STAT 630 Problem Set 9 Nikko Cleri November 17, 2021

Question I.

6.5.1

(a) We have shown previously that

$$L = \frac{1}{\sqrt{2\pi\sigma^2}} e^{\frac{(x_i - \mu)^2}{2\sigma^2}}$$

$$l = -\frac{n}{2} (\log 2\pi + \log \sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2$$

$$\frac{\partial l}{\partial \sigma^2} = \frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (x_i - \mu)^2$$

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \mu_0)^2$$

The Fisher information is

$$I(\theta) = -E_{\theta} \left[\frac{\partial^2}{\partial \theta^2} \log(f_{\theta}(X)) \right]$$

so we have for

$$nI(\sigma^2) = -nE_{\theta} \left[\frac{\partial}{\partial \sigma^2} \left(\frac{n}{2\sigma^2} + \frac{1}{2(\sigma^2)^2} \sum_{i=1}^n (x_i - \mu)^2 \right) \right]$$
$$= \frac{-n}{2\sigma^4} + \frac{n\sigma^2}{\sigma^6}$$
$$= \frac{n}{2\sigma^4}$$

(b) The asymptotic distribution is

$$\hat{\sigma}^2 \sim N(\sigma^2, 2\frac{\sigma^4}{n})$$

which does agree with the CLT.

Question II.

(a)

$$Y = -\log(X_i) \sim Exp(a)$$

The sum of exponentially distributed random variables of parameter a goes like Gamma(n, a). The expectation of \hat{a} is

$$\int -\frac{n}{y} f_Y(y) \, \mathrm{d}y$$
$$= \frac{n}{n-1} a$$

so the Bias is

$$\operatorname{Bias}_{\alpha}(\hat{\alpha} = \frac{n}{n-1}a - a$$

which is 0 in the large n limit.

(b) The Fisher information is

$$I(\hat{a}) = -E_{\hat{a}} \left[\frac{\partial^2}{\partial a^2} \log f \right]$$
$$= \frac{n}{a^2}$$

and the asymptotic normal distribution is $N(a, \frac{a^2}{n})$

Question III.

(a) The Fisher information is

$$I = -E_{\theta} \left[\frac{\partial^{2}}{\partial \theta^{2}} \log(f_{\theta}(X)) \right]$$
$$= -E_{\theta} \left(\frac{-T_{i}}{\lambda^{2}} \right)$$
$$= \frac{1}{\lambda}$$
$$nI = \frac{n}{\lambda}$$

The asymptotic distribution is $N(\lambda, \frac{\lambda}{n})$. The variance of \bar{T} is $\frac{n\lambda}{n^2} = \frac{\lambda}{n}$

(b) The log likelihood function is

$$l = \sum_{i=1}^{n} x_i \log \left(\lambda^{-1/2}\right)^{-2} - n((\lambda^{-1/2})^{-2}) + \sum_{i=1}^{n} x_i!$$

The MLE is

$$\hat{\lambda^{-1/2}} = \sqrt{\frac{n}{\sum_{i=1}^{n} x_i}}$$

The delta method gives

$$\sqrt{n} \left[\hat{\lambda^{-1/2}} - \hat{\lambda^{-1/2}} \right] \to N(0, \frac{1}{4\lambda^2 n})$$

(c) The results from R give $\lambda \approx 1.56$ and $\theta = 0.800$, with a standard error of 0.026.

Question IV.

6.3.1

The mean of the data is $\bar{x} = 4.88$. From the appendix,

$$c = Z_{0.975}$$

= 1.96

so we have

$$P(4.88 - 1.96\sqrt{\frac{0.5}{10}} < \mu < 4.88 + 1.96\sqrt{\frac{0.5}{10}}) = 0.95$$
 (1)

gives the 95 percent confidence interval (4.44, 5.32). The observed mean lies within the interval.

6.3.2

If we now drop the knowledge of the variance, we now have a t-test

$$c = t_{0.975}(9) = 2.262$$

and

$$S = \sqrt{\frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2}$$
$$= 0.69$$

which gives the confidence interval (4.38, 5.38), which contains the sample mean.

Question V.

6.3.8

The Wald confidence interval is (for $Z_0.95 = 1.645$)

$$0.62 \pm 1.645 \left(\frac{0.62(1 - 0.62)}{250}\right)^{1/2}$$

gives the interval (0.5695, 0.6705). The score interval is

$$\left(\hat{\theta} + \frac{Z^2}{2n} \pm Z \left(\frac{\hat{\theta}(1-\hat{\theta})}{n} + \frac{Z^2}{4n^2}\right)^{1/2}\right) \frac{1}{1 + \frac{Z^2}{n}}$$

which gives the interval (0.5685, 0.6690). Both the Wald and score intervals include the sample mean.

Question VI.

6.5.4

(a) The sample mean is $\hat{\lambda} = \bar{x} = 9.65$. The Wald interval for 95 percent confidence is

$$\bar{x} \pm Z \sqrt{\frac{\bar{x}}{n}}$$

which gives the interval (8.288, 11.012).

- (b) The asymptotic pivot $\frac{\hat{\lambda}-\lambda}{\sqrt{\lambda/n}} \sim N(0,1)$, from the notes, this gives the interval (6.71, 17.18).
- (c) From R: The Wald interval had 246 of 10000 sample means lie within the interval, while the score interval had all 10000 sample means lie within the interval.

Question VII.

6.5.5

From R we find the mean of the sample as $\bar{x}=1627$ and has variance $S^2=672476$. Using the MOME $\hat{\theta}=2/\bar{x}$, we have the interval using the score method

$$\frac{\hat{\theta}}{1 \pm \frac{Z_{0.95}}{\sqrt{n}}}$$

to give the interval (0.00093, 0.0018)

6.5.6 This gives the same result except with $\hat{\theta} = 1/\bar{x}$. This gives the interval (0.00047, 0.00090)