## 1 Bounding maxima

**Theorem 1.1** (Kolmogorov's Maximal Inequality).  $(X_i)_{i=1}^n$  independent,  $\mathbb{E}X_i = 0$ ,  $\mathbb{E}X_i^2 < \infty$ .  $S_m = \sum_{i=1}^m X_i$ .  $S_n^+ = \max_{1 \le m \le n} |S_m|$  Then  $P(S_n^+ \ge x) \le \frac{\mathbb{E}S_n^2}{x^2}$ , x > 0.

The proof uses a general trick related to martingales of considering stopping times.

*Proof.* Fix x. Event  $\{S_n^+ \ge x\} = \bigcup_{k=1}^n A_k$  where  $A_k = \{|S_k| \ge x, |S_i| < x, \text{ all } 1 \le i \le k\}$ . Note  $(S_k, A_k)$  independent of  $S_n - S_k$ .

Notice  $S_n = S_k + (S_n - S_k)$  so we can write

$$\mathbb{E}S_n^2 \ge \sum_{k=1}^n \mathbb{E}[S_n 1_{A_k}] \tag{1.1}$$

$$=\sum_{k=1}^{n}\left(\mathbb{E}(S_k^2 1_{A_k}) + 2\mathbb{E}(\underbrace{S_k 1_{A_k}(S_n - S_k)}_{\mathbb{E}(S_n - S_k) = 0}) + \underbrace{\mathbb{E}((S_n - S_k)^2 1_{A_k}}_{\geq 0}\right)$$
(1.2)

$$\geq \sum_{k=1}^{n} \mathbb{E}(S_k^2 \mathbf{1}_{A_k}) \tag{1.3}$$

$$\geq \sum_{k=1}^{n} \mathbb{E}(x^2 1_{A_k}) \tag{1.4}$$

$$= x^2 P(\bigcup_{k=1}^n A_k) (1.5)$$

$$= x^2 P(|S_n^+| \ge x) \tag{1.6}$$

where we have used independence of  $S_k 1_{A_k}$  and  $(S_n - S_k)$  in  $\ref{eq:special}$ , and  $|S_k| \ge x$  on  $A_k$  in  $\ref{eq:special}$ ??.

## 2 Almost sure convergence

" $\sum_{i=1}^{\infty} x_i$  converges" means  $\lim_{n\to\infty} \sum_{i=1}^{n} x_i$  exists and is finite. This is equivalent to the Cauchy criterion:

$$\sup_{n \ge K} \left| \sum_{i=K+1}^{n} x_i \right| \to 0 \text{ as } K \to \infty$$
 (2.1)

Thus,  $\sum_{i=1}^{\infty} X_i$  converges a.s. means

$$P(\omega: \lim_{N \to \infty} \sum_{i=1}^{N} X_i(\omega) \text{ exists, finite}) = 1$$
 (2.2)

**Theorem 2.1.**  $(X_i)$  independent,  $\mathbb{E}X_i = 0$ ,  $\sigma_i^2 = Var|X_i| < \infty$ . If  $\sum_{i=1}^{\infty} \sigma_i^2 < \infty$ , then  $\sum_{i=1}^{\infty} X_i$  converges a.s.

## Comment:

$$\operatorname{Var}(\sum_{i}^{n} X_{i}) = \sum_{i=1}^{n} \sigma_{i}^{2}$$
(2.3)

$$\operatorname{Var} \sum_{i=1}^{\infty} = \sum_{i=1}^{\infty} \sigma_i^2 \le \infty(*)$$
 (2.4)

$$\implies \sum_{i=1}^{\infty} X_i \text{ is finite a.s.}$$
 (2.5)

**Exercise**: Given theorem, show (\*)

*Proof.* Define  $M_k = \sup_{n \ge k} \left| \sum_{i=k+1}^n X_i \right|$ . By Cauchy criterion, suffices to show  $M_k \stackrel{\text{a.s.}}{\to} 0$  as  $k \to \infty$ .

$$P\left(\sup_{k < n \le N} \left| \sum_{i=k+1}^{n} X_i \right| \ge \epsilon \right) \le \epsilon^{-2} \operatorname{Var}\left(\sum_{i=k+1}^{N} X_i\right)$$
 (2.6)

$$= \epsilon^{-2} \sum_{i=k+1}^{N} \operatorname{Var}(X_i)$$
 (2.7)

(2.8)

As  $N \to \infty$ 

$$P(M_k > \epsilon) \le \epsilon^{-2} \sum_{i=1}^{\infty} \sigma_i^2$$
 (2.9)

$$P(w_k > \epsilon) \le P(M_k > \epsilon/2) \le 4\epsilon^{-2} \sum_{i=1}^{\infty} \sigma_i^2$$
 (2.10)

where  $w_k = \sup_{n_1 > n_1 > k} \left| \sum_{i=n_1+1}^{n_2} X_i \right|$ . Note  $M_k \le w_k \le 2M_k$  and  $w_k \downarrow$  as  $k \uparrow$ . As  $k \to \infty$ ,  $w_k \downarrow_{\text{a.s.}} w_\infty$ 

$$P(w_{\infty} > \epsilon) = 0 \tag{2.11}$$

$$\implies w_{\infty} \stackrel{\text{a.s.}}{=} 0$$
 (2.12)

$$\implies w_k \downarrow_{\text{a.s.}} 0$$
 (2.13)

$$\implies M_k \stackrel{\text{a.s.}}{\to} 0 \tag{2.14}$$

**Lemma 2.2** (Kronecker).  $(x_n) \in \mathbb{R}^{\omega}$ .  $S_n = \sum_{i=1}^n x_i$ .  $0 < a_n \uparrow \infty$  as  $n \uparrow \infty$ . If  $\sum_i \frac{x_i}{a_i}$  converges, then  $\frac{s_n}{a_n} \to 0$ .

*Proof.* Exercise.

**Corollary 2.3.**  $(X_i)$  independent,  $\mathbb{E}X_i = 0$ ,  $\mathbb{E}X_i^2 < \infty$ . If  $0 < a_n \uparrow \infty$  as  $n \uparrow \infty$  and if  $\sum_n \frac{\mathbb{E}X_n^2}{a_n^2} < \infty$ , then  $\frac{S_n}{a_n} \to 0$  a.s..

*Proof.* Previous theorem implies  $\sum_{n} \frac{X_n}{a_n}$  converges a.s.. Lemma implies  $\frac{S_n}{a_n} \stackrel{\text{a.s.}}{\to} 0$ .

**Specialization**: Suppose also  $\mathbb{E}X_n^2 \sim cn^{2\alpha}$ ,  $\alpha > 0$ . Take  $a_n^2 = n^{1+2\alpha+\epsilon}$  ( $\epsilon > 0$  small). Then corrolary implies  $\frac{S_n}{n^{1/2+\alpha+\epsilon}} \stackrel{\text{a.s.}}{\to} 0$  TODO: Check the 1/2.

**Specialization**: Suppose  $\sup_n \mathbb{E} X_n^2 < \infty$ . Take  $a_n^2 = n(\log n)^{1+2\epsilon}$ . The corollary implies  $\frac{S_n}{\sqrt{n(\log n)^{1+\epsilon}}} \stackrel{\text{a.s.}}{\to} 0$ .

Implicitly from CLT: If  $(X_i)$  i.i.d., then

$$\frac{S_n}{\sqrt{n}} \stackrel{\text{a.s.}}{\to} 0 \tag{2.15}$$

Law of iterated log.

**Theorem 2.4** (Strong Law of Large Numbers (SLLN)). Let  $(X_i)$  iid with  $\mathbb{E}|X_i| < \infty$ ,  $S_n := \sum_{i=1}^n X_i$ . Then  $\frac{S_n}{n} \stackrel{a.s.}{\to} \mathbb{E} X$  as  $n \to \infty$ .

*Proof.* Truncate, center, apply corollary  $(Z \ge 0)$ .  $\mathbb{E} Z^k = \int_0^\infty k z^{k-1} P(Z \ge z) dz \approx \int_0^\infty x^k f(x) dx$  (Truncate): Define  $Y_k = X_k 1_{|X_k| \le k}$ , so  $Y_k$  are no longer iid. However

$$\sum_{k} P(Y_k \neq X_k) = \sum_{k=1}^{\infty} P(|X| > k) \le \int_{0}^{\infty} P(|x| > x) dx = \mathbb{E}|X| \le \infty$$
 (2.16)

By Borel Cantelli 1,  $P(Y_k = X_k \text{ e.v.}) = 1$ . Thus, suffices to prove  $\frac{1}{n} \sum_{k=1}^n Y_k \stackrel{\text{a.s.}}{\to} \mathbb{E} X$ . (Center): Define  $X'_k = Y_k - \mathbb{E} Y_k$ . Claim:

$$\sum_{k} \frac{\operatorname{Var}(X_{k}')}{k^{2}} < \infty \tag{2.17}$$

To show the claim:

$$\mathbb{E}Y_k^2 = \int_0^\infty 2y P(|Y_k| > y) dy \tag{2.18}$$

$$= \int_0^\infty 2y P(k \ge |X_k| \ge y) 1_{y \le k} dy$$
 (2.19)

$$\leq \int_0^\infty 2y P(|X_k| \geq y) 1_{y \leq k} dy \tag{2.20}$$

$$\sum_{k} \frac{\operatorname{Var} X_n'}{k^2} \le \sum_{k} \frac{\mathbb{E} Y_k^2}{k^2} \tag{2.21}$$

$$\leq \sum_{k} \frac{1}{k^2} \int_0^\infty 2y P(|X| \geq y) 1_{y \leq k} dy \tag{2.22}$$

$$= \int_0^\infty \left( \underbrace{\sum_k \frac{1}{k^2} 1_{y \le k}}_{G(y)} 2y \right) P(|X| \ge y) dy \tag{2.23}$$

**Claim**:  $G(y) \le 4$  for all  $0 < y < \infty$ . True for  $y \le 1$ . Take y > 1

$$\frac{1}{k^2} \le \int_{k-1}^k \frac{1}{x^2} dx \tag{2.24}$$

$$\sum_{k} \frac{1}{k^2} 1_{y \le k} = \sum_{k \ge \lceil y \rceil} \frac{1}{k^2} \le \int_{\lceil y \rceil - 1}^{\infty} \frac{1}{x^2} dx = \frac{1}{\lceil y \rceil - 1}$$
 (2.25)

$$\implies G(y) \le \frac{2y}{\lceil y \rceil - 1} \le 4 \tag{2.26}$$

Hence

$$\sum_{k} \frac{\operatorname{Var} X_{n}'}{k^{2}} \le 4 \int_{0}^{\infty} P(|X| > y) dy = 4\mathbb{E}|X|$$
(2.27)

Apply corollary to  $X'_n$ 

$$\frac{1}{n} \sum_{i=1}^{n} X_i' \stackrel{\text{a.s.}}{\to} 0 \tag{2.28}$$

$$\frac{1}{n} \sum_{i}^{n} (Y_i - \mathbb{E}Y_i) \stackrel{\text{a.s.}}{\to} 0 \tag{2.29}$$

Note  $\mathbb{E}Y_i = \mathbb{E}X1_{|X| \leq ?} \to \mathbb{E}X$  by dominated convergence, so

$$\frac{1}{n} \sum_{i}^{n} (\mathbb{E}Y_i - \mathbb{E}X) \stackrel{\text{a.s.}}{\to} 0 \tag{2.30}$$

Adding ?? with ?? yields

$$\frac{1}{n} \sum_{i=1}^{n} (Y_i - \mathbb{E}X) \stackrel{\text{a.s.}}{\to} 0 \tag{2.31}$$

$$\frac{1}{n} \sum_{i}^{n} (Y_{i} - \mathbb{E}X) \stackrel{\text{a.s.}}{\to} 0$$

$$\frac{1}{n} \sum_{i}^{n} Y_{i} \stackrel{\text{a.s.}}{\to} \mathbb{E}X$$
(2.31)