

## Identifying minimal failure-causing schemas in the presence of multiple faults

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Combinatorial testing (CT) has been proven effective in revealing the failures caused by the interaction of factors that affect the behavior of a system. The theory of Minimal Failure-Causing Schema (MFS) has been proposed to isolate the root cause of a failure after CT. Most algorithms that aim to identify MFS focus on handling a single fault in the System Under Test (SUT). However, we argue that multiple faults are more common in practice, under which masking effects may be triggered so that some failures cannot be observed. The traditional MFS theory lacks a mechanism to handle such effects; hence, they may incorrectly isolate the MFS. To address this problem, we propose a new MFS model that takes into account multiple faults. We first formally analyse the impact of the multiple faults on existing MFS identifying algorithms, especially in situations where masking effects are triggered by multiple faults. We then develop an approach that can assist traditional algorithms to better handle multiple faults. Empirical studies were conducted using several kinds of open-source software, which showed that multiple faults with masking effects do negatively affect traditional MFS identifying approaches and that our approach can help to alleviate these effects.

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## 1. INTRODUCTION

With the increasing complexity and size of modern software, many factors, such as input parameters and configuration options, can affect the behaviour of the SUT. The failures caused by the interaction of these factors can make software testing challenging, especially when the interaction space is large. In the worst case, we need to examine every possible interaction of these factors as each interaction may cause unique

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Table I. MS word example

id	Highlight	Status bar	Bookmarks	Smart tags	Outcome
1	On	On	On	On	PASS
2	Off	Off	On	On	PASS
3	Off	On	Off	Off	Fail
4	On	Off	Off	On	PASS
5	On	Off	On	Off	PASS

failure [Song et al. 2012]. While exhaustive testing achieves maximal test coverage, it is impractical and uneconomical. One remedy for this problem is Combinatorial Testing (CT)<sup>1</sup>, which systematically samples the interaction space and selects a relatively small set of test cases that cover all valid interactions, with the number of factors involved in each interaction no more than a prior fixed integer, i.e., the *strength* of the interaction. Many works in CT aim to construct the smallest set of test cases [Cohen et al. 1997; Bryce et al. 2005; Cohen et al. 2003; Lei et al. 2008], which is also called *covering array*.

Once failures are detected by a covering array, the failure-inducing interactions in the failing test cases should be isolated. This task is important as it can facilitate debugging efforts by reducing the scope of code that is needed for inspection. [Ghandehari et al. 2012]. However, information from a covering array sometimes is not sufficient to identify the location and number of the failure-inducing interactions [Colbourn and McClary 2008]. Thus, additional information is needed. Consider the following example [Bach and Schroeder 2004], Table I presents a two-way covering array for testing an MS-Word application in which we want to examine various interactions of options for ‘Highlight’, ‘Status Bar’, ‘Bookmarks’ and ‘Smart tags’. Assume the third test case failed. We can get six two-way suspicious interactions that may be responsible for this failure. They are respectively (Highlight: Off, Status Bar: On), (Highlight: Off, Bookmarks: Off), (Highlight: Off, Smart tags: Off), (Status Bar: On, Bookmarks: Off), (Status Bar: On, Smart tags: Off), and (Bookmarks: Off, Smart tags: Off). Without additional information, it is difficult to figure out the specific interactions in this suspicious set that caused the failure. In fact, considering that the higher strength interactions could also be failure-inducing interactions, e.g., (Highlight: Off, Status Bar: On, Smart tags: Off), the problem becomes more complicated.

To address this problem, prior work [Nie and Leung 2011a] specifically studied the properties of the failure-inducing interactions in the SUT, based on which additional test cases were generated to identify them. Other approaches to identify the failure-inducing interactions in the SUT include building a tree model [Yilmaz et al. 2006], adaptively generating additional test cases according to the outcome of the last test case [Zhang and Zhang 2011], ranking suspicious interactions based on some rules [Ghandehari et al. 2012], and using graphic-based deduction [Martínez et al. 2008], among others.

Most existing approaches mainly focus on the ideal scenario in which SUT only contains one fault, under which the outcomes of test cases can be simply categorized into failure or pass. However, in this paper, we argue that SUT with multiple faults is the more common testing scenario in practice, which will result in many distinguished failures as outcomes of test cases, and moreover, this affects the effectiveness of Failure-inducing Interactions Identifying (FII) approaches. One main impact of multiple faults on FII approaches is the masking effect. A masking effect [Dumlu et al. 2011; Yilmaz et al. 2014] occurs when some failures prevent test cases from checking interactions that are supposed to be tested. Take the Linux command *Grep* for example. We no-

<sup>1</sup>Another term for CT is Combinatorial Interaction Testing, which is abbreviated as CIT. In this paper, they are uniformly cited as Combinatorial testing (CT).

ticed that there are two different failures reported in the bug tracker system. The first<sup>2</sup> claims that Grep incorrectly matches unicode patterns with ‘\<\>’, while the second<sup>3</sup> claims an incompatibility between option ‘-c’ and ‘-o’. When we put these two scenarios into one test case, only one failure will be observed, which means the other failure is masked by the observed failure. This effect will prevent test cases from executing normally, leading to incorrect judgment of the correlation between the interactions checked in the test case and the failure that has been masked. This effect was firstly noted by Dumlu and Yilmaz in [Dumlu et al. 2011], in which they found that the masking effects in CT can prevent traditional covering array in testing some interactions.

As masking effect can negatively affect the performance of FII approaches, a natural question is how this effect biases the results of these approaches. In this paper, we formalize the process of identifying the failure-inducing interactions under the circumstances in which masking effects exist in the SUT and try to answer this question. One insight from the formal analysis is that we cannot completely avoid the impact of masking effects even if we do exhaustive testing. Even worse, either ignoring the masking effects or treating different failures as the same failure is detrimental to the FII process.

To address this concern, we propose a strategy to alleviate this impact by adopting a divide and conquer framework. With this framework, FII approaches is scheduled to separately handle each failure in the SUT. Specifically, for a particular failure, FII approaches only focus on the test cases that either pass or trigger the same failure under analysis. Test cases that triggered other different failures will be replaced with some newly generated test cases. In this way, FII approaches can properly work with little interference from the negative masking effects.

The key to our strategy is to search for a test case that does not trigger *unexpected* failures, i.e., failures different from the one under analysis. To guide the search process, a natural idea is to take some characteristics from the existing test cases, and make the characteristics of the newly searched test case as disparate from the test cases which triggered unexpected failures as possible. To reach this target, we define the *suspiciousness* between a factor and the failure. The higher the *suspiciousness* a factor is related to a particular failure, the greater the likelihood that the factor will trigger this fault. We then use the integer linear programming (ILP) technique to find a test case which has the least *suspiciousness* with unexpected failures.

To evaluate the effectiveness of our approach, we applied our strategy on the FII approach FIC\_BS [Zhang and Zhang 2011]. The subjects used were two open-source software systems found in the developers’ forum in the Source-Forge community. Through studying their bug reports in the bug tracker system as well as their user’s manuals, we built a testing model which can reproduce the reported bugs with given test cases. We then compared the FII approach augmented with our strategy to the original FII approach. We further empirically studied the performance of the important component of our strategy – searching satisfied test cases. To conduct this study, we compared our approach with the augmented FII approach by randomly searching satisfied test cases. We finally compared our approach with the only existing masking handling technique – FDA-CIT [Yilmaz et al. 2014]. Our studies showed that our replacing strategy as well as the searching test case component achieved a better performance than the traditional approaches when the subject encountered multiple faults, especially when these faults can induce masking effects.

The main contributions of this paper are:

<sup>2</sup><http://savannah.gnu.org/bugs/?29537>

<sup>3</sup><http://savannah.gnu.org/bugs/?33080>

```

public float foo(int a, int b, int c, int d){
    //step 1 will cause an exception when b == c
    float x = (float)a / (b - c);

    //step 2 will cause an exception when c < d
    float y = Math.sqrt(c - d);

    return x+y;
}

```

Fig. 1. A simple program *foo* with four input parameters

- We formally analysed the relationships between failure-inducing interactions and test sets. (Section 3)
- We provide an unified framework for explaining how traditional FII approaches work. (Section 3.3)
- We studied the impact of the masking effects caused by multiple faults on FII approaches. (Section 4)
- We proposed an efficient test case replacement strategy to alleviate the impact of these effects (Section 5).
- We conducted several empirical studies and showed that our strategy can assist FII approaches to achieve better performance in identifying failure-inducing interactions in the SUT with masking effects. (Section 8)

## 2. MOTIVATING EXAMPLE

We constructed a small example to illustrate the motivation of our approach. Assume a method *foo* has four input parameters: *a*, *b*, *c*, and *d*. The four parameter types are all integers and their values are:  $v_a = \{7, 11\}$ ,  $v_b = \{2, 4, 5\}$ ,  $v_c = \{4, 6\}$ ,  $v_d = \{3, 5\}$ . The code of *foo* is shown in Figure 1.

There are two potential failures of *foo*: first, in step 1 we can get an *Arithmetic Exception* when *b* is equal to *c*, i.e.,  $b = 4$  and  $c = 4$ , that causes a division by zero. Second, another *Arithmetic Exception* will be triggered in step 2 when  $c < d$ , i.e.,  $c = 4$  and  $d = 5$ , taking square root of a negative number. So the expected failure-inducing interactions in this example should be  $(-, 4, 4, -)$  and  $(-, -, 4, 5)$ .

FII approaches do not consider the code detail; instead, they apply black-box testing, i.e., feed inputs to the programs and execute them to observe the result. The basic proposition of FII approaches is that the failure-inducing interactions for a particular failure can only appear in those test cases that trigger this fault. FII approaches often aim at using as few test cases as possible to get the same (or approximate) result as exhaustive testing, so the results derived from an exhaustive testing set are the best that FII approaches can achieve. Here, we will show how exhaustive testing works to identify the failure-inducing interactions for the program.

We first generate every possible test case listed in the column “test case” of Table II. The execution results are listed in the result column of Table II. In this column, *PASS* means that the program runs without any exception. *Ex 1* indicates that the program triggered an exception corresponding to step 1 and *Ex 2* indicates the program triggered an exception corresponding to step 2. From the data listed in Table II, we can determine that  $(-, 4, 4, -)$  must be the failure-inducing interaction of *Ex 1* as all the test cases that triggered *Ex 1* contain this interaction. Similarly, interactions  $(-, 2, 4, 5)$  and  $(-, 5, 4, 5)$  must be the failure-inducing interactions of *Ex 2*. We list these interactions and the corresponding exceptions in Table III.

Table II. test cases and their corresponding result

id	test case	result	id	test case	result
1	(7, 2, 4, 3)	PASS	13	(11, 2, 4, 3)	PASS
2	(7, 2, 4, 5)	Ex 2	14	(11, 2, 4, 5)	Ex 2
3	(7, 2, 6, 3)	PASS	15	(11, 2, 6, 3)	PASS
4	(7, 2, 6, 5)	PASS	16	(11, 2, 6, 5)	PASS
5	(7, 4, 4, 3)	Ex 1	17	(11, 4, 4, 3)	Ex 1
<b>6</b>	<b>(7, 4, 4, 5)</b>	<b>Ex 1</b>	<b>18</b>	<b>(11, 4, 4, 5)</b>	<b>Ex 1</b>
7	(7, 4, 6, 3)	PASS	19	(11, 4, 6, 3)	PASS
8	(7, 4, 6, 5)	PASS	20	(11, 4, 6, 5)	PASS
9	(7, 5, 4, 3)	PASS	21	(11, 5, 4, 3)	PASS
10	(7, 5, 4, 5)	Ex 2	22	(11, 5, 4, 5)	Ex 2
11	(7, 5, 6, 3)	PASS	23	(11, 5, 6, 3)	PASS
12	(7, 5, 6, 5)	PASS	24	(11, 5, 6, 5)	PASS

Table III. Identified failure-inducing interactions and their corresponding Exception

Failure-inducing interaction	Exception
(-, 4, 4, -)	Ex 1
(-, 2, 4, 5)	Ex 2
(-, 5, 4, 5)	Ex 2

Note that in this example we did not get the expected result with traditional FII approaches. The failure-inducing interactions for Ex 2 are (-,2,4,5) and (-,5,4,5), respectively, instead of the expected interaction (-, -,4,5). So why did we fail to get the (-, -,4,5)? The reason lies in *test case 6* (7,4,4,5) and *test case 18* (11,4,4,5). These two test cases contain the interaction (-, -,4,5), but they did not trigger Ex 2; instead, Ex 1 was triggered.

Now consider the source code of *foo*. If Ex 1 is triggered, it will stop executing the remaining code and report the exception result. In other words, Ex 1 may mask Ex 2. Let us re-examine the interaction (-, -,4,5). If we suppose that *test case 6* and *test case 18* should trigger Ex 2 if they did not trigger Ex 1, then we can conclude that (-, -,4,5) should be the failure-inducing interaction of Ex 2, which is identical to the expected one. However, we cannot get this result, unless we fix the code that triggers Ex 1 and re-execute all the test cases.

So in practice, *when we lack resources to execute all the test cases repeatedly or can only do black-box testing, a more economical and efficient approach to alleviate the masking effect on FII approaches is desired.*

### 3. FORMAL MODEL OF MINIMAL FAILURE-CAUSING SCHEMA

This section presents some definitions and propositions for a formal description of failure-inducing interactions and test sets.

#### 3.1. Minimal Failure-causing Schemas in CT

Assume that the behaviour of a SUT is influenced by  $k$  parameters, and each parameter  $p_i$  has  $a_i$  discrete values from the finite set  $V_i$ , i.e.,  $a_i = |V_i|$  ( $i = 1, 2, \dots, k$ ). In practice, these parameters can represent many factors, such as input variables, run-time options, building options, etc. Next we will give some formal definitions, some of them (Definitions 3.1, 3.3, 3.4) were originally defined in [Nie and Leung 2011b].

**Definition 3.1.** A *test case* of a SUT is an array of  $k$  values, one for each parameter of the SUT, which is denoted as a  $k$ -tuple  $(v_1, v_2, \dots, v_k)$ , where  $v_1 \in V_1, v_2 \in V_2 \dots v_k \in V_k$ .

For the example in Section 2,  $(a = 7, b = 2, c = 4, d = 3)$  is a test case, which is actually a group of values being assigned to each input parameter.

*Definition 3.2.* A *failure* is an abnormal execution of a test case.

In CT, such a *failure* can be a thrown exception, compilation error, assertion failure or constraint violation. In this paper, failures are classified according to the specific *fault* information. For example, if failures have the same exception traces information, they are treated as the same failure. The main point in this paper is to study the impact of multiple different *failures* on failure-inducing interactions identification.

To facilitate our discussion, we introduce the following assumptions that will be used throughout this paper:

ASSUMPTION 1. *The execution result of a test case is deterministic.*

This assumption is the most common assumption of CT fault diagnosis. It indicates that the outcome of executing a test case is reproducible and will not be affected by some random events.

ASSUMPTION 2. *Different failures in the SUT can be distinguished by various information such as exception traces, state conditions, or the like.*

This assumption indicates that the testers can detect different failures during testing. As different failures will complicate fault diagnosis tasks, distinguishing them is the first step to resolve them.

ASSUMPTION 3. *All the tests are valid in the SUT, i.e., there is no inter-option constraints in the SUT.*

Here *option* indicates the parameter of a SUT. This assumption shows that we can always determine the outcome after executing a test case, i.e., either pass or fail with some exceptions.

In practice, these assumptions will not always be satisfied. In Section 7, we will discuss their impacts on the theories and approach proposed in this paper, as well as the measures to alleviate them. Particularity for the third assumption, we will show that one type of inter-option constraint is essentially equal to a specific failure. Thus we can directly use the approach in this paper to solve it.

Now let us consider the condition that some failures are triggered by some test cases. It is then desirable to determine the cause of these failures and hence some parameter values of the failing test cases must be analysed.

*Definition 3.3.* For the SUT, the  $\tau$ -tuple  $(-, v_{b_1}, \dots, v_{b_\tau}, \dots)$  is called a  $\tau$ -degree *schema* ( $0 < \tau \leq k$ ) when some  $\tau$  parameters have fixed values and the others can take on their respective allowable values, represented as “-”.

When  $\tau = k$ ,  $\tau$ -degree *schema* is actually a test case. Furthermore, if every fixed value in a schema is in a test case, we say this test case *contains* the schema.

For example,  $(-, 4, 4, -)$  in Table III is a 2-degree schema. And the test case  $(7, 4, 4, 3)$  contains this schema.

*Definition 3.4.* Let  $c_1$  be a  $m$ -degree schema,  $c_2$  be an  $n$ -degree schema in the SUT, and  $m < n$ . If all the fixed parameter values in  $c_1$  are also in  $c_2$ , then  $c_2$  *subsumes*  $c_1$ . In this case, we can also say that  $c_1$  is a *sub-schema* of  $c_2$ , and  $c_2$  is a *super-schema* of  $c_1$ , denoted as  $c_1 \prec c_2$ .

For example, in the motivating example, the 2-degree schema  $(-, 4, 4, -)$  is a sub-schema of the 3-degree schema  $(-, 4, 4, 5)$ , that is,  $(-, 4, 4, -) \prec (-, 4, 4, 5)$ .

According to definition 3.4, it is obvious that the subsuming relationship ‘ $\prec$ ’ is transitive, i.e., if  $c_1 \prec c_2$ ,  $c_2 \prec c_3$ , then  $c_1 \prec c_3$ .

**Definition 3.5.** If for any test case, as long as it contains the schema  $c$ , it will trigger a particular fault  $F$ . Then we call  $c$  the *failure-causing schema* of  $F$ . Additionally, if none of the sub-schema of  $c$  is the *failure-causing schema* of  $F$ , we then call  $c$  the *Minimal Failure-causing Schema* (MFS) of  $F$ .

In fact, MFS is identical to the failure-inducing interactions discussed previously. Identifying MFS helps to focus on the root cause of a failure and thus facilitates the debugging efforts. Note that all the failures discussed in this paper are option-related [Yilmaz et al. 2006]. That is, all the failures are caused by the interactions of the parameter values in the SUT. Another noteworthy thing is that, in practice, the MFS definition (Section 3.5) may not be valid. This is because in some specific cases, e.g., the masking problems which we will discuss later, we may not determine whether some test cases will trigger the particular failure of  $F$ , or not.

Some notations used later are listed below for convenient reference:

- $k$  : The number of parameters that influence the SUT.
- $V_i$  : The set of discrete values that the  $i$ th parameter of the SUT can take.
- $T^*$  : The exhaustive set of test cases for the SUT. For a SUT with  $k$  parameters, and each parameter can take  $|V_i|$  values, the number of test cases in  $T^*$  is  $\prod_{i=1}^k |V_i|$ . Note that some test cases may be invalid if there exists constraints among the parameters.
- $T$  : A set of test cases. (Similarly for  $T_i, T_j, \dots$ )
- $\bar{T}$  : The complementary test set of  $T$ , i.e.,  $\bar{T} \cup T = T^*, \bar{T} \cap T = \emptyset$ .
- $A \setminus B$  : The set of elements that belong to set  $A$  but not to  $B$ . For example  $T_i \setminus T_j$  indicates the set of test cases that belong to set  $T_i$ , but not to  $T_j$ .
- $L$  : The number of faults contained in the SUT.
- $F_m$  : The  $m$ th fault in the SUT ( $1 \leq m \leq L$ ); A failure which is classified to *fault*  $F_m$  is called *failure* of  $F_m$ .
- $T_{F_m}$  : All the test cases that can trigger the failure of  $F_m$  in the SUT.
- $\mathcal{T}(c)$  : All the test cases that contain the schema  $c$  in the SUT. Based on the definition of MFS, we know that if schema  $c$  is an MFS of  $F_m$ , then  $\mathcal{T}(c) \subseteq T_{F_m}$ . Note that, there may be multiple MFS for  $F_m$ .
- $\mathcal{I}(t)$  : All the schemas that are contained in the test case  $t$ , e.g.,  $\mathcal{I}((111)) = \{(1 - -)(-1 -)(- - 1)(11 -)(1 - 1)(-11)(111)\}$ .
- $\mathcal{I}(T)$  : All the schemas that are contained in test set  $T$ , i.e.,  $\mathcal{I}(T) = \bigcup_{t \in T} \mathcal{I}(t)$ .
- $\mathcal{S}(T)$  : All the schemas that are only contained in test set  $T$  (Referred to as *Special schemas*);  $\mathcal{S}(T) = \mathcal{I}(T) \setminus \mathcal{I}(\bar{T})$ .
- $\mathcal{C}(T)$  : A set of the minimal schemas that are only contained in test set  $T$  (Referred to as *Minimal schemas*);  $\mathcal{C}(T) = \{c | c \in \mathcal{S}(T) \text{ and } \nexists c' \prec c, s.t., c' \in \mathcal{S}(T)\}$ .

### 3.2. Relations between schemas and test sets

**PROPOSITION 3.6** (SMALLER SCHEMA  $c$  HAS A LARGER  $\mathcal{T}(c)$ ). *For schemas  $c_1, c_2$ , if  $c_1 \prec c_2$ , then all the test cases that contain  $c_2$  must also contain  $c_1$ , i.e.,  $\mathcal{T}(c_2) \subseteq \mathcal{T}(c_1)$ .*

We omit the proof of this proposition as it is quite obvious. Suppose a SUT with four binary parameters, which can be denoted as SUT(2<sup>4</sup>). Table IV illustrates an example of the Proposition 3.6. The left column lists the schema  $c_2 = (0, 0, -, -)$  as well as all the test cases in  $\mathcal{T}(c_2)$ , while the right column lists the schema  $c_1 = (0, -, -, -)$  and  $\mathcal{T}(c_1)$ . We can see that when  $c_1 \prec c_2$ ,  $\mathcal{T}(c_2) \subseteq \mathcal{T}(c_1)$ .

**PROPOSITION 3.7** (PROPERTY OF  $\mathcal{S}(T)$ : SPECIAL SCHEMA SET OF TEST SET  $T$ ). *For any test set  $T$  of the SUT,  $\bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c) = T$ .*

Table IV. Example of Proposition 3.6

$c_2$		$c_1$	
$(0, 0, -, -)$		$(0, -, -, -)$	
$\mathcal{T}(c_2)$		$\mathcal{T}(c_1)$	
$(0, 0, 0, 0)$		$(0, 0, 0, 0)$	
$(0, 0, 0, 1)$		$(0, 0, 0, 1)$	
$(0, 0, 1, 0)$		$(0, 0, 1, 0)$	
$(0, 0, 1, 1)$		$(0, 0, 1, 1)$	
		$(0, 1, 0, 0)$	
		$(0, 1, 0, 1)$	
		$(0, 1, 1, 0)$	
		$(0, 1, 1, 1)$	

PROOF. As  $\mathcal{S}(T) = \mathcal{I}(T) \setminus \mathcal{I}(\bar{T})$ ,  $\forall c \in \mathcal{S}(T)$ ,  $c \in \mathcal{I}(T)$  and  $c \notin \mathcal{I}(\bar{T})$ . Then  $\forall t \in \mathcal{T}(c)$ ,  $t$  contains  $c$ , indicating that  $t \in T$ . Hence,  $\mathcal{T}(c) \subseteq T$ . Then  $\bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c) \subseteq T$ .

On the other hand,  $\forall t \in T$ ,  $\exists c' \in \mathcal{I}(t)$ , such that  $c' \notin \mathcal{I}(\bar{T})$  (at least it holds when  $c' = t$ ). Hence,  $c' \in \mathcal{S}(T)$ . Obviously  $t \in \mathcal{T}(c') \subseteq \bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c)$ . Therefore,  $T \subseteq \bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c)$ .  $\square$

**PROPOSITION 3.8 (PROPERTY OF  $\mathcal{C}(T)$ ): MINIMAL SCHEMA SET OF TEST SET  $T$ ).**

For any test set  $T$  of the SUT,  $\bigcup_{c \in \mathcal{C}(T)} \mathcal{T}(c) = T$ .

PROOF. As  $\mathcal{C}(T) = \{c | c \in \mathcal{S}(T) \text{ and } \nexists c' \prec c, s.t., c' \in \mathcal{S}(T)\}$ , indicating that  $\mathcal{C}(T) \subseteq \mathcal{S}(T)$ . It is then obviously  $\bigcup_{c \in \mathcal{C}(T)} \mathcal{T}(c) \subseteq \bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c)$ . Hence, we just need to prove that  $\bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c) \subseteq \bigcup_{c \in \mathcal{C}(T)} \mathcal{T}(c)$ .

$\forall t \in \bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c)$ ,  $\exists c \in \mathcal{S}(T)$ , s.t.,  $t \in \mathcal{T}(c)$ . According to the definition of  $\mathcal{C}(T)$ ,  $\exists c' \in \mathcal{C}(T)$ , s.t.,  $c' = c$  or  $c' \prec c$ . Correspondingly  $\mathcal{T}(c') = \mathcal{T}(c)$ , or  $\mathcal{T}(c) \subseteq \mathcal{T}(c')$  by Proposition 3.6. Hence,  $t \in \mathcal{T}(c') \subseteq \bigcup_{c \in \mathcal{C}(T)} \mathcal{T}(c)$ .

Therefore,  $\bigcup_{c \in \mathcal{C}(T)} \mathcal{T}(c) = \bigcup_{c \in \mathcal{S}(T)} \mathcal{T}(c) = T$ .  $\square$

Table V gives an example of  $T^*$ ,  $T$ ,  $\bar{T}$ ,  $\mathcal{S}(T)$  and  $\mathcal{C}(T)$  in  $\text{SUT}(2^3)$ . We can find that all the schemas in  $\mathcal{S}(T)$  and  $\mathcal{C}(T)$  are only contained in test set  $T$ , and for any  $t \in T$ , it contains at least one schema in  $\mathcal{S}(T)$  and  $\mathcal{C}(T)$ . Additionally,  $\mathcal{C}(T)$  is a minimal schema set which filters those super schemas in  $\mathcal{S}(T)$ .

Table V. Example of the special and minimal schemas

$T^*$	$T$	$\bar{T}$	$\mathcal{S}(T)$	$\mathcal{C}(T)$
$(0, 0, 0)$	$(0, 0, 0)$		$(0, 0, -)$	$(0, 0, -)$
$(0, 0, 1)$	$(0, 0, 1)$		$(0, -, 0)$	$(0, -, 0)$
$(0, 1, 0)$	$(0, 1, 0)$		$(0, 0, 0)$	
$(0, 1, 1)$		$(0, 1, 1)$	$(0, 0, 1)$	
$(1, 0, 0)$		$(1, 0, 0)$	$(0, 1, 0)$	
$(1, 0, 1)$		$(1, 0, 1)$		
$(1, 1, 0)$		$(1, 1, 0)$		
$(1, 1, 1)$		$(1, 1, 1)$		

Let  $T_{F_m}$  denotes the set of all the test cases triggering failure of  $F_m$ , then  $\mathcal{C}(T_{F_m})$  actually is the MFS set of  $F_m$ .

According to the definition of  $\mathcal{C}(T)$ , one observation is  $\mathcal{C}(T) \subseteq \mathcal{S}(T)$ , and for any schema in  $\mathcal{S}(T)$ , it either belongs to  $\mathcal{C}(T)$ , or is the super schema of one element of  $\mathcal{C}(T)$ , i.e.,  $\forall c \in \mathcal{S}(T)$ ,  $\exists c' \in \mathcal{C}(T)$ , s.t.,  $c' = c$ , or  $c' \prec c$ .

**PROPOSITION 3.9 (MINIMAL SCHEMA OF THE SUBSET OF TEST SET  $T$ ).** For any test set  $T$  and schema  $c$  of the SUT, if  $\mathcal{T}(c) \subseteq T$ , then  $c \in \mathcal{S}(T)$ .



PROOF. Assume  $c \notin \mathcal{S}(T)$ , i.e.,  $c \notin \mathcal{I}(T) \setminus \mathcal{I}(\bar{T})$ , then  $c \in \mathcal{I}(\bar{T})$ . It indicates that  $\exists t \in \bar{T}, t \in \mathcal{T}(c)$ , which contradicts that  $\mathcal{T}(c) \subseteq T$ . Therefore,  $c \in \mathcal{S}(T)$ .  $\square$

Table VI shows an example of this proposition for  $\text{SUT}(2^3)$ . In this table, the test set  $\mathcal{T}(c)$  of schema  $c$  is the subset of test set  $T$ . As a result, the special schema set  $\mathcal{S}(T)$  of  $T$  contains this schema  $c = (0, 0, -)$ .

Table VI. Example of a minimal schema of the subset of a test set

$c$	$\mathcal{T}(c)$	$T$	$\mathcal{S}(T)$
(0, 0, -)	(0, 0, 0)	(0, 0, 0)	(0, -, -)
	(0, 0, 1)	(0, 0, 1)	(0, -, 0)
		(0, 1, 0)	(0, -, 1)
		(0, 1, 1)	(0, 0, -)
			(0, 1, -)
			(0, 0, 0)
			(0, 0, 1)
			(0, 1, 0)
			(0, 1, 1)

Based on Proposition 3.9, as long as  $\mathcal{T}(c) \subseteq T$  for any schema  $c$  and any test set  $T$  in the SUT,  $c$  either belongs to  $\mathcal{C}(T)$  or is the super-schema of some schema in  $\mathcal{C}(T)$ . Considering a more general scenario, i.e., two test sets  $T_1, T_2$  with  $T_2 \subseteq T_1$ , we then can easily obtain the relationships between  $\mathcal{C}(T_1)$  and  $\mathcal{C}(T_2)$  according to Proposition 3.9.

**PROPOSITION 3.10 (MINIMAL SCHEMAS IN THE SMALLER TEST SET).** *For  $T_1$  and  $T_2$  of the SUT with  $T_2 \subseteq T_1$ ,  $\forall c_2 \in \mathcal{C}(T_2)$ ,  $\exists c_1 \in \mathcal{C}(T_1)$ , s.t., either  $c_1 = c_2$  or  $c_1 \prec c_2$ .*

PROOF.  $\forall c_2 \in \mathcal{C}(T_2)$ ,  $\mathcal{T}(c_2) \subseteq T_2 \subseteq T_1$ . According to Proposition 3.9,  $c_2 \in \mathcal{S}(T_1)$ . By definitions of  $\mathcal{S}(T)$  and  $\mathcal{C}(T)$ ,  $\exists c_1 \in \mathcal{C}(T_1)$ , s.t.,  $c_1 = c_2$ , or  $c_1 \prec c_2$ .  $\square$

**PROPOSITION 3.11 (MINIMAL SCHEMAS IN THE LARGER TEST SET).** *For  $T_1$  and  $T_2$  of the SUT,  $T_2 \subseteq T_1$ . Then  $\forall c_1 \in \mathcal{C}(T_1)$ ,  $\exists c_2 \in \mathcal{C}(T_2)$ , s.t., (1)  $c_1 = c_2$ , or (2)  $c_1 \prec c_2$ , or (3)  $\nexists c_2 \in \mathcal{C}(T_2)$ , s.t.,  $c_2 \prec c_1$  or  $c_2 = c_1$ , or  $c_1 \prec c_2$ .*

PROOF. Assume that  $\exists c_1 \in \mathcal{C}(T_1)$ , s.t.,  $\exists c_2 \in \mathcal{C}(T_2), c_2 \prec c_1$ . According to Proposition 3.10, as  $T_2 \subseteq T_1$ , so  $\forall c_2 \in \mathcal{C}(T_2)$ ,  $\exists c'_1 \in \mathcal{C}(T_1)$ , s.t., either  $c'_1 = c_2$  or  $c'_1 \prec c_2$ . Combining them, we can get that  $c'_1 \prec c_1$ . This is a obvious contradiction, as  $c_1$  and  $c'_1$  are both minimal schemas in  $\mathcal{C}(T_1)$ .

Hence, apart from the impossible relationship  $c_2 \prec c_1$ , then  $\forall c_1 \in \mathcal{C}(T_1)$ ,  $\exists c_2 \in \mathcal{C}(T_2)$ , s.t., (1)  $c_1 = c_2$ , or (2)  $c_1 \prec c_2$ , or (3)  $\nexists c_2 \in \mathcal{C}(T_2)$ , s.t.,  $c_2 \prec c_1$  or  $c_2 = c_1$ , or  $c_1 \prec c_2$ .  $\square$

We need to note the third case, i.e.,  $\nexists c_2 \in \mathcal{C}(T_2)$ , s.t.,  $c_1 \prec c_2$  or  $c_1 = c_2$ , or  $c_2 \prec c_1$ . We refer to this case as  $c_1$  is *irrelevant* to  $\mathcal{C}(T_2)$ . Furthermore, we can also say a schema is *irrelevant* to another schema if these two schemas are neither identical nor subsuming each other.

We illustrate these scenarios with examples in Table VII for  $\text{SUT}(2^3)$ . There are two parts in this table, with each part showing two test sets:  $T_1$  and  $T_2$ , which have  $T_2 \subseteq T_1$ . In the left part, the schemas in  $\mathcal{C}(T_2)$ : (0, 0, -) and (0, -, 0), both are the super-schemas of the one in  $\mathcal{C}(T_1)$ : (0, -, -). While in the right part, the schemas in  $\mathcal{C}(T_2)$ : (0, 0, -) and (0, -, 0) are both also in  $\mathcal{C}(T_1)$ . Furthermore, one schema in  $\mathcal{C}(T_1)$ : (1, 1, -) is irrelevant to  $\mathcal{C}(T_2)$ .

In summary, these propositions provide the foundation of MFS identification (Propositions 3.7 and 3.8), and more meaningful, they clarify the relationships between the

Table VII. Minimal schemas of two subsuming test set

$T_2$	$T_1$	$T_2$	$T_1$
(0, 0, 0)	(0, 0, 0)	(0, 0, 0)	(0, 0, 0)
(0, 0, 1)	(0, 0, 1)	(0, 0, 1)	(0, 0, 1)
(0, 1, 0)	(0, 1, 0)	(0, 1, 0)	(0, 1, 0)
	(0, 1, 1)		(1, 1, 0)
			(1, 1, 1)
$\mathcal{C}(T_2)$	$\mathcal{C}(T_1)$	$\mathcal{C}(T_2)$	$\mathcal{C}(T_1)$
(0, 0, -)	(0, -, -)	(0, 0, -)	(0, 0, -)
(0, -, 0)		(0, -, 0)	(0, -, 0)
			(1, 1, -)

minimal schemas of two different test sets (Propositions 3.11 and 3.10), which is the key to explain the impact of masking effects later.

### 3.3. Identify the MFS

In this section, we will study how traditional MFS identification approaches, i.e., FII approaches, identify the MFS. Let  $T_F$  indicates all the failing test cases and  $T_P$  all the passing test cases (Note that  $T_F \cap T_P = \emptyset$  and  $T_F \cup T_P = T^*$ ). Then the MFS for the SUT is  $\mathcal{C}(T_F)$ . Theoretically, to accurately figure out the MFS in the SUT, we need to exhaustively execute each possible test case, and collect the failing test cases  $T_F$ . This is impossible in practice, especially when the testing space is very large.

As it is impractical to get all the MFS, the appropriate is to determine some MFS of them. With the following proposition, we can learn that with existing subset of the test cases, we can also identify some of them.

**PROPOSITION 3.12 (EQUAL MINIMAL SCHEMAS).** *For  $T_1$  and  $T_2$ ,  $T_1 \subset T_2$ . If  $\exists c \in \mathcal{C}(T_1)$ , such that  $c \in \mathcal{C}(T_2)$ . Then, for any test set  $T_3$ , such that  $T_1 \subset T_3 \subset T_2$ , we have  $c \in \mathcal{C}(T_3)$ .*

**PROOF.** According to Proposition 3.11,  $\forall c \in \mathcal{C}(T_1)$ , exists  $c' \in \mathcal{C}(T_3)$ , such that either  $c' \prec c$  or  $c' = c$ . Let us assume that  $c' \prec c$ . As there exists  $c'' \in \mathcal{C}(T_2)$ , such that  $c = c''$ , then  $c' \prec c''$ . This is contradiction as  $T_3 \subset T_2$ . Hence, we have the assumption  $c' \prec c$  does not hold, indicating that  $c' = c$ .  $\square$

For example, Table VIII shows the minimal schemas for  $T_1$ ,  $T_2$ ,  $T_3$  and  $T'_3$ , where  $T_1 \subset T_3 \subset T_2$  and  $T_1 \subset T'_3 \subset T_2$ . We can find that  $\mathcal{C}(T_1)$ , i.e., (0, 0, -) is identical to that one in  $\mathcal{C}(T_2)$ , which also in  $\mathcal{C}(T_3)$  and  $\mathcal{C}(T'_3)$ .

Table VIII. Identical minimal schemas of subsuming test sets

$T_1$	$T_2$	$T_3$	$T'_3$
0 0 0	0 0 0	0 0 0	0 0 0
0 0 1	0 0 1	0 0 1	0 0 1
	1 1 1	1 1 1	1 1 0
	1 1 0		
$\mathcal{C}(T_1)$	$\mathcal{C}(T_2)$	$\mathcal{C}(T_3)$	$\mathcal{C}(T'_3)$
0 0 -	0 0 -	0 0 -	0 0 -
	1 1 -	1 1 1	1 1 0

Based on Proposition 3.12, let  $T_{F-known}$  be the given test cases which are known to be failing, and  $T_{P-known}$  be the known passing test cases.  $T_{unknown} = T^* \setminus (T_{F-known} \cup T_{P-known})$  be the test cases that are not known to be failing or not, we have the following lemma:

**LEMMA 3.13 (DETERMINE SOME MFS).**  $\forall c \in \mathcal{C}(T_{F-known})$ , if  $\exists c' \in \mathcal{C}(T_{F-known} \cup T_{unknown})$ , such that  $c = c'$ , then  $c$  must be MFS.

The proof of this lemma is obvious. This is because whatever the outcomes of the test cases  $T_{unknown}$  will be, the actual failing test cases (let it to be  $T_{F-actual}$ ), must have  $T_{F-known} \subset T_{F-actual} \subset (T_{F-known} \cup T_{unknown})$ . Based on Proposition 3.12, this lemma holds. With this lemma, we can determine some MFS in the SUT without executing all the test cases. However, to identify a MFS  $c$ , there still needs many test cases to execute; at least, we should execute all the test cases in  $\mathcal{T}(c)$ . For example, for SUT(2<sup>8</sup>), to identify the MFS (1, 1, -, -, -, -, -), we should make sure all the test cases contain this schema must fail, of which the number is  $|T_{F-known}| = 2^{8-2} = 64$ . This number grows exponentially with number of parameters in the SUT. Hence, traditional FII approaches will take some assumptions to further reduce the number of needed test cases to identify the MFS.

Generally, with proper assumptions, apart from the test cases which have been executed, we can infer some of the remaining test cases to be passing or failing. Formally, let  $T_{F-infer}$  refer to the test cases that can be inferred to be failing and  $T_{P-infer}$  the inferred passing test cases. With the test cases whose outcomes can be inferred, the test cases which are known to failing is increased to be  $T_{F-known} \cup T_{F-infer}$ , and the unknown test cases decreased to be  $T_{unknown} \setminus (T_{F-infer} \cup T_{P-infer})$ . Based on Lemma 3.13,  $\forall c \in \mathcal{C}(T_{F-known} \cup T_{F-infer})$ , if it is identical to some schema in  $\mathcal{C}((T_{F-known} \cup T_{F-infer}) \cup (T_{unknown} \setminus (T_{F-infer} \cup T_{P-infer})))$ , then  $c$  must be MFS.

The assumptions and corresponding inferring process vary with different FII approaches. To specifically describe the inferring processes of different FII approaches is beyond this paper. Next, we will only focus on describing the FIC\_BS approach [Zhang and Zhang 2011]—an efficient adaptive MFS identification approach.

The algorithm of FIC\_BS is listed in the appendix, as well as the corresponding interpretation. We will give a simple example to illustrate how FIC\_BS works. For SUT(2<sup>4</sup>), assume the test case (1, 1, 1, 1) failed, then the FIC\_BS approach can be illustrated in Table IX. In this table, to identify the MFS in  $t_0$ , FIC\_BS generated 4 more test cases, i.e.,  $t_1, t_2, t_3$  and  $t_4$ . Each of these test cases consists of two parts: the same part in the original failing test case, we called the *fixed* part against  $t_0$ ; and the part with values different from the original failing one, we called them the *mutated* part against  $t_0$ . For example, for test case  $t_1$ , the *fixed* part (against  $t_0$ ) is (-, -, 1, 1) and the *mutated* part is (0, 0, -, -). Note that these mutated parts do not necessarily have to be value (0, 0, -, -). We just need make them different from what are in  $t_0$  (This is important to the approach we later propose in the paper). The MFS obtained by FIC\_BS is (-, 1, -, 1). This conclusion is based on the following assumption.

Table IX. FIC\_BS running example

ID	Test Case				outcome
$t_0$	1	1	1	1	fail
$t_1$	0	0	1	1	pass
$t_2$	0	1	1	1	fail
$t_3$	0	1	0	0	pass
$t_4$	0	1	0	1	fail

**ASSUMPTION 4.** *None of the parameter values of the mutated part of each additional test cases is the part of any MFS.*

We later refer to the parameter values which is not the part of MFS as *safe values*. Hence, this assumption indicates that the mutated part of each additional test cases are all safe values. Based on this assumption, we have the following two lemmas:

**LEMMA 3.14.**  $\forall t \in T_{unknown}$ , if  $\exists t' \in T_{P-known}$ , such that  $t$  is obtained by mutating some parameter values of  $t'$ , and the mutated part are all safe values, then  $t$  must be a passing test case.

**PROOF.** Assume that  $t$  is a failing test case. As  $\{t\} \subset T_F$ ,  $\exists c \in \mathcal{C}(T_F)$ , such that  $c \prec t$  or  $c = t$  (Based on Proposition 3.9). In other words,  $\forall e \in c$ , it has  $e \in t$ . As  $c$  has no parameter value that is safe value ( $c$  is a MFS), so  $\forall e \in c$ , it has  $e \in (t \setminus mutate)$ , where  $mutate$  indicates the mutated part of  $t$  against  $t'$ . However, as  $t'$  also contain  $(t \setminus mutate)$ , so  $t'$  should be a failing test case, which is contradiction. Hence,  $t$  must be a passing test case.  $\square$

For example, if  $(0, 0, 0, 0)$  is known to be a passing test case, and  $(-, -, -, 1)$  is safe value, then test case  $(0, 0, 0, 1)$  must be a passing test case.

**LEMMA 3.15.**  $\forall t \in T_{unknown}$ , if  $\exists t' \in T_{F-known}$ , such that  $t$  is obtained by mutating some parameter values of  $t'$ , and the mutated part are all safe values, then  $t$  must be a failing test case.

This lemma can be proved according to Lemma 3.14. As a example, assume  $(0, 0, 0, 0)$  is known to be a failing test case, and  $(-, -, -, 0)$  is safe value, then test case  $(0, 0, 0, 1)$  must be a failing test case.

Based on these two lemmas, we can easily infer some unknown test cases to be failing or not. In table IX, FIC\_BS assumes that  $(0, -, -, -)$ ,  $(-, 0, -, -)$ ,  $(-, -, 0, -)$  and  $(-, -, -, 0)$  are *safe values*. Then those inferred test cases are listed in Table X with column ‘State’ *infer*. For example, the mutated part of  $t_2$  against the known passing test case  $t_4$  is the  $(-, -, 0, -)$ , which is safe value, so  $t_2$  should be inferred to be passing; the mutated part of the known failing test case  $t_{16}$  against the unknown test case  $t_{14}$  is still the safe value  $(-, -, 0, -)$ , so  $t_{14}$  should be inferred to be failing.

Table X. Test outcomes for MFS(- 1 - 1)

ID	Test Case				Outcome	State
<b><math>t_1</math></b>	<b>0</b>	<b>0</b>	<b>0</b>	<b>0</b>	<b>pass</b>	<b>infer</b>
<b><math>t_2</math></b>	<b>0</b>	<b>0</b>	<b>0</b>	<b>1</b>	<b>pass</b>	<b>infer</b>
<b><math>t_3</math></b>	<b>0</b>	<b>0</b>	<b>1</b>	<b>0</b>	<b>pass</b>	<b>infer</b>
$t_4$	0	0	1	1	pass	known
$t_5$	0	1	0	0	pass	known
$t_6$	0	1	0	1	fail	known
$t_7$	0	1	1	0	-	unknown
$t_8$	0	1	1	1	fail	known
$t_9$	1	0	0	0	-	unknown
$t_{10}$	1	0	0	1	-	unknown
$t_{11}$	1	0	1	0	-	unknown
$t_{12}$	1	0	1	1	-	unknown
$t_{13}$	1	1	0	0	-	unknown
<b><math>t_{14}</math></b>	<b>1</b>	<b>1</b>	<b>0</b>	<b>1</b>	<b>fail</b>	<b>infer</b>
$t_{15}$	1	1	1	0	-	unknown
$t_{16}$	1	1	1	1	fail	known

After these test cases having been inferred, we can find that  $\mathcal{C}(T_{F-known} \cup T_{F-infer}) = \mathcal{C}(\{t_6, t_8, t_{14}, t_{16}\}) = (-, 1, -, 1)$ , and  $\mathcal{C}((T_{F-known} \cup T_{F-infer}) \cup (T_{unknown} \setminus (T_{F-infer} \cup T_{P-infer}))) = \mathcal{C}(\{t_6, t_8, t_{14}, t_{16}, t_7, t_9, t_{10}, t_{11}, t_{12}, t_{13}, t_{15}\}) = (1, -, -, -), (-, 1, 1, -), (-, 1, -, 1)$ . Hence, we can derive that  $(-, 1, -, 1)$  must be MFS.

This example illustrates that the FII approach does not execute all the test cases, but with proper assumptions, the schemas identified by FII can be approximate to or even the same as the exactly correct MFS in the SUT. Other FII approaches can also be built into our formal model for MFS identification, but this will not be discussed

here as it is not the point of this paper. Even though each FII approach tries to identify the MFS as accurately as possible, masking effects will negatively affect the observed outcomes and the inferred results. We next discuss the masking problem.

#### 4. MASKING EFFECT

As discussed before,  $\mathcal{C}(T_{F_m})$  is the MFS set of fault  $F_m$  in theory. When considering the masking effects between multiple faults, however, this formula is not correct.

*Definition 4.1.* A *masking effect* occurs when a test case  $t$  contains an MFS of a particular fault, but it does not trigger the expected fault because other unexpected event, such as the other fault or unaccounted control dependency, was triggered ahead of it that prevents  $t$  from being normally checked.

Taking the masking effects into account, when identifying the MFS of a specific fault  $F_m$ , we should not ignore those test cases which did not trigger  $F_m$  but should have triggered it. We call these test cases  $T_{mask(F_m)}$ . Hence, the MFS of fault  $F_m$  should be  $\mathcal{C}(T_{F_m} \cup T_{mask(F_m)})$ .

Going back to the motivating example in Section 2, as *test case 6* and *test case 18* should trigger Ex 2 in case they did not trigger Ex 1,  $T_{mask(F_2)}$  is  $\{(7,4,4,5), (11,4,4,5)\}$ . Hence, the MFS of Ex2 is  $\mathcal{C}(T_{F_2} \cup T_{mask(F_2)})$ , which is  $(-, -, 4, 5)$  instead of the incorrect schema set  $\{(-, 2, 4, 5), (-, 5, 4, 5)\}$ .

In practice with masking effects, however, it is not possible to correctly identifying the MFS, unless we fix some bugs in the SUT and re-execute the test cases to figure out  $T_{mask(F_m)}$ .

For traditional FII approaches, without the knowledge of  $T_{mask(F_m)}$ , two common strategies can be adopted to deal with the multiple faults problem, i.e., *regarded as same failure* and *distinguishing failures*. The former strategy treats all types of faults as one fault, and hence all the failures are considered the same, while the latter distinguishes the failures but with no special consideration of the masking effects, i.e., if a test case fails with a particular type of fault, this strategy presumes it does not contain other types of faults.

##### 4.1. Regarded as same failure strategy

This is the most common strategy. With this strategy, the minimal schemas are the set  $\mathcal{C}(\bigcup_{i=1}^L T_{F_i})$ , where  $L$  is the number of all the faults in the SUT. Obviously,  $T_{F_m} \cup T_{mask(F_m)} \subseteq \bigcup_{i=1}^L T_{F_i}$ . By Proposition 3.11, some schemas obtained by this strategy may be the sub-schemas of some of the actual MFS, or be irrelevant to the actual MFS.

As an example, consider the test cases in Table XI. Assume we need to characterize the MFS of *fault 1*. All the test cases that triggered *fault 1* are listed in column  $T_{F_1}$ ; similarly, we list the test cases that triggered failures of other faults in column  $T_{mask(F_1)}$  and  $T_{non\_mask}$ , respectively, in which the former masked *fault 1*, while the latter did not. Actually the MFS of *fault 1* should be  $(1,1,-,-)$  and  $(-,1,1,1)$  as we listed them in the column ‘actual MFS of *fault 1*’. However, when we use the *regarded as same failure* strategy, the minimal schemas obtained will be  $(-, -, 0, 0)$ ,  $(1,1,-,-)$ ,  $(-, -, 1, 1)$ , in which  $(-, -, 0, 0)$  is irrelevant to the actual MFS of *fault 1*, and  $(-, -, 1, 1)$  is a sub-schema of the actual MFS  $(-, 1, 1, 1)$ .

##### 4.2. Distinguishing failures strategy

Distinguishing the failures by the exception traces or error code can help identify the MFS related to a particular failure. Yilmaz [Yilmaz et al. 2014] proposed the *multiple-class* failure characterizing method instead of the *ternary-class* approach to make the

Table XI. masking effects for exhaustive testing

$T_{F_1}$	$T_{mask(F_1)}$	$T_{non\_mask}$
(1, 1, 1, 1)	(1, 1, 0, 0)	(0, 1, 0, 0)
(1, 1, 1, 0)	(0, 1, 1, 1)	(0, 0, 0, 0)
(1, 1, 0, 1)		(1, 0, 0, 0)
		(1, 0, 1, 1)
		(0, 0, 1, 1)
actual MFS of <i>Fault 1</i>	regarded as same failure	distinguishing failures
$\mathcal{C}(T_{F_1} \cup T_{mask(F_1)})$	$\mathcal{C}(T_{F_1} \cup T_{mask(F_1)} \cup T_{non\_mask})$	$\mathcal{C}(T_{F_1})$
(1, 1, -, -)	(-, -, 0, 0)	(1, 1, -, 1)
(-, 1, 1, 1)	(1, 1, -, -)	(1, 1, 1, -)
	(-, -, 1, 1)	

characterizing process more accurate. Besides, other approaches can also be easily extended using this strategy for testing the SUT with multiple faults.

This strategy focuses on identifying the set of  $\mathcal{C}(T_{F_m})$ . As  $T_{F_m} \cup T_{mask(F_m)} \supseteq T_{F_m}$ , by Proposition 3.11, some schemas obtained by this strategy may be the super-schema of some actual MFS. Moreover, some MFS may be irrelevant to the schemas obtained by this strategy, which means that this strategy will *ignore* these actual MFS.

For the simple example shown in Table XI, when using this strategy, we will get the minimal schemas (1, 1, -, 1) and (1, 1, 1, -), which are both super schemas of the actual MFS (1,1,-,-). Furthermore, no schemas obtained by this strategy have any relationship with the actual MFS (-,1,1,1), which means it was ignored.

It is noted that the motivating example in section 2 actually adopted this strategy. As a result, the schemas identified for Ex 2: (-,2,4,5), (-,3,4,5) are the super-schemas of the correct MFS(-,-,4,5).

#### 4.3. Masking effects for FII approaches

Based on previous analysis, even though exhaustive testing is conducted to obtain  $T_{F_m}$ , we cannot determine the MFS set because of the masking effects. For traditional MFS identification approaches, i.e., FII approaches, masking effects can make problems worse. This is because FII approaches, with masking effects, not only unable determine  $T_{mask(F_m)}$  as discussed before, but also make wrong inferring outcomes of some test cases without executing, i.e., wrong  $T_{F-infer}$  and  $T_{P-infer}$  as mentioned in Section 3.3.

Consider a FII example that still use the FIC\_BS method [Zhang and Zhang 2011]. Similar to the example listed in Table IX, assume the test case (1, 1, 1, 1) failed with  $F_1$  (The MFS is still assumed to be (-, 1, -, 1) as Table IX). Then assume that an known MFS (-, 1, 0, -) that trigger  $F_2$ , where fault  $F_2$  can mask  $F_1$ . Next we will illustrate how FIC\_BS works using strategies the *regarded as one strategy* and *distinguishing failures strategy*, respectively.

First, for the first strategy, i.e., *regarded as one strategy*, we list how it works in Fig. 2, as well as the masking effects it suffers. With this strategy, it just needs three additional test cases to determine the MFS in the original failing test case  $t$ , i.e.,  $t'_1$ ,  $t'_2$ , and  $t'_3$  (listed in the left part of Fig. 2). The right part of Fig. 2 shows the MFS it obtained is (-, 1, -, -), which is obviously wrong (See the correct result in Table X, which should be (-, 1, -, 1), but this strategy treat the sub-schema (-, 1, -, -) of it to be MFS). The middle part of Fig. 2 explains why this strategy can get this result. Specifically, it treats all failures as the same failure, i.e., failure of  $F_1$ . Hence, test case  $t_5$  in the middle of Fig. 2 (Marked *Strategy* in the column 'State') is determined to trigger  $F_1$ . Here, the test case  $t_5$  corresponds to the test case  $t'_3$  in the left part of Fig.2. According to the Assumption 3.14 and 3.15, test cases  $t_1$ ,  $t_2$ ,  $t_3$ ,  $t_6$ ,  $t_7$ ,  $t_{13}$ ,  $t_{14}$ , and  $t_{15}$  are determined to be pass or fail with  $F_1$ , respectively. After then, we can determine that (-, 1, -, -) should

be the MFS. This is because,  $\mathcal{C}(T_{F_1}) = \mathcal{C}(\{t_5, t_6, t_7, t_8, t_{13}, t_{14}, t_{15}, t_{16}\}) = (-, 1, -, -)$ , which has been included in  $\mathcal{C}(T_{F_1} \cup T_{unknown}) = \mathcal{C}(\{t_5, t_6, t_7, t_8, t_9, t_{10}, t_{11}, t_{12}, t_{13}, t_{14}, t_{15}, t_{16}\}) = (1, -, -, -), (-, 1, -, -)$ .

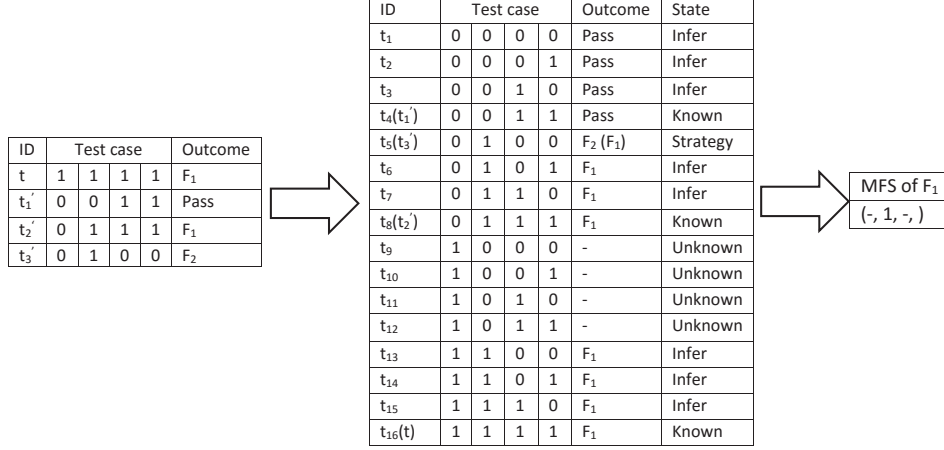


Fig. 2. Masking effects of strategy Regarded as same failure

For the second strategy, i.e., *distinguishing failures*, we list the result in Fig. 3. Different from the first strategy, it treat the different failure other than failure of  $F_1$  as pass. As a result, test case  $t_5$  and  $t_6$  in the middle part of Fig. 3 are determined to be pass. Then combining the inferred test cases, it finally get the MFS  $(-, 1, 1, 1)$ , which is also not correct (This schema is the super-schema of the original MFS of  $F_1$   $(-, 1, -, 1)$ ).

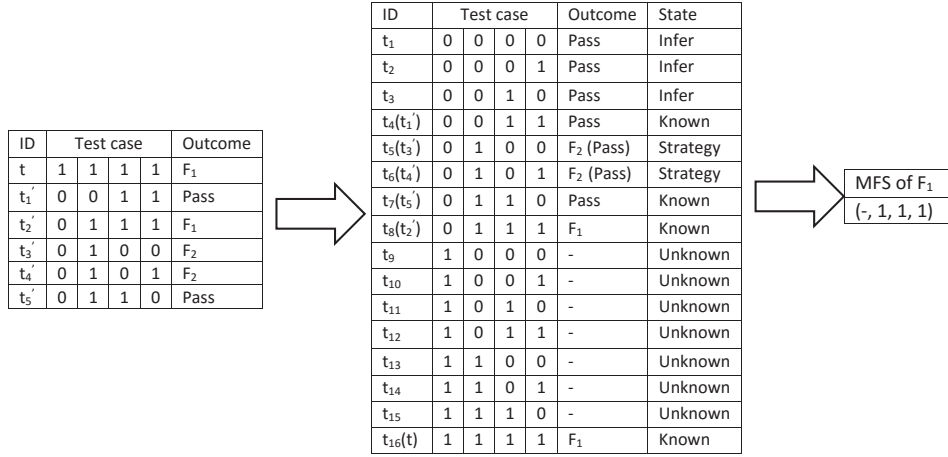


Fig. 3. Masking effects of strategy Distinguishing failures

Such deviations for these two strategies are caused by the reduction of the test cases which are known to be failing with specific fault. Here, test cases (0, 1, 0, 0) and (0, 1, 0, 1) both triggered other faults, among which the second test case (0, 1, 0, 1) should

trigger  $F_1$  if fault  $F_2$  was not triggered first, while the first test case not. Hence, neither strategy *Regarded as same failure* nor *Distinguishing failures* can make a perfect determination of the result of these two test cases. Besides this, the incorrect determination of these test cases negatively affect the inference of FII approach—FIC\_BS. For example, with strategy *Regarded as same failure*, it wrongly infer the test case  $t_{15}$  (1, 1, 1, 0) to be fail with  $F_1$  (See the middle part of Fig. 2), which should be marked as ‘Unknown’ in Table X. With strategy *Distinguishing failures*, however, it marked test case  $t_{14}$  (1, 1, 0, 1) as ‘Unknown’ (See the middle part of Fig. 3), which should be inferred to be fail with  $F_1$  in Table X.

As masking effect has significant impact on the FII approaches, alleviating this negative effect is desired to improve the quality of the identified MFS.

## 5. TEST CASE REPLACING STRATEGY

The main reason that the FII approach fails to work properly is that it cannot determine  $T_{mask(F_m)}$ . In other words, if the test case triggers other unexpected failures which are different from the currently analysed  $F_m$ , it cannot figure out whether this test case will trigger  $F_m$  because of the masking effects. So to limit the impact of this effect on the FII approach, it is important to reduce the number of test cases that trigger other different failures, as it can reduce the probability that the expected failure may be masked by other different failures.

For the exhaustive testing model, i.e.,  $\mathcal{C}(T_{F_m})$ , as all the test cases will be used to identify the MFS, there is no room left to improve the performance unless we fix other different failures and re-execute all the test cases. However, if only a subset of all test cases is used to identify the MFS (which is how the traditional FII approach works), it is important to make the right selection to limit the size of  $T_{mask(F_m)}$  to be as small as possible.

### 5.1. Replacing test cases that trigger unexpected failures

The basic idea is to pick the test cases that trigger other faults and generate new test cases to replace them. The regenerated test cases should either pass in the execution or trigger  $F_m$ .

Normally, when we replace the test case that triggers an unexpected failure with a new test case, we should keep some part of the original test case. We call this part the *fixed part*, and mutate the other part with different values from the original one. For example, if a test case (1,1,1,1) triggered an unexpected failure, and the fixed part is (-,-,1,1). Then, we can replace it with a test case (0,0,1,1) which may either pass or trigger the same failure as currently analysed.

The *fixed part* can vary for different FII approaches. For example, for the OFOT [Nie and Leung 2011a] algorithm, the parameter values are the fixed part except for the one that needs to be validated, while for the FIC\_BS [Zhang and Zhang 2011] approach, the fixed parts are dynamically changed, depending on the outcome of the execution of last generated test case.

This replacement process may need to be executed multiple times for one fixed part as it may not always be possible to find a test case that coincidentally satisfied our requirement. One replacement method is randomly choosing test cases until the satisfied test case is found. While this method may be simple and straightforward, however, it also may require tries. So to handle this problem and reduce the cost, we propose a replacement approach by computing the *suspiciousness* of the test case with the other faults, and then selecting the test case from a group of candidate test cases that has the least *suspiciousness* with other faults.

To explain the *suspiciousness* notion, we first introduce the *suspiciousness* between a parameter value  $o$  and a particular fault. We use  $all(o)$  to represent the number of



executed test cases that contain this parameter value, and  $f_i(o)$  to indicate the number of test cases that trigger the failure of  $F_i$ ,  $1 \leq i \leq L$  and contain this parameter value. Then, the *suspiciousness* between a parameter value and a particular fault, i.e.,  $Sp(o, F_i)$ , is  $\frac{f_i(o)}{all(o)+1}$ . This heuristic formula is based on the idea that if a parameter value frequently appears in the test cases that trigger a particular failure, then it is more likely to be the inducing factor that triggers this fault. We add 1 in the denominator for two reasons: (1) avoid division by zero when the parameter value has never appeared before, (2) reduce the bias when a parameter value rarely appears in the test set but coincidentally appears in a failing test case with a particular fault.

Table XII gives an example to compute the suspiciousness between parameter values and particular faults. The left part of this table gives 4 executed test cases with their outcomes, in which faults  $F_1$  and  $F_2$  are triggered. The right part shows the suspiciousness between each parameter value and these two faults. Specifically, consider the parameter  $P_1$  taking value 0. There are three test cases  $t_1$ ,  $t_2$ , and  $t_4$  that contain this parameter value, and only  $t_2$  triggered  $F_1$ . Hence,  $Sp(P_1 = 0, F_1) = \frac{f_1(o)}{all(o)+1} = \frac{1}{3+1} = 1/4$ .

Table XII. Suspiciousness example

ID	Test cases executed	Outcomes	Suspiciousness for $F_1$		
			$P_1$	$P_2$	$P_3$
$t_1$	(0, 0, 0)	$Pass$	0	1/4	0
$t_2$	(0, 1, 1)	$F_1$	1	0	1/4
$t_3$	(1, 1, 0)	$F_2$		1/4	1/3
$t_4$	(0, 1, 1)	$F_2$			
			Suspiciousness for $F_2$		
			$P_1$	$P_2$	$P_3$
			0	1/4	0
			1	1/2	1/3

With the *suspiciousness* associated with a parameter value, we then define the *suspiciousness* between a test case  $t$  and a particular fault  $F_i$  as:

$$Sp(t, F_i) = \frac{1}{k} \sum_{o \in t} Sp(o, F_i) \quad (EQ1)$$

where  $k$  is the number of parameters in  $t$ , and  $o$  is the specific parameter value in  $t$ . The *suspiciousness* between a test case and a fault is actually the average *suspiciousness* between each parameter value in the test case and this fault. For example, considering an unexecuted test case (1, 0, 0). With the suspiciousness of all the parameter values listed Table XII,  $Sp((1, 0, 0), F_1) = 1/3 \times (0 + 0 + 0) = 0$  and  $Sp((1, 0, 0), F_2) = 1/3 \times (1/2 + 0 + 1/3) = 5/18$ .

For a selected test case, we want its ability to trigger failures of other faults to be as small as possible, such that the masking effects can be alleviated. In practice, the *suspiciousness* varies between test cases and different faults. As a result, we cannot always find a test case that, for any fault, the *suspiciousness* between this test case and that fault is the least.

Table XIII illustrates such a scenario for SUT(2<sup>4</sup>). Suppose the FII approach is analysing MFS for fault  $F_1$ , and needs to replace test case (0,0,0,0) with fixed part (0, -, -, -) that triggers failures of other faults, i.e.,  $F_2$  or  $F_3$ . This table lists five candidate test cases, with their suspiciousness with  $F_2$  and  $F_3$  given in the corresponding columns. It is obvious that  $t_1$  has the least suspiciousness with fault  $F_3$  and  $t_2$  has the least suspiciousness with  $F_2$ . These two test cases, however, should not be selected for their suspiciousness with another fault is too high. Instead,  $t_3$  is a good choice as both its suspiciousness with  $F_2$  and  $F_3$  is not high (The higher one is just 0.4).

Table XIII. Select minimal maximal suspiciousness

ID	Candidate test cases	$Sp(t, F_2)$	$Sp(t, F_3)$	$Max Sp(t, F_m), m = 2, 3$
$t_1$	(0, 0, 0, 1)	0.7	0.2	$Sp(t_1, F_2) : 0.7$
$t_2$	(0, 0, 1, 0)	0.2	0.6	$Sp(t_2, F_3) : 0.6$
<b><math>t_3</math></b>	<b>(0, 0, 1, 1)</b>	<b>0.4</b>	<b>0.3</b>	<b><math>Sp(t_3, F_2) : 0.4</math></b>
$t_4$	(0, 1, 0, 0)	0.3	0.5	$Sp(t_4, F_3) : 0.5$
$t_5$	(0, 1, 0, 1)	0.5	0.3	$Sp(t_5, F_2) : 0.5$

With this in mind, we have to settle for a test case, such that the most likely fault (except for the one that is currently analysed) it can trigger should be the least likely to be triggered when compared with that of other test cases. In other words, we need to find a test case, so that the maximal *suspiciousness* between this test case and the most likely fault it can trigger is minimal. Formally, we should choose a test case  $t$ , s.t.,

$$\min_{t \in R} \max_{1 \leq i \leq L \& i \neq m} Sp(t, F_i) \quad (\text{EQ2})$$

where  $L$  is the number of faults, and  $m$  is the current analysed fault.  $R$  is the set of all possible test cases that contain the *fixed* part except those that have been tested. As *fixed* part is a set of parameter values which can be deemed as a schema, then obviously  $R = \mathcal{T}(\text{fixed}) \setminus T_{executed}$ , where  $\mathcal{T}(\text{fixed})$  is all the test cases that contain this fixed part and  $T_{executed}$  represents those executed test cases. Additional test cases need to be selected in set  $R$ , so that FII approaches can work properly.

The complete process of replacing a test case with a new one while keeping some fixed part is depicted in Algorithm 1.

---

**ALGORITHM 1:** Replacing test cases triggering unexpected failures

---

**Input:** fault  $F_m$ , all the candidate test cases  $R$ , the suspiciousness matrix  $Sp$

**Output:**  $t_{new}$  the regenerate test case

---

```

1 while not  $R$  is empty do
2    $t_{new} \leftarrow \min_{t \in R} \max_{1 \leq i \leq L \& i \neq m} Sp(t, F_i);$ 
3    $result \leftarrow execute(t_{new});$ 
4    $R \leftarrow R \setminus t_{new};$ 
5    $update\_SP(t_{new});$ 
6   if  $result == PASS$  or  $result == F_m$  then
7     return  $t_{new};$ 
8   else
9     continue;
10  end
11 end
12 return null

```

---

The inputs to this algorithm consist of the fault  $F_m$  under analysis, the candidate test cases  $R = \mathcal{T}(\text{fixed}) \setminus T_{executed}$  and the suspiciousness matrix  $Sp$ , which records the suspiciousness between each factor  $o$  and each fault  $F_i$ , i.e.,  $Sp(o, F_i)$  ( $1 \leq i \leq L$ ). The output of this algorithm is a test case  $t_{new}$  which either triggers the expected  $F_m$  or passes.

The outer loop of this algorithm (lines 1 - 11) contains three parts:

The first part (lines 2 - 3) generates and executes a new test case which is supposed to be least likely to trigger faults different from  $F_m$ . The new test case is generated according to EQ2. In our implementation, we use the solver introduced in [Berkelaar et al. 2004], which is a mixed Integer Linear Programming (MILP) solver suitable for satisfaction and optimization problems. (Note that a simpler linear search algorithm

can also be applied in our approach to find a proper test case. However, as this problem to search a proper test case is related to Integers (a test case consists of discrete parameters with discrete values), so we use Integer Linear Programming (ILP) technique instead.) When a test case is generated, we remove it from the candidate test case set  $R$  to avoid redundancy (line 4).

The second part (line 5) updates the suspiciousness matrix ( $Sp$ ) for each parameter value that is involved in this newly generated test case (line 5). Specifically, for a particular parameter value  $o$ , the number of executed test cases that contain  $o$ , i.e.,  $all(o)$ , increases by 1. Additionally, if this test case triggers failure of  $F_i$  ( $1 \leq i \leq L$ ), then the number of test cases that contain  $o$  and trigger failure of  $F_i$ , i.e.,  $f_i(o)$ , increases by 1. At last, the suspiciousness value will be re-computed according to formula  $Sp(o, F_i) = \frac{f_i(o)}{all(o)+1}$  ( $1 \leq i \leq L$ ).

The last part (lines 6 - 10) checks whether the newly generated test case is as expected. Specifically, if the test case passes or triggers the same failure of fault -  $F_m$ , a satisfied test case is obtained (line 6) and returned (line 7). Otherwise, we will repeat the process, i.e., generate a new test case and check again (lines 8 - 9).

## 6. ILLUSTRATION OF THE APPROACH

This section will completely illustrate the MFS identification approach that combines the traditional MFS identification procedure with test cases replacement strategy.

### 6.1. MFS identification with replacement strategy

Algorithm 2 shows the procedure of the approach. The inputs to this algorithm are the fault  $F_m$  that is currently focused on and an original failing test case  $t$ . The output of this algorithm is to give the MFS for fault  $F_m$  in the test case  $t$ . Variable  $SP$  is to record the suspicious value between each parameter value with each fault.  $T_{Unknown}$  is the set of test cases that is not yet determined to be fail with fault  $F_m$  or not, and  $T_{F-Known}$  is the set of test cases that are deemed to trigger the fault  $F_m$ . This algorithm loops until some MFS are determined (line 25 - 27). This is based on Lemma 3.13 in Section 3.3. Specifically, by computing the minimal schemas of  $T_{F-Known}$ , i.e.,  $\mathcal{C}(T_{F-Known})$ , and minimal schemas of  $T_{F-Known} \cup T_{Unknown}$ , i.e.,  $\mathcal{C}(T_{F-Known} \cup T_{Unknown})$ , respectively, we can determine  $\mathcal{C}(T_{F-Known})$  are MFS if they are contained in  $\mathcal{C}(T_{F-Known} \cup T_{Unknown})$ .

In each iteration of this algorithm, it will generate one additional test case (line 7) to execute. This test case is randomly selected from a candidate set  $R$  which has the same fixed part (line 6). The fixed part is different in each iteration (See Algorithm 3 in the Appendix for detail). Then the approach will execute this test case (line 8) and update the suspicious matrix (line 9) according to the execution result. If the result is pass or fail with the same fault  $F_m$ , then the generated test case is a proper test case for MFS identification; otherwise, we should run the replacement strategy (line 10 - 15) based on Algorithm 1. Note that if we cannot find a proper test case by Algorithm 1, we need to randomly select one test case in  $R$  (line 13), and regard it as fail with  $F_m$  (line 14). Next we will infer the result of some unknown test cases according to the result of the generated test case  $t_{next}$  (line 17 - 24). This inferring procedure is exactly based on Lemma 3.15 and 3.14. After that, we should updated the  $T_{F-known}$  and  $T_{Unknown}$  (line 19 - 20, line 23), and re-check if the ending condition is satisfied (line 27).

### 6.2. A case study using the replacement strategy

In this section, we will give a case study to illustrate the approach. Suppose we have to test a system with eight parameters, each of which has three options, i.e., SUT( $3^8$ ). When we execute the test case  $t_0 = (0, 0, 0, 0, 0, 0, 0, 0)$ , a failure of fault- $e1$  is triggered.

**ALGORITHM 2:** MFS identification with Replacing test cases strategy

---

**Input:** fault  $F_m$ , original failing test case  $t$   
**Output:** MFS return the determined MFS

```

1  $Sp \leftarrow \text{init\_Suspicious\_Matrix}(t)$ ;
2  $T_{Unknown} \leftarrow T_{ALL} \setminus t$ ;
3  $T_{F-Known} \leftarrow \{t\}$ ;
4  $MFS \leftarrow \text{empty}$ ;
5 while true do
6    $R \leftarrow \mathcal{T}(\text{current\_fixed\_part})$ ;
7    $t_{next} \leftarrow \text{random\_pick}(R)$ ;
8    $result \leftarrow \text{execute}(t_{next})$ ;
9    $\text{update\_SP}(t_{next})$ ;
10  if  $result \neq \text{PASS}$  and  $result \neq F_m$  then
11     $t_{next} \leftarrow \text{Replacing}(F_m, Sp, R)$ ;
12    if  $t_{next} == \text{null}$  then
13       $t_{next} \leftarrow \text{random\_pick}(R)$ ;
14       $result \leftarrow F_m$ 
15    end
16  end
17  if  $result == F_m$  then
18     $T_{infer-F} \leftarrow \text{InferringFailing}(t_{next}, SV)$ ;
19     $T_{F-Known} \leftarrow T_{F-Known} \cup t_{next} \cup T_{infer-F}$ ;
20     $T_{Unknown} \leftarrow T_{Unknown} \setminus (t_{next} \cup T_{infer-F})$ ;
21  else
22     $T_{infer-P} \leftarrow \text{InferringPassing}(t_{next}, SV)$ ;
23     $T_{Unknown} \leftarrow T_{Unknown} \setminus (t_{next} \cup T_{infer-P})$ ;
24  end
25   $MFS \leftarrow \mathcal{C}(T_{F-Known})$ ;
26   $M_{Possible} \leftarrow \mathcal{C}(T_{F-Known} \cup T_{Unknown})$ ;
27  if  $MFS \subseteq M_{Possible}$  then
28    break;
29  end
30 end
31 return  $M_{Candidate}$ 

```

---

Furthermore, there are two more potential faults,  $e_2$  and  $e_3$ , that may be triggered during the testing and they will mask the desired fault  $e_1$ . Next, we will use FIC\_BS [Zhang and Zhang 2011] with replacement strategy to identify the MFS of  $e_1$ . The process is shown in Figure 4. In this figure, there are two main columns. The left main column indicates the executed test cases during testing as well as the executed results, with each executed test case corresponding to a specific label,  $t_1 - t_8$ , at the left. The underline part for each test case is the *fixed* part according to FIC\_BS [Zhang and Zhang 2011]. The right main column lists the suspiciousness matrix when a test case triggers  $e_2$  or  $e_3$ . The executed test case, shown in bold, indicates the one that triggers the failure of other faults and should be replaced in the next iteration.

The completed MFS identifying process listed in Figure 4 works as follows: firstly the original FII approach determines which *fixed* part needed to be tested in each iteration. Then an extra test case will be generated to fill in the remaining part. After executing the extra test case, if the result of the execution is normal, i.e., did not trigger any unexpected faults ( $e_2, e_3$ ), then the original FII process will continue until the MFS is identified. Otherwise, the replacement strategy starts when an unexpected failure of other faults is triggered. The replacement process will mutate the parameter values that are not in the *fixed* part according to EQ2. After the replacement process,

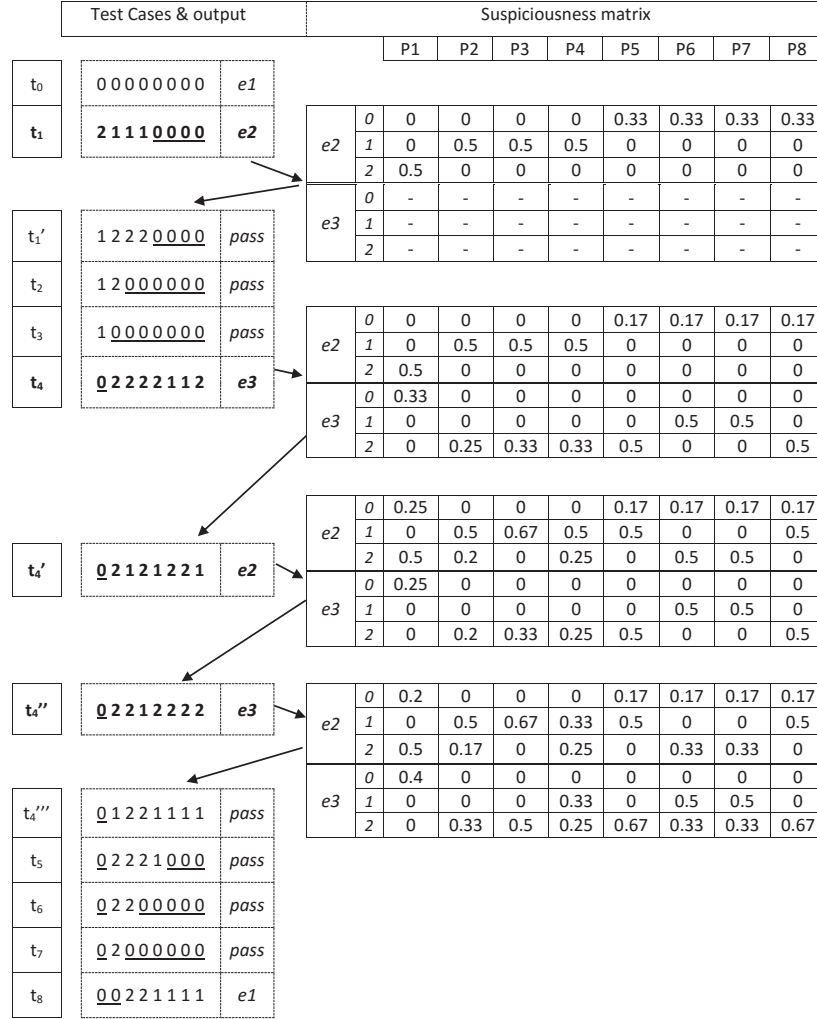


Fig. 4. A case study using our approach

the control for the MFS identifying process will be passed back to the original FII approach. Next we will specifically explain how the replacement works with an example in Figure 4.

From Figure 4, for the test case that triggered  $e2$  – (2, 1, 1, 1, 0, 0, 0, 0) (in this case, the fixed part of the test case is (-, -, -, -, 0, 0, 0, 0), in which the last four parameter values are the same as the original test case  $t_0$ ), we generate the related matrix at left. Each element in this matrix is computed according to EQ1. All the suspiciousness with  $e3$  is labeled with a short slash as there is no test case triggering this fault in this iteration. After this matrix has been determined, we can obtain the optimal test case with the ILP solver, which is  $t_1'$  – (1, 2, 2, 2, 0, 0, 0, 0), with its suspiciousness 0.167, which is smaller than all other test cases.

This replacement process is started each time a new test case that triggered another fault until we finally get one proper test case. Sometimes we could not find a satisfied replacing test case in just one trial like  $t_1$  to  $t_1'$ . When this happened, we needed to

repeat searching the proper test case. For example, for  $t_4$  which triggered  $e_3$ , we tried three times— $t'_4, t''_4, t'''_4$  to finally get a satisfied  $t'''_4$  which passes the testing. Note that the *suspiciousness* matrix continues to change as the test case is generated and executed so that we can adaptively find an optimal one.

With the replacement test cases, the FII approach can properly work. At last, the MFS identified for fault  $e_1$  is (0,0, -, -, -, -, -) (Test cases passed when we mutated the first two parameter values of the original failing cases). Note that, we did not show the theoretic process to obtain the MFS (0, 0, -, -, -, -, -). This is because we need to list all the possible test cases ( $3^8 = 6561$  test cases in total) to illustrate this process, which is not applicable in this paper.

Another notable point is that, in this example, test cases that trigger  $e_2$  and  $e_3$  do not contribute to the MFS identification. In real-world scenario, however, it is appealing to iteratively start MFS identification for  $e_2$  and  $e_3$  based on these test cases.

## 7. DISCUSSION ABOUT THE INFLUENCE OF THE ASSUMPTIONS

In section 3.1, we introduced three assumptions that can simplify our formal modeling and proposed approach, and in section 3.3, we introduced the safe value assumption which can reduce the number of test cases to identify MFS. In this section, we will discuss the influence of them on our propositions and approach, as well as some measures to alleviate their impacts.

### 7.1. Deal with non-deterministic problem

The first assumption is that the outcomes of all the tests are deterministic. In practice, re-executing the same test case may result in different outcomes. For example, if the program using some random variables, different runs of the test case will assign different values to these random variables. As a result, the control flow of the program may be changed and hence the outcome will be different. We called this type of failure the non-deterministic failure [Yilmaz et al. 2006; Fouché et al. 2009]. Non-deterministic failure will complicate the MFS identification, and even worse, it may lead to an unreliable result of the MFS identification. Consider the the following example in Table XIV, the only difference between the left with the right test set is the outcome of test case (0, 1, 1). This subtle difference result in the different results of the MFS identification. Hence, if there is one or more test cases which have non-deterministic executing results, the MFS identification is not reliable.

Table XIV. MFS of two similar test sets

$T_{fail}$	$T_{pass}$	MFS	$T_{fail}$	$T_{pass}$	MFS
(0, 0, 0)		(0, 0, -)	(0, 0, 0)		(0, -, -)
(0, 0, 1)		(0, -, 0)	(0, 0, 1)		
(0, 1, 0)			(0, 1, 0)		
	(0, 1, 1)		(0, 1, 1)		
	(1, 0, 0)			(1, 0, 0)	
	(1, 0, 1)			(1, 0, 1)	
	(1, 1, 0)			(1, 1, 0)	
	(1, 1, 1)			(1, 1, 1)	

Inspiring the idea that using multiple same-way covering arrays to identify the MFS [Fouché et al. 2009], one potential solution to alleviate this non-deterministic failure is by adding redundancy, i.e., through re-executing one test case to obtain the relatively stable outcome. However, this measure will increase the overall cost of the MFS identification, so the tradeoff between cost and the quality of MFS identification should be further studied.

### 7.2. Deal with failures distinguishing problem

The second assumption is that that different errors in the software can be easily distinguished by information such as exception traces, state conditions, or the like. If we cannot directly distinguish them, our approach does not work. This is because we cannot determine which test case should be replaced and should be replaced with what. In such case, one potential solution is to use the clustering techniques to classify the failures according to available information [Zheng et al. 2006; Jones et al. 2007; Podgurski et al. 2003]. If we cannot classify them because we do not have enough information (e.g., the black box testing) or it is too costly, we believe the only approach is to take the *regarded as same failure* strategy. With this strategy, we must aware that the MFS identified are likely to be sub-schemas or irrelevant schemas of the actual MFS.

### 7.3. Deal with inter-option constraints

The third assumption is that all the test cases are valid, i.e., SUT does not has any inter-option constraint. This is a very strong assumption. In this section, we will try to remove this assumption. In fact, with constraints has been known in prior, many studies in CT has proposed approaches to avoid them when generating test cases [Cohen et al. 2007a; 2007b; 2008; Calvagna and Gargantini 2008; Yu et al. 2013; Petke et al. 2013; Yu et al. 2015]. Generally, with constraints known in prior, not all all test cases in the SUT are valid. We set those valid test cases, i.e., do not contain any of the constraints as  $T_{valid}$ . Hence, the test cases we generated for MFS identification or replacement are limited to those in  $T_{valid}$ . Note that this can also make MFS identification unavailable. For example, in Table XV, assume test case(0, 1, 1) is invalid. Then without any more information, we can determine the MFS (See two different results in Table XIV).

Table XV. MFS identification with constraints

$T_{fail}$	$T_{pass}$	$T_{invalid}$
(0, 0, 0)		
(0, 0, 1)		
(0, 1, 0)		
		(0, 1, 1)
	(1, 0, 0)	
	(1, 0, 1)	
	(1, 1, 0)	
	(1, 1, 1)	

Another type of constraint is those that do not be found at first, but triggered later in testing[Yilmaz et al. 2014]. The same as the approach proposed in [Yilmaz et al. 2014], we can just treat them as one type of failure, and then use MFS identification approach to identify them when encounter those invalid test cases.

### 7.4. Deal with safe values

The last assumption is the safe value, i.e., each parameter has one value that is not the part of any MFS. With this assumption, based on Lemma 3.15 and 3.14, we can infer the result of many test cases without executing them. On the contrary, if this assumption does not hold, those test cases may be inferred to the wrong results. As a result, the MFS that identified by FII approaches are not accurate.

In theory, to solve this problem, we should exhaustively execute all the test cases of which the results could be inferred. This is a costly process, and hence, some approaches have been proposed to handle such problem. Martinez [Martinez et al. 2008; 2009] proposed two approaches without the safe value assumption, but both of them

either needs to know the number of the MFS or the degrees of them in prior. Then either through generating higher-way covering array or adaptively generating additional test case, they can determine the MFS. We [Niu et al. 2013] have also proposed one approach to alleviate this problem. In that work, we repeatedly check the same schema to reduce the impacts of non-safe values problem.

## 8. EMPIRICAL STUDIES

To investigate the impact of masking effects on FII approaches in real software testing scenarios and to evaluate the performance of our approach in handling this effect, we conducted several empirical studies. Each of the studies focuses on addressing one particular issue, as follows:

**Q1:** Do masking effects exist in real software that contains multiple faults?

**Q2:** How well does our approach perform compared to traditional approaches?

**Q3:** Is the ILP-based test case searching technique efficient compared to the random selection?

**Q4:** Compared to another masking effects handling approach FDA-CIT [Yilmaz et al. 2014], does our approach have any advantages ?

### 8.1. The existence and characteristics of masking effects

In the first study, we surveyed three kinds of open-source software systems to gain an insight into the existence of multiple faults and their effects. The software under study were HSQLDB, JFlex and Grep. The first is a database management software written in pure Java, the second is a lexical analyser generator, and the last one is a command-line utility for searching plain-text data sets for lines matching a regular expression. The reason that we chose these systems is because they contain different versions and are all highly configurable so that the options and their interactions can affect their behaviour. Additionally, they all have a developer community so that we can easily obtain the real bugs reported in the bug tracker forum. Table XVI lists the program, the versions surveyed, number of lines of uncommented code, number of classes in the project, and the bug's id <sup>4</sup> for each of the software.

Table XVI. Software under survey

software	versions	LOC	classes	bugs' id
HSQLDB	2.0rc8	139425	495	#981 & #1005
	2.2.5	156066	508	#1173 & #1179
	2.2.9	162784	525	#1286 & #1280
JFlex	1.4.1	10040	58	#87 & #80
	1.4.2	10745	61	#98 & #93
Grep	2.6.3	27046	156	#7600 & #29537
	2.22	48101	297	#33080 & #28588

*8.1.1. Study setup.* We first looked through the bug tracker forum and focused on the bugs which are caused by the options interactions. For each of them, we derived its MFS by analysing the bug description report and the associated test file which can reproduce the bug. For example, through analysing the source code of the test file of

<sup>4</sup><http://sourceforge.net/p/hsqldb/bugs>  
<http://sourceforge.net/p/jflex/bugs>  
<http://savannah.gnu.org/bugs/>



bug#981 for HSQLDB, we found the failure-inducing interaction for this bug is (*pre-parestatement, placeHolder, Long string*). These three parameter values together form the condition that triggers the bug. The analysed result was later regarded as the “prior MFS”.

We further built the testing scenario for each version of the software listed in Table XVI. The testing scenario is constructed so that we can reproduce different failures by controlling the inputs to the test file. For each version, the source code of the testing file as well as other detailed information is available at <http://gist.nju.edu.cn/doc/multi/>.

Next, we built the input model which consists of the options related to the failure-inducing interactions and additional options that are commonly used. The detailed models information is shown in Table XVII for HSQLDB, JFlex and Grep, respectively. Each table is organised into three groups: (1) *common options*, which lists the options as well as their values under which every version of this software can be tested; (2) *specific options*, under which only the specific version can be tested; and (3) *configure space*, which depicts the input model for each version of the software, presented in the abbreviated form  $\#values\#number\ of\ parameters \times \dots$ , e.g.,  $2^9 \times 3^2 \times 4^1$  indicates the software has 9 parameters that can take 2 values, 2 parameters 3 values, and only one parameter 4 values.

Table XVII. Input models of HSQLDB, JFlex and Grep

<i>HSQLDB</i>		<i>JFlex</i>		<i>Grep</i>		
common options	values	common options	values	common options	values	
Server Type	3	generation	3	-E	2	
existed form	2	charset	4	-i	2	
resultSetTypes	3	public	2	-V or -color	3	
resultSetConcurrencys	2	apiprivate	2	versions specific options values		
resultSetHoldabilitys	2	cup	2	2.6.3	ascii	2
StatementType	2	caseless	2		command	2
sql.enforce_strict_size	2	char	2		word	2
sql.enforce_names	2	line	2		charset	2
sql.enforce_refs	2	column	2		-A or -B or -C	3
versions specific options values		notunix	2		value	4
2.0rc8	more	yyeof	2	2.22	only-matching	2
	placeHolder	versions specific options values			-count	2
	cursorAction	1.4.1	hasReturn			
2.2.5	multiple		normal			
	placeHolder	1.4.2	lookAhead			
2.2.9	duplicate		type			
	defailure_commit		standalone			
versions Config space		versions Config space				
2.0rc8	$2^9 \times 3^2 \times 4^1$	1.4.1	$2^{10} \times 3^2 \times 4^1$			
2.2.5	$2^8 \times 3^3$	1.4.2	$2^{11} \times 3^2 \times 4^1$			
2.2.9	$2^8 \times 3^3$					

We then generated the exhaustive test set consisting of all possible interactions of these options. For each of them, we executed the prepared testing file. We recorded the output of each test case to observe whether there were test cases containing prior MFS that did not produce the corresponding bug. Later we refer to those test cases that contain the MFS but did not trigger the expected failure as the *masked* test cases.

**8.1.2. Results and discussion.** Table XVIII lists the results of our survey. Column “all tests” gives the total number of test cases executed. Column “failure” indicates the number of test cases that failed during testing. Specifically, we give the specific number of failing test cases of each fault (labeled in the form #n). Note that the faults we listed here include some uncontrolled dependencies, so there are more faults than Table XVI. Column “masking” indicates the specific number of test cases that are masked by each

Table XVIII. Number of failures and their masking effects

software	versions	all tests	failure			masking	total
HSQldb	2cr8	18432	#1(2304)	#2(1152)	#3(1152)	#1>#2#3(768)	768 (16.7%)
-	2.2.5	6912	#1(1728)	#2(1728)	-	#1>#2(576)	576 (16.7%)
-	2.2.9	6912	#1(2304)	#2(768)	#3(384)	#1>#2#3(960) #2>#3(768)	1728 (50%)
JFlex	1.4.1	36864	#1(12288)	#2(12288)	-	#1>#2(6144)	6144 (25%)
-	1.4.2	73728	#1(18432)	#2(18432)	-	#1>#2(6144)	6144 (16.7%)
Grep	2.6.3	384	#1(128)	#2(64)	#3(72)	#1>#2#3(80) #2>#3(16)	96 (36.4%)
-	2.22	576	#1(192)	#2(64)	#3(80)	#1>#2#3(80) #2>#3(16)	6144 (28.6%)

fault. In this column, we use the form ( $\#m > \#n\#n'...$ ) to indicate that fault  $\#m$  masks faults ( $\#n\#n'...$ ). The last column “total” shows the number of masked test cases in total (for all the faults). The percentage in the parentheses in this column indicates the proportion of masked test cases and the failing test cases.

We observed that for each version of the software under analysis listed in Table XVIII, test cases with masking effects do exist, i.e., test cases containing MFS did not trigger the corresponding bug. In fact, there are 768 out of 4608 test cases (about 16.7%) in hsqldb with 2cr8 version. This rate is about 16.7%, 50%, 25%, 16.7%, 36.4%, and 28.6% respectively, for the remaining software versions.

So the answer to **Q1** is that in practice, when SUTs have multiple faults, masking effects do exist widely.

It is notable that in Yilmaz’s [Yilmaz et al. 2014] paper, a similar study about the existence of the masking effects has been conducted. The main difference between that work and ours is that their work quantifies the impact of the masking effects as the number of  $\tau$ -degree schemas that only appear in the test cases that triggered failures of other faults. Here, the  $\tau$ -degree schemas can be either MFS or not. Our work, however, quantifies the masking effects as the number of test cases that are masked by different failures. These test cases should contain some MFS, i.e., they should have triggered the expected failure if they did not trigger any other different failure. The reason that we quantify the masking effects in such way is because our work seeks to overcome the masking effects in the MFS identifying process. As the test cases which contain the MFS but do not produce the corresponding failure will significantly affect the MFS identifying results, their number can better reflect the impact of the masking effects on the FII approach.

## 8.2. Comparing our approach with traditional approaches

The second study aims to compare the performance of our approach with traditional approaches in identifying MFS under the impact of masking effects. To conduct this study, we need to apply our approach and traditional algorithms to identify MFS in a variety of software and evaluate their results. The seven versions of software in Table XVI used as test objects are far from the requirement for a general evaluation. However, to construct real testing objects for evaluations is time-consuming. This is because we must carefully study the detail of that software as well as the bug tracker reports. To compromise, we synthesized 5 more testing objects. These synthesized objects are five small programs which can directly return outputs when executed with given inputs (details of them are also available at <http://gist.nju.edu.cn/doc/multi/>). These synthesized objects are created such that their testing models are similar to those real ones (We increase the number of faults in these synthesized objects to show the performance of these approaches with more faults).

Table XIX lists the testing model for both the real and synthetic testing objects. In this table, column ‘Object’ indicates the SUT under test. For the real SUT listed in

Table XIX. The testing models used in the case study

Object	Model	Faults	MFS of each fault
H2cr8	$2^9 \times 3^2 \times 4^1$	$e_1 > e_2 > e_3$	$(51, 60, 70)_{e_1}, (51, 82, 92)_{e_2}, (51, 82, 91)_{e_2}, (51, 83, 92)_{e_3}, (51, 83, 91)_{e_3}$
H2.2.5	$2^8 \times 3^3$	$e_1 > e_2$	$(61, 70)_{e_1}, (52)_{e_2}$
H2.2.9	$2^8 \times 3^3$	$e_1 > e_2 > e_3$	$(60)_{e_1}, (01, 51, 70)_{e_2}, (00, 51, 70)_{e_2}, (51, 70)_{e_3}$
J1.4.1	$2^{10} \times 3^2 \times 4^1$	$e_1 > e_2$	$(00)_{e_1}, (10)_{e_2}$
J1.4.2	$2^{11} \times 3^2 \times 4^1$	$e_1 > e_2$	$(10, 21)_{e_1}, (01)_{e_2}$
G2.6.3	$2^5 \times 3^1 \times 4^1$	$e_1 > e_2 > e_3$	$(00)_{e_1}, (11, 21)_{e_2}, (30, 41)_{e_3}, (31, 41)_{e_3}, (32, 41)_{e_3}$
G2.22	$2^4 \times 3^2 \times 4^1$	$e_1 > e_2 > e_3$	$(00)_{e_1}, (10, 23)_{e_2}, (11, 23)_{e_2}, (30, 40)_{e_3}$
syn1	$2^7 \times 3^2 \times 4^1$	$e_4 > e_3 > e_2 > e_1$	$(20, 70)_{e_1}, (31, 51)_{e_2}, (40)_{e_2}, (60, 72)_{e_3}, (61, 82)_{e_4}$
syn2	$2^4 \times 3^2 \times 4^1$	$e_3 > e_2 > e_1$	$(20, 30)_{e_1}, (20, 51)_{e_1}, (40, 61)_{e_2}, (31, 60)_{e_2}, (22, 43)_{e_3}$
syn3	$2^4 \times 3^3 \times 4^1$	$e_4 > e_3 > e_2 > e_1$	$(00, 10)_{e_1}, (11, 21)_{e_2}, (21, 31)_{e_2}, (41, 71)_{e_3}, (52, 62)_{e_4}$
syn4	$2^7 \times 3^2 \times 4^1$	$e_5 > e_4 > e_3 > e_2 > e_1$	$(00)_{e_1}, (11, 30)_{e_2}, (21)_{e_3}, (40, 50)_{e_4}, (60)_{e_5}$
syn5	$3^7$	$e_4 > e_3 > e_2 > e_1$	$(00)_{e_1}, (20, 30)_{e_2}, (21, 41)_{e_2}, (12, 22)_{e_3}, (02, 60)_{e_4}$

Table XVI, we label the seven software as  $H2cr8$ ,  $H2.2.5$ ,  $H2.2.9$ ,  $J1.4.1$ ,  $J1.4.2$ ,  $G2.6.3$  and  $G2.22$  respectively. While for the synthesized ones, we label them in the form of ‘syn+ id’. Column ‘Model’ presents the input space for each testing object. Column ‘Faults’ shows the different faults in the software and their masking relationships. In this column, ‘>’ means the left fault will mask the right fault, i.e., if the a failure of the left fault is triggered, then the failure of the right fault will not be triggered. Furthermore, ‘>’ is transitive so that the left fault can mask all the faults in the right. For example, for the  $H2cr8$  object, we can find three faults :  $e_1$ ,  $e_2$ , and  $e_3$ . By using the formula  $e_1 > e_2 > e_3$ , we indicate that fault  $e_2$  will mask  $e_3$ ; and  $e_1$  will mask both  $e_2$  and  $e_3$ . Here for the simplicity of the experiment, we did not build more complex testing scenarios such as the masking effects can be in the form  $e_1 > e_2$ ,  $e_2 > e_3$ ,  $e_3 > e_1$  or even  $e_1 > e_2$ ,  $e_2 > e_1$ . The last column shows the MFS of each fault. The MFS is presented in an abbreviated form  $\{\#index_{\#value}\}_{fault}$ , e.g., for the object  $H2cr8$ ,  $(51, 60, 70)_{e_1}$  actually means  $(-, -, -, -, -, 1, 0, 0, -, -, -, -)$  is the MFS of fault  $e_1$ .

**8.2.1. Study setup.** After preparing the objects under testing, we then applied our approach (FIC\_BS with replacement strategy) to identify the MFS. Specifically, for each SUT we selected each test case that failed during testing and fed it into our FII approach as the input. Then, after the identifying process was completed, we recorded the identified MFS and the extra test cases needed. For the traditional FIC\_BS approach, we designed the same experiment. But as the objects being tested have multiple faults for which the traditional FIC\_BS can not be applied directly, we adopted two traditional strategies on the FIC\_BS algorithm, i.e., *regarded as same failure* and *distinguishing failures* as described in Section 4.3. The purpose of recording the generated additional test cases is to quantify the additive cost of our approach.

We next compared the identified MFS of each approach with the prior MFS to quantify the degree that each suffers from masking effects. There are five metrics used in this study, listed as follows:

- (1) *Accurate number* : the number of identified MFS which are actual prior MFS.
- (2) *Super number*: the number of identified MFS that are the super schemas of some prior MFS.
- (3) *Sub number* : the number of identified MFS that are the sub schemas of some prior MFS.
- (4) *Ignored number* : the number of schemas that are in the prior MFS, but irrelevant to the identified MFS.

- (5) *Irrelevant number* : the number of schemas in the identified MFS that are irrelevant to the prior MFS.

Among these five metrics, the *accurate number* directly indicates the effectiveness of the FII approaches, since to identify as many actual MFS as possible is the target for every FII approach. Metrics *ignored number* and *irrelevant number* indicate the extent of deviation for the FII approaches; specifically, the former indicates how much information about the MFS will miss, while the latter indicates how serious the distraction would be due to the “useless” schemas identified by the FII approach. *Super number* and *sub number* are the metrics in between, i.e., to identify some schemas that is *super* or *sub* schemas of the actual MFS is better than identifying *irrelevant* ones or ignoring some MFS, but it is worse than identifying the schema that is identical to some actual MFS. This is intuitive, as given the *super* / *sub* schemas, we just need to *remove* / *add* some elements of the original schemas to get the actual MFS. While for the *irrelevant* or *ignore* schemas, however, more efforts will be needed (e.g., both *adding* and *removing* operations will be needed to revise the irrelevant schemas to the actual MFS).

Besides these specific metrics, we also define a composite metric to measure the overall performance of each approach. The composite metric *aggregate* is defined as follows:

$$\text{Aggregate} = \frac{\text{accurate} + \text{related}(\text{super}) + \text{related}(\text{sub})}{\text{accurate} + \text{super} + \text{sub} + \text{irrelevant} + \text{ignored}}$$

In this formula, *accurate*, *super*, *sub*, *irrelevant*, and *ignored* represent the value of specific metric. To refine the evaluation of different *super* / *sub* schemas, we design a *related* function which gives the similarity between the schemas (either *super* or *sub*) and the real MFS, so that we can quantify the specific effort for changing a *super* / *sub* schema to the real MFS. The similarity between two schemas  $c_1$  and  $c_2$  is computed as:

$$\text{Similarity}(c_1, c_2) = \frac{\text{number of same elements in } c_1 \text{ and } c_2}{\max(\text{Degree}(c_1), \text{Degree}(c_2))}$$

For example, the similarity of (- 1 2 - 3) and (- 2 2 - 3) is  $\frac{2}{3}$ . This is because (- 1 2 - 3) and (- 2 2 - 3) have the same third and last elements, and both of them are 3-degree.

The *related* function is the summation of similarity of all the *super* or *sub* schemas with their corresponding MFS.

**8.2.2. Results and discussion.** Figure 5 depicts the results of the second case study. There are seven sub-figures in this figure, i.e., Figure 5(a) to Figure 5(g). They indicate the results of the number of accurate MFS each approach identified, the number of identified schemas which are the sub-schema / super-schema of some prior MFS, the number of ignored prior MFS, the number of identified schemas which are irrelevant to all the prior MFS, the aggregate value, and the extra test cases each algorithm needed, respectively.

For each sub-figure, there are four polygonal lines, each of which shows the results for one of the four strategies: *regarded as same failure*, *distinguishing failures*, *replacement strategy based on ILP searching*, *replacement strategy based on random searching* (The last one will be discussed in the next case study). Specifically, each point in the polygonal line indicates the specific result of a particular strategy for the corresponding testing object. For example in Figure 5(a), the point marked with ‘■’ at (1,0.25) indicates that the approach using *regarded as same failure* strategy identified 0.25 ac-



Fig. 5. Result of the evaluation for second case study

curate MFS on average for one failing test case of the testing object–HSQLDB 2cr8. The raw data for this experiment can be found in Table XXI of the Appendix. Note that all these data are average values, i.e., the average performance of these approaches when identifying the MFS in each failing test case.

**Accurate number:** Figure 5(a) shows the average number of accurate schemas that each approach achieved. It appears that *ILP* performed the best among the three approaches. In fact, for most testing objects (testing objects 1, 4, 5, 6, 7, 9, 10, 11 and 12), *ILP* either obtained the most number of accurate MFS, or tie with the most number of accurate MFS. The second best approach is *distinguishing failures*, which obtained better results than *regarded as same failure* for testing objects 1, 6, 8, 9, 10, 11, and 12.

**Sub number & super number:** Figure 5(b) and 5(c) depicts the results for *sub number* and *super number*, respectively. These two figures firstly showed a clear trend for strategies *regarded as same failure* and *distinguishing failures*, i.e., the former identified more sub schemas of actual MFS than the latter, while the latter identified more super schemas of actual MFS than the former. This is consistent with our formal analysis in Section 4.1 and Section 4.2.

The performance of our strategy *ILP* for these two metrics are in between. Specifically, *ILP* identify more sub schemas than strategy *distinguishing failures* but fewer than *regarded as same failure*; and *ILP* identify more super schemas than *regarded as same failure*, but fewer than *distinguishing failures*.

**Ignore number & irrelevant number:** The results of the two negative performance metrics are given in Figure 5(d) and 5(e), respectively. One observation is that, comparing with strategy *regarded as same failure*, *distinguishing failures* obtained fewer irrelevant schemas, but ignored more actual MFS. This is also consistent with formal analysis in Section 4. Note that for metric *irrelevant number*, strategy *regarded as same failure* obtained significantly much more irrelevant schemas than the remaining strategies, which make it hard to distinguish each other if we post them together into one figure. Hence, we draw this figure with only three polygonal lines (for *distinguishing failures*, *ILP*, and *random* respectively). Besides this, we offer a smaller figure in Figure 5(e), which includes the strategy *regarded as same failure*, to depicts the comparison of all these approaches.

The second observation is that *ILP* did a good job at reducing the scores for these two negative metrics. Specifically, for *ignored number*, our approach performed better than strategy *distinguishing failures* at the most testing objects (1, 3, 7, 10, 11, 12) in Figure 5(d), but is not as good as strategy *regarded as same failure*. In fact, strategy *regarded as same failure* has a significant advantage at reducing the number of ignored MFS as it tends to associate the failures with all the failing test cases. However, when we consider the *irrelevant number*, we can find that our approach is the best among all three strategies (better than *distinguishing failures* at testing objects 1, 3, 6, 7, 11, 12 in Figure 5(e), and better than strategy *regarded as same failure* for almost all the testing objects). We believe this improvement is caused by our test cases replacing strategy, as it can increase the test cases that are useful for identifying the MFS and decrease those useless test cases.

**Aggregative for the five metrics:** The composite results are given in Figure 5(f). This metric gives an overall evaluation of the quality of the identified schemas. From this figure, we can find that *ILP* performed the best, next the *distinguishing failures*, the last is the *regarded as same failure* (See the testing objects 1, 3, 4, 6, 7, 10, 11, and 12 in Figure 5(f)).

It is as expected that *ILP* performed better than *distinguishing failures* as it is actually the refinement version of latter. In fact, *ILP* also make the failures distinguished from each other. The main difference between *ILP* and *distinguishing failures* is that

the former has to replace the test cases that triggered any failure other than the currently analysed one while the latter will not change the generated test cases.

It is a bit of surprise to find, however, that strategy *distinguishing failures* performed better than *regarded as same failure* at almost all the testing objects. This result cannot be derived from the formal analysis. We believe the reason is that the masking effects are *monotonic* in these testing objects we constructed for evaluation. That is, it can only appear that bug *A* always mask bug *B*, but cannot have the circumstances that bug *A* can mask bug *B* and bug *B* can also mask bug *A*. This condition is favorable for the *distinguishing failures* strategy. For example, assume bug *A* masks bug *B*; then when we identify the MFS of bug *A*, the *distinguishing failures* is the correct strategy, as if there is a test case trigger the bug *B*, then it must not trigger the bug *A* (otherwise, bug *B* will not be triggered). Hence, there is a probability that is up to 50 that *distinguishing failures strategy* makes the correct operation.

**Test cases:** The number of test cases generated for identifying the MFS indicates the cost of FII approach. The result is shown in Figure 5(g). We can find that strategy *ILP* generated more test cases than the other strategies. In specific, the gap between the *ILP* and other two strategies ranged from about 2 to 5. This is acceptable when comparing to all the test cases that each approach needed. The increase in test cases for our approach is necessary, as additional test cases must be generated when some test cases cannot help to identify the MFS of the currently analysed failure. As for strategies *distinguishing failures* and *regarded as same failure*, there is no significant difference between the number of test cases generated.

Above all, we draw three conclusions, which help to answer **Q2**:

1) *Distinguishing failures* strategy obtained more *super number* and *ignored number* than *regard one failure* strategy, while the latter identified more *sub number* and *irrelevant number* than the former. This result is consistent with the previous formal analysis in Section 4.

2) Considering the quality of the MFS each approach identified, we can find that our *ILP* approach achieves the best performance, followed by the strategy *distinguishing failures*.

3) Although our approach need more test cases than the other two strategies, it is acceptable.

### 8.3. Evaluating the ILP-based test case searching method

The third empirical study aims to evaluate the efficiency of the ILP-based test case searching component of our approach. To conduct this study, we implemented an FII approach which is also augmented by the *replacing test cases* strategy, with test case randomly replaced.

**8.3.1. Study setup.** The setup of this case study is based on the second case study, and uses the same SUT model as shown in Table XIX. We apply the new random searching based FII approach to identify the MFS in the prepared SUTs. To avoid the bias coming from the randomness, we repeat the new approach 30 times to identify the MFS in each failing test case. We then compute the average additional test cases as well as other metrics listed in section 8.2.1 for the random-based approach. As the approach involves the random replacement, we conducted t-test on the original 30 groups of data to test whether the difference between the random-based approach and the ILP-based approach is of statistical significance or not.

**8.3.2. Results and discussion.** The evaluation of this random-based approach is also shown in Figure 5, in which the polygonal line marked with ‘x’ in each sub-figure indicates the results. The raw data can also be found in the Row ‘R’ of Table XXI in the appendix. It is noted that for the random approach we just give the average re-

sult of 30 tries. We offer the p-value in the Row ‘P(I,R)’ of Table XXI. Note that A p-value smaller than 0.05 indicates that the performance of these two approaches are statistically different with each other with 95% confidence.

Compared to the ILP-based approach, we can firstly observe that there is little distinction between them in terms of the metrics super-schemas and sub-schemas. For metric accurate number, the ILP-based approach performs slightly better (see testing objects 1, 6, 7, 11, where ILP obtained more number of accurate MFS), and for metrics ignore and irrelevant number, the ILP-based approach also has a slightly better performance, i.e., ILP obtained smaller number of these schemas which is of no-use for testing. Combining these, we can learn that the ILP-based approach has a similar performance of the random-based approach at the quality of identified MFS (The former may perform slightly better at some testing objects). This is consistent to Figure 5(f).

The similar quality of MFS identification of these two approaches can also be embodied with the p-values we obtained. From Table XXI in the Appendix, we can learn that, for all these metrics except for ‘testNum’ (which will be discussed later), there exists some p-values that are *NaN* (Not a number). This is because for these testing objects, the results we obtained are the same for all the 30 tries. The reason why the results are the same is because the only random part of the approach is the test case replacement. This random process normally does not affect the determination of whether a schema is failure-inducing or not, as no matter which test case we choose, the replacement has the same ending condition, i.e., to find some test case which trigger the same failure or pass (The only exception is that different test case for replacement may introduce some newly MFS of the same failure, which will interfere the MFS identification process [Zhang and Zhang 2011]). As a result, the MFS identified will be the same for the 30 tries for most cases. Then as the random part normally does not the quality of the MFS identification, it is conceivable that these two approaches has a similar performance as they both use the *test case replacement* strategy.

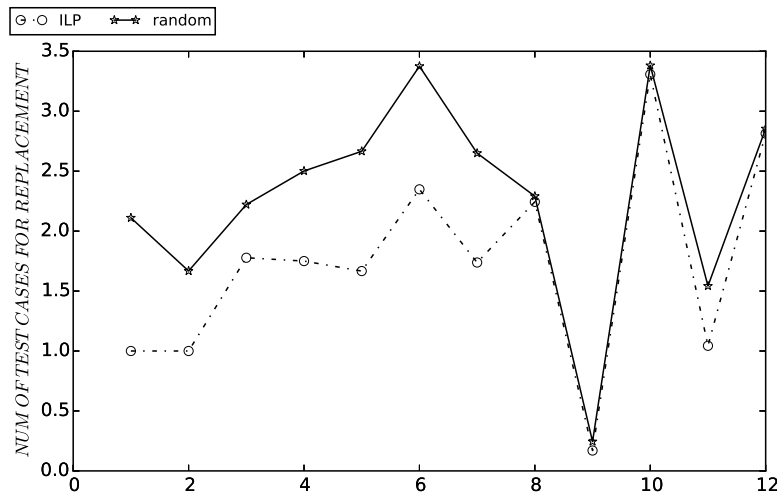


Fig. 6. The comparison of the number of replacement test cases

Secondly, when considering the cost, we find that the ILP-based approach performs better, which can reduce on average 1 to 2 test cases compared to the random-based



Table XX. The average test cases needed for one replacement

	H2cr8	H2.2.5	H2.2.9	J1.4.1	J1.4.2	G2.6.3	G2.22	syn1	syn2	syn3	syn4	syn5
ILP	1	1	1.77	1.75	1.66	2.34	1.73	2.24	0.16	3.3	1.04	2.81
Rand	2.11	1.66	2.22	2.5	2.66	3.37	2.64	2.29	0.24	3.38	1.54	2.85
$\mathcal{P}$	1.6E-60	1.1E-51	1.3E-51	2.5E-67	3.9E-70	8.0E-37	1.5E-42	2.5E-28	2.1E-28	1.1E-28	4.3E-45	3.2E-16

procedure (See testing objects of Figure 5(g)). This conclusion is statistically significant, as we can learn that the p-value of these two approaches for this metric is extremely small, which are ranged from 1.74E-26 to 1.26E-71 (See Row ‘testNum’ of Table XXI in the Appendix).

To be precise, we next compared the test cases only used for replacement of the two approaches. By this we can eliminate the interference of other test cases for identifying MFS, and focus on the performance of the key part of these two approaches — replacement strategy. Figure 6 shows the result. We can learn that on average, for each call of the replacement algorithm, it can save about 1 test case. This may be trivial. But considering the masking status in Table XVIII, this cost reduction is significant for MFS identifying in practice. We listed the original data and p-values on Table XX. It can be easily learned that the difference between these two approaches is also statistically significant, as the Row ‘ $\mathcal{P}$ ’ shows all these p-values are very small.

In summary, the answer for **Q3** is that searching for a satisfied test case affects the performance of our approach, especially regarding the number of extra test cases, and the ILP-based test cases can handle the masking effects at a relatively smaller cost than the random-based approach.

#### 8.4. Comparison with Feedback driven combinatorial testing

The *FDA-CIT* [Yilmaz et al. 2014] approach can handle masking effects so that the generated covering array can cover all the  $\tau$ -degree schemas without being masked by the MFS. There is an integrated FII approach in the FDA-CIT, which has two versions, i.e., *ternary-class* and *multiple-class*. In this paper, we use the multiple-class version for our comparative approach, as Yilmaz claims that it performs better than the former [Yilmaz et al. 2014].

The FDA-CIT process starts with generating a covering array (In [Yilmaz et al. 2014], this is a test case-aware covering array [Yilmaz 2013]). After executing the test cases in this covering array, it records the outcome of each test case and then applies the classification tree method on the test cases to characterize the MFS of each fault. It then labels these MFS as the schemas that can trigger masking effects. Later, if the interaction coverage is not satisfied (here the interaction coverage criteria is different from the traditional covering array [Yilmaz et al. 2014]), it will re-generate a covering array that aims to cover these schemas that were masked by these MFS and then repeat the previous steps.

The main target of FDA-CIT is to guarantee that the generated test cases should cover all the  $\tau$ -degree schemas. To achieve this goal, FDA-CIT needs to repeatedly identify the schemas that can trigger the masking effects. So to make the two approaches (FDA-CIT and ILP) comparable, we need to collect all the MFS that FDA-CIT characterized in each iteration and then compare them with the MFS identified by our approach.

**8.4.1. Study setup.** As FDA-CIT used a post-analysis (classification tree) technique on covering arrays, we first generated 2 to 4 ways covering arrays. The covering array generating method is based on augmented simulated annealing [Cohen et al. 2003], as it can be easily extended with constraint dealing and seed injection [Cohen et al. 2007b], which is needed by the FDA-CIT process. As different test cases will influence the results of the characterization process, we generated 30 different 2 to 4 way ini-

tial covering arrays. Then before we feed them to the FDA-CIT approach, we extend the corresponding  $\tau$ -way covering array to  $\tau+1$ -way covering array. This is because FDA-CIT needs higher-way covering arrays to identify MFS [Yilmaz et al. 2014]. The classification tree method which is integrated into FDA-CIT is implemented by Weka J48 [Hall et al. 2009], of which the configuration options are set as follows: the  $n$ -fold of cross validation is set to 10, the accuracy cutoff is set to be 1 and the default confidence factor is 0.25. After running FDA-CIT, we recorded the MFS identified, and by comparing them with prior actual MFS, we can evaluate the quality of the identified schemas according to the metrics mentioned in the previous case study.

Besides the FDA-CIT, we also applied our ILP-based approach to the initial  $\tau$ -way covering array. Specifically, for each failing test case (except test cases which contain existed identified MFS) in the covering array, we separately applied our approach to identify the MFS of that case. In fact, we can reduce the number of extra test cases if we utilize the other test cases in the covering array [Li et al. 2012]), but we did not utilize the information to simplify the experiment. Similarly, we then recorded the MFS that are identified by our approach, and evaluate them according to the corresponding metrics. In addition, we recorded the overall test cases (including the initially generated covering array) that this approach needed and compared the magnitude of these test cases with that of FDA-CIT.

**8.4.2. Result and discussion.** The result is shown in Figure 7. There are three sub-figures of Figure 7, which describe the results of the experiments based on 2-way, 3-way and 4-way covering arrays, respectively. In each sub-figure, there are 7 columns, showing the outcomes for the previous mentioned 6 metrics and one more metric (Column *Testcase*), which indicates the overall test cases that each approach needed. Each column has two bars, which indicate the results for approach FDA-CIT, and ILP respectively.

Note that in Figure 7, the results for each metric is the average evaluation for all the results of the experiments on the 12 testing objects in Table XIX. The raw results (as well as the p-values) for each testing object are listed in Table XXII in the appendix. The raw data is organised the same way as Table XXI.

From Figure 7, we have the following observations:

First, for every metric in our study, the performance of each approach is relatively stable against the change of degree  $\tau$ . The only exceptions are metrics *ignore number* and *irrelevant number*. With the increasing of  $\tau$ , the value of ignore number and irrelevant number of approach *FDA-CIT* decrease rapidly, while the performance of *ILP* is relatively steady when compared to *FDA-CIT*. Apart from these two metrics, the relationships between the two approaches of other metrics are stable. Take for example the metric *accurate number*. No matter what  $\tau$  is (2, 3 or 4), *ILP* always obtained the more accurate schemas than *FDA-CIT*. This observation indicates that the difference between the performance of these approaches is not dependent on the characteristics of the covering array, but instead on the approaches themselves.

From the perspective of p-values, we can also get the same observation. For all the metrics except *ignore number* and *irrelevant number*, the p-values are relatively small. Note that there also exists some test objects that the p-values are *NaN*. This is because that the MFS of these testing objects are relative simple (Only 2 - 3 number of MFS). As a result, both the two approaches can obtain the same schemas for the 30 tries. Besides the *NaN* values, other values of these metrics indicates that the difference between these two approaches are statically significant. As for metrics *ignore number* and *irrelevant number*, the p-values is relatively large. For example, for the 2-way experiment, some p-values of metric *ignore* are around 0.3 (Column '*syn1*') and 0.1

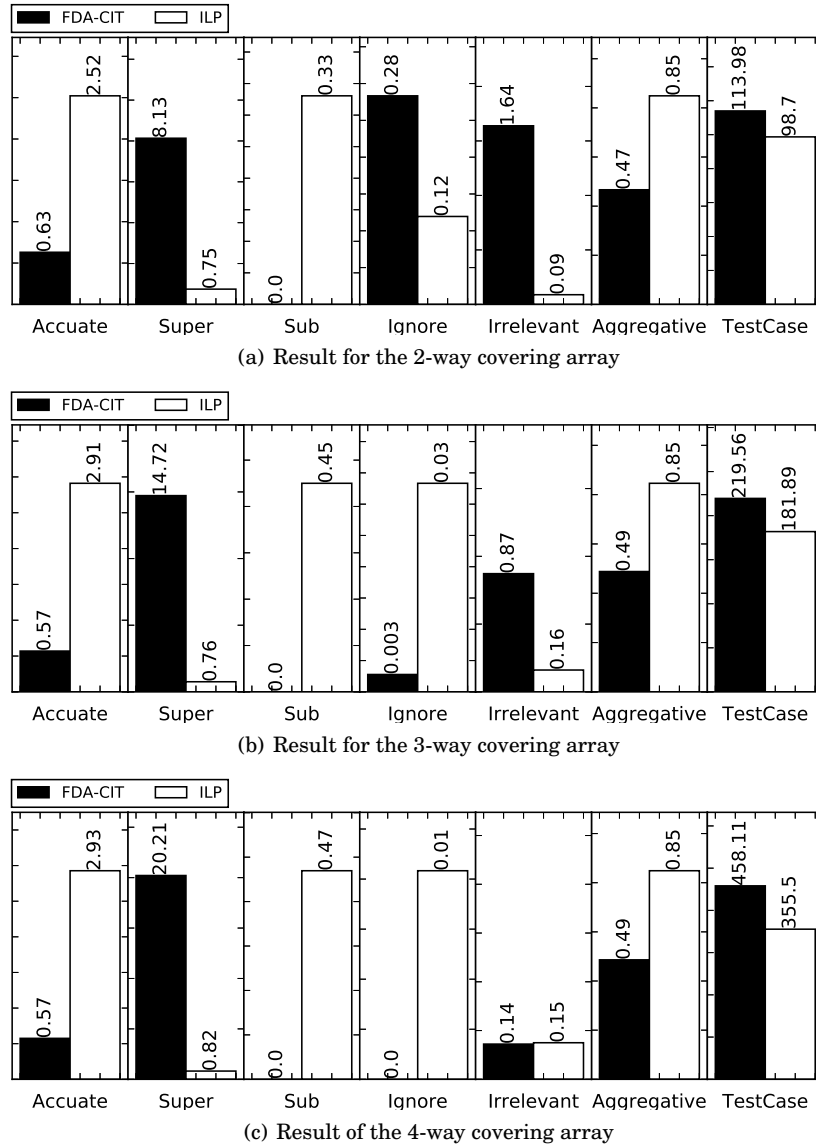


Fig. 7. Three approaches augmented with the replacing strategy

(Column ‘syn3’), which shows that there are no significant different between the two approaches for this metric.

Second, comparing to *ILP*, the schemas identified by *FDA-CIT* tend to be super schemas rather than sub schemas of the original MFS. We believe this result is due to the use of the classification tree analysis. Note that the way that CTA constructing decision trees usually work top-down, by choosing a parameter value at each step that best splits the set of test cases [Rokach and Maimon 2005]. After that, one path (conjunction of nodes from the root to one leaf in the tree) in this tree is deemed as an MFS. Consequently, the MFS identified by *FDA-CIT* tends to share some common nodes (here, a node represents a parameter value), e.g. the root node, which result in

the schemas identified by *FDA-CIT* may contain some unnecessary parameter value. This induces the so-called ‘over fitting’ problem. As a result, the schemas identified by *FDA-CIT* tend to be the super schemas of the actual MFS.

Third, in terms of the quality of the MFS identified, we can clearly find that our approach performed better than *FDA-CIT*. This is manifested in that our approach obtained more accurate schemas (the differences are statically significant for all the 3 degrees). We believe this gap is mainly caused by the *FII* approach. However, this result does not mean that *FIC\_BS* is better than the classification tree method under all conditions. In fact, the classification tree method has its own advantage, i.e., it is flexible and does not need to generate additional test cases. As a result, *FDA-CIT* can be applied on some complex testing scenarios, e.g., occasional failures, and incompatibilities between test cases and configurations [Yilmaz et al. 2014].

At last, when considering the cost of these two approaches, it is obviously that our approach needs smaller amount of test cases. However, note that this conclusion is depending on the fact that *FDA-CIT* is conducted on  $\tau+1$  way covering arrays, hence, *FDA-CIT* can get better performance when using  $\tau$ -way covering array.

Above all, we can conclude three points in this experiment, which provide answer to **Q4**:

1) The degree  $\tau$  of the covering array does not affect the overall performance of approaches, but for approach *FDA-CIT*, it may get better performance at reducing the *ignore number* and *irrelevant number* when using higher- $\tau$ -way covering array.

2) When taking account into the quality of MFS identification, *ILP* approach performs better.

3) Without additional measures, *FDA-CIT* is a better choice when handling some complex testing scenarios.

It is noted that *FDA-CIT*’s primary concern is to avoid masking effects and give every  $\tau$ -degree a fair chance to be tested, not to perform fault characterization. *FDA-CIT* is also an established technique which can work with non-deterministic failures and in the presence of inter-option constraints. So even though our approach *ILP* has shown a better performance at MFS identification, it is appealing and effective to use *FDA-CIT* to guide the test generation when taking account into the masking effects. In fact, considering *FDA-CIT* can work with other fault characterization approaches, it may be also appealing to combine these two approaches to MFS identification.

## 8.5. Threats to validity

**8.5.1. internal threats.** There are two threats to internal validity. First, the characteristics of the actual MFS in the SUT can affect the *FII* results. This is because the magnitude and location of the MFS can make the *FII* approaches generate different test cases. As a result, it can make the observed failing test cases and inferred failing test cases different. In the worst case, the *FII* approach happens to identify the exact actual MFS, then our test case replacing strategy is of no use. In this paper, we used 12 testing objects, in which 5 are real software systems with real faults and 10 synthetic ones with injected faults. To reduce the influence caused by different characteristics of the MFS, we need to build more testing objects and injected additional types of faults for a more comprehensive study of our approach.

The second threat is that we just applied our test case replacing strategy on one *FII* approach – *FIC\_BS* [Zhang and Zhang 2011]. Although we believe the test case replacing strategy can also improve the quality of the identified MFS for other *FII* approaches when the testing object is suffering from masking effects, the extent to which their results can be refined may vary for different *FII* approaches. For example, for *FIC\_BS* [Zhang and Zhang 2011] used in this paper, there are about  $(v-1)$  to  $(v-1)^{k-1}$  ( $k$  is the number of parameters in a test case,  $v$  is the number of values

each parameter can take) candidate test cases that can be replaced when one test case triggered failure of other fault, while for OFOT [Nie and Leung 2011a], there are  $(v-1)$  candidates. As a result, FIC\_BS can have a higher chance than OFOT to find a satisfied test case. To learn the difference between the improvement of different FII approaches when applying our test case replacing strategy, we need to try more FII approaches in the future.

**8.5.2. external threats.** One threat to external validity comes from the real software we used. In this paper we have only surveyed two types of open-source software with five different versions, of which the program scale is medium-sized. This may impact the generality of our results.

The second threat comes from the possible masking relationships between multiple faults in the real software. In this paper, we just focus on the condition that the masking effects are transitive, i.e., if fault  $A$  masks  $B$ , and fault  $B$  masks  $C$ , then fault  $A$  must mask fault  $C$ . In practice, the relationships between multiple faults may be more complicated. One possible scenario is that two faults are in a loop, for which they can even mask each other in a particular condition. Such a case will make our formal analysis invalid and will significantly complicate the relationships between schemas and their corresponding test cases. A new formal model should be proposed to handle that type of masking effects.

The third threat is that all the failures in the experiments are option-related. In practice, some failures may not related to the parameters you modeled in the SUT. For example, consider the Internet Explorer option-compliant testing problem [Nie et al. 2013]. Initially we may not properly model the options we should test, as a result, it can happen the explorer will always crash no matter we change which option in the initial model. This will cause the MFS identification non-sense, as all the test cases fail during testing. To solve this problem, one potential solution is to re-model the options we should test, as the error may be related to other options in the SUT (For example, those options which are set to default value). Or alternately, we should try other testing techniques to assist original CT. For example, if the error is related to the internal code instead of those configuration options, we may try programming slicing technique [Weiser 1981] or spectrum-based approaches [Naish et al. 2011]. In such case, the parameters that we model for the SUT should not only limited to be configuration options or simple inputs, but also those variables, predicates, or other logical structures and data in the software we test.

## 9. RELATED WORKS

Shi and Nie presented an approach for failure revealing and failure diagnosis in CT [Shi et al. 2005], which first tests the SUT with a covering array, then reduces the value schemas contained in the failing test case by eliminating those appearing in the passing test cases. If the failure-causing schema is found in the reduced schema set, failure diagnosis is completed with the identification of the specific input values which caused the failure; otherwise, a further test suite based on SOFOT is developed for each failing test case, and the schema set is then further reduced, until no more faults are found or the fault is located. Based on this work, Wang proposed an AIFL approach which extended the SOFOT process by adaptively mutating factors in the original failing test cases in each iteration to characterize failure-inducing interactions [Wang et al. 2010].

Nie et al. introduced the notion of Minimal Failure-causing Schema(MFS) and proposed the OFOT approach which is an extension of SOFOT that can isolate the MFS in the SUT [Nie and Leung 2011a]. This approach mutates one value for that parameter, hence generating a group of additional test cases each time to be executed. Compared

with SOFOT, this approach strengthens the validation of the factor under analysis and can also detect the newly imported faulty interactions.

Delta debugging [Zeller and Hildebrandt 2002] is an adaptive divide-and-conquer approach to locate interaction failure. It is very efficient and has been applied to real software environment. Zhang et al. also proposed a similar approach that can efficiently identify the failure-inducing interactions that has no overlapped part [Zhang and Zhang 2011]. Later, Li improved the delta-debugging based approach by exploiting useful information in the executed covering array [Li et al. 2012].

Colbourn and McClary proposed a non-adaptive method [Colbourn and McClary 2008]. Their approach extends a covering array to the locating array to detect and locate interaction failures. Martinez proposed two adaptive algorithms. The first one requires safe value as the assumption and the second one removes this assumption when the number of values of each parameter is equal to 2 [Martínez et al. 2008; 2009]. Their algorithms focus on identifying faulty tuples that have no more than 2 parameters.

Ghandehari et al. defined the suspiciousness of tuple and suspiciousness of the environment of a tuple [Ghandehari et al. 2012]. Based on this, they ranked the possible tuples and generated the test configurations. They further utilized the test cases generated from the inducing interaction to locate the fault [Ghandehari et al. 2013].

Yilmaz proposed a machine learning method to identify inducing interactions from a combinatorial testing set [Yilmaz et al. 2006]. They constructed a classification tree to analyze the covering arrays and detect potential faulty interactions. Beside this, Fouché [Fouché et al. 2009] and Shakya [Shakya et al. 2012] made some improvements in identifying failure-inducing interactions based on Yilmaz's work.

Our previous work [Niu et al. 2013] proposed an approach that utilizes the tuple relationship tree to isolate the failure-inducing interactions in a failing test case. One novelty of this approach is that it can identify the overlapped faulty interaction. This work also alleviates the problem of introducing new failure-inducing interactions in additional test cases.

In addition to the studies that aim at identifying the failure-inducing interactions in test cases, there are others that focus on working around the masking effects.

Constraints handling become more and more popular in CT these years. A constraint is an invalid interaction that should not appear in the test case. It can be deemed as the masking effect which are known in prior [Yilmaz et al. 2014]. Cohen [Cohen et al. 2007a; 2007b; 2008] studied the impact of the constraints that render some generated test cases invalid in CT. They also proposed an approach that integrates the incremental SAT solver with the covering arrays generating algorithm to avoid those invalid interactions. Further study was conducted [Petke et al. 2013] to show that with consideration of constraints, higher-strength covering arrays with early failure detection are practical.

Besides, there are additional works that aim to study the impact of constraints for CT [Garvin et al. 2011; Bryce and Colbourn 2006; Calvagna and Gargantini 2008; Grindal et al. 2006; Yilmaz 2013]. Among them, [Bryce and Colbourn 2006] distinguished the constraints into two types: *hard* and *soft*, which the former cannot be included in the test case, while the latter can be permitted, but not desirable. [Grindal et al. 2006] comprehensively compared the performance of four strategies at handling the constraints in the covering array. [Calvagna and Gargantini 2008] proposed an heuristic strategy to handle the constraints. It can support an ad-hoc inclusion or exclusion of interactions such that the user can customize output of the covering array. [Garvin et al. 2011] refined the simulated annealing algorithm to efficiently construct the covering array while considering the constraints. [Yilmaz 2013] introduced the test

case-specific constraints; differing from the system-wide constraints, these constraints can only be triggered in some specific test cases.

Chen et al. addressed the issues of shielding parameters in combinatorial testing and proposed the Mixed Covering Array with Shielding Parameters (MCAS) to solve the problem caused by shielding parameters [Chen et al. 2010]. The shielding parameters can disable some parameter values to expose additional interaction errors, which can be regarded as a special case of masking effects.

Dumlu and Yilmaz proposed a feedback-driven approach to work around the masking effects [Dumlu et al. 2011]. Specifically, they first used classification tree to classify the possible failure-inducing interactions and eliminate them. Then they generate new test cases to detect possible masked interaction in the next iteration. They further extended their work [Yilmaz et al. 2014] by proposing a multiple-class CTA approach to distinguish failures in the SUT. In addition, they empirically studied the impact of masking effects on both ternary-class and multiple-class CTA approaches.

These works can be categorized into 3 groups according to their relationships with our work. First, we discuss the works that aim to identifying the MFS in the SUT. Our work also focuses on identifying the MFS, but instead of single fault, our work considers the impact of multiple faults on the FII approaches, and based on this, a test case replacement strategy is proposed that can assist these FII approaches in reducing the negative effects. Second, the works that aim to deal with the constraints. As discussed before the constraints can be deemed as a special masking effect. Our work differs from them in that the masking effects handled in this paper are those that can be dynamically triggered; that is, we did not know them in prior. Another difference between our work with these constraints handling works is that their target is to avoid the constraints when generating covering array. However, our work aims to remove the masking effects of the FII approaches. Last, the work that is most similar to our work [Yilmaz et al. 2014], which also considered the masking effects that are dynamically appeared in test cases. But different from our work, it mainly focused on reducing the masking effects in the covering array, so that the covering array can support a comprehensive validation of all the  $\tau$ -degree schemas. The approach used to reduce this negative effect is to use the FII approach to identify the schemas that can trigger this effect in each iteration. Our approach, however, addresses the masking effects that happened in these FII approaches themselves, and our approach alleviates the masking effects by augmenting the FII approaches with a test case replacement strategy.

## 10. CONCLUSIONS

Masking effects of multiple faults in the SUT can bias the results of traditional failure-inducing interactions identifying approaches. In this paper, we formally analysed the impact of masking effects on FII approaches and showed that the two traditional strategies, i.e., *regarded as same failure* and *distinguishing failures*, are both inefficient in handling such impact. We further presented a test case replacement strategy for FII approaches to alleviate such impact.

In our empirical studies, we extended FIC\_BS [Zhang and Zhang 2011] with our strategy. The comparison between our approach and traditional approaches was performed on several open-source software. The results indicated our strategy assists the traditional FII approach in achieving better performance when facing masking effects in the SUT. We also empirically evaluated the efficiency of the test case searching component by comparing it with the random searching based FII approach. The results showed that the ILP-based test case searching method can perform more efficiently. Last, we compared our approach with existing technique for handling masking effects – FDA-CIT [Yilmaz et al. 2014], and observed that our approach achieved a more pre-

cise result which can better support debugging, though our approach required more test cases than FDA-CIT.

As for the future work, we need to do more empirical studies to make our conclusions more general. Our current experiments focus on medium-sized software. We would like to extend our approach to more complicated, large-scaled testing scenarios. Another promising work in the future is to integrate the white-box testing technique into the FII approaches. We believe gaining insight into source code can help figure out the relationships between multiple faults, and hence facilitate the FII approaches obtaining more accurate results. And last, because the extent to which the FII suffers from masking effects varies with different algorithms, combining these different FII approaches would be desired in the future to further improve identifying MFS of multiple faults.

## REFERENCES

- James Bach and Patrick Schroeder. 2004. Pairwise testing: A best practice that isn't. In *Proceedings of 22nd Pacific Northwest Software Quality Conference*. Citeseer, 180–196.
- Michel Berkelaar, Kjell Eikland, and Peter Notebaert. 2004. lp\_solve 5.5, Open source (Mixed-Integer) Linear Programming system. Software. (May 1 2004). <http://lpsolve.sourceforge.net/5.5/> Last accessed Dec, 18 2009.
- Renée C Bryce and Charles J Colbourn. 2006. Prioritized interaction testing for pair-wise coverage with seeding and constraints. *Information and Software Technology* 48, 10 (2006), 960–970.
- Renée C Bryce, Charles J Colbourn, and Myra B Cohen. 2005. A framework of greedy methods for constructing interaction test suites. In *Proceedings of the 27th international conference on Software engineering*. ACM, 146–155.
- Andrea Calvagna and Angelo Gargantini. 2008. A logic-based approach to combinatorial testing with constraints. In *Tests and proofs*. Springer, 66–83.
- Baiqiang Chen, Jun Yan, and Jian Zhang. 2010. Combinatorial testing with shielding parameters. In *Software Engineering Conference (APSEC), 2010 17th Asia Pacific*. IEEE, 280–289.
- David M. Cohen, Siddhartha R. Dalal, Michael L Fredman, and Gardner C. Patton. 1997. The AETG system: An approach to testing based on combinatorial design. *Software Engineering, IEEE Transactions on* 23, 7 (1997), 437–444.
- Myra B Cohen, Charles J Colbourn, and Alan CH Ling. 2003. Augmenting simulated annealing to build interaction test suites. In *Software Reliability Engineering, 2003. ISSRE 2003. 14th International Symposium on*. IEEE, 394–405.
- Myra B Cohen, Matthew B Dwyer, and Jiangfan Shi. 2007a. Exploiting constraint solving history to construct interaction test suites. In *Testing: Academic and Industrial Conference Practice and Research Techniques-MUTATION, 2007*. IEEE, 121–132.
- Myra B Cohen, Matthew B Dwyer, and Jiangfan Shi. 2007b. Interaction testing of highly-configurable systems in the presence of constraints. In *Proceedings of the 2007 international symposium on Software testing and analysis*. ACM, 129–139.
- Myra B Cohen, Matthew B Dwyer, and Jiangfan Shi. 2008. Constructing interaction test suites for highly-configurable systems in the presence of constraints: A greedy approach. *Software Engineering, IEEE Transactions on* 34, 5 (2008), 633–650.
- Myra B Cohen, Peter B Gibbons, Warwick B Mugridge, and Charles J Colbourn. 2003. Constructing test suites for interaction testing. In *Software Engineering, 2003. Proceedings. 25th International Conference on*. IEEE, 38–48.
- Charles J Colbourn and Daniel W McClary. 2008. Locating and detecting arrays for interaction faults. *Journal of combinatorial optimization* 15, 1 (2008), 17–48.
- Emine Dumlu, Cemal Yilmaz, Myra B Cohen, and Adam Porter. 2011. Feedback driven adaptive combinatorial testing. In *Proceedings of the 2011 International Symposium on Software Testing and Analysis*. ACM, 243–253.
- Sandro Fouché, Myra B Cohen, and Adam Porter. 2009. Incremental covering array failure characterization in large configuration spaces. In *Proceedings of the eighteenth international symposium on Software testing and analysis*. ACM, 177–188.
- Brady J Garvin, Myra B Cohen, and Matthew B Dwyer. 2011. Evaluating improvements to a meta-heuristic search for constrained interaction testing. *Empirical Software Engineering* 16, 1 (2011), 61–102.



- Laleh Sh Ghandehari, Yu Lei, David Kung, Raghu Kacker, and Richard Kuhn. 2013. Fault localization based on failure-inducing combinations. In *Software Reliability Engineering (ISSRE), 2013 IEEE 24th International Symposium on*. IEEE, 168–177.
- Laleh Shikh Gholamhossein Ghandehari, Yu Lei, Tao Xie, Richard Kuhn, and Raghu Kacker. 2012. Identifying failure-inducing combinations in a combinatorial test set. In *Software Testing, Verification and Validation (ICST), 2012 IEEE Fifth International Conference on*. IEEE, 370–379.
- Mats Grindal, Jeff Offutt, and Jonas Mellin. 2006. Handling constraints in the input space when using combination strategies for software testing. (2006).
- Mark Hall, Eibe Frank, Geoffrey Holmes, Bernhard Pfahringer, Peter Reutemann, and Ian H Witten. 2009. The WEKA Data Mining Software: An Update; SIGKDD Explorations, Volume 11, Issue 1. (2009).
- James A Jones, James F Bowring, and Mary Jean Harrold. 2007. Debugging in parallel. In *Proceedings of the 2007 international symposium on Software testing and analysis*. ACM, 16–26.
- Yu Lei, Raghu Kacker, D Richard Kuhn, Vadim Okun, and James Lawrence. 2008. IPOG/IPOG-D: efficient test generation for multi-way combinatorial testing. *Software Testing, Verification and Reliability* 18, 3 (2008), 125–148.
- Jie Li, Changhai Nie, and Yu Lei. 2012. Improved Delta Debugging Based on Combinatorial Testing. In *Quality Software (QSIC), 2012 12th International Conference on*. IEEE, 102–105.
- Conrado Martínez, Lucia Moura, Daniel Panario, and Brett Stevens. 2008. Algorithms to locate errors using covering arrays. In *LATIN 2008: Theoretical Informatics*. Springer, 504–519.
- Conrado Martínez, Lucia Moura, Daniel Panario, and Brett Stevens. 2009. Locating errors using ELAs, covering arrays, and adaptive testing algorithms. *SIAM Journal on Discrete Mathematics* 23, 4 (2009), 1776–1799.
- Lee Naish, Hua Jie Lee, and Kotagiri Ramamohanarao. 2011. A model for spectra-based software diagnosis. *ACM Transactions on software engineering and methodology (TOSEM)* 20, 3 (2011), 11.
- Changhai Nie and Hareton Leung. 2011a. The minimal failure-causing schema of combinatorial testing. *ACM Transactions on Software Engineering and Methodology (TOSEM)* 20, 4 (2011), 15.
- Changhai Nie and Hareton Leung. 2011b. A survey of combinatorial testing. *ACM Computing Surveys (C-SUR)* 43, 2 (2011), 11.
- Changhai Nie, Henry Leung, and Kai-Yuan Cai. 2013. Adaptive combinatorial testing. In *Quality Software (QSIC), 2013 13th International Conference on*. IEEE, 284–287.
- Xintao Niu, Changhai Nie, Yu Lei, and Alvin TS Chan. 2013. Identifying Failure-Inducing Combinations Using Tuple Relationship. In *Software Testing, Verification and Validation Workshops (ICSTW), 2013 IEEE Sixth International Conference on*. IEEE, 271–280.
- Justyna Petke, Shin Yoo, Myra B Cohen, and Mark Harman. 2013. Efficiency and early fault detection with lower and higher strength combinatorial interaction testing. In *Proceedings of the 2013 9th Joint Meeting on Foundations of Software Engineering*. ACM, 26–36.
- Andy Podgurski, David Leon, Patrick Francis, Wes Masri, Melinda Minch, Jiayang Sun, and Bin Wang. 2003. Automated support for classifying software failure reports. In *Software Engineering, 2003. Proceedings. 25th International Conference on*. IEEE, 465–475.
- Lior Rokach and Oded Maimon. 2005. Top-down induction of decision trees classifiers-a survey. *Systems, Man, and Cybernetics, Part C: Applications and Reviews, IEEE Transactions on* 35, 4 (2005), 476–487.
- Kiran Shakya, Tao Xie, Nuo Li, Yu Lei, Raghu Kacker, and Richard Kuhn. 2012. Isolating Failure-Inducing Combinations in Combinatorial Testing using Test Augmentation and Classification. In *Software Testing, Verification and Validation (ICST), 2012 IEEE Fifth International Conference on*. IEEE, 620–623.
- Liang Shi, Changhai Nie, and Baowen Xu. 2005. A software debugging method based on pairwise testing. In *Computational Science-ICCS 2005*. Springer, 1088–1091.
- Charles Song, Adam Porter, and Jeffrey S Foster. 2012. iTree: Efficiently discovering high-coverage configurations using interaction trees. In *Proceedings of the 2012 International Conference on Software Engineering*. IEEE Press, 903–913.
- Ziyuan Wang, Baowen Xu, Lin Chen, and Lei Xu. 2010. Adaptive interaction fault location based on combinatorial testing. In *Quality Software (QSIC), 2010 10th International Conference on*. IEEE, 495–502.
- Mark Weiser. 1981. Program slicing. In *Proceedings of the 5th international conference on Software engineering*. IEEE Press, 439–449.
- Cemal Yilmaz. 2013. Test case-aware combinatorial interaction testing. *Software Engineering, IEEE Transactions on* 39, 5 (2013), 684–706.
- Cemal Yilmaz, Myra B Cohen, and Adam A Porter. 2006. Covering arrays for efficient fault characterization in complex configuration spaces. *Software Engineering, IEEE Transactions on* 32, 1 (2006), 20–34.

- Cemal Yilmaz, Emine Dumlu, Myra B Cohen, and Adam Porter. 2014. Reducing Masking Effects in Combinatorial Interaction Testing: A Feedback Driven Adaptive Approach. *Software Engineering, IEEE Transactions on* 40, 1 (Jan 2014), 43–66.
- Linbin Yu, Feng Duan, Yu Lei, Raghu N Kacker, and D Richard Kuhn. 2015. Constraint handling in combinatorial test generation using forbidden tuples. In *Software Testing, Verification and Validation Workshops (ICSTW), 2015 IEEE Eighth International Conference on*. IEEE, 1–9.
- Linbin Yu, Yu Lei, Mehra Nourozborazjany, Raghu N Kacker, and D Rick Kuhn. 2013. An efficient algorithm for constraint handling in combinatorial test generation. In *Software Testing, Verification and Validation (ICST), 2013 IEEE Sixth International Conference on*. IEEE, 242–251.
- Andreas Zeller and Ralf Hildebrandt. 2002. Simplifying and isolating failure-inducing input. *Software Engineering, IEEE Transactions on* 28, 2 (2002), 183–200.
- Zhiqiang Zhang and Jian Zhang. 2011. Characterizing failure-causing parameter interactions by adaptive testing. In *Proceedings of the 2011 International Symposium on Software Testing and Analysis*. ACM, 331–341.
- Alice X Zheng, Michael I Jordan, Ben Liblit, Mayur Naik, and Alex Aiken. 2006. Statistical debugging: simultaneous identification of multiple bugs. In *Proceedings of the 23rd international conference on Machine learning*. ACM, 1105–1112.

## Online Appendix to: Identifying minimal failure-causing schemas in the presence of multiple faults

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### A. THE DETAIL ALGORITHMS OF FIC\_BS

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#### ALGORITHM 3: The algorithm of FIC\_BS

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**Input:** Failing test case  $t_{original}$ , Safe values of parameter  $SV$

**Output:** The  $MFS$  in the  $t_{original}$

```
1  $C_{free} \leftarrow NULL$  ;
2  $MFS \leftarrow NULL$  ;
3 while true do
4    $C_{cand} \leftarrow t_{original} \setminus C_{free} \setminus MFS$  ;
5   while  $|C_{cand}| > 1 \ \&\& \text{ContainMFS\_factor}(C_{cand})$  do
6      $(C_{low}, C_{high}) \leftarrow \text{BinaryPartition}(C_{cand})$  ;
7      $testCase \leftarrow \text{Mutate}(t_{original}, SV, C_{low} \cup C_{free})$  ;
8     if  $\text{Run}(testCase) == PASS$  then
9        $C_{cand} \leftarrow C_{low}$  ;
10    else
11       $C_{cand} \leftarrow C_{high}$  ;
12       $C_{free} \leftarrow C_{low} \cup C_{free}$  ;
13    end
14  end
15  if  $C_{cand} == NULL$  then
16    break ;
17  end
18   $MFS \leftarrow MFS \cup C_{cand}$  ;
19 end
20 return  $MFS$  ;
```

---

Here,  $C_{free}$  indicates the parameter values in  $t_{original}$  that is not related to the failure.  $C_{cand}$  are the set of parameter values in  $t_{original}$  that is needed to check in one iteration (Note that  $C_{cand}$  must contain at least one factor of  $MFS$  (line 5)). This algorithm consists of two loops. The outer loop (line 3 - 19) is to repeatedly search failure-inducing parameter values in  $t_{original}$  and append them in  $MFS$  (line 18), until none of failure-inducing parameter value is found (line 15). The inner loop (line 5 - 14) focus on finding one failure-inducing parameter value in  $C_{cand}$  (Those parameters which have been determined to be part of  $MFS$  or in  $C_{free}$  is omitted (line 4)). To reach this target, the inner loop repeatedly use binary search technique to reduce the scope of parameter values ( $C_{cand}$ ) that need to check (line 5 - 14). In each iteration,  $C_{cand}$  is split into two equally size part, i.e.,  $C_{low}$  and  $C_{high}$  (line 6). Then a new test case is generated by

mutating the  $C_{free}$  part and  $C_{low}$  part of the original failing test case  $t_{original}$  (line 7). For example, assume the original failing test case is (1, 1, 1, 1). The  $C_{free}$  is the second parameter value, i.e., (-, 1, -, -) and  $C_{low}$  part is the third parameter value (-, -, 1, -). Then a new test case may be generated like (1, 0, 0, 1). Note that the mutated values should be one of safe values set ( $SV$ ), and the parameter values that is not mutated in this iteration is called *fixed* part. If the newly generated test case passed (This indicates that MFS was broken. As  $C_{free}$  is not related to the failure, hence, at least one element in  $C_{low}$  is failure-inducing), then the scope of failure-inducing elements will be reduced to  $C_{low}$  (line 8 - 9). Otherwise, it will be reduced to  $C_{high}$  part (line 11) (This is because MFS was not broken, indicating that MFS is in  $C_{high}$ ). In this case,  $C_{low}$  will be appended to the  $C_{free}$  part (line 12), as it does not related to the MFS. The inter loop ends when  $C_{cand}$  has just one element (one failure-inducing parameter value is found, and should appended to MFS (line 18)), or is empty (no more failure-inducing parameter value can be found, the algorithm ends (line 15- 17)).

## B. THE DETAIL OF THE EXPERIMENTS

Table XXI. Result of the evaluation

		HSQL1	HSQL2	HSQL3	JFlex1	Jflex2	Grep1	Grep2	syn1	syn2	syn3	syn4	syn5
accurate	R <sup>1</sup>	0.25	0.83	0.72	1	0.83	0.59	0.7	0.39	0.88	0.55	0.54	0.28
	D <sup>2</sup>	0.66	0.83	0.72	0.75	0.83	0.62	0.7	0.55	0.96	0.76	0.71	0.65
	I <sup>3</sup>	0.83	0.83	0.66	1	0.83	0.78	0.92	0.48	0.98	0.93	0.78	0.76
	R <sup>4</sup>	0.54	0.83	0.66	1	0.83	0.76	0.85	0.55	0.98	0.93	0.64	0.76
	p(I,R) <sup>4</sup>	2.34E-57	NaN	NaN	NaN	NaN	3.29E-15	2.61E-27	1.87E-38	NaN	NaN	8.04E-42	NaN
super	R	0	0	0	0	0	0	0	0	0	0	0	0
	D	0.08	0.16	0.16	0.25	0.16	0.19	0.2	0.32	0.03	0.23	0.08	0.19
	I	0.16	0.16	0.11	0	0.16	0.06	0.02	0.32	0.01	0.06	0	0.07
	R	0.02	0.16	0.11	0	0.16	0.04	0.02	0.32	0.01	0.06	0	0.07
	p(I,R)	2.32E-56	NaN	NaN	NaN	NaN	2.92E-10	NaN	NaN	NaN	NaN	NaN	NaN
sub	R	0.58	0.16	0.27	0	0.16	0.45	0.32	0.41	0.03	0.61	0.44	0.91
	D	0.16	0	0.05	0	0	0.06	0.04	0.11	0	0.02	0.13	0.12
	I	0	0	0.22	0	0	0.06	0.04	0.18	0	0.06	0.21	0.15
	R	0.16	0	0.22	0	0	0.07	0.07	0.11	0	0.06	0.22	0.15
	p(I,R)	7.84E-47	NaN	NaN	NaN	NaN	1.25E-06	5.15E-12	1.87E-38	NaN	NaN	1.15E-19	NaN
ignore	R	0.16	0	0	0	0	0	0	0.25	0.07	0.02	0.05	0.08
	D	0.08	0	0.05	0	0	0.12	0.04	0.06	0	0.07	0.13	0.19
	I	0	0	0	0	0	0.09	0	0.06	0	0.03	0.07	0.15
	R	0.27	0	0	0	0	0.12	0.04	0.06	0	0.03	0.2	0.15
	p(I,R)	7.23E-56	NaN	NaN	NaN	NaN	2.15E-12	1.25E-22	NaN	NaN	NaN	5.90E-40	NaN
irrelevant	R	0.16	0.66	0.66	5.75	3.83	1.84	1.42	0.58	0.07	2.22	0.95	2.2
	D	0.08	0	0.05	0	0	0.12	0.04	0.27	0	0.29	0.12	0.03
	I	0	0	0	0	0	0.09	0	0.27	0	0.29	0.08	0
	R	0.27	0	0	0	0	0.12	0.04	0.27	0	0.29	0.2	0
	p(I,R)	7.23E-56	NaN	NaN	NaN	NaN	2.15E-12	1.25E-22	NaN	NaN	NaN	1.16E-36	NaN
aggregate	R	0.49	0.77	0.58	0.41	0.52	0.39	0.5	0.45	0.91	0.35	0.52	0.3
	D	0.84	0.91	0.86	0.87	0.91	0.78	0.85	0.72	0.99	0.75	0.77	0.76
	I	0.95	0.91	0.88	1	0.91	0.85	0.96	0.69	0.99	0.82	0.81	0.81
	R	0.66	0.91	0.88	1	0.91	0.82	0.9	0.73	0.99	0.82	0.67	0.81
	p(I,R)	2.00E-60	NaN	NaN	NaN	NaN	1.65E-17	1.59E-28	1.87E-38	NaN	NaN	2.64E-40	NaN
testNum	R	8.12	8.66	9.16	23.5	20.5	10.24	9.5	9.9	13.03	14.51	10.4	12.75
	D	11.91	7.66	8.61	6.5	9	7.09	6.47	13.02	13.69	9.96	8.49	8.33
	I	17	10.16	12	8	11.66	11.13	12.04	18.03	14.69	17.18	11.39	21.1
	R	17.72	11.29	13.13	9.67	13.11	13.56	14.26	18.45	14.83	17.66	12.49	20.47
	p(I,R)	7.19E-47	5.07E-50	8.32E-54	1.26E-71	8.87E-73	6.97E-40	2.75E-37	5.28E-39	1.74E-26	6.31E-42	5.79E-38	3.86E-44

<sup>1</sup> *O* denotes the strategy regarded as same failure.<sup>2</sup> *D* denotes the strategy distinguishing failures.<sup>3</sup> *I* denotes the replacement strategy based on ILP searching.<sup>4</sup> *R* denotes the replacement strategy based on randomly searching.<sup>4</sup> *p(I, R)* denotes the p-value of a test of significance is the probability that the results of ILP and Random are equal

Table XXII. Comparison with FDA-CIT

			HSQL1	HSQL2	HSQL3	JFlex1	Jflex2	Grep1	Grep2	syn1	syn2	syn3	syn4	syn5
2-way	accurate	Fda-cit	1.63	0	1	0	0	0.7	1	0.3	0.83	0.13	0.96	1
		ILP	3.66	2	2	2	2	2.2	3.93	1.23	1.56	4.03	1	4.63
		p-value	4.41E-12	NaN	NaN	NaN	NaN	1.12E-17	4.18E-28	5.95E-05	4.38E-08	1.77E-26	3.26E-01	4.14E-27
	super	Fda-cit	3.23	5.16	4.83	4	5.03	9.26	8.6	6.5	0.53	19.16	12.06	19.2
		ILP	1.33	0.36	0.26	0	0.4	0.8	0.03	1.6	0	0.53	3.63	0
		p-value	2.95E-05	1.83E-41	3.47E-20	NaN	1.25E-34	1.04E-16	2.93E-18	4.53E-12	3.65E-02	9.08E-18	4.86E-13	1.7E-26
	sub	Fda-cit	0	0	0	0	0	0	0	0	0	0	0	0
		ILP	0	0	1	0	0	0.66	0.03	0.9	0	0.63	0	0.73
		p-value	NaN	NaN	NaN	NaN	NaN	2.14E-08	3.26E-01	1.31E-07	NaN	8.72E-08	NaN	8.05E-10
	ignore	Fda-cit	1.2	0	0.2	0	0	0.09	0.06	0.93	0.63	0.06	0.2	0
		ILP	0.3	0	0	0	0	0.46	0	0.23	0	0.03	0.4	0
		p-value	1.78E-04	NaN	8.31E-02	NaN	NaN	2.02E-02	3.26E-01	7.9E-06	8.72E-08	5.62E-01	1.20E-01	NaN
	irrelevant	Fda-cit	2.53	0	0.1	0	0	2	1.93	3.1	0.7	4.63	1.56	3.1
		ILP	0	0	0	0	0	0.46	0	0.56	0	0	0	0.03
		p-value	1.48E-13	NaN	8.31E-02	NaN	NaN	2.05E-07	2.22E-07	1.59E-09	4.89E-07	2.32E-10	1.02E-09	1.61E-07
3-way	aggregate	Fda-cit	0.46	0.51	0.61	0.5	0.52	0.47	0.48	0.36	0.58	0.36	0.31	0.41
		ILP	0.88	0.93	0.82	1	0.93	0.63	0.98	0.59	1	0.86	0.6	0.92
		p-value	1.29E-16	6.32E-26	1.14E-13	NaN	8.79E-23	2.42E-09	4.88E-35	1.04E-07	1.04E-07	2.14E-30	1.13E-25	3.06E-43
	testNum	Fda-cit	72.33	62.13	61.76	73	72	106.36	162.53	83.76	77.76	204.53	120.16	271.4
		ILP	132.86	57.13	63.7	61.93	73.53	74.06	92.43	103.56	92.06	161.53	153.8	117.73
		p-value	1.65E-16	2.68E-01	6.66E-01	1.95E-26	4.13E-01	1.05E-03	1.24E-04	5.75E-03	6.61E-09	2.95E-02	1.99E-02	4.47E-06
	accurate	Fda-cit	0.86	0	1	0	0	0.96	1	0	1	0	1	1
		ILP	4.63	2	2	2	2	2.3	3.73	3.13	3.43	4.03	1	4.66
		p-value	4.14E-35	NaN	NaN	NaN	NaN	5.14E-15	1.89E-19	2.93E-21	7.95E-14	3.04E-22	NaN	1.7E-27
	super	Fda-cit	8.8	5	5.96	4	5	12.53	13.66	19.66	23.6	38.9	21.43	18.1
		ILP	1.33	0.4	0.43	0	0.4	0.83	0.1	1.2	0	0.43	3.96	0
		p-value	2.01E-14	7.79E-30	3.94E-37	NaN	7.79E-30	2.5E-21	3.5E-22	5.22E-25	9.59E-25	1.74E-30	2.69E-25	3.38E-53
	sub	Fda-cit	0	0	0	0	0	0	0	0	0	0	0	0
		ILP	0	0	1	0	0	0.9	0.13	0.9	1.2	0.66	0	0.6
		p-value	NaN	NaN	NaN	NaN	NaN	4.88E-16	4.34E-02	1.31E-07	8.42E-11	2.14E-08	NaN	3.15E-07
	ignore	Fda-cit	0	0	0	0	0	0	0	0	0	0	0.03	0
		ILP	0	0	0	0	0	0.1	0	0.2	0	0	0.09	0
		p-value	NaN	NaN	NaN	NaN	NaN	8.31E-02	NaN	3.14E-02	NaN	NaN	3.10E-01	NaN
	irrelevant	Fda-cit	0.8	0	0	0	0	1.2	0.96	1.23	2.16	2.53	1.13	0.43
		ILP	0	0	0	0	0	0.46	0	1.06	0.36	0	0	0.03
		p-value	7.43E-06	NaN	NaN	NaN	NaN	1.05E-04	8.56E-07	5.51E-01	2.42E-11	5.4E-10	1.78E-06	0.066167
	aggregate	Fda-cit	0.66	0.52	0.63	0.5	0.52	0.48	0.48	0.38	0.49	0.38	0.27	0.48
		ILP	0.94	0.93	0.81	1	0.93	0.66	0.96	0.67	0.77	0.86	0.64	0.93
		p-value	5.71E-30	5.47E-22	1.68E-35	NaN	5.47E-22	3.63E-17	5.37E-29	2.67E-10	6.57E-16	2.42E-24	2.53E-35	9.19E-38
	testNum	Fda-cit	182.36	130	129.13	190.03	186.96	185.73	228.8	243.8	359.8	310.16	171.8	316.13
		ILP	219.3	114.2	122.53	143.13	156.13	113.66	142.93	217.73	267.83	235.5	205.43	244.23
		p-value	2.69E-08	1.32E-11	8.66E-04	7.77E-54	2.31E-21	4.26E-04	2.10E-03	1.59E-04	9.18E-35	1.21E-02	2.55E-03	6.07E-10
4-way	accurate	Fda-cit	0.9	0	1	0	0	1	1	0	1	0	1	1
		ILP	5	2	2	2	2	2.16	3.76	3.2	3.5	3.86	1	4.63
		p-value	1.59E-34	NaN	NaN	NaN	NaN	1.59E-16	8.79E-21	7.47E-23	1.56E-12	5.52E-21	NaN	4.14E-27
	super	Fda-cit	8.9	5	6	4	5	22	19.26	25.7	29.23	66.96	32.46	18
		ILP	1.5	0.36	0.36	0	0.33	0.76	0.16	1.5	0	0.5	4.3	0
		p-value	5.78E-16	3.95E-30	1.43E-32	NaN	1.71E-30	1.71E-43	6.47E-37	8.4E-28	9.14E-25	4.1E-36	3.6E-28	NaN
	sub	Fda-cit	0	0	0	0	0	0	0	0	0	0	0	0
		ILP	0	0	1	0	0	1	0.13	0.93	1.13	0.7	0	0.73
		p-value	NaN	NaN	NaN	NaN	NaN	NaN	4.34E-02	1.35E-07	8.04E-09	4.54E-09	NaN	1.67E-08
	ignore	Fda-cit	0	0	0	0	0	0	0	0	0	0	0	0
		ILP	0	0	0	0	0	0	0	0.09	0	0	0	0
		p-value	NaN	NaN	NaN	NaN	NaN	NaN	NaN	1.84E-01	NaN	NaN	NaN	NaN
	irrelevant	Fda-cit	0	0	0	0	0	0	0	0.5	0	1.03	0.2	0
		ILP	0	0	0	0	0	0.56	0	1.1	0.13	0	0	0
		p-value	NaN	NaN	NaN	NaN	NaN	1.03E-06	NaN	1.42E-02	4.34E-02	2.28E-06	3.14E-02	NaN
	aggregate	Fda-cit	0.7	0.52	0.63	0.5	0.52	0.44	0.47	0.37	0.5	0.36	0.26	0.49
		ILP	0.94	0.93	0.82	1	0.94	0.64	0.95	0.67	0.81	0.85	0.64	0.93
		p-value	4.17E-40	2.36E-22	2.29E-32	NaN	8.75E-23	1.59E-18	4.52E-25	3.47E-11	1.54E-15	3.86E-24	2.02E-39	3.52E-30
	testNum	Fda-cit	421.8	295.13	295.73	451.66	456.4	241.16	358.56	490.2	975.93	562.76	298.7	649.3
		ILP	403.93	237.6	244.16	331.86	352.66	174.46	243.86	378.33	662.9	401.56	304.3	530.36
		p-value	4.32E-03	2.33E-36	4.39E-27	2.29E-64	1.62E-41	8.32E-04	2.41E-03	2.02E-19	2.46E-04	4.99E-06	7.41E-01	1.48E-19