

Department of Statistics and Data Science at SUSTech

MAT7035: Computational Statistics

Tutorial 8: Monte Carlo Integration

A. Numerical integration methods

A.1 Aim

- Suppose that we are interested in evaluating the integral

$$\mu = E\{h(X)\} = \int_{\mathbb{X}} h(x) \cdot f(x) \, dx,$$

where $h(\cdot) \geq 0$ is a function and $f(\cdot)$ is the pdf of a r.v. X with support \mathbb{X} .

A.2 Classical Monte Carlo integration

- Let $\{X^{(i)}\}_{i=1}^m \stackrel{\text{iid}}{\sim} f(\cdot)$, then

$$\bar{\mu}_m = \frac{1}{m} \sum_{i=1}^m h(X^{(i)})$$

is called the (classical) Monte Carlo integration of μ .

A.3 Riemannian simulation

- Let $\{X^{(i)}\}_{i=1}^{m+1} \stackrel{\text{iid}}{\sim} f(\cdot)$, $\{X_{(i)}\}_{i=1}^{m+1}$ are the order statistics of $\{X_{(i)}\}_{i=1}^{m+1}$, then

$$\hat{\mu}^R = \sum_{i=1}^m h(X_{(i)}) f(X_{(i)}) [X_{(i+1)} - X_{(i)}]$$

is called the Riemannian sum estimator of μ .

A.4 The importance sampling method

— Let $H(x) \triangleq h(x)f(x)$ be defined on \mathbb{X} .

- If we could find an easy-sampling density function $g(\cdot)$ with support \mathbb{X} , we can write

$$\mu = \int_{\mathbb{X}} H(x) \, dx = \int_{\mathbb{X}} \frac{H(x)}{g(x)} \cdot g(x) \, dx = \int_{\mathbb{X}} w(x) \cdot g(x) \, dx,$$

where $w(x) \triangleq H(x)/g(x)$ is called ratio function.

- Let $X^{(1)}, \dots, X^{(m)} \stackrel{\text{iid}}{\sim} g(x)$, then μ can be estimated by

$$\tilde{\mu}_m = \frac{1}{m} \sum_{i=1}^m \mathbf{w}(X^{(i)}),$$

which is called the *importance sampling* (IS) estimator.

Example T8.1 (Classical Monte Carlo integration method). Let $\gamma = -0.5$. Use the classical Monte Carlo integration method to compute

$$\mu = \frac{1}{\sqrt{\pi}} \int_{-\infty}^{\infty} \frac{1 + e^x}{1 + \gamma} e^{-0.5x^2} \, dx.$$

Solution: Let

$$h(x) = \sqrt{2} \frac{1 + e^x}{1 + \gamma},$$

then

$$\mu = \int_{-\infty}^{\infty} h(x) \cdot \phi(x) \, dx \quad \text{and} \quad \bar{\mu}_m = \frac{1}{m} \sum_{i=1}^m h(X^{(i)}),$$

where $\{X^{(i)}\}_{i=1}^m$ are i.i.d. samples from the standard normal density $\phi(x)$.

R code:

```
> MC.in1 <- function(m)
{   # Function name: MC.in1
    # ----- Input -----
    # m = the number of samples
    # ----- Onput -----
    # mu = numerical integration by
    #      Classical MC method
    # -----
```

```

g <- -0.5
x <- rnorm(m,0,1)
h <- sqrt(2)*(1+exp(x))/(1+g)
mu<- mean(h)
return(mu)
}

```

Results:

m	$\bar{\mu}_m$
10^3	7.306
10^5	7.511
10^7	7.493

Example T8.2 (Classical Monte Carlo integration and Riemannian simulation). Use the classical Monte Carlo integration and Riemannian simulation to compute

$$\mu = \int_0^1 \frac{4}{1+x^2} dx.$$

Solution: Let $h(x) = 4(1+x^2)^{-1}$ and $f(x) = 1$ for $x \in (0, 1)$, then the integral is rewritten as

$$\mu = \int_0^1 h(x)f(x) dx.$$

We have

$$\bar{\mu}_m = \frac{1}{m} \sum_{i=1}^m h(X^{(i)}) \quad \text{and} \quad \mu_m^R = \sum_{i=1}^{m-1} h(X_{(i)})f(X_{(i)})[X_{(i+1)} - X_{(i)}],$$

where $\{X^{(i)}\}_{i=1}^m \stackrel{\text{iid}}{\sim} U(0, 1)$ and $\{X_{(i)}\}_{i=1}^m$ are the order statistics of $\{X^{(i)}\}_{i=1}^m$.

R code:

```

> MC.in2 <- function(m)
{
  # Function name: MC.in2
  # ----- Input -----
  # m      = the number of samples

```

```

# ----- Onput -----
# m      = (m[1],m[2])
# m[1] = Classical MC Method
# m[2] = Riemannian Simulation
# -----
mu <- rep(0, 2)
x  <- runif(m,0,1)
x  <- sort(x)
h  <- 4/(1+x^2)
mu[1] <- 1/m*sum(h)
mu[2] <- sum(h[2:m]*diff(x))
return(mu)
}

```

Results:

m	$\bar{\mu}_m$	μ_m^R
10^3	3.15126	3.12872
10^5	3.13665	3.14124
10^7	3.14049	3.14152

Remark: For an integral in a finite interval, the uniform distribution on the interval, i.e., a constant pdf $f(\cdot)$ is a good choice.

Example T8.3 (Classical Monte Carlo integration and the importance sampling method).
Compute the value of normal cdf

$$\mu = \Phi(t) = \int_{-\infty}^t \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx,$$

for $t = 0$ by the classical Monte Carlo integration and the importance sampling method with the standard Cauchy density $g(x) = [\pi(1 + x^2)]^{-1}$ as the proposal density.

Solution: (a) Note that

$$\mu = \int_{-\infty}^{\infty} I(x \leq t) \cdot \phi(x) dx,$$

where $I(\cdot)$ is the indicator function, and $\phi(x)$ is the pdf of the standard normal distribution.

Then the classical Monte Carlo estimator is

$$\bar{\mu}_m = \frac{1}{m} \sum_{i=1}^m I(X^{(i)} \leq t),$$

where $\{X^{(i)}\}_{i=1}^m \stackrel{\text{iid}}{\sim} N(0, 1)$.

(b) Use the importance sampling method, we have

$$\begin{aligned} \mu &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} I(x \leq t) dx \\ &= \int_{-\infty}^{\infty} \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} I(x \leq t)}{g(x)} \cdot g(x) dx \\ &= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} I(x \leq t) \pi(1+x^2) \cdot g(x) dx \\ &= \sqrt{\frac{\pi}{2}} \int_{-\infty}^{\infty} e^{-\frac{x^2}{2}} (1+x^2) I(x \leq t) \cdot g(x) dx, \end{aligned}$$

so that

$$\tilde{\mu}_m = \frac{1}{m} \sqrt{\frac{\pi}{2}} \sum_{i=1}^m e^{-\frac{(X^{(i)})^2}{2}} [1 + (X^{(i)})^2] I(X^{(i)} \leq t),$$

where $\{X^{(i)}\}_{i=1}^m$ are i.i.d. samples from the standard Cauchy density. We choose $t = 0$, which simply makes $\Phi(t) = 0.5$, to test the algorithms.

[R code:](#)

```
> MC.in3 <- function (m)
{ # Function name: MC.in3
  # ----- Input -----
  # m = the number of samples
  # ----- Output -----
  # mu = numerical integration by
  #     Classical MC Method
  # -----
  t <- 0
```

```

      x <- rnorm(m,0,1)
      h <- (x<=t)
      mu <- mean(h)
      return(mu)
}

> MC.in4 <- function (m)
{   # Function name: MC.in4
    # ----- Input -----
    # m = the number of samples
    # ----- Onput -----
    # mu = numerical integration by
    #      Importance Sampling Method
    # -----
    t <- 0
    x <- rcauchy(m,0,1)
    w <- exp(-x^2/2)*(1+x^2)*(x<=t)
    mu <- sqrt(pi/2)*mean(w)
    return(mu)
}

```

Results:

m	$\bar{\mu}_m$	$\tilde{\mu}_m$
10^3	0.5170	0.4828
10^5	0.5043	0.4988
10^7	0.5001	0.4996