

STA258: Statistics with Applied Probability

Formula Sheet for Final Exam

Probability Distributions

Distribution	Distribution function	Support	Mean	Variance	MGF
Bernoulli	$f(x) = p^x(1-p)^{1-x}$	$x = 0, 1$	p	$p(1-p)$	$1 - p + pe^t$
Binomial	$f(x) = \binom{n}{x} p^x (1-p)^{n-x}$	$x = 0, 1, \dots, n$	np	$np(1-p)$	$(1 - p + pe^t)^n$
Poisson	$f(x) = \frac{\lambda^x e^{-\lambda}}{x!}$	$x = 0, 1, 2, \dots$	λ	λ	$e^{\lambda(e^t - 1)}$
Normal	$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$-\infty \leq x \leq +\infty$	μ	σ^2	$e^{\mu t + \frac{1}{2}\sigma^2 t^2}$
Gamma	$f(x) = \frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} e^{-\frac{x}{\beta}}$	$x > 0$	$\alpha\beta$	$\alpha\beta^2$	$(1 - \beta t)^{-\alpha}, t < \frac{1}{\beta}$
Chi-Square	$f(x) = \frac{1}{2^{k/2}\Gamma(k/2)} x^{(k/2)-1} e^{-x/2}$	$x > 0$	k	$2k$	$(1 - 2t)^{-k/2}, t < \frac{1}{2}$

Transformations

Let X be a normally distributed random variable with a mean μ and variance σ^2 .

$$Z = \frac{x - \mu}{\sigma}$$

Let \bar{x} represent the sample mean of n independent observations from a distribution with finite mean μ and finite variance σ^2 . For normal populations or sufficiently large n :

$$Z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}}$$

Let \bar{x} and s^2 represent the sample mean and sample variance of n independent observations from a distribution with finite mean μ and some unknown finite variance. For normal populations or sufficiently large n :

$$T = \frac{\bar{x} - \mu}{s/\sqrt{n}}$$

Continuity Correction for the Normal Approximation

Binomial Probability	Continuity Correction	Normal Approximation
$P(X = x)$	$P(x - 0.5 \leq X \leq x + 0.5)$	$P\left(\frac{x - 0.5 - \mu}{\sigma} \leq Z \leq \frac{x + 0.5 - \mu}{\sigma}\right)$
$P(X \leq x)$	$P(X \leq x + 0.5)$	$P\left(Z \leq \frac{x + 0.5 - \mu}{\sigma}\right)$
$P(X < x)$	$P(X \leq x - 0.5)$	$P\left(Z \leq \frac{x - 0.5 - \mu}{\sigma}\right)$
$P(X \geq x)$	$P(X \geq x - 0.5)$	$P\left(Z \geq \frac{x - 0.5 - \mu}{\sigma}\right)$
$P(X > x)$	$P(X \geq x + 0.5)$	$P\left(Z \geq \frac{x + 0.5 - \mu}{\sigma}\right)$
$P(a \leq X \leq b)$	$P(a - 0.5 \leq X \leq b + 0.5)$	$P\left(\frac{a - 0.5 - \mu}{\sigma} \leq Z \leq \frac{b + 0.5 - \mu}{\sigma}\right)$

Some Relationships Between Random Variables

Let $Z_1, Z_2, \dots, Z_n \stackrel{\text{iid}}{\sim} N(0, 1)$. Then

$$V = \sum_{i=1}^n Z_i^2 \sim \chi_n^2$$

Let $X_1, X_2, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$. Let \bar{x} and S^2 represent the sample mean and variance of X_1, X_2, \dots, X_n respectively. Then

$$\frac{(n-1)S^2}{\sigma^2} = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \bar{X})^2 \sim \chi_{n-1}^2$$

Let $Z \sim N(0, 1)$ and let $W \sim \chi_n^2$. Then

$$T = \frac{Z}{\sqrt{\frac{W}{n}}} \sim t_n$$

Let $X_1, X_2, \dots, X_{n_x} \stackrel{\text{iid}}{\sim} N(\mu_x, \sigma^2)$ with sample variance S_x^2 , and let $Y_1, Y_2, \dots, Y_{n_y} \stackrel{\text{iid}}{\sim} N(\mu_y, \sigma^2)$ with sample variance S_y^2 . with the X_i 's and the Y_j 's all being independent. Then

$$F = \frac{\left(\frac{n_x - 1}{\sigma^2}\right) S_x^2 / (n_x - 1)}{\left(\frac{n_y - 1}{\sigma^2}\right) S_y^2 / (n_y - 1)} \sim F_{n_x - 1, n_y - 1}$$

Inference Procedures for the Population Mean μ

One-Sample

Inference Procedure for μ when σ is known

Confidence interval for μ :

$$\bar{x} \pm Z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

To test $H_0 : \mu = \mu_0$, test statistic is:

$$Z^* = \frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}}$$

Inference Procedure for μ when σ is not known

Confidence interval for μ :

$$\bar{x} \pm t_{(\alpha/2, n-1)} \frac{s}{\sqrt{n}}$$

To test $H_0 : \mu = \mu_0$, test statistic is:

$$t^* = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$$

Two-Sample

Inference Procedure for Difference between μ_1 and μ_2 when σ_1 and σ_2 are known

Confidence interval for $\mu_1 - \mu_2$:

$$(\bar{x}_1 - \bar{x}_2) \pm Z_{\alpha/2} \sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$$

To test $H_0 : \mu_1 - \mu_2 = \mu_D$, test statistic is:

$$Z^* = \frac{(\bar{x}_1 - \bar{x}_2) - \mu_D}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

Inference Procedure for Difference between μ_1 and μ_2 when σ_1 and σ_2 are not known

Case 1: Unequal Variances

Confidence interval for $\mu_1 - \mu_2$:

$$(\bar{x}_1 - \bar{x}_2) \pm t_{(\alpha/2, d)} \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}$$

To test $H_0 : \mu_1 - \mu_2 = \mu_D$, test statistic is:

$$t^* = \frac{(\bar{x}_1 - \bar{x}_2) - \mu_D}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}}$$

where d is the smaller of $n_1 - 1$ and $n_2 - 1$.

Case 2: Equal Variances

Confidence interval for $\mu_1 - \mu_2$:

$$(\bar{x}_1 - \bar{x}_2) \pm t_{(\alpha/2, n_1+n_2-2)} s_p \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}$$

To test $H_0 : \mu_1 - \mu_2 = \mu_D$, test statistic is:

$$t^* = \frac{(\bar{x}_1 - \bar{x}_2) - \mu_D}{s_p \sqrt{\left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

Pooled Sample Variance:

$$s_p^2 = \frac{(n_1 - 1)s_1^2 + (n_2 - 1)s_2^2}{n_1 + n_2 - 2}$$

Inference Procedures for the Population Proportion p

One-Sample

Confidence interval for p :

$$\hat{p} \pm Z_{\alpha/2} \sqrt{\frac{\hat{p}(1 - \hat{p})}{n}}$$

To test $H_0 : p = p_0$, test statistic is:

$$Z^* = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1 - p_0)}{n}}}$$

Two-Sample

Confidence interval for $p_1 - p_2$

$$(\hat{p}_1 - \hat{p}_2) \pm Z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1 - \hat{p}_1)}{n_1} + \frac{\hat{p}_2(1 - \hat{p}_2)}{n_2}}$$

To test $H_0 : p_1 - p_2 = p_D$, test statistic is:

$$Z^* = \frac{(\hat{p}_1 - \hat{p}_2) - p_D}{\sqrt{\hat{p}(1 - \hat{p}) \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}$$

where

$$\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$$

ANOVA

For an analysis of k groups:

Source	DF	SS	MS	F
Treatment	$k - 1$	$SSTrt$	$MSTrt = \frac{SSTrt}{k - 1}$	$F = \frac{MSTrt}{MSE}$
Error	$n - k$	SSE	$MSE = \frac{SSE}{n - k}$	—
Total	$n - 1$	$SSTotal$	—	—

Simple Linear Regression

$$\begin{array}{l}
 S_{xx} = \sum_{i=1}^n (x_i - \bar{x})^2 \\
 = \left(\sum_{i=1}^n x_i^2 \right) - \frac{\left(\sum_{i=1}^n x_i \right)^2}{n}
 \end{array}
 \quad \left| \quad
 \begin{array}{l}
 S_{yy} = \sum_{i=1}^n (y_i - \bar{y})^2 \\
 = \left(\sum_{i=1}^n y_i^2 \right) - \frac{\left(\sum_{i=1}^n y_i \right)^2}{n}
 \end{array}
 \quad \left| \quad
 \begin{array}{l}
 S_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) \\
 = \left(\sum_{i=1}^n x_i y_i \right) - \frac{\left(\sum_{i=1}^n x_i \right) \left(\sum_{i=1}^n y_i \right)}{n}
 \end{array}$$

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}} \qquad \bar{y} = \hat{\beta}_1 \bar{x} + \hat{\beta}_0$$

$$\text{Sum Square of Regression (SSR)} = \sum_{i=1}^n (\hat{y}_i - \bar{y})^2$$

$$\text{Sum Square Error (SSE)} = \sum_{i=1}^n (y_i - \hat{y}_i)^2 = S_{yy} - \hat{\beta}_1 S_{xy}$$

$$\text{Sum Square Total } SSTotal = \sum_{i=1}^n (y_i - \bar{y})^2$$

$$\text{Coefficient of Correlation} \quad r = \frac{S_{xy}}{\sqrt{S_{xx} S_{yy}}}$$

$$\text{Coefficient of Determination} \quad r^2 = \frac{SSR}{SSTotal} = 1 - \frac{SSE}{S_{yy}}$$

$$\text{Residuals} \quad e_i = y_i - \hat{y}_i$$

$$\begin{array}{l}
 \text{Estimate of Variance} \\
 \text{of Random Error}
 \end{array}
 \quad s^2 = \frac{\sum_{i=1}^n e_i^2}{n-2} = \frac{\sum_{i=1}^n (y_i - \hat{y}_i)^2}{n-2} = \frac{S_{yy} - \hat{\beta}_1 S_{xy}}{n-2}$$

Inference Procedures on the Slope β_1

Confidence interval for β_1 :

$$\hat{\beta}_1 \pm t_{(\alpha/2, n-2)} \frac{s}{\sqrt{S_{xx}}}$$

To test $H_0 : \beta_1 = 0$, test statistic is:

$$t^* = \frac{\hat{\beta}_1}{s/\sqrt{S_{xx}}}$$