ECON675: Assignment 2

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1 Question 1: Kernel Density Estimation

1.1 Density derivatives

I follow the derivation in Hansen's notes. We are interested in estimating

$$f^{(s)}(x) = \frac{d^s}{dx^s} f(x).$$

The natural estimator is

$$\hat{f}^{(s)}(x) = \frac{d^s}{dx^s} \hat{f}(x)$$

Now, we know that $\hat{f}(x) = \frac{1}{nh} \sum_{i} K\left(\frac{X_i - x}{h}\right)$. Thus,

$$\hat{f}^{(1)}(x) = \frac{-1}{nh^2} \sum_{i=1}^n K^{(1)} \left(\frac{X_i - x}{h} \right),$$

$$\hat{f}^{(2)}(x) = \frac{1}{nh^3} \sum_{i=1}^n K^{(2)} \left(\frac{X_i - x}{h} \right),$$

:

$$\hat{f}^{(s)}(x) = \frac{(-1)^s}{nh^{1+s}} \sum_{i=1}^n K^{(s)} \left(\frac{X_i - x}{h} \right).$$

Now,

$$\mathbb{E}[\hat{f}^{(s)}(x)] = \frac{1}{n} \sum_{i=1}^{n} \mathbb{E}\left[\frac{(-1)^{s}}{h^{1+s}} K^{(s)} \left(\frac{X_{i} - x}{h}\right)\right]$$

$$= \mathbb{E}\left[\frac{(-1)^{s}}{h^{1+s}} K^{(s)} \left(\frac{X_{i} - x}{h}\right)\right], \text{ since } X_{i} \text{ are iid.}$$

$$= \int_{-\infty}^{\infty} \frac{(-1)^{s}}{h^{1+s}} K^{(s)} \left(\frac{z - x}{h}\right) f(z) dz$$

Next, we want to use integration by parts: $\int u dv = uv - \int v du$. Define

$$dv = \frac{(-1)^s}{h^s} \frac{1}{h} K^{(s)} \left(\frac{z - x}{h} \right) \implies v = \frac{(-1)^s}{h^s} K^{(s-1)} \left(\frac{z - x}{h} \right)$$

And

$$u = f(z) \implies du = f^{(1)}(z)$$

Thus,

$$\mathbb{E}[\hat{f}^{(s)}(x)] = \left[\frac{(-1)^s}{h^s} K^{(s-1)} \left(\frac{z-x}{h}\right) f^{(1)}(z)\right]_{-\infty}^{\infty} - \int_{-\infty}^{\infty} \frac{(-1)^s}{h^s} K^{(s-1)} \left(\frac{z-x}{h}\right) f^{(1)}(z) dz.$$

$$= -\int_{-\infty}^{\infty} \frac{(-1)^s}{h^s} K^{(s-1)} \left(\frac{z-x}{h}\right) f^{(1)}(z) dz.$$

Repeating this s times give

$$\mathbb{E}[\hat{f}^{(s)}(x)] = (-1)^s \int_{-\infty}^{\infty} \frac{(-1)^s}{h} K\left(\frac{z-x}{h}\right) f^{(s)}(z) dz$$
$$= \int_{-\infty}^{\infty} \frac{1}{h} K\left(\frac{z-x}{h}\right) f^{(s)}(z) dz$$

Next, use the following change of variables: $u = \frac{z-x}{h}$, which implies $z = x + hu \implies dz = hdu$. Thus,

$$\mathbb{E}[\hat{f}^{(s)}(x)] = \int_{-\infty}^{\infty} K(u)f^{(s)}(x+hu)du \tag{1}$$

The next step is to take a Taylor expansion of $f^{(s)}(x + hu)$ around x + hu = x, which is valid if $h \to 0$. We get

$$f^{(s)}(x+hu) = f^{(s)}(x) + f^{(s+1)}(x)hu + \frac{1}{2}f^{(s+2)}(x)h^2u^2 + \dots + \frac{1}{P!}f^{(s+P)}(x)h^Pu^P + o(h^P).$$

Substituting this expression back into (1), integrating over each term, and using the fact that $\int_{-\infty}^{\infty} K(u)du = 1$ and the notation

$$\mu_{\ell}(K) = u^{\ell}K(u)$$

gives

$$\mathbb{E}[\hat{f}^{(s)}(x)] = f^{(s)}(x) + f^{(s+1)}(x)h\mu_1(K) + \frac{1}{2}f^{(s+2)}(x)h^2\mu_2(K) + \dots + \frac{1}{P!}f^{(s+P)}(x)h^P\mu_P(K) + o(h^P).$$

Finally, noting that since K is a P-order kernel, $\mu_{\ell}(K) = 0$ for all $\ell < P$, gives the desired result

$$\mathbb{E}[\hat{f}^{(s)}(x)] = f^{(s)}(x) + \frac{1}{P!} f^{(s+P)}(x) h^P \mu_P(K) + o(h^P). \tag{2}$$

Next we consider the variance of the derivative estimator.

$$\mathbb{V}[\hat{f}^{(s)}(x)] = \mathbb{V}\left[\frac{(-1)^s}{nh^{1+s}} \sum_{i=1}^n K^{(s)}\left(\frac{X_i - x}{h}\right)\right]$$
$$= \frac{1}{nh^{2+2s}} \mathbb{V}\left[K^{(s)}\left(\frac{X_i - x}{h}\right)\right],$$

since $\{X_i\}$ are iid there are no covariance terms and each term has the same variance. Continuing,

$$\mathbb{V}[\hat{f}^{(s)}(x)] = \frac{1}{nh^{2+2s}} \left\{ \mathbb{E}\left[K^{(s)}\left(\frac{X_i - x}{h}\right)^2\right] - \mathbb{E}\left[K^{(s)}\left(\frac{X_i - x}{h}\right)\right]^2 \right\}$$

$$= \frac{1}{nh^{2+2s}} \mathbb{E}\left[K^{(s)}\left(\frac{X_i - x}{h}\right)^2\right] - \frac{1}{n} \mathbb{E}\left[\frac{1}{h^{1+s}}K^{(s)}\left(\frac{X_i - x}{h}\right)\right]^2 \tag{3}$$

Now, from above we know that

$$\mathbb{E}\left[\frac{1}{h^{1+s}}K^{(s)}\left(\frac{X_i - x}{h}\right)\right] = f^{(s)}(x) + \frac{1}{P!}f^{(s+P)}(x)h^P\mu_P(K) + o(h^P)$$
$$= f^{(s)}(x) + o(1)$$

since the remainder goes to zero as $h \to 0$. Thus, the second term in (3) is $O(\frac{1}{n})$; i.e. the same order as 1/n. Furthermore $O(\frac{1}{n})$ is of smaller order than $O(\frac{1}{nh^{1+2s}})$ since $h \to 0$ and $n \to \infty$. Accordingly, we can write

$$\mathbb{V}[\hat{f}^{(s)}(x)] = \frac{1}{nh^{2+2s}} \mathbb{E}\left[K^{(s)} \left(\frac{X_i - x}{h}\right)^2\right] + o\left(\frac{1}{nh^{1+2s}}\right),$$

Thus,

$$\mathbb{V}[\hat{f}^{(s)}(x)] = \frac{1}{nh^{1+2s}} \int_{-\infty}^{\infty} \frac{1}{h} K^{(s)} \left(\frac{z-x}{h}\right)^2 f(z) dz + o\left(\frac{1}{nh^{1+2s}}\right)$$

Again we use the change of variables $u = \frac{z-x}{h}$ so that

$$\mathbb{V}[\hat{f}^{(s)}(x)] = \frac{1}{nh^{1+2s}} \int_{-\infty}^{\infty} K^{(s)}(u)^2 f(x+hu) du + o\left(\frac{1}{nh^{1+2s}}\right)$$

With the usual Taylor expansion of f(x + hu) we can write

$$V[\hat{f}^{(s)}(x)] = \frac{1}{nh^{1+2s}} \int_{-\infty}^{\infty} K^{(s)}(u)^{2} (f(x) + O(h)) du + o\left(\frac{1}{nh^{1+2s}}\right)$$

$$= \frac{f(x)}{nh^{1+2s}} \int_{-\infty}^{\infty} K^{(s)}(u)^{2} du + o\left(\frac{1}{nh^{1+2s}}\right)$$

$$= \frac{1}{nh^{1+2s}} f(x) \vartheta_{s}(K) + o\left(\frac{1}{nh^{1+2s}}\right),$$

where $\vartheta_s(K) = \int_{-\infty}^{\infty} K^{(s)}(u)^2 du$ as required.

1.2 Optimal bandwidth

We have

$$AIMSE[h] = \int_{-\infty}^{\infty} \left[\left(h^P \mu_P(K) \cdot \frac{f^{(P+s)}(x)}{P!} \right)^2 + \frac{1}{nh^{1+2s}} \vartheta_s(K) f(x) \right] dx$$
$$= h^{2P} \left(\frac{\mu_P(K)}{P!} \right)^2 \vartheta_{s+P}(f) + \frac{1}{nh^{1+2s}} \vartheta_s(K),$$

since f(x) integrates to 1 and where $\vartheta_{s+P}(f) = \int (f^{(P+s)}(x))^2 dx$. Thus,

$$\frac{d}{dh} \text{AIMSE}[h] = 2Ph^{2P-1} \left(\frac{\mu_P(K)}{P!}\right)^2 \vartheta_{s+P}(f) - (1+2s) \frac{1}{nh^{2+2s}} \vartheta_s(K) = 0$$

$$\implies 2Ph^{1+2P+2s} \left(\frac{\mu_P(K)}{P!}\right)^2 \vartheta_{s+P}(f) = (1+2s) \frac{1}{n} \vartheta_s(K),$$

which gives the optimal bandwidth

$$h^* = \left[\frac{1+2s}{2Pn} \left(\frac{P!}{\mu_P(K)}\right)^2 \frac{\vartheta_s(K)}{\vartheta_{s+P}(f)}\right]^{\frac{1}{1+2P+2s}}.$$

A fully data-driven method for estimating h^* is cross-validation. This procedure attempts to directly estimate the mean-squared error, and then choose the bandwidth which minimizes this estimate. From the lecture notes the cross-validation bandwidth is the value h which minimizes the criteria

$$\hat{h}_{CV} = \arg\min_{h} CV(h) = \frac{1}{n^2 h} \sum_{i=1}^{n} \sum_{j=1}^{n} (K * K) \left(\frac{X_i - X_j}{h} \right) - \frac{2}{n} \sum_{i=1}^{n} \hat{f}_{(i)}(X_i)$$

where $\hat{f}_{(i)}(x_i)$ is the density estimate computed without observation X_i .

1.3 Monte Carlo experiment

(a) First, we want to compute the theoretically optimal bandwidth for s = 0, n = 1000, using the Epanechnikov kernel (P = 2), with the following Gaussian DGP:

$$x_i \sim 0.5\mathcal{N}(-1.5, -1.5) + 0.5\mathcal{N}(1, 1)$$

From Table 1 in Hansen's notes, $\mu_2(K) = 1/5$ and $\vartheta(K) = 3/5$ for the Epanechnikov kernel. Thus, the only other ingredient we need is $\vartheta_2(f) = \int [f^{(2)}(x)]^2 dx$ for the above DGP.