and  $g(X_{k+1}, \dots, X_n)$  are independent for arbitrary measurable functions  $f$  and  $g$ .

**Exercise 2.7.** Formalize and verify. “Hereditary property of independence.”

**Exercise 2.8.** To show that events  $\{A_i\}_{i=1}^n$  are independent, suffices to show

$$P(\cap_{i \in \mathcal{I}} A_i) = \prod_{i \in \mathcal{I}} P(A_i) \quad \forall \mathcal{I} \subset \{1, 2, \dots, n\} \quad (2.5)$$

### 3 Real-valued Random Variables

Let  $X_i, Y_i$  be real-valued random variables.

Know that  $X_n \rightarrow X$  a.s. means  $P(\{\omega : X_n(\omega) \rightarrow X(\omega) \text{ a.s. } n \rightarrow \infty\}) = 1$ .

**Definition 3.1** (Convergence in probability).  $X_n \rightarrow_p X$  means

$$\lim_{n \rightarrow \infty} P(|X_n - X| > \epsilon) = 0, \quad \forall \epsilon > 0 \quad (3.1)$$

**Definition 3.2** (Convergence in  $L^p$ ). For  $1 \leq p < \infty$ , say  $X_n \rightarrow X$  in  $L^p$  or  $X_n \xrightarrow{L^p} X$  to mean

$$\mathbb{E}|X_n - X|^p = \|X_n - X\|_p^p \rightarrow 0 \text{ as } n \rightarrow \infty \quad (3.2)$$

(and  $\mathbb{E}|X_n|^p < \infty$  for all  $n$ ).

**Lemma 3.3.** If  $X_n \rightarrow X$  in  $L^p$ , then  $X_n \rightarrow_p X$ .

*Proof.* Use general form of Markov’s inequality  $\phi(x) = |x|^p$  applied to  $X_n - X$ .

$$\implies P(\|X_n - X\| > \epsilon) \leq \frac{\mathbb{E}|X_n - X|^p}{\epsilon^p} \rightarrow 0 \text{ as } n \rightarrow \infty \quad (3.3)$$

□

**Definition 3.4** (Variance). If  $\mathbb{E}X^2 < \infty$ , define the *variance*

$$\text{Var}(X) = \mathbb{E}X^2 - (\mathbb{E}X)^2 = \mathbb{E}(X - \mathbb{E}X)^2 \quad (3.4)$$

**Proposition 3.5.** If  $(X_i)_{i=1}^n$  independent, then  $\text{Var}(\sum_i X_i) = \sum_i \text{Var}(X_i)$

**Definition 3.6.** If  $\mathbb{E}X_i^2 \leq \infty$ ,  $\mathbb{E}X_1 X_2 = (\mathbb{E}X_1)(\mathbb{E}X_2)$ , say  $X_1$  and  $X_2$  are *uncorrelated*

Independence  $\implies$  pairwise independent  $\implies$  uncorrelated

**Theorem 3.7** ( $L^2$  weak law of large numbers). Given  $X_i, i \geq 1$ , suppose  $\sup \mathbb{E}X_i^2 \leq c$  for some constant  $c$  and suppose uncorrelated. Write  $\mu_i = \mathbb{E}X_i$ ,  $S_n = \sum_{i=1}^n X_i$ ,  $\bar{\mu}_n = \frac{1}{n} \sum_{i=1}^n \mu_i$ .

Then  $\frac{S_n}{n} - \bar{\mu}_n \rightarrow 0$  in  $L^2$  as  $n \rightarrow \infty$ .

*Proof.*

$$\frac{1}{n}\mathbb{E}S_n = \bar{\mu}_n \tag{3.5}$$

$$\text{Var}(S_n) = \sum_{i=1}^n \text{Var}(X_i) \leq cn \tag{3.6}$$

$$\text{Var}\left(\frac{1}{n}S_n\right) \leq \frac{c}{n} \tag{3.7}$$

$$\mathbb{E}\left(\frac{S_n}{n} - \bar{\mu}_n\right)^2 = \text{Var}\left(\frac{S_n}{n}\right) \leq \frac{c}{n} \rightarrow 0 \text{ as } n \rightarrow \infty \tag{3.8}$$

□