Due date: May 14, 2014

For this assignment you can do it in a group of maximum two people. Objectives of this assignment:

- To learn about Monte Carlo methods.
- To gain experience in applying MCMC techniques.

Altogether there are six questions. You will need to email your programs to the TA.

- 1. Use Monte Carlo to evaluate each of the following integrals:
 - (a) $\int_0^1 x^2 dx$
 - (b) $\int_0^1 \int_{-2}^2 x^2 \cos(xy) dx dy$
 - (c) $\int_0^\infty \frac{3}{4} x^4 e^{-x^3/4} dx$
- 2. Let

$$I = \frac{1}{\sqrt{2\pi}} \int_{1}^{2} e^{-x^{2}/2} dx.$$

Estimate I using importance sampling. Take g to be $N(1.5, \nu^2)$ with $\nu = 0.1, 1$ and 10. Plot a histogram of the values you are averaging to see if there are any extreme values.

3. We will approximate the following integral using both the simple Monte Carlo integration and the control variate method:

$$I = \int_0^1 \frac{1}{1+x} dx.$$

- (a) Set $h(x) = \frac{1}{1+x}$ and let U_1, \ldots, U_n be iid Unif[0,1]. Estimate I using $\hat{I}_{MC} = \frac{1}{n} \sum_i h(U_i)$ with n = 1500. The exact value for I is $\ln 2$.
- (b) Next introduce c(x) = 1 + x as a control variate, and estimate I with

$$\hat{I}_{CV} = \frac{1}{n} \sum_{i=1}^{n} h(U_i) - b \left[\frac{1}{n} \sum_{i=1}^{n} c(U_i) - E\{c(U)\} \right].$$

Note: you will need to analytically calculate $E\{c(U)\}$, and estimate the optimal value for b. Also use n=1500.

- (c) Estimate and compare the variances of \hat{I}_{MC} and \hat{I}_{CV} .
- (d) Can you design a new estimator for I that has a smaller variance than \hat{I}_{CV} ?
- 4. Consider a common application in statistics: three different treatments are to be compared by applying them to randomly selected experimental units. This, of course, usually leads us to "analysis of variance" using a model such as $y_{ij} = \mu + \alpha_i + e_{ij}$ with standard meanings of these symbols and the usual assumptions about the random component e_{ij} in the model. Suppose that instead of the usual assumptions, we assume that the e_{ij} have independent and identical double exponential distributions centered on zero.
 - (a) Describe how you would perform a Monte Carlo test instead of the usual ANOVA test. Be clear in stating the alternative hypothesis.
 - (b) Describe some other computer-intensive test that you could use even if you make no assumptions about the distribution of e_{ij} .

- 5. In the zero-inflated Poisson (ZIP) model, random data X_1, \ldots, X_n are assumed to be of the form $X_i = R_i Y_i$, where the Y_i 's have a Poisson(λ) distribution and the R_i 's have a Bernoulli(p) distribution, all independent of each other. Given an outcome $\mathbf{x} = (x_1, \ldots, x_n)$, the objective is to estimate both λ and p. Consider the following hierarchical Bayes model:
 - $p \sim \text{Uniform}(0,1)$ (prior for p),
 - $(\lambda|p) \sim \text{Gamma}(a,b)$ (prior for λ).
 - $(r_i|p,\lambda) \sim \text{Bernoulli}(p)$ independently (from the model above),
 - $(x_i|\mathbf{r},\lambda,p) \sim \text{Poisson}(\lambda r_i)$ independently (from the model above),

where a and b are known parameters, and $\mathbf{r} = (r_1, \dots, r_n)$. It follows that

$$f(\boldsymbol{x}, \boldsymbol{r}, \lambda, p) = \frac{b^a \lambda^{a-1} e^{-b\lambda}}{\Gamma(a)} \prod_{i=1}^n \frac{e^{-\lambda r_i} (\lambda r_i)^{x_i}}{x_i!} p^{r_i} (1-p)^{1-r_i}.$$

We wish to sample from the posterior pdf $f(\lambda, p, r|x)$ using the Gibbs sampler.

- (a) Generate a random sample of size n = 100 for the ZIP model using parameters p = 0.3 and $\lambda = 2$.
- (b) Show that
 - i. $(\lambda | p, \mathbf{r}, \mathbf{x}) \sim \text{Gamma}(a + \sum_{i} x_i, b + \sum_{i} r_i),$
 - ii. $(p|\lambda, \boldsymbol{r}, \boldsymbol{x}) \sim \text{Beta}(1 + \sum_{i} r_i, n + 1 \sum_{i} r_i),$
 - iii. $(r_i|\lambda, p, \boldsymbol{x}) \sim \text{Bernoulli}\left(\frac{pe^{-\lambda}}{pe^{-\lambda} + (1-p)I_{\{x_i=0\}}}\right)$.
- (c) Implement the Gibbs sampler: generate a large (dependent) sample from the posterior distribution and use this to construct 95% Bayesian confidence intervals for p and λ using the data in (a). Compare these with the true values. Try different values of a and b, but you can start with a = b = 1.
- 6. Independence-Metropolis-Hastings Algorithm is an importance-sampling version of MCMC. We draw the proposal from a fixed distribution g. Generally, g is chosen to be an approximation to f. The acceptance probability becomes

$$r(x,y) = \min \left\{ \frac{f(y)}{f(x)} \frac{g(x)}{g(y)}, 1 \right\}.$$

A random variable Z has a inverse Gaussian distribution if it has density

$$f(z) \propto z^{-3/2} \exp\left\{-\theta_1 z - \frac{\theta_2}{z} + 2\sqrt{\theta_1 \theta_2} + \log\sqrt{2\theta_2}\right\}, z > 0,$$

where $\theta_1 > 0$ and $\theta_2 > 0$ are parameters. It can be shown that

$$E(Z) = \sqrt{\frac{\theta_2}{\theta_1}}$$
 and $E\left(\frac{1}{Z}\right) = \sqrt{\frac{\theta_1}{\theta_2}} + \frac{1}{2\theta_2}$.

Let $\theta_1 = 1.5$ and $\theta_2 = 2$. Draw a sample of size 1,000 using the independence-Metropolis-Hastings algorithm. Use a Gamma distribution as the proposal density. To assess the accuracy, compare the mean of Z and 1/Z from the sample to the theoretical means. Try different Gamma distributions to see if you can get an accurate sample.