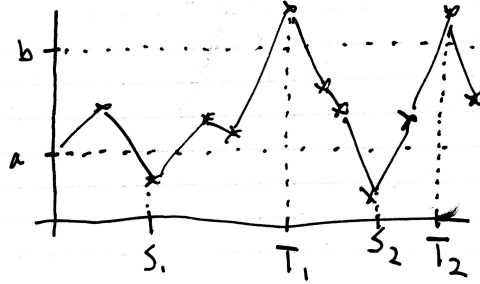


1 Upcrossing Inequality

Take any \mathbb{R} -valued $(X_n, n \geq 0)$ and any $a < b \in \mathbb{R}$.

Let $S_1 = \min\{n : X_n \leq a\}$, $T_1 = \min\{n : X_n \geq b\}$, $S_2 = \min\{n > T_1 : X_n \leq a\}$, $T_2 = \min\{n > S_2 : X_n \geq b\}, \dots$

Definition 1.1. Define $U_n = U_n[a, b] = \max\{k : T_k \leq n\}$ to be the *number of upcrossings* over $[a, b]$ completed by time n .



Theorem 1.2 (Upcrossing Inequality). Suppose (X_n) is a sub-MG. Then

$$(b - a)\mathbb{E}U_n \leq \mathbb{E}(X_n - a)^+ - \mathbb{E}(X_0 - a)^+ \quad (1.1)$$

$$\leq \mathbb{E}X_n^+ + |a| \quad (1.2)$$

Proof. For the second inequality, note $(x - a)^+ \leq x^+ + |a|$ so $\mathbb{E}(X - a)^+ \leq \mathbb{E}X^+ + |a|$.

(Trick) When $X_n \geq a \forall n$, we will prove

$$(b - a)\mathbb{E}U_n \leq \mathbb{E}X_n^+ - \mathbb{E}X_0^+ \quad (1.3)$$

For general MG (X_n) , apply the result to $\max(X_n, a) - a$, which is a sub-MG.

Use the “buy low, sell high” strategy: buy 1 share at S_i (low) and sell at T_i (high). Consider $Y = H \cdot X$, where $H_n = \sum_i 1_{S_i < n \leq T_i}$. (H_n) is predictable and $H_n \geq 0$, so (by Durrett 2.7) (Y_n) is a sub-MG.

$$Y_n = \sum_{i=1}^{U_n} \underbrace{(X_{T_i} - X_{S_i})}_{\text{profit}} + \underbrace{(X_n - X_{S_{U_n+1}})1_{n > S_{U_n+1}}}_{\substack{\text{value of stock if buy at } S_{U_n+1} \\ \text{and sell at time } n < T_{U_n+1}}} \quad (1.4)$$

$$\geq (b - a)U_n + 0 \quad (1.5)$$

$$\mathbb{E}Y_n \geq (b - a)\mathbb{E}U_n \quad (1.6)$$

Consider the opposite strategy K : $K_n = 1 - H_n$ ("Buy high, sell low"). (K_n) is also predictable, $K_n \geq 0$, so

$$(X_n - Y_n) = X_n - (H \cdot X)_n = \sum_{i=1}^n \Delta_i^X + X_0 - \sum_{i=1}^n H_i \Delta_i^X = \sum_{i=1}^n \underbrace{(1 - H_i)}_{=K_i} \Delta_i^X + X_0 = (K \cdot X)_n + X_0 \quad (1.7)$$

is a sub-martingale and

$$\mathbb{E}[X_0 - Y_0] \overset{0}{\leq} \mathbb{E}[X_n - Y_n] \quad (1.8)$$

$$\mathbb{E}X_0 \leq \mathbb{E}X_n - \mathbb{E}Y_n \quad (1.9)$$

$$(b - a)\mathbb{E}U_n \leq \mathbb{E}X_n - \mathbb{E}X_0 \quad (1.10)$$

□

2 Martingale convergence

Theorem 2.1 (Martingale Convergence Theorem (MCT)). *If (X_n) is a sub-MG, $\sup_n \mathbb{E}X_n^+ < \infty$, then $X_n \xrightarrow{a.s.} X_\infty$ for some X_∞ with $\mathbb{E}|X_\infty| < \infty$.*

Proof. $U_n[a, b] \uparrow U_\infty[a, b]$ so by monotone convergence theorem, upcrossing inequality, and the assumption $\mathbb{E}X_n^+ < \infty$

$$\mathbb{E}U_\infty[a, b] = \lim_n \mathbb{E}U_n[a, b] \leq \frac{\sup_n \mathbb{E}X_n^+ + |a|}{b - a} < \infty \quad (2.1)$$

This implies that $U_\infty[a, b] < \infty$ a.s., hence

$$P(U_n[a, b] < \infty \forall a, b \in \mathbb{Q}, a < b) = 1 \quad (2.2)$$

For reals (x_n) , if $\limsup_n x_n > \liminf_n x_n$, then $U_\infty[a, b] = \infty$ for some $a < b$.

Since $U_\infty[a, b] < \infty$ for all rational $a < b$, the contrapositive implies $\limsup x_n = \liminf x_n \in [-\infty, \infty]$. Therefore, $X_n \rightarrow X_\infty$ a.s..

We have so far $X_\infty \in [-\infty, \infty]$, but we would like $\mathbb{E}|X_\infty| < \infty$. To show this, recall Fatou's Lemma: If $Y_n \geq 0$, then $\mathbb{E} \liminf_n Y_n \leq \liminf_n \mathbb{E}Y_n$.

$X_n^+ \rightarrow X_\infty^+$ a.s. implies (by Fatou's Lemma) that $\mathbb{E}X_\infty^+ \leq \liminf_n \mathbb{E}X_n^+ < \infty$. Also

$$\mathbb{E}X_n^- = \mathbb{E}X_n^+ - \mathbb{E}X_n \leq \mathbb{E}X_n^+ - \mathbb{E}X_0 \quad (2.3)$$

since (X_n) a sub-MG means $\mathbb{E}X_0 \leq \mathbb{E}X_n$. As $X_n^- \rightarrow X_\infty^-$ a.s., by Fatou's Lemma

$$\mathbb{E}X_\infty^- \leq \liminf_n \mathbb{E}X_n^- \leq \sup_n \mathbb{E}X_n^+ - \mathbb{E}X_0 < \infty \quad (2.4)$$

Since $\mathbb{E}X_\infty^+ < \infty$ and $\mathbb{E}X_\infty^- < \infty$, we have $\mathbb{E}|X_\infty| < \infty$. □

Corollary 2.2. If (X_n) is a super-MG, $X_n \geq 0$ a.s., then $X_n \xrightarrow{a.s.} X_\infty$ and $0 \leq \mathbb{E}X_\infty \leq \mathbb{E}X_0$.

Proof. Apply MCT to $-X_n$, so $X_n \xrightarrow{a.s.} X_\infty$. Use Fatou's Lemma: $\mathbb{E}X_\infty \leq \liminf_n \mathbb{E}X_n \leq \mathbb{E}X_0$. \square

Example 2.3 (WARNING: MCT does not imply $\mathbb{E}X_n \rightarrow \mathbb{E}X_\infty$). Consider a simple random walk $X_0 = 1$, stopped at $T = \min\{n : X_n = 0\}$. Let $Y_n = X_{\min(T,n)}$. Then $Y_n \rightarrow 0 = Y_\infty$ a.s., but $\mathbb{E}Y_n = 1 \forall n$ which differs from $\mathbb{E}Y_\infty = 0$.

This is similar to *uniform convergence*: if continuous f_n converge uniformly to f , then f is continuous. Not necessarily true for pointwise convergence.

For sequences of expectations to converge, we need *uniform integrability*.

3 Facts about Uniform (Equi-)Integrability

Consider \mathbb{R} -valued RVs (Y_α)

Definition 3.1. A family (Y_α) is *uniformly integrable (UI)* if

$$\lim_{b \rightarrow \infty} \sup_\alpha \mathbb{E} \left[|Y_\alpha| 1_{|Y_\alpha| \geq b} \right] = 0 \quad (3.1)$$

“We have integrability *uniformly over all* RVs Y_α in the family”

If $\mathbb{E}|Y| < \infty$, then $\lim_{b \rightarrow \infty} \mathbb{E} \left[|Y| 1_{|Y| > b} \right] = 0$

Facts relating to UI (see Durrett or Billingsley)

1. If $\sup_\alpha \mathbb{E}|Y_\alpha|^q < \infty$ for some $q > 1$, then (Y_α) is UI, which implies that $\sup_\alpha \mathbb{E}|Y_\alpha| < \infty$
2. If $Y_n \xrightarrow{a.s.} Y_\infty$ and (Y_n) is UI, then $\mathbb{E}|Y_\infty| < \infty$ and $\mathbb{E}|Y_n - Y_\infty| \rightarrow 0$ (i.e. $Y_n \rightarrow Y_\infty$ in L^1)
3. If $Y_n \rightarrow Y_\infty$ in L^1 , then (Y_n) is UI.
4. If $\mathbb{E}|Y| < \infty$, the family $\{\mathbb{E}[Y | \mathcal{G}] : \forall \mathcal{G} \subset \mathcal{F}\}$ is UI.

Theorem 3.2. For a MG (not sub-MG!) (X_n) , TFAE:

(a) (X_n) is UI

(b) X_n converges in L^1

(c) There exists a RV X_∞ with $\mathbb{E}|X_\infty| < \infty$ such that $X_k = \mathbb{E}[X_\infty | \mathcal{F}_k] \forall k$

If any of these conditions hold, then $\exists X_\infty$ such that $X_n \rightarrow X_\infty$ both a.s. and in L^1 .

Proof. (c) \implies (a), item 4..

(a) implies (by item 1.) that $\sup_n \mathbb{E}|X_n| < \infty$, which by MCT implies X_n converges to some X_∞ a.s., which implies by 2. that $X_n \rightarrow X_\infty$ in L^1 i.e. (b).

Given (b), $X_n \rightarrow X_\infty$ in L^1 , which means that $\mathbb{E}|X_n - X_\infty| \rightarrow 0$ with $\mathbb{E}|X_\infty| < \infty$. We need to prove that $\mathbb{E}X_\infty 1_A = \mathbb{E}X_k 1_A$ for any $A \in \mathcal{F}_k$. Fix A and k . By the MG property, for $n > k$, $\mathbb{E}[X_n | \mathcal{F}_k] = X_k$ so $\mathbb{E}X_n 1_A = \mathbb{E}X_k 1_A$. Hence

$$|\mathbb{E}X_\infty 1_A - \mathbb{E}X_n 1_A| \leq \mathbb{E}|X_\infty - X_n| \rightarrow 0 \quad (3.2)$$

as $n \rightarrow \infty$, so $|\mathbb{E}X_\infty 1_A - \mathbb{E}X_k 1_A| = 0$. \square

Using UI with MCT leads to a convergence property for conditional expectations.

Theorem 3.3 (Levy 0-1 Law). *Let $(Y_n)_{n \geq 0}$ be any process, \mathcal{F}_n the natural filtration $\sigma(Y_k, 0 \leq k \leq n)$. Let Z be any RV with $\mathbb{E}|Z| < \infty$ and $Z \in \mathcal{F}_\infty$.*

Then $X_n = \mathbb{E}[Z | \mathcal{F}_n]$ is a UI martingale, so theorem 3.2 implies $X_n \rightarrow X_\infty$ both a.s. and in L^1 . In terms of Z

$$\mathbb{E}[Z | \mathcal{F}_n] \rightarrow \mathbb{E}[Z | \mathcal{F}_\infty] \quad \text{as } n \rightarrow \infty \quad (3.3)$$

In fact, $X_\infty = Z$ because

$$\mathbb{E}[X_\infty | \mathcal{F}_n] = X_n = \mathbb{E}[Z | \mathcal{F}_n] \quad \text{MG property} \quad (3.4)$$

$$0 = \mathbb{E}[X_\infty - Z | \mathcal{F}_n] = X_\infty - Z \quad X_\infty - Z \text{ is } \mathcal{F}_\infty\text{-meas} \quad (3.5)$$

Remark 3.4. In particular, take $Z = 1_A$ for some event A . Then

$$P(A | \mathcal{F}_n)(\omega) \xrightarrow{\text{a.s.}} 1_A(\omega) \quad (3.6)$$

for all $A \in \sigma(Y_n, n \geq 0)$.

In English:

If we are learning gradually all the information that determines the outcome of an event, then we will become gradually certain what the outcome will be.

For independent (Y_n) , let $A \in \mathcal{F}_\infty$ be some tail event. Then

$$P(A | \mathcal{F}_n)(\omega) = P(A) \xrightarrow{\text{a.s.}} 1_A \quad \text{as } n \rightarrow \infty \quad (3.7)$$

which implies that 1_A is a constant a.s. and hence $P(A) \in \{0, 1\}$. This is Kolmogorov's zero-one law.