# Latent Networks Models Game of Thrones

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#### Overview

- Introduction
- Latent Network Model
- 3 Expectaction Maximimation
  - Unweighted Network Model
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#### Latent Network Model

Using Hoff's work, we model the presence of an edge given our latent variables as

$$logit \mathbb{P}(Y_{ij} = 1|Z) = ||Z_i - Z_j|| + \epsilon_{ij}$$

where  $||Z_i - Z_i||$  is the latent distance between nodes i and j and

$$Z_i \stackrel{ind}{\sim} \sum_{k=1}^G \lambda_k \mathsf{MVN}_d(\mu_k, \sigma_k^2 I_d)$$

#### Latent Network Model: Priors

$$Y_{ij}|Z_i,Z_j \stackrel{ind}{\sim} \operatorname{Bern}\left[\operatorname{logit}^{-1}(\|Z_i-Z_j\|)\right]$$
 $Z_i|K_i=k_i \stackrel{ind}{\sim} MVN(\mu_{k_i},\sigma_{k_i}^2I_d)$ 
 $K \stackrel{iid}{\sim} \operatorname{Multinoulli}(G,\lambda)$ 
 $\lambda_k \stackrel{iid}{\sim} \frac{1}{G}$ 
 $\mu_k \stackrel{iid}{\sim} \operatorname{MVN}_d(0,I_2)$ 
 $\sigma_k^2 \stackrel{iid}{\sim} \operatorname{Inv}\chi_1^2$ 

### Latent Network Model: Likelihood

$$\begin{split} \mathcal{L}(Z,\theta;Y) &= \prod_{i < j} \mathbb{P}(Y_{ij}|Z_i,Z_j) \mathbb{P}(Z_i|K_i,\mu_{k_i},\sigma_{k_i}^2) \mathbb{P}(Z_j|K_j,\mu_{k_j},\sigma_j^2) \\ &\times \mathbb{P}(K_i|\lambda_i) \mathbb{P}(\lambda_i) \mathbb{P}(\mu_{k_i}) \mathbb{P}(\sigma_{k_i}^2) \mathbb{P}(K_j) \mathbb{P}(\mu_{k_j}) \mathbb{P}(\sigma_{k_j}^2) \\ &\propto \prod_{i < j} \left( \text{logit}^{-1} (\|Z_i - Z_j\|) \right)^{Y_{ij}} \left( 1 - \text{logit}^{-1} (\|Z_i - Z_j\|) \right)^{1 - Y_{ij}} \\ &\times \frac{1}{(\sigma_{k_i}^2)^{1/2}} \exp \left\{ -\frac{1}{2\sigma_{k_i}^2} (Z_i - \mu_{k_i})^T (Z_i - \mu_{k_i}) \right\} \frac{1}{(\sigma_{k_j}^2)^{1/2}} \\ &\times \exp \left\{ -\frac{1}{2\sigma_{k_j}^2} (Z_j - \mu_{k_j})^T (Z_j - \mu_{k_j}) \right\} \exp \left\{ -\frac{1}{2} \mu_{k_i}^T \mu_{k_i} \right\} \\ &\times \exp \left\{ -\frac{1}{2} \mu_{k_j}^T \mu_{k_j} \right\} \times \frac{1}{(\sigma_{k_i}^2)^2} \exp \left\{ -\frac{1}{\sigma_{k_i}^2} \right\} \frac{1}{(\sigma_{k_j}^2)^2} \exp \left\{ -\frac{1}{\sigma_{k_j}^2} \right\} \\ &\times \lambda_i \times \lambda_j \end{split}$$

## Unweighted Network Model

Let  $Y_{ij}$  indicate whether there is an edge  $E_{ij}$  between nodes i and j.

$$Y_{ij}|p_{ij} \stackrel{\textit{ind}}{\sim} \mathsf{Bern}(p_{ij}) \ p_{ij} \stackrel{\textit{iid}}{\sim} \mathsf{Beta}(lpha,eta)$$

where  $p_{ij} \equiv 2 - 2 * logit^{-1}(d_{ij})$ .

Then the log-likelihood for this model can be written as

$$\begin{split} I(p,\alpha,\beta;Y) &= \sum_{i < j} Y_{ij} \log \left( \frac{p_{ij}}{1 - p_{ij}} \right) + \log(1 - p_{ij}) + \log \Gamma(\alpha + \beta) \\ &- \log \Gamma(\alpha) - \log \Gamma(\beta) + (\alpha - 1) \log p_{ij} + (\beta - 1) \log(1 - p_{ij}) \end{split}$$

## Unweighted Network Model: E-Step

We can take the expectation of the log-likelihood given the data  ${\bf Y}$  and parameters  $\theta=(\alpha,\beta)$ 

$$\begin{split} Q(\theta; \theta^{(t)}) &= \sum_{i < j} (Y_{ij} + \alpha - 1) \mathbb{E}_{p_{ij} | Y_{ij}, \theta^{(t)}} [\log p_{ij}] \\ &+ (\beta - Y_{ij}) \mathbb{E}_{p_{ij} | Y_{ij}, \theta^{(t)}} [\log (1 - p_{ij})] + \log \Gamma(\alpha + \beta) \\ &- \log \Gamma(\alpha) - \log \Gamma(\beta) \end{split}$$

Since  $p_{ij}|Y_{ij}, \theta \propto \textit{Beta}(\alpha + Y_{ij}, \beta + 1 - Y_{ij})$ , we define the following

$$\begin{split} \pi_{ij} &\equiv \mathbb{E}_{p_{ij}|Y_{ij},\theta} \big[\log p_{ij}\big] = \Psi \Big(\alpha + Y_{ij}\Big) - \Psi \Big(\alpha + \beta + 1\Big) \\ \eta_{ij} &\equiv \mathbb{E}_{p_{ij}|Y_{ij},\theta} \big[\log (1-p_{ij})\big] = \Psi \Big(\beta + 1 - Y_{ij}\Big) - \Psi \Big(\alpha + \beta + 1\Big) \end{split}$$

# Unweighted Network Model: M-Step

To maximize,  $Q(\theta, \theta^{(t)})$  with respect to  $\theta$ , we see that  $\theta^{(t+1)}$  must satisfy the following:

$$\Psi\left(\alpha^{(t+1)} + \beta^{(t)}\right) - \Psi\left(\alpha^{(t+1)}\right) = -\frac{\sum_{i < n} \mathbb{E}_{p_{ij}|Y_{ij},\theta^{(t)}}\left[\log p_{ij}\right]}{\binom{n}{2}} \qquad (\alpha_U)$$

$$\Psi\left(\alpha^{(t+1)} + \beta^{(t+1)}\right) - \Psi(\beta^{(t+1)}) = -\frac{\sum_{i < n} \mathbb{E}_{p_{ij}|Y_{ij},\theta^{(t)}}\left[\log(1 - p_{ij})\right]}{\binom{n}{2}} \left(\beta_{U}\right)$$

Here, we use Newton-Raphson to obtain both  $\alpha^{(t+1)}$  and  $\beta^{(t+1)}$ .

# Unweighted Network Model: Pseudocode

#### Algorithm 1: EM for simplified latent network unweighted model

```
1 LNM EM (G, tol);
   Input: Graph G
                 Tolerance tol
   Output: Nuisance Parameters \alpha^*, \beta^*
                 Latent Probability Estimates \hat{p}
                 Latent Distance Estimates \hat{d}
2 Initialize Q^{(0)} repeat
        E: calculate \pi^{(t)}, \eta^{(t)};
3
        M: update \alpha^{(t+1)} using (\alpha_U);
4
        update \beta^{(t+1)} using (\beta_{II}):
5
        calculate Q(\theta, \theta^{(t+1)})
6
7 until \left|\frac{Q(\theta^{(t+1)}, \theta^{(t)}) - Q(\theta^{(t)}, \theta^{(t)})}{Q(\theta^{(t)}, \theta^{(t)})}\right| < tol;
8 return \alpha^*, \beta^*, \hat{p} = e^{\pi^*}, \hat{d} = logit^{-1}(1 - \frac{e^{\pi^*}}{2}); where \alpha^*, \beta^*, \pi^* are
     converged values
```

## Unweighted Network Model: Distances

## Unweighted Network Model: Probabilities

# Weighted Network Model

Let  $Y_{ij}$  be the weight on edge  $E_{ij} \in \mathbf{E}$ .

$$Y_{ij}|\lambda_{ij} \stackrel{ind}{\sim} \mathsf{Pois}(\lambda_{ij}) \ \lambda_{ij} \stackrel{iid}{\sim} \mathsf{Gamma}(\alpha, \beta)$$

Then the log-likelihood for this model can be written as

$$\begin{split} I(\lambda, \alpha, \beta; Y) &= \sum_{i < j} \Big\{ \log \lambda_{ij} (Y_{ij} + \alpha - 1) - \lambda_{ij} (1 + \beta) \\ &- \log (Y_{ij}!) + \alpha \log(\beta) - \log \Gamma(\alpha) \Big\} \end{split}$$

# Weighted Network Model: E-Step

Taking an expectation of this log-likelihood given the data  ${\bf Y}$  and parameters  $\theta=(\alpha,\beta)$ 

$$\begin{split} Q(\theta; \theta^{(t)}) &= \sum_{i < j} \left\{ (Y_{ij} + \alpha - 1) \mathbb{E}_{\lambda_{ij} | Y_{ij}, \theta^{(t)}} \big[ \log \lambda_{ij} \big] \right. \\ &- (1 + \beta) \mathbb{E}_{\lambda_{ij} | Y_{ij}, \theta^{(t)}} \big[ \lambda_{ij} \big] - \log(Y_{ij}!) + \alpha \log(\beta) - \log \Gamma(\alpha) \right\} \end{split}$$

Seeing as  $\lambda_{ij}|Y_{ij}, heta \propto \mathsf{Gamma}(lpha + Y_{ij}, eta + 1)$  we can define

$$egin{aligned} \pi_{ij} &\equiv \mathbb{E}_{\lambda_{ij} \mid Y_{ij}, heta} \left[ \lambda_{ij} 
ight] = rac{lpha + Y_{ij}}{1 + eta} \ \eta_{ij} &\equiv \mathbb{E}_{\lambda_{ij} \mid Y_{ij}, heta} \left[ \log \lambda_{ij} 
ight] = \log(1 + eta) + \Psi(lpha + Y_{ij}) \end{aligned}$$

## Weighted Network Model: M-Step

Maximizing  $Q(\theta, \theta^{(t)})$  with respect to  $\theta$ , we see that  $\theta^{(t+1)}$  must satisfy

$$\beta^{(t+1)} = \frac{\binom{n}{2}}{\sum_{i < j} \pi_{ij}} \alpha^{(t+1)}$$

$$\Psi(\alpha^{(t+1)}) = \frac{\sum_{i < j} \eta_{ij} + \binom{n}{2} \log(\beta^{(t+1)})}{\binom{n}{2}}$$

We first update  $\beta^{(t+1)}$  using  $\alpha^{(t)}$  then use Netwon-Raphson to attain  $\theta^{(t+1)}$ .

# Weighted Network Model: Psuedo-Code

#### Algorithm 2: EM for simplified latent network weighted model

```
1 LNM EM (G, tol);
   Input: Graph G
                 Tolerance tol
   Output: Nuisance Parameters \alpha^*, \beta^*
                 Latent Mean Estimates \hat{\lambda}
                 Latent Distance Estimates \hat{d}
2 Initialize Q^{(0)} repeat
        E: calculate \pi^{(t)}. \eta^{(t)}:
3
        M: update \beta^{(t+1)} using (\beta_W);
4
        update \alpha^{(t+1)} using (\alpha_W);
        calculate Q(\theta, \theta^{(t+1)})
6
7 until \left|\frac{Q(\theta^{(t+1)},\theta^{(t)})-Q(\theta^{(t)},\theta^{(t)})}{Q(\theta^{(t)},\theta^{(t)})}\right| < tol;
8 return \alpha^*, \beta^*, \hat{\lambda} = \pi^*, \hat{d} = \frac{1}{\pi^*}; where \alpha^*, \beta^*, \pi^* are converged values
```

## Weighted Network Model: Distance Estimates

## Weighted Network Model: Distance Density

## Weighted Network Model: $\lambda$ Estimates

# Weighted Network Model: Spectral Clustering