

# PRACTICE EXAM THREE

## SOLUTIONS

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- ① Requirement:  $M \geq 175$  ← requirement satisfied  
 $H_0: M \geq 175$        $H_1: M < 175$  ← requirement not satisfied  
Type I error = Reject  $H_0$  when  $H_0$  is true  
= Conclude not satisfied when it really is satisfied.

Type II error = Accept  $H_0$  when  $H_0$  is false = Conclude is satisfied when it really is not

Type II error is the more serious error because the company would proceed to use/sell Shipment not knowing the requirement is not satisfied. Because it is type II error that is more serious we need to swap  $H_0, H_1$  to assign Type I error as the more serious. What does it mean to swap them

$$M_0 = 175$$

- ②  $H_0: M \leq 175$   
 $H_1: M > 175$  ← requirement is not satisfied  
 $\alpha = 0.05$  Test form C uses a lower bound C.I to do the test. This will be a t-test.  $\bar{x} = 178, n = 20, v = 19$   
 $t_{0.05, 19} = 1.729$   
 $(\bar{X} - t_{0.05, 19} (S.E(\bar{X})), \infty)$

$$S^2 = \frac{\sum_{i=1}^{20} (x_i - \bar{x})^2}{n-1} = \frac{304}{20-1} = 16$$

$$S = \sqrt{16} = 4$$

$$S.e(\bar{x}) = \frac{S}{\sqrt{n}} = \frac{4}{\sqrt{20}} = \frac{4}{4.4721} = 0.8944$$

$$(178 - 1.729)(0.8944), \infty)$$

$$(178 - 1.5464, \infty) = (176.45, \infty)$$

95% confident that  $\mu$  is in the C.I  
since 175 is not in C.I  $\Rightarrow$  Reject  $H_0$

Conclude that requirement is satisfied

c Largest  $\mu_0$  for which we can reject  $H_0$   
has to be outside C.I

d  $T = 4.1$ , we a p-value to test your hypotheses in part a

So you do a Z-test

$$S.e(\bar{x}) = \frac{T}{\sqrt{n}} = \frac{4.1}{\sqrt{20}} = \frac{4.1}{4.4721} = 0.9168$$

$$Z^* = \frac{\bar{x} - \mu_0}{S.e} = \frac{178 - 175}{0.9168} = \frac{3}{0.9168} = 3.27$$

Test form c

$$P = 1 - \Phi(Z^*) = 1 - \Phi(3.27)$$

$$P = 1 - 0.99946 = 0.00054$$

$$\alpha = 0.05 \text{ so } P < \alpha$$

Reject  $H_0$ .

Bushing i	1	2	3	4	5	6	7	8
Student (x)	10.2	9.8	6.5	7.6	8.2	9.9	8.6	9.1
Student (y)	10.0	10.1	6.9	7.5	8.2	9.7	8.3	8.8
d	0.2	-0.3	-0.4	0.1	0	0.2	0.3	0.3

Paired case  $\rightarrow$  same spanner bushing for both students.

$$H_0: \mu_x = \mu_y \Rightarrow \mu_x - \mu_y = 0$$

$$H_1: \mu_x \neq \mu_y \quad \mu_x - \mu_y \neq 0$$

add the difference between x and y

$$\bar{d} = \frac{0.2 - 0.3 - 0.4 + 0.1 + 0 + 0.2 + 0.3 + 0.3}{8}$$

$$\bar{d} = \frac{0.4}{8} = 0.05$$

$$\sum d_i^2 = 0.04 + 0.09 + 0.16 + 0.01 + 0 + 0.04 + 0.09 + 0.09$$

$$s_D^2 = \frac{(\sum d_i^2) - \frac{(\sum d)^2}{n}}{n-1} = \frac{0.52 - \frac{(0.4)^2}{8}}{7} = \frac{0.52 - \frac{0.16}{8}}{7} = \frac{0.52 - 0.02}{7}$$

$$s_D^2 = \frac{0.5}{7} = 0.0714$$

$$s_D = \sqrt{s_D^2} = 0.2673$$

$$S.E\{\bar{d}\} = \frac{s_d}{\sqrt{n}} = \frac{0.2673}{\sqrt{8}} = \frac{0.2673}{2.8284}$$

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$$t^* = \frac{\bar{d} - 0}{S.E} = \frac{0.05}{0.0945} = 0.529$$

$$\alpha = 0.10 \Rightarrow \frac{\alpha}{2} = 0.05 \Rightarrow t_{0.05, 7} = 1.895$$

$$|t^*| = 0.529 < 1.895$$

Accept  $H_0$  or Fail to Reject  $H_0$

It means that you cannot conclude that  $\mu_x$  is different from  $\mu_y$ . It could be the same.

$$b) 95\% C.I., \alpha = 0.05, \alpha/2 = 0.025$$

$$t_{0.025, 7} = 2.365$$

$$\bar{d} \pm (2.365)(S.E(\bar{d}))$$

$$0.05 \pm (2.365)(0.0945)$$

$$0.05 \pm 0.2235$$

$$(-0.1735, 0.273)$$

Lower limit

Upper limit

③  $\mu_1 = \text{before}$      $\mu_2 = \text{after}$

Two means independent samples

i)  $\sigma_1, \sigma_2$  known

$\sigma_1 = \sigma_2$  unknown  $\rightarrow$  Pooled

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This is case 2 because the variances are the same.

Test  $H_1: \mu_1 > \mu_2$

$$H_0: \mu_1 \leq \mu_2 ; H_0: \mu_1 - \mu_2 \leq 0$$

$$H_1: \mu_1 > \mu_2 ; H_1: \mu_1 - \mu_2 > 0$$

This is from C lower bound C.I

Assumed pooled  $\sigma_1 = \sigma_2 = \sigma$  - common estimate the  $S_p = 0.3$

$$\alpha = 0.01 \text{ degree of freedom} = n_1 + n_2 - 2 \\ = 22 + 20 - 2 = 40$$

$$t_{0.01, 40} = 2.423$$

$$((\bar{x}_1 - \bar{x}_2) - t_{0.01, 40} \text{ (se } (\bar{x}_1 - \bar{x}_2), \infty))$$

$$S.e = S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} = 0.3 \times \sqrt{\frac{1}{22} + \frac{1}{20}}$$

$$= 0.0927 \\ ((9.83 - 9.67) - (2.423)(0.0927), \infty) \\ ((0.16) - 0.2246), \infty \\ (-0.0646, \infty)$$

0 is inside C.I so we fail to reject  $H_0$  FTR  $H_0$

$$\textcircled{4} \quad A: n_1 = 10 \quad \bar{x}_1 = 30.1 \quad S_1^2 = 0.17 \\ B: n_2 = 11 \quad \bar{x}_2 = 69.8 \quad S_2^2 = 0.19$$

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$$95\% \text{ C.I} \quad \frac{\sigma_1^2}{\sigma_2^2} \quad \alpha = 0.05 \\ \alpha/2 = 0.025$$

$$\left( \frac{s_1^2}{s_2^2} f_{1-0.025, n_2-1, n_1-1}, \frac{s_1^2}{s_2^2} f_{0.025, n_2-1, n_1-1} \right)$$

$$\left( \frac{s_1^2}{s_2^2} f_{0.975, 10, 9}, \frac{s_1^2}{s_2^2} f_{0.025, 10, 9} \right)$$

$$\left( \frac{s_1^2}{s_2^2} \mid \frac{s_1^2}{s_2^2} f_{0.025, 9, 10}, \frac{s_1^2}{s_2^2} f_{0.025, 10, 9} \right)$$

$$(0.2367, 3.543)$$

$$\text{Testing } H_0: \sigma_1^2 = \sigma_2^2 \Rightarrow \sigma_1^2 / \sigma_2^2 = 1$$

$$H_1: \sigma_1^2 \neq \sigma_2^2 \Rightarrow \sigma_1^2 / \sigma_2^2 \neq 1$$

Since 1 is inside C.I we fail to reject  $H_0$ . So we cannot conclude different

$$\textcircled{5} \quad n = 15, \quad s^2 = 0.08 \\ \text{Claim: } \sigma^2 \leq 0.05$$

$$H_0: \sigma^2 \leq 0.05$$

$$H_1: \sigma^2 > 0.05$$

Test form C, use CI lowerbound

$\chi^2$

$$\left( \frac{(n-1)s^2}{\chi^2_{\alpha, n-1}}, \infty \right) = \left( \frac{(n-1)s^2}{\chi^2_{0.01, 15-1}} , \infty \right)$$

$$\left( \frac{(15-1)(0.08)}{29.141}, \infty \right)$$

$$(0.0384, \infty)$$

We can clearly see that 0.05 is inside CI Fail to reject  $H_0$   
We cannot reject the claim.

$$\therefore \sigma_0^2 = 0.0384$$

The producer is to claim that  $\sigma^2 \leq 0.0384$  still inside the CI

6 We need to know  $S_{xx}$ ,  $S_{yy}$ ,  $S_{xy}$   
 $b_1 = \frac{\sum x_i y_i - \frac{(\sum x_i)(\sum y_i)}{n}}{\sum x_i^2 - \frac{(\sum x_i)^2}{n}}$

$$= \frac{S_{xy}}{S_{xx}}$$

$$\frac{871.77 - \frac{(92.25)(135)}{15}}{589.1875 - \frac{(92.25)^2}{15}}$$

$$= 871.77 - \frac{12453.75}{15}$$

$$589.1875 - \frac{8510.0625}{15}$$

$$\frac{871.77 - 830.25}{589.1875 - 567.3375} = \frac{41.52}{21.85}$$

$$b_1 = 1.9$$

$S_{xx}$

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$$\begin{aligned}
 b_0 &= \frac{1}{n} (\sum y_i - b_1 \sum x_i) \\
 &= \frac{1}{15} (135 - (1.9)(92.25)) \\
 &= \frac{1}{15} (135 - 175.275) \\
 &= \frac{1}{15} (-40.275) \\
 b_0 &= -2.685
 \end{aligned}$$

$$\begin{aligned}
 \hat{y} &= b_0 + b_1 x \\
 \hat{y} &= -2.685 + 1.9 x
 \end{aligned}$$

↳ When you have this type of question  
 it is asking  $R^2$  coefficient of determination, which is the explained

$$R^2 = \frac{SSR}{SST} \text{ or } 1 - \frac{SSE}{SST}$$

$$SSR = b_1^2 S_{xx} \text{ or } b_1 S_{xy}$$

$$\begin{aligned}
 &= (1.9)^2 (21.85) \\
 &= (3.61)(21.85) \\
 &= 78.8785
 \end{aligned}$$

$$SST = Syy = \sum y_i^2 - \frac{(\sum y_i)^2}{n}$$

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$$= 1312.4 - \frac{(135)^2}{15}$$

$$1312.4 - \frac{18225}{15}$$

$$1312.4 - 1215 = 97.4$$

$$R^2 = \frac{78.8785}{97.4} = 0.8098$$

c  $H_0: \beta_1 = 0$

$H_1: \beta_1 \neq 0$  ← Slope is Significant

using test statistic  $t$

$$t^* = \frac{b_1 - 0}{SE}$$

$$SE\{b_1\} = \sqrt{\frac{MSE}{S_{xx}}}$$

$$SSE = SST - SSR = 97.4 - 78.8785 \\ = 18.5215$$

$$MSE = \frac{SSE}{n-2} = \frac{18.5215}{15-2} = 1.4247$$

$$SE\{b_1\} = \sqrt{\frac{1.4247}{21.85}} = 0.255$$

$$t^* = \frac{b_1 - 0}{\text{se}} = \frac{1.9}{0.255} = 7.44$$

I need to compare against the  
 $t_{0.025, 13} = 2.160$

$$|7.44| > 2.160 \rightarrow \text{Reject } H_0$$

Conclude Slope is significant

2  $P = \text{true proportion immunized}$

$$\hat{P} = \frac{1125}{1500} = 0.75$$

two-sided 95% CI for  $p$

$$\alpha = 0.05 \rightarrow \frac{\alpha}{2} = 0.025$$

$$Z_{0.025} = 1.96$$

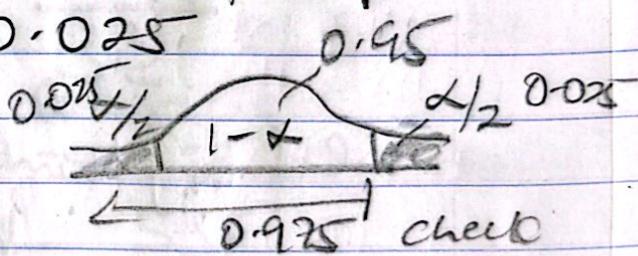


Table A.3 for the 0.975 inside the Z-side.

$$\hat{P} \pm Z_{0.025} \sqrt{\frac{\hat{P}(1-\hat{P})}{n}}$$

$$0.75 \pm (1.96) \sqrt{\frac{(0.75)(1-0.75)}{1500}}$$

$$0.75 \pm (1.96) \sqrt{\frac{0.1875}{1500}}$$

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$$0.75 \pm (1.96) \sqrt{0.000125}$$

$$0.75 \pm (1.96) (0.0118)$$

$$0.75 \pm 0.02191$$

$$(0.728, 0.772)$$

$$\text{C.I length} \leq 0.03 \Rightarrow e = \frac{0.03}{2} \\ = 0.015$$

$$n \geq \left( \frac{\Sigma 0.025}{e} \right)^2 \hat{p} (1 - \hat{p})$$

$$= \left( \frac{1.96}{0.015} \right)^2 (0.75) (0.25)$$

$$= (130.66)^2 (0.75) (0.25)$$

$$= (17073.78) (0.75) (0.25)$$

$$= 3201.33$$

round up to 3202.

For Cleveland we have no data  
we will use  $\hat{p} = 0.5$

$$n \geq \left( \frac{1.96}{0.015} \right)^2 (0.5) (0.5)$$

$$\approx (17073.78) (0.5)^2$$

$$\approx 4268.45$$

$$\approx 4269$$

b)  $P_B = \text{Baltimore}$   
 $P_C = \text{Cleveland}$

$$\hat{P}_B = 0.75$$

$$\hat{P}_C = \frac{1960}{1200} = 0.8$$

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Test  $P_B < P_C$

$$H_0: P_B \geq P_C$$

$$H_1: P_B < P_C \leftarrow \text{Baltimore} < \text{Cleveland}$$

$$\text{use } \hat{P} = \frac{1125 + 960}{1500 + 1200} = 0.7722$$

$$\alpha = 0.01 \rightarrow Z_{0.01} = 2.326$$

using a test statistic so we need  $Z^*$

$$Z^* = \frac{\hat{P}_B - \hat{P}_C}{\text{se}}$$

$$\begin{aligned} \text{se}\{\hat{P}_B - \hat{P}_C\} &= \sqrt{\hat{P}(1-\hat{P})\left(\frac{1}{n_1} + \frac{1}{n_2}\right)} \\ &= \sqrt{(0.7722)(0.2278)} \left( \frac{1}{1500} + \frac{1}{1200} \right) \\ &= 0.01624 \end{aligned}$$

$$\begin{aligned} Z^* &= \frac{\hat{P}_B - \hat{P}_C}{\text{se}} = \frac{0.75 - 0.8}{0.01624} \\ &= -3.08 \end{aligned}$$

Reject  $H_0$  if  $Z^* < -Z_{0.01}$

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$-3.08 < -2.326 \Rightarrow$  Reject  $H_0$   
We conclude that  $p_B < p_C$

ii What is the p-value of your test?

$$P = \Phi(2^*) = \Phi(-3.08)$$

$$= 0.00103$$