

## Bias in estimating expected ratio of two gamma variables

### Problem

Suppose  $X \sim \Gamma(\lambda_x, \alpha_x)$  and  $Y \sim \Gamma(\lambda_y, \alpha_y)$ . That is,

$$\begin{aligned} f_X(x) &= \frac{1}{\Gamma(\alpha_x)} \cdot x^{\alpha_x-1} e^{-\lambda_x x} \lambda_x^{\alpha_x} \\ f_Y(y) &= \frac{1}{\Gamma(\alpha_y)} \cdot y^{\alpha_y-1} e^{-\lambda_y y} \lambda_y^{\alpha_y} \end{aligned}$$

We also take  $X$  and  $Y$  independent.

We want to estimate  $\theta = E(X)/E(Y)$ .

We define as estimator  $T = \bar{X}/\bar{Y}$ .

### The exact value of the expected ratio of two Gamma r.v.'s

If  $X \sim \Gamma(\lambda_x, \alpha_x)$  and  $Y \sim \Gamma(\lambda_y, \alpha_y)$ , then we can compute  $\mu = E(X/Y)$  exactly. Indeed, because of independency

$$\mu = E(X) \cdot E(1/Y)$$

and

$$\begin{aligned} E\left(\frac{1}{Y}\right) &= \int_0^\infty \frac{1}{y} f_Y(y) dy \\ &= \int_0^\infty \frac{1}{\Gamma(\alpha_y)} \cdot y^{\alpha_y-2} e^{-\lambda_y y} \lambda_y^{\alpha_y} dy \\ &= \frac{\lambda_y}{\alpha_y - 1} \cdot \int_0^\infty \frac{1}{\Gamma(\alpha_y - 1)} \cdot y^{\alpha_y-2} e^{-\lambda_y y} \lambda_y^{\alpha_y-1} dy \\ &= \frac{\lambda_y}{\alpha_y - 1} \cdot 1 \end{aligned}$$

The last integral must be equal to one, as the integrand is the density of a  $\Gamma(\lambda_y, \alpha_y - 1)$  variable.

Hence

$$\mu = \frac{\alpha_x}{\lambda_x} \cdot \frac{\lambda_y}{\alpha_y - 1}$$

Note that  $\mu \neq \theta$ .

### The distributions of the sample means

If  $X \sim \Gamma(\lambda_x, \alpha_x)$ , then  $n_x \bar{X} \sim \Gamma(\lambda_x, n_x \alpha_x)$  and  $\bar{X} \sim \Gamma(n_x \lambda_x, n_x \alpha_x)$

We thus have

$$E(T) = \frac{\alpha_x}{\lambda_x} \cdot \frac{n_y \lambda_y}{n_y \alpha_y - 1} = \frac{\alpha_x}{\lambda_x} \cdot \frac{\lambda_y}{\alpha_y - 1/n_y}$$

The bias is than

$$\begin{aligned}
 E(T) - \theta &= \frac{\alpha_x}{\lambda_x} \cdot \frac{\lambda_y}{\alpha_y - 1/n_y} - \frac{\alpha_x}{\lambda_x} \cdot \frac{\lambda_y}{\alpha_y} = \frac{\alpha_x}{\lambda_x} \cdot \frac{\lambda_y}{\alpha_y} \cdot \left( \frac{1}{1 - 1/(n_y \alpha_y)} - 1 \right) \\
 &= \theta \cdot \frac{1}{n_y \alpha_y - 1}
 \end{aligned}$$