4 Random Vector and Matrices

- Expectation: Let **Y** and **X** be $p \times 1$ random vectors. The expected value of

$$\boldsymbol{Y} = \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_p \end{pmatrix} \text{ is given by } E(\boldsymbol{Y}) = \begin{pmatrix} E(Y_1) \\ E(Y_2) \\ \vdots \\ E(Y_p) \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \\ \vdots \\ \mu_p \end{pmatrix} = \boldsymbol{\mu}$$

 $ext{Note: } E(m{X}+m{Y})=E(m{X})+E(m{Y})$ (where E(Yi)= μ i is obtained as E(Yi)= $\int y_i f_i(y_i) dy_i$)

- Covariance Matrix:

$$\Sigma = Cov(\mathbf{Y}) = E\{[\mathbf{Y} - E(\mathbf{Y})][\mathbf{Y} - E(\mathbf{Y})]'\} = \begin{pmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1p} \\ \sigma_{21} & \sigma_{22} & \cdots & \sigma_{2p} \\ \vdots & \vdots & & \vdots \\ \sigma_{p1} & \sigma_{p2} & \cdots & \sigma_{pp} \end{pmatrix}$$

Note: 10. The $\sum_{i=1}^{\infty}$ is symmetric since bij = bji

2°. In many applications, \lesssim is assumed to be positive definite

—Define the E of a random matrix ≥ as the matrix of expected values:

$$E(\mathcal{Z}) = E \begin{pmatrix} Z_{11} Z_{12} \cdots Z_{1n} \\ Z_{21} Z_{22} \cdots Z_{2n} \\ \vdots \\ Z_{n1} Z_{n2} \cdots Z_{nn} \end{pmatrix} = \begin{pmatrix} E(Z_{11}) & E(Z_{12}) \cdots & E(Z_{1n}) \\ E(Z_{11}) & E(Z_{22}) \cdots & E(Z_{2n}) \\ \vdots \\ E(Z_{n1}) & E(Z_{n2}) \cdots & E(Z_{nn}) \end{pmatrix}$$

Theorem 3.6(a). If α is a px1 vector of constant and y is a px1 random vector with mean vector k, then. $\mu_a = E(\alpha' y) = \alpha' E(y) = \alpha' k$

Theorem 3.6 (b) Suppose that X is a random vector, X is a random matrix. A and X are vectors of constants, and X and X are matrices of constants.

* sample mean and sample covariance matrix.

get some observations
$$\chi_i = \left(\begin{array}{c} \chi_{i1} \\ \vdots \\ \chi_{ip} \end{array} \right) \ i = 1 \cdots n$$

$$\Rightarrow \hat{\lambda} = \overline{\chi} = \frac{1}{1} \stackrel{\text{P}}{\approx} \chi$$

$$\hat{\mathbf{x}} = \mathbf{x} = \frac{1}{12} \mathbf{x} (\mathbf{x} - \mathbf{x}) (\mathbf{x} - \mathbf{x})'$$

$$Pf: = E \left[(AI - E(AI)) \cdot (AX - E(AI))' \right]$$

$$= E \left[A(X - E(X)) \cdot (X - E(X))'A' \right]$$

$$= A E \left[(Y - E(Y)) \cdot (Y - E(Y))' \right] A'$$

$$= A COV(X) A'$$

- Let \boldsymbol{A} be a constant matrix, then

$$Cov(\mathbf{AY}) = \mathbf{A}[Cov\mathbf{Y}]\mathbf{A}'$$

- Let $\boldsymbol{A}, \boldsymbol{B}$ be constant matrices, then

Pf: = E [(
$$AX - E(AX)$$
) ($EX - E(BX)$)]
= E [$A(X - E(X)) \cdot (X - E(X))B'$]
= $ACOV(X \cdot X)B'$

$$Cov(\boldsymbol{A}\boldsymbol{X}, \boldsymbol{B}\boldsymbol{Y}) = \mathbf{A}Cov(\boldsymbol{X}, \boldsymbol{Y})\boldsymbol{B}'$$

- Generalized variance: overall measure of variability can be defined as the determinant of Ξ .
 - i.e. Generalized variance $= |\Sigma|$

- Correlation matrices

$$\mathbf{\Omega} = \begin{pmatrix} 1 & \rho_{12} & \cdots & \rho_{1p} \\ \rho_{21} & 1 & \cdots & \rho_{2p} \\ \vdots & \vdots & & \vdots \\ \rho_{p1} & \rho_{p2} & \cdots & 1 \end{pmatrix}$$

where

$$\rho_{ij} = \frac{\sigma_{ij}}{\sqrt{\sigma_{ii}}\sqrt{\sigma_{jj}}}$$

for $i \neq j$.

- Partitioned random vectors

$$\begin{split} \boldsymbol{V} &= \begin{pmatrix} \boldsymbol{Y} \\ \boldsymbol{X} \end{pmatrix} \\ \boldsymbol{\mu} &= E(\boldsymbol{V}) = E \begin{pmatrix} \boldsymbol{Y} \\ \boldsymbol{X} \end{pmatrix} = \begin{pmatrix} E(\boldsymbol{Y}) \\ E(\boldsymbol{X}) \end{pmatrix} = \begin{pmatrix} \boldsymbol{\mu}_{\boldsymbol{Y}} \\ \boldsymbol{\mu}_{\boldsymbol{X}} \end{pmatrix} \\ \boldsymbol{\Sigma} &= Cov(\boldsymbol{V}) = Cov \begin{pmatrix} \boldsymbol{Y} \\ \boldsymbol{X} \end{pmatrix} = \begin{pmatrix} \boldsymbol{\Sigma}_{\boldsymbol{Y}\boldsymbol{Y}} & \boldsymbol{\Sigma}_{\boldsymbol{Y}\boldsymbol{X}} \\ \boldsymbol{\Sigma}_{\boldsymbol{X}\boldsymbol{Y}} & \boldsymbol{\Sigma}_{\boldsymbol{Y}\boldsymbol{X}} \end{pmatrix} \\ &= E \left[\begin{pmatrix} \vec{\gamma} - E(\vec{\gamma}) \\ \vec{\gamma} - E(\vec{\gamma}) \end{pmatrix} \begin{pmatrix} \vec{\gamma} - E(\vec{\gamma}) \vec{\gamma} & \ell \vec{\chi} - E(\vec{\lambda}) \vec{\gamma} \end{pmatrix} \right] \\ &= E \left[\begin{pmatrix} \vec{\gamma} - E(\vec{\gamma}) \\ \vec{\gamma} - E(\vec{\gamma}) \end{pmatrix} \begin{pmatrix} (\vec{\gamma} - E(\vec{\gamma}) \vec{\gamma}) & \ell \vec{\chi} - E(\vec{\lambda}) \vec{\gamma} \end{pmatrix} \right] \end{split}$$

 Σ^{-1} (use formulation discussed in ch3)

- Let \mathbf{Y} be a random vector with mean $\boldsymbol{\mu} = E(\mathbf{Y})$ and $\boldsymbol{\Sigma} = Cov(\mathbf{Y})$, then $E(\mathbf{Y}'A\mathbf{Y}) = tr(A\boldsymbol{\Sigma}) + \boldsymbol{\mu}'A\boldsymbol{\mu}$ where \mathbf{A} is a symmetric matrix.

metric matrix.

Since $E(\chi-\mu)A(\chi-\mu) = E(\chi'A\chi) - \mu'A\mu = E(\chi'A\chi) - \mu'A\mu + \mu'A\mu +$

E(YAY) = E(Y-H)'A(X-K)+ M'AM

E(X-K)'A(X-K) — quadratic form

tr(AB) = tr(BA)

= tr [EIT-HJAIT-M)]

= E tr[LT-MYA (X-K)]

= E tr[A(X-K)(X-K)']

= tr { A E[LX-K) LX-KY]]

= tr AZ

- MGF: The moment generating function of a random vector \boldsymbol{Y} is given by

$$M_{\boldsymbol{Y}}(\mathbf{t}) = E(e^{\mathbf{t}'\boldsymbol{Y}})$$

where
$$\mathbf{t} = \begin{pmatrix} \mathbf{t}_1 \\ \mathbf{t}_2 \\ \vdots \\ \mathbf{t}_n \end{pmatrix}$$
 if the expectation exists

for $-h < t_i < h$ where h > 0 and $i = 1, \dots, n$

- Theorem

Let $g_1(\mathbf{Y}_1), \dots, g_m(\mathbf{Y}_m)$ be m functions of the random vectors $\mathbf{Y}_1, \dots, \mathbf{Y}_m$, respectively. If $\mathbf{Y}_1, \dots, \mathbf{Y}_m$ are mutually independent, then g_1, \dots, g_m are mutually independent.