After-hours Incentives and Emergency Department Visits: Evidence from Ontario

Rose Anne Devlin * Koffi Ahoto Kpelitse † Lihua Li ‡ Nirav Mehta §

Sisira Sarma ¶

August 1, 2019

Abstract

One important component of the primary care reform in Ontario, Canada is to incentivize physicians to work "after hours" in order to improve access to core primary care services and potentially reduce visits to hospital emergency departments. Empirically, evidence on this link is ambiguous. We suggest reasons for this ambiguity, and then harness rich administrative data from Ontario to carefully investigate if and why after-hours incentives affect ED usage. The data cover visits to physicians' offices and ED visits from 2003 to 2007, a period with exogenous changes in after-hours incentives. We find strong evidence that nonurgent ED visits are reduced as a result of these incentives.

Keywords: primary care; physician incentives

JEL classification: I11, I12, I18, H51

^{*}Department of Economics, University of Ottawa, Ottawa, ON

[†]Department of Epidemiology & Biostatistics, University of Western Ontario, London, ON

[‡]Institute for Clinical Evaluative Sciences, Toronto, ON

[§]Corresponding Author; Department of Economics, University of Western Ontario, London, ON

[¶]Department of Epidemiology & Biostatistics, University of Western Ontario, London, ON and Institute for Clinical Evaluative Sciences, Toronto, ON

1 Introduction

Avoidable, or nonurgent, emergency department (ED) visits are commonplace in developed countries (Carret et al., 2009). About 39% of Canadians reported that their ED visits could have been avoided if primary care were available (Schoen et al., 2005). Aside from contributing to overcrowding and delaying care for patients in urgent need, the use of EDs for nonurgent health problems contributes to higher health care costs (Campbell et al., 2005; Mehrotra et al., 2009; Thygeson et al., 2008) and lower continuity of care, adversely affecting health outcomes, especially for patients with chronic conditions (Dunnion and Kelly, 2005; Stiell et al., 2003; Vinker et al., 2004).

Primary care is publicly funded in all jurisdictions in Canada with no direct cost to the patient for physician and hospital visits. While no direct financial cost is borne, ED visits entail long waits before a physician is seen, especially for nonurgent cases. All else being equal, patients would likely prefer to be treated by their family doctor for non-urgent health problems.

Reforms to primary care were introduced in the early 2000s across various jurisdictions in Canada (Health Canada, 2007; Sweetman and Buckley, 2014). Common across these reforms was a move from traditional fee-for-service (FFS) remuneration towards pay-for-performance incentives for preventive care and chronic disease management (Hutchison et al., 2011;). Several new types of non-FFS primary care delivery models featuring these financial incentives have been introduced in Canada's most populous province, Ontario, since 2004 (Hutchison et al., 2011). By 2010, more than two-thirds of Ontario's family physicians had joined one of these models, with Family Health Organizations (FHO) and Family Health Groups (FHG) being the two most popular choices (Henry et al., 2012).

One of the goals of these new models is to reduce ED visits by increasing access to primary health care services outside of regular working hours. Each physician practicing in these new models is required to provide a minimum of one three-hour session per week either weeknights after 5 pm or on weekends or statutory holidays. In return, these physicians receive an after-hours premium, initially of 10% when first introduced in 2003 and increasing to 15% in April 2005, 20% in April 2006 and then to 30% in September 2011 (Sweetman and Buckley, 2014). The main goal of our paper is to determine whether this policy was successful at reducing ED utilization.

Pay for performance schemes have been studied in several contexts in Ontario (Kantarevic and Kralj, 2013; Li et al., 2014). A handful of studies examine the link between improved after-hours access to primary care and ED visits outside of Canada, with mixed findings. This includes the implementation of an after-hours clinic or cooperative (Buckley et al.,

2010; Pickin et al., 2004), the extension of primary care practice opening hours (Dolton and Pathania, 2016; Harris et al., 2011; Lippi Bruni et al., 2016; Lowe et al., 2005), the reorganization of after-hours care (van Uden and Crebolder, 2004; van Uden et al., 2005), and after-hours financial incentives (Franco et al., 1997; Piehl et al., 2000).

At least two reasons explain why increased after-hours services may not reduce ED visits unambiguously: Firstly, ED visits would fall only if patients have conditions that are otherwise treatable by primary care physicians; no amount of after-hours care would reduce visits to the ED for the most urgent-need patients. Secondly, while ED visits may fall because of increased after-hours services, they may rise if regular-hours services are reduced.

We use a rich, longitudinal data set that allows us to control for factors that would normally comprise unobserved heterogeneity, and exploit exogenous variation in the strength of after-hours incentives, to help establish the impact of these incentives on ED visits. The large number of physicians in this data set also allows us to estimate this impact by different subgroups of physicians, such as those with sicker patients. As expected, regular- and after-hours services move in opposite directions in response to stronger after-hours incentives. We find that after-hours services reduce nonurgent ED visits. Most nonurgent ED reductions come from practices with below-median co-morbidity, suggesting that after-hour incentive reduces ED visits from healthier patients, allowing more time for more urgent visits.

2 Data and Variables

Several administrative databases held at the Institute for Clinical Evaluative Sciences (ICES) provide the data for this study. The Physician Database contains characteristics of primary care physicians; the Corporate Provider Database provides physicians' model type, effective date of eligibility for billing under the Ontario Health Insurance Plan ("OHIP", the public insurer/payer) and physician group size. The Client Agency Program Enrollment Database allows us to match physicians with enrolled patients. If a physician was affiliated with more than one practice type, the most recent one joined was selected. Only physicians who put in a claim for after-hour incentives to the public insurer (OHIP) are in our sample. The billing codes eligible for the after-hours premium correspond to fairly basic services that most physicians provide—like minor assessments, primary mental health care, counseling, and annual physical examinations. Knowing this, we expected an appreciable effect of premium changes on ED utilization. We exclude part-time physicians, defined as having fewer than

 $^{^{1}}$ The OHIP billing codes used to compute after-hour costs are A001, A003, A004, A007, A008, A888, K005, K013, K017, K033, K030, Q050, K130, K131, and K132; the after-hours premium codes are Q12 and Q16.

500 patients or 500 visits in any given year. We focus on data spanning 2003-2007, with 1,321 unique physicians and 6,605 physician-year observations comprising the balanced panel we use for our analysis.²

Patient visits to a physician were identified through OHIP billing claims. For each physician, total annual office visits were derived as the sum of patient visits.³ The total number of annual office visits minus the total number of annual after-hour visits defines regular visits for each physician. Group size sums up the number of primary care physicians with the same group number.

Emergency department visits come from the National Ambulatory Care Reporting System which classifies them into urgent and nonurgent based on the Canadian Triage and Acuity Scale (CTAS) (Beveridge et al., 1999). A triage level of 1 (resuscitation), 2 (emergent), or 3 (urgent) is urgent, while 4 (less-urgent/semi-urgent) or 5 (nonurgent) was not; we aggregate these into three mutually exclusive categories, where group 1 ("very urgent") contains ED visits with a CTAS score of 1 or 2, group 2 ("urgent") contains ED visits with a CTAS score of 4 or 5. The Aggregated Diagnosis Group (ADG) reflecting the health status of each patient is based on their diagnosis codes from the hospital Discharge Abstract Database (DAD) and OHIP, using the Johns Hopkins Adjusted Clinical Group case-mix adjustment system. There are 32 diagnosis groups, which we sum so that each patient has a score between 1 and 32. For each physician, we calculate the average ADG of her patients.

Ontario's health registry database (the Registered Persons Database) provides patients' age, sex and postal codes which we use to obtain deprivation and rurality indices around the patient's residence ("census dissemination area"). The deprivation index is organized into quintiles, where 1 is least marginalized and 5 is most marginalized—our measure is the percent of physician's patients from the fourth and fifth quintiles (i.e., the most deprived areas)—and individuals with a rurality index of 40 or higher are considered to reside in rural areas (Kralj, 2000; Matheson et al., 2012).

Table 1 reports descriptive statistics of regular-hours, after-hours, and ED utilization by year. Between 2003 and 2007, the number of regular-hours visits per 1,000 patients decreased, while after-hours visits per 1,000 patients increased and then fell near the end. Correspondingly, (deflated) total costs per 1,000 patients decreased substantially, even though (deflated)

²We have data covering 2003-2013. However, beginning in 2008, a large number of FPs switched to the Family Health Organization model, a capitation-based payment system found to change physician practice patterns (Zhang and Sweetman, 2018). As these changes may confound our analysis of the impact of afterhours incentives on ED visits, we restricted our analysis to the years 2003-2007.

³Patient visits are defined as the combination of billing codes and the serve date in OHIP (i.e. the date the service(s) were provided to patients). Multiple billing codes in the same date are defined as a one visit, but two billing codes in two separate dates are defined as two visits.

after-hours costs almost doubled.⁴ During the same period, the number of very urgent ED visits increased, urgent ED visits increased sharply initially and then stayed roughly constant, and the number of nonurgent ED visits also increased sharply and then gradually decreased. Table 2 presents statistics pooled over the sample period and shows that after-hours visits comprise about 11% of physicians' total visits on average.

3 Empirical Framework

The net effect of increasing the after-hours premium on ED visits is ambiguous. We start by estimating the regression⁵

$$\mu_{it} = \rho_{\mu,\pi} \pi_t + Z'_{it} \beta_{\mu,\pi} + \epsilon_{it,\mu,\pi}, \tag{1}$$

where μ_{it} can be total, nonurgent, urgent, or very urgent ED visits per 1,000 patients for physician i in year t; π_t is the after-hours premium in year t; Z_{it} includes a time trend, physician's age, physician's age squared, proportion of female physicians and foreign graduates in the physician's practice, group size, average age of patients, average ADG score of patients, proportion of patients living in deprived areas, and proportion of patients living in rural areas; and $\epsilon_{it,\mu,\pi}$ represents the error term, which may include a fixed-effect component for the physician. However, the net effect of increasing the after-hours premium on overall costs may be ambiguous, even if nonurgent ED visits decrease due to the after-hours premium increase. We thus re-estimate the model using total costs as the dependent variable.

To examine how physician behavior is affected by incentivizing after-hours access, we estimate a model of services provided by physician i during year t:

$$x_{it} = \rho_{x,\pi} \pi_t + Z'_{it} \beta_{x,\pi} + \epsilon_{it,x,\pi}, \tag{2}$$

where x_{it} can be either regular- or after-hours services, measured by either visits per 1,000 patients or (deflated) costs per 1,000 patients (i.e., the value of services in 2004 prices); π_t is the premium level in year t; Z_{it} contains the variables from equation (1); and $\epsilon_{it,x,\pi}$ represents the error term, which may include a fixed-effect component for the physician. As fee-for-service physicians do not receive after-hours premiums, we focus our analysis on physicians who have switched into a scheme that incentivizes after-hours services (recall that

⁴Note that by "after-hours costs" we refer to the value of after hours services, excluding the after-hours premium.

⁵We adopt the notational convention of subscripting regression coefficients with the dependent variable first and then the primary regressor of interest.

we restrict our analysis to go until 2007, before the large transition to FHO occurred).

Next, we estimate how ED visits vary with respect to the value of after-hours services in 2004 prices, using the regression model

$$\mu_{it} = \rho_{\mu,x} x_{it} + Z'_{it} \beta_{\mu,x} + \epsilon_{it,\mu,x}, \tag{3}$$

where μ_{it} can be total, nonurgent, urgent, or very urgent ED visits per 1,000 patients for physician i in year t; x_{it} represents the value of after-hours services in 2004 prices; Z_{it} contains the variables from equation (1); and $\epsilon_{it,\mu,x}$ represents the error term, which may include a fixed-effect component for the physician.

One might wonder whether physicians would engage in "gaming" in response to the financial incentives. They could, for instance, encourage their patients to come during afterhours in order to gain the premium. If all changes in after-hours services were driven by gaming behavior there would be no change in the provision of primary care and, thus, no effect on ED visits. Our empirical approach takes into account this potential for gaming, as we estimate the net effect of premium increases on ED utilization. Evidence of an effect of after-hours premium increases on ED visits would suggest such gaming is not the dominant force at play. From a practical perspective, however, the regulatory framework affords very limited scope for gaming. Physicians are required to post their after-hours availability to patients, and the incentives are applicable only to enrolled patients seen during posted after-hours sessions. Physicians are free to provide services to non-enrolled patients but these are not eligible for after-hours incentives.

4 Empirical Findings

Effect of after-hours incentives on ED utilization: Table 3 provides the estimated coefficients from the OLS and fixed effects models of ED visits (total, and split by urgency group) on the after-hours premium as well as physician and practice characteristics (equation (1)). From the OLS results in the left panel, the estimated coefficients on the after-hours premium (premium) are negative across the board. Not only do total ED visits fall, but so too do the visits of all groups (with nonurgent visits being statistically insignificant). The last result is surprising insofar as one would expect that nonurgent visits would decrease after the introduction/strengthening of after-hours incentives for regular physicians. The OLS procedure does not control for time-invariant physician heterogeneity which affects the reliability of these estimates. Indeed, when we take account of physician level fixed effects, these results change quite significantly. From the right-hand side of Table 3, we

find that the after-hours premium exerts a a negative and statistically significant impact for the nonurgent group (specification (6)) and a positive and statistically significant impact on more urgent ED visits (specifications (8) and (7)). Together, these findings support the idea that the after-hours premium encouraged those patients with nonurgent conditions to seek care at their physician's office, thus reducing their crowding out of more urgent ED visits. We also see that female physicians tend to have fewer ED visits, as do international medical graduates (IMG) and those physicians who practice in larger practices (group size). Physicians with younger patients (avg. age), fewer deprived patients (avg. deprived), and fewer rural patients (avg. rural) also have fewer ED visits. Physicians with higher mean patient ADGs (avg. ADG) have more very urgent ED visits and fewer nonurgent ED visits.

Effect of after-hours incentives on services: If increasing the after-hours premium reduces ED utilization by changing physician behavior, this should show up in the data. Office visits are one measure of physician behavior, but may not be ideal because the amount of services provided per visit may change if the after-hours premium changed. The rich data used in this study enable us to use two measures of physician after-hours services: visits and the value of services provided. Value of services provided are a good measure of services rendered if service prices capture input amounts. We construct price-adjusted values, "after-hours values" and "regular values", by correcting for the change in the prices of services (in 2004 values), excluding the after-hours bonuses (which were, e.g., 10% in 2004). We use these deflated values as our primary measure of services in our empirical analysis.

Table 4 presents the fixed-effects results from regressing measures of services on the after-hours premium (equation (2)). Specification (1) regresses after-hours visits on the premium, and finds a positive and significant estimated coefficient: a higher premium is correlated with more after-hours visits. Specification (2) uses the (deflated) value of after-hours services—i.e., the cost, adjusted for changes in the premium and in the prices of services—as the measure of services rendered, and finds a strong positive relationship between the premium amount and services. Specification (3) examines how the after-hours premium affects regular visits, and reveals the expected negative, significant, effect. Specification (4) examines how the (deflated) value of all services co-vary with the premium, and finds a negative, significant effect. As total costs include regular and after-hours costs, this implies that the regular costs measure of services provided during regular hours must have decreased. In short, both measures of services provided after hours increase in response to increases in response to premium increases.

Effect of after-hours services on ED utilization: The fixed-effects estimates of the impact of the after-hours premium on ED visits indicate that they fell for the nonurgent group (Table 3), while after-hours physician services, not surprisingly, increased (Table 4). Table 5 summarizes the relationship between after-hours costs and ED utilization, using OLS (first row) and fixed-effects (second row) estimates of equation (3). The OLS results are presented for comparison purposes, but we focus on the fixed-effects ones. The fixed-effects estimates show that, while increases in after-hours services do not decrease overall ED visits (specification (1)), they do significantly decrease the least-urgent ED visits (specification (2)); they are also associated with small (but statistically significant) increases in urgent ED visits (specification (3)) but are not associated with any change in very urgent ED visits (specification (4)).

The fixed-effects models control for time-invariant physician heterogeneity and provide evidence that increases in after-hours services reduce nonurgent ED visits. We can provide even stronger evidence supporting the link between after-hours services and ED visits by exploiting variation in the after-hours premium over time as an instrument, and using a fixed-effects instrumental variables estimation approach. Because changes in the after-hours premium over time are exogenous⁶, the premium is arguably a valid instrument in the ED utilization equations. Given the higher wage rate faced by physicians during after-hours sessions as well as regulations governing the minimum number of these sessions, we expect the after-hours premium to be highly correlated with after-hours services. This is confirmed by the large F-statistics in the first-stage regression (see Table 6). Moreover, there is no a priori reason to expect the after-hours premium awarded family physicians to affect ED utilization other than indirectly through the increased provision of incentivized services during after-hours.

Specification (1) of Table 6 shows there is no significant estimated impact of after-hours services on total ED visits. However, specification (2) shows that there is a significant reduction stemming from increases of after-hours services on the least-urgent ED visits. Estimates of the effect on urgent and very urgent ED visits are positive and statistically significant, yet smaller in magnitude than the effect on nonurgent ED visits. These results corroborate those of the fixed-effects models, namely: increasing the after-hours premium increases physicians' provision of after-hours services, which leads to a reduction in nonurgent ED visits. The advantage of the instrumental variables estimates is that they isolate physician behavior as the channel affecting ED utilization.

Up to now, we see that increasing the after-hours premium reduces ED visits for nonur-

⁶Even if the policy to increase after-hours premiums was made in response to there being an undesirably large number of ED visits, it is exogenous to the actions of any given physician.

gent patients by way of increasing physicians' after-hours services. Now we want to determine if this affects overall health-system costs. Using values for ED visits (2002 dollars) from the Ministry of Health⁷, we estimate

$$\$\mu_{it} = \rho_{\$\mu,x}\$x_{it} + Z'_{it}\beta_{\$\mu,x} + \epsilon_{it,\$\mu,x},\tag{4}$$

where $\$\mu_{it}$ is the cost of ED visits per 1,000 patients and $\$x_{it}$ is the cost of after-hours incentive per 1,000 patients. Note the cost of the after-hours incentive used in this specification includes the increase in premium. Here, we are primarily interested in the estimate of $\rho_{\$\mu,x}$, the change in ED costs with respect to after-hours costs.

Table 7 presents fixed-effects instrumental variables estimates of after-hours services on ED costs. The increase in after-hours services (due to the higher after-hours premium) results in reductions in the least urgent ED costs (specification (2)). By dividing the estimated ED cost savings by the estimated reduction in nonurgent ED visits,

-\$0.6442 per 1,000 patients per physician per year / -0.00432 visits per 1,000 patients per physician per year, we see a cost savings of about \$149 per reduced ED visit. This calculation exploits changes in costs and visits induced by variations in the after-hours premium, and as such represents the mean costs of treating inframarginal patients at the nonurgent ED. The increase in costs due to more urgent care visits is of a much smaller magnitude than the cost savings found for nonurgent ED visits. Overall, our results suggest that the after-hours policy contributed to a reduction in ED costs.

Finally, it is likely that reductions in ED visits are not uniformly distributed over all practices. In particular, in practices with healthier patients, increased after-hours services may allow patients with less-urgent conditions to be treated by their physicians rather than the ED, as opposed to practices with high-morbidity patients who may need to go to the ED irrespective of the availability of after-hours sessions. To examine this possibility, we re-run our prior fixed-effects regressions for two subsamples of practices: those with below-median average patient ADG and those with above-median average patient ADG.

The first row of Table 8 reports the change in ED visits with respect to after-hours services (measured by after-hours costs), split by whether practices have low and high mean ADGs. The results from specifications (3) and (4) indicate a significant negative relationship with respect to value of after-hours services and nonurgent ED visits; the estimated relationship for urgent ED visits is positive, but much smaller in magnitude. In the second row, we find that the response of nonurgent ED costs with respect to after-hours services is -0.109 for practices with below-median mean ADG, while it is not significantly different from zero at

⁷Specifically, ED costs were derived by multiplying the resource intensity weight available from NACR database with the average cost case by the cost per weighted case. This is the standard procedure commonly used to calculate costs at the population level (Wodchis et al., 2013).

practices with above-median mean ADG. The finding that changes in ED costs are driven by physicians with relatively healthy (and, therefore, easier to treat) patients is consistent with the idea that premium changes affect ED utilization via physician behavior.

5 Concluding Remarks

Our results show that the after-hours incentive in Ontario's primary care setting resulted in more nonurgent patients seeing their family physicians, reducing nonurgent ED visits. Our paper has several strengths. We provide empirical evidence using novel health administrative data from Ontario that after-hours incentives reduce ED visits. Moreover, we are able to investigate the impact of after-hours incentives on ED utilization at the intensive margin. We uncover evidence that after-hours incentives reduce nonurgent ED utilization, stemming largely from practices with healthier patients, rather than those with sicker ones.

We noted earlier that, although we have data from 2003-2013, we focus our analysis on 2003-2007. In addition to an increase in the after-hours premium from 20% to 30%, there were several additional changes in the institutional environment from 2008-2013, including the introduction of additional pay for performance incentives and that many physicians switched to FHO and team settings, which was found to reduce services in FHO baskets (Zhang and Sweetman, 2018). We conducted a robustness analysis covering this later period, and found several qualitatively similar results, which we present in an Online Appendix. Briefly, similar to the 2003-2007 period, Online Appendix Table 9 shows that increases in the premium are associated with statistically significant reductions in nonurgent ED visits (specifications (2) and (6)). Online Appendix Table 10 shows that there are significant reductions in regular visits and costs (specifications (3) and (4)), although specifications (1) and (2) do not show significant increases in after-hours services, as measured by either visits or costs. The within-physician (fixed-effects) estimates presented in the second row of Online Appendix Table 11 show that increases in after-hours services reduce nonurgent ED visits (specification (2)), similar to the results from 2003-2007. Finally, similar to our results from 2003-2007, Online Appendix Table 12 shows significant reductions in ED costs associated with nonurgent visits (specification (2)).

One interesting direction for future work concerns the after-hours premium itself. For example, a constant after-hours premium may not be the most cost-effective, and one that differs with disease severity may lead to further health-system savings. Examining the optimal non-linear incentive structure would be fruitful. Additionally, another point for future research is whether after-hours incentives work differently in retrospective and perspective payment systems.

Like other studies using administrative data, we are limited as to the variables available for empirical analyses. The lack of socio-economic information on physicians and their patients, including information on family income, constrains the work. Such data would enhance future research in this area.

Acknowledgments An earlier version of this paper was presented at the 49th Annual Conference of the Canadian Economics Association, the 16th Annual Canadian Health Economists' Study Group meeting, and the 2017 international Health Economics Association Congress and we are grateful to various comments of participants and discussants. We thank Alex Kopp, Sue Schultz, Rick Glazier, Salimah Shariff and Amit Garg at the Institute for Clinical Evaluative Sciences (ICES) for numerous help in moving forward with the ICES data. Funding for this research by the Canadian Institutes of Health Research operating grant (MOP–130354) is gratefully acknowledged. This study was undertaken at the ICES Western site. ICES is funded by an annual grant from the Ontario Ministry of Health and Long-Term Care (MOHLTC). Core funding for ICES Western is provided by the Academic Medical Organization of Southwestern Ontario (AMOSO), the Schulich School of Medicine and Dentistry (SSMD), Western University, and the Lawson Health Research Institute (LHRI). The opinions, results and conclusions are those of the authors and are independent from the funding sources. No endorsement by ICES, AMOSO, SSMD, LHRI, CIHR, or the MOHLTC is intended or should be inferred.

References

- [1] Afilalo, J., Marinovich, A., Afilalo, M., Colacone, A., Léger, R., Unger, B., Giguère, C., 2004. Nonurgent emergency department patient characteristics and barriers to primary care. Academic Emergency Medicine, 11(12), 1302–1310.
- [2] Beveridge, R., Clarke, B., Janes, L., 1999. Canadian Emergency Department Triage and Acuity Scale Implementation Guidelines. Canadian Journal of Emergency Medicine 1, S2–S20.
- [3] Billings, J., Parikh, N., Mijanovich, T., 2000. Emergency department use: the New York story. Issue Brief 435, 1-5.
- [4] Broadway, B., Kalb, G., Li, J. and Scott, A., 2017. Do Financial Incentives Influence GPs' Decisions to Do After-hours Work? A Discrete Choice Labour Supply Model. Health Economics.
- [5] Buckley, D.J., Curtis, P.W., McGirr, J.G., 2010. The effect of a general practice after-hours clinic on emergency department presentations: a regression time series analysis. The Medical Journal of Australia 192, 448–451.
- [6] Campbell, M.K., Silver, R.W., Hoch, J.S., Ostbye, T., Stewart, M., Barnsley, J., Hutchison, B., Mathews, M., Tyrrell, C., 2005. Re-utilization outcomes and costs of minor

- acute illness treated at family physician offices, walk-in clinics, and emergency departments. Canadian Family Physician 51, 83.
- [7] Carret, M.L.V., Fassa, A.C.G., Domingues, M.R., 2009. Inappropriate use of emergency services: a systematic review of prevalence and associated factors. Cadernos de saúde pública 25, 7–28.
- [8] Dolton, P., Pathania Vikram, 2016. Can increased care access reduce demand for emergency care? Evidence from England's 7-day GP opening. Journal of Health Economics 49, 193-208.
- [9] Dunnion, M.E., Kelly, B., 2005. From the emergency department to home. Journal of Clinical Nursing 14, 776–785.
- [10] Franco, S.M., Mitchell, C.K., Buzon, R.M., 1997. Primary care physician access and gatekeeping: a key to reducing emergency department use. Clinical Pediatrics 36, 63–68.
- [11] Glazier, R., Zagorski, B.M., Rayner, J., 2012. Comparison of Primary Care Models in Ontario by Demographics, Case Mix and Emergency Department Use, 2008/09 to 2009/10.
- [12] Glazier, R.H., Moineddin, R., Agha, M.M., Zagorski, B., Hall, R., Manuel, D.G., Sibley, L.M., Kopp, A., 2008. The Impact of Not Having a Primary Care Physician Among People with Chronic Conditions ICES Investigative Report.
- [13] Harris, M.J., Patel, B., Bowen, S. 2011. Primary care access and its relationship with emergency department utilisation: an observational, cross-sectional, ecological study. The British Journal of General Practice 61, e787-793.
- [14] Health Canada, 2007. The Primary Health Care Transition Fund: A Legacy for Change. National Conference Report.
- [15] Henry, S.E., Glazier, R.H., Bhatia, R.S., Dhalla, I.A., Al., D.A., 2012. Payments to Ontario Physicians from Ministry of Health and Long-Term Care Sources, 1992/93 to 2009/10. ICES Investigative Report, Toronto: Institute for Clinical Evaluative Sciences.
- [16] Hutchison, B., Levesque, J.-F.-F., Strumpf, E., Coyle, N., 2011. Primary health care in Canada: systems in motion. Milbank Quarterly 89, 256–288.

- [17] Kantarevic, J., Kralj, B., 2013, Link between Pay for Performance Incentives and Physician Payment Mechanisms: Evidence from the Diabetes Management Incentive in Ontario. Health Economics 22(12).
- [18] Kralj, B., 2000. Measuring rurality for purposes of health-care planning: an empirical measure for Ontario. Ontario Medical Review.
- [19] Li, J.; Hurley, J.; DeCicca, P., Buckley, G. Physician Response To Pay-For-Performance: Evidence From A Natural Experiment Health Economics, Wiley Online Library, 2014, 23, 962–978.
- [20] Lippi Bruni, M., Mammi, I., Ugolini, C., 2016. Does the extension of primary care practice opening hours reduce the use of emergency services? Journal of Health Economics 50, 144–155.
- [21] Lowe, R.A., Localio, A.R., Schwarz, D.F., Williams, S. et al., 2005. Association between primary care practice characteristics and emergency department use in a Medicaid managed care organization. Medical Care 43, 792–800.
- [22] Matheson, F.I., Dunn, J.R., Smith, K.L.W., Moineddin, R., Glazier, R.H., 2012. Development of the Canadian Marginalization Index: a new tool for the study of inequality. Canadian Journal of Public Health 103, S12–6.
- [23] McGuire, T.G., 2000. Physician Agency. Handbook of Health Economics, 1, 461–536.
- [24] Mehrotra, A., Liu, H., Adams, J.L., Wang, M.C., Lave, J.R., Thygeson, N.M., Solberg, L.I., McGlynn, E.A., 2009. Comparing costs and quality of care at retail clinics with that of other medical settings for 3 common illnesses. Annals of Internal Medicine 151, 321–328.
- [25] Pickin, D. M., O'Cathain, A., Fall, M., Morgan, A. B., Howe, A., Nicholl, J. P., 2004. The impact of a general practice co-operative on accident and emergency services, patient satisfaction and GP satisfaction. Family Practice 21, 180-182.
- [26] Piehl, M.D., Clemens, C.J., Joines, J.D., 2000. "Narrowing the Gap": decreasing emergency department use by children enrolled in the Medicaid program by improving access to primary care. Archives of Pediatrics & Adolescent Medicine 154, 791–795.
- [27] Ricketts, T.C., 2011. The health care workforce: will it be ready as the boomers age? A review of how we can know (or not know) the answer. Annual Review of Public Health 32, 417–430.

- [28] Sarma, S., Thind, A., Chu, M.-K., 2011. Do new cohorts of family physicians work less compared to their older predecessors? The evidence from Canada. Social Science & Medicine 72, 2049–2058.
- [29] Schoen, C., Osborn, R., Huynh, P.T., Doty, M., Zapert, K., Peugh, J., Davis, K., 2005. Taking the pulse of health care systems: experiences of patients with health problems in six countries. Health Affairs Suppl Web, W5509–5525.
- [30] Stiell, A., Forster, A.J., Stiell, I.G., van Walraven, C., 2003. Prevalence of information gaps in the emergency department and the effect on patient outcomes. Canadian Medical Association Journal 169, 1023–1028.
- [31] Sweetman, A., Buckley, G., 2014. Ontario's Experiment with Primary Care Reform. University of Calgary, School of Public Policy Research papers 7, 1–35.
- [32] Thygeson, M., Van Vorst, K.A., Maciosek, M. V., Solberg, L., 2008. Use And Costs Of Care In Retail Clinics Versus Traditional Care Sites. Health Affairs 27, 1283–1292.
- [33] Van Uden, C.J.T., Crebolder, H.F.J.M., 2004. Does setting up out of hours primary care cooperatives outside a hospital reduce demand for emergency care? Emergency Medicine Journal 21, 722–723.
- [34] Van Uden, C.J.T., Winkens, R.A.G., Wesseling, G., Fiolet, H.F.B.M., van Schayck, O.C.P., Crebolder, H.F.J.M., 2005. The impact of a primary care physician cooperative on the caseload of an emergency department: the Maastricht integrated out-of-hours service. Journal of General Internal Medicine 20, 612–617.
- [35] Vinker, S., Kitai, E., Or, Y., Nakar, S., 2004. Primary care follow up of patients discharged from the emergency department: a retrospective study. BMC Family Practice 5, 16.
- [36] Wodchis WP, Bushmeneva K, Nikitovic M, McKillop I. Guidelines on person-level costing using administrative databases in Ontario. Working Paper Series. Vol 1. Toronto: Health System Performance Research Network; 2013.
- [37] Zhang, X., Sweetman, A., 2018. Blended capitation and incentives: Fee codes inside and outside the capitated basket. Journal of Health Economics 60, 16–29.

Table 1: Sample means, by year (2003-2007)

Year	2003	2004	2005	2006	2007
Premium	0.10	0.10	0.15	0.20	0.20
Obs.	1321	1321	1321	1321	1321
Below variables are per 1,00	0 patients				
Regular-hours visits	4129.85	3096.45	2968.89	2933.39	2751.81
After-hours visits	224.94	402.10	408.39	379.91	350.26
Total costs (deflated)	173668.80	142387.30	140541	134611.60	125896.30
After-hours costs (deflated)	6295.60	11251.11	11 860.40	11073.03	10228.15
Very urgent ED visits	35.67	43.40	47.79	50.37	54.65
Urgent ED visits	138.58	147.07	151.62	152.62	152.89
Nonurgent ED visits	138.58	183.78	177.37	176.03	173.65

Note: "very urgent", "urgent", and "nonurgent" ED visits respectively correspond to CTAS groups 1, 2, and 3.

Table 2: Sample means and standard deviations, pooled over all years

Variable	Mean	Std. Dev.	
Age	48.96	8.61	
Female	0.32	0.46	
IMG	0.08	0.26	
Group size	23.18	28.33	
Roster size	1773.75	666.15	
Avg. age	38.47	5.29	
Avg. ADG	3.18	0.38	
Avg. deprived	23.12	13.02	
Avg. rural	12.50	22.20	
Premium	0.12	0.05	

Below variables are per 1,000 patients

	Mean	Std. Dev.
Regular-hours visits	3176.08	1006.39
After-hours visits	353.12	319.02
Very urgent ED visits	46.38	21.83
Urgent ED visits	148.56	49.30
Nonurgent ED visits	169.88	119.86
After-hours costs (deflated)	10 141.66	9317.56
Total costs (deflated)	143421	47742.82
Very urgent ED costs	16876.07	8254.99
Urgent ED costs	35443.63	11456.86
Nonurgent ED costs	25882.34	18176.57
Total ED costs	78210.01	27569.96
Obs.	6605	

Note: "very urgent", "urgent", and "nonurgent" ED categories respectively correspond to CTAS groups $1,\,2,\,$ and 3.

Table 3: Change in emergency department visits, as function of after-hours premium and other covariates

	ent		Err.	4.239	0.771		0.007			9000	0.210	1.461	0.027	0.107	1528.8	
	Very urgent	8	Coef. St. Err.	11.839 4	3.391 0		0.004 0			0.010 0	0.547 0	1.964 1	0.0960.0	-0.046 0	-6791.7 13	9099
	nt			6.648 1	1.277		0.012 (0.011 (0.450	2.486	0.059	0.154 -	2530.7 -6	
	Urgent	(7)	Coef. St. Err.	43.297	-0.087		0.017			-0.013	0.464	7.897	0.171	0.298	225.3	6605
FE	gent		St. Err.	21.046	3.838		0.036			0.028	1.501	7.747	0.566	0.504	7588.3	
	Nonurgent	(9)	Coef.	-56.890	0.084		-0.008			-0.079	0.987	8.728	5.103	-0.177	-150.7	6605
	visits		St. Err.	17.137	3.310		0.028			0.026	1.653	8.302	0.159	1.083	6545.4	
	Total visits	(2)	Coef.	32.307	-10.556		0.135			0.048	1.208	8.862	0.382	3.580	21072.0	6605
	rgent		St. Err.	8.732	0.370	0.455	0.005	0.960	1.677	0.011	0.109	1.317	0.046	0.018	739.6	
	Very urgent	(4)	Coef.	-16.591	4.472	-0.518	0.005	-7.684	-7.419	-0.025	1.058	6.276	0.447	-0.148	-8969.0	6605
	ant		St. Err.	16.837	0.698	1.108	0.011	2.449	3.698	0.025	0.245	3.073	0.096	0.064	1397.5	
∞	Urgent	(3)	Coef.	-33.918	2.990	-2.899	0.022	-6.057	-13.093	-0.184	2.312	0.950	1.300	0.303	-5874.5	6605
OLS	rgent		Coef. St. Err.	27.732	1.166	2.011	0.020	4.263	6.597	0.039	0.469	5.652	0.267	0.249	2339.4	
	Nonurgent	(2)		-1.353	1.786	-1.919	0.008	4.550	-21.445	-0.177	2.195	-20.765	2.662	3.061	-3449.7	6605
	visits		St. Err.	33.860	1.437	2.932	0.029	6.584	9.915	0.060	0.705	8.329	0.260	0.303	2879.5	
	Total visits	(1)	Coef.	-126.261	4.120	-7.493	0.055	-13.213	-38.437	-0.331	5.207	-12.750	3.648	3.979	-7918.8	6605
				Premium	Year	Age	Age sq.	Female	IMG	Group size	Avg. age	Avg. ADG	Avg. deprived	Avg. rural	Constant	Obs.

Note: All regressions include the full set of control variables, Z_{it} . Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 1321 clusters.

Table 4: Fixed-effects regressions of regular- and after-hours services on after-hours premium and other characteristics

	After-hours visits		After-ho	urs costs	Regulai	visits	Total costs	
	(1)		(2	(2)		(3))
	Coef.	Std. Err	. Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.
Premium	436.649	69.423	13 166.930	2048.579	-2223.516	244.539	-58298.540	10 185.040

Note: All regressions include the full set of control variables, Z_{it} . Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 6605 observations and 1321 clusters. Full results are in Online Appendix Table 13.

Table 5: Change in emergency department visits, with respect to after-hours services

	Total visits		Nonu	rgent	Urg	ent	Very urgent	
	(1)		(2	2)	(3	3)	(4)	
	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	$\operatorname{Std.}$ Err.	Coef.	Std. Err.
OLS	-0.00280	0.00027	-0.00172	0.00020	-0.00082	0.00010	-0.00006	0.00006
FE	0.00019	0.00015	-0.00127	0.00025	0.00018	0.00006	0.00006	0.00004

Note: All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 6605 observations and 1321 clusters.

Table 6: IV Estimate of effect of after-hours services on ED visits

	Total visits		Nonu	rgent	Ur	gent	Very urgent		
	(1)		(2)		(3)		(4)		
	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	
After-hours costs	0.00246	0.00134	-0.00432	0.00165	0.00329	0.00069	0.00090	0.00036	
F-stat	41.20								

Note: The Stata command xtivreg2 was used to estimate these regressions. All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Standard errors are clustered at physician level; there are 6605 observations and 1321 clusters.

Table 7: IV estimate of effect of after-hours services on ED costs

	Total cost		Nonurg	gent cost	Urgent cost		Very u	rgent cost
	(1)		((2)		(3)		(4)
	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.
After-hours costs	-0.1979	0.2452	-0.6442	0.1908	0.2258	0.1202	0.2196	0.1320
F-stat	41.20							

Note: "very urgent", "urgent", and "nonurgent" ED categories respectively correspond to CTAS groups 1, 2, and 3. The Stata command xtivreg2 was used to estimate these regressions. All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Standard errors are clustered at physician level; there are 6605 observations and 1321 clusters.

Table 8: Changes in ED visits and costs, with respect to after-hours services

		igh		. Err.	00 047	16627
		Avg. ADG high	(8)	Std	0.0 90	96 0.0
	Very urgent	Avg. A		Coef.	-0.000 00	-0.00299
,	Very	G low	_	td. Err.	0.000.0	0.02239
		Avg. ADG low	(7)	Coef. S	0.000.0	0.01833
		Avg. ADG low Avg. ADG high		td. Err.	70 000.0	0.02129
		vg. AD0	(9)	sef. S) 90 00) 89 40
	Urgent	A		č Z	1 0.00	1 0.00
,	\supset	DG low	5)	Std. Err	0.00014	0.03354
		Avg. A	<u> </u>	Coef.	0.00049	0.08699
	d high		Std. Err.	0.00031	0.01553	
-	gent	Avg. ADG high	(4)	Coef. S	-0.00114	0.01807
,	Nonurgent	vg. ADG low		Std. Err.	0.000 52	0.04067
		Avg. A Γ	(3)	Coef. Std. Err.	$ \frac{2D \text{ visits}}{2D \text{ visits}} - 0.00006 \ \ 0.00035 \ \ 0.00019 \ \ 0.00015 \ \ -0.00112 \ \ 0.00052 \ \ -0.00114 \ \ 0.00031 \ \ 0.00049 \ \ 0.00014 \ \ 0.00006 \ \ 0.00007 \ \ 0.00009 \ \ 0.00006 \ \ -0.000006 \ \ 0.000004 \ \ 0.000000 \ \ 0.000000 \ \ 0.000000 \ \ 0.0000000 \ \ 0.0000000 \ \ 0.0000000 \ \ 0.00000000$	-0.10850
		Avg. ADG high	2)	Std. Err.	0.00015	0.03462
	J.	Avg. AI	3	Coef.	0.00019	0.02277
[Total	G low		Std. Err.	0.00035	0.06217
		Avg. ADG low	(1)	Coef.	-0.000.06	-0.00299
					ED visits -	$ED\ costs\ -0.00299\ 0.06217\ 0.02277\ 0.03462\ -0.10850\ 0.04067\ 0.01807\ 0.01553\ 0.08699\ 0.03354\ 0.00768\ 0.02129\ 0.01833\ 0.02239\ -0.002996\ 0.016627$

Note: "Very urgent", "urgent", and "nonurgent" ED categories respectively correspond to CTAS groups 1, 2, and 3. "Avg. ADG low" and "high" respectively refer to practices where mean patient ADG is below and above the sample median mean practice ADG. All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 6605 observations and 1321 clusters.

Table 9: ONLINE APPENDIX: Change in emergency department visits, as function of after-hours premium and other covariates (2007-2013)

	Very urgent	(8)	Coef. St. Err. Coef. St. Err. Coef. St. Err. Coef. St. Err.	$12.505 6.246 -3.985 3.047 \begin{vmatrix} -19.386 & 7.266 & -22.103 & 5.776 & 3.463 & 3.537 & 1.599 & 2.132 \end{vmatrix}$	26,229
	Jrgent	7)	St. Err. (3.537	
FE	Urg		Coef.	3.463	26,229
	Nonurgent		St. Err.	5.776	
	Nonu	9)	Coef.	-22.103	26,229
	visits		St. Err.	7.266	
	Total visits	(2)	Coef.	-19.386	26,229
	ırgent	(1	St. Err.	3.047	
	ent Very urgent	5)	Coef.	-3.985	26,229
			Coef. St. Err. Coef. St. Err.	6.246	
S	Urgen	(3	Coef.	-12.505	26,229
OLS	Vonurgent		St. Err.	11.186	
	Nonu	(7)	Coef. St. Err. Coef. St. Err.	-47.570	26,229
	visits		St. Err.	15.930	
	Total visits	(1	Coef.	-66.367	26,229
				Premium -66.367 15.930 -47.570 11.186	Obs.

Note: "Very urgent", "urgent", and "nonurgent" ED categories respectively correspond to CTAS groups 1, 2, and 3. All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 26229 observations and 3747 clusters.

Table 10: ONLINE APPENDIX: Fixed-effects regressions of regular- and after-hours services on after-hours premium and other characteristics (2007-2013)

Afte	After-hours visits		After-hou	After-hours costs		visits	Total costs		
	(1)		(2	(2)		(3)		(4)	
Co	ef. Sto	d. Err.	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	
Premium 25.1	14 3	8.243	-1969.060	1134.828	-1108.476	185.159	-60146	6361.081	

Note: All regressions include the full set of control variables, Z_{it} . Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 26229 observations and 3747 clusters.

Table 11: ONLINE APPENDIX: Change in emergency department visits, with respect to after-hours services (2007-2013)

	Total visits		Nonu	rgent	Urg	ent	Very urgent	
	(1)		(2)		(3)		(4)	
	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.
OLS	-0.00220	0.00018	-0.00106	0.00009	-0.00086	0.00007	-0.00028	0.00003
FE	-0.00028	0.00008	-0.00021	0.00006	-0.00008	0.00004	0.00001	0.00002

Note: "Very urgent", "urgent", and "nonurgent" ED categories respectively correspond to CTAS groups 1, 2, and 3. All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 26229 observations and 3747 clusters.

Table 12: ONLINE APPENDIX: Changes in ED costs, with respect to after-hours services (2007-2013)

Tota	Total cost		Nonurgent cost		Urgent cost		Very urgent cost	
((1)	(:	2)	(3)		(4)	
Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	
OLS -0.4254	0.0361	-0.1507	0.0137	-0.1845	0.0167	-0.0894	0.0100	
FE -0.0376	0.0155	-0.0356	0.0088	-0.0123	0.0089	0.0106	0.0068	

Note: "Very urgent", "urgent", and "nonurgent" ED categories respectively correspond to CTAS groups 1, 2, and 3. All regressions include the full set of control variables, Z_{it} . After-hours services are measured using deflated after-hours costs. Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 26229 observations and 3747 clusters.

Table 13: ONLINE APPENDIX: Fixed-effects regressions of regular- and after-hours services on after-hours premium and other characteristics

	After-hours visits (1)		After-hours costs (2)		Regular visits (3)		Total costs (4)	
	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.	Coef.	Std. Err.
Premium	436.649	69.423	13 166.930	2048.579	-2223.516	244.539	-58298.540	10 185.040
Year	8.000	12.475	374.644	363.682	-392.729	56.335	-15179.900	2220.844
Age sq.	-0.132	0.120	-4.175	3.515	2.422	0.538	85.532	20.591
Group size	0.432	0.126	13.230	3.866	0.037	0.458	18.283	21.566
Avg. age	-9.187	3.898	-262.267	116.235	41.953	29.158	1736.373	1079.795
Avg. ADG	-15.247	23.388	-473.423	731.220	268.685	139.306	7911.881	5508.647
Avg. deprived	5.114	0.685	133.383	19.264	-21.123	2.768	-588.098	91.350
Avg. rural	-3.938	1.517	-111.980	44.178	12.102	6.401	553.658	306.966
Constant	-15088.580	24719.630	-722637.500	720612.300	782735.400	111712.100	3.030e7	4402387
Obs.	6605		6605		6605		6605	

Note: Fixed-effects are at physician level. Standard errors are clustered at physician level; there are 6605 obesrvations and 1321 clusters.