Generative models: Discriminant Analysis, Naïve Bayes

Nicolas Keriven CNRS, IRISA, Rennes

(material from Florent Chatelain, Olivier Michel)

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Reminder on classification problem

Variable terminology

- ▶ observed data referred to as input variables, predictors or features ← usually denoted as x
- ▶ data to predict referred to as output variables, or responses ← usually denoted as y

Classification task

y are categorical data (discrete qualitative variables) that take values in a discrete set \mathcal{Y} , e.g.

- ightharpoonup email $\in \{ ext{spam}, ext{ham} \}$,
- ▶ handwritten digits $\in \{0, ..., 9\}$

Given a feature vector $x \in \mathbb{R}^d$, build a function f(x) that takes as input the feature vector x and predicts its value for $y \in \mathcal{Y}$

Try to minimize the misclassification rate $\mathcal{R}(f) = \mathbb{P}(f(x) \neq y)$ (aka expected risk for the 0-1 loss)

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Bayes rule for classification

Classification problem with K classes : $y \in \mathcal{Y} = \{1, \dots, K\}$.

Probability of class y = k given X = x

Bayes rule:

$$\mathbb{P}(Y=k|X=x) = \frac{\mathbb{P}(Y=k)p(x|Y=k)}{p(x)} = \frac{\mathbb{P}(Y=k)p(x|Y=k)}{\sum_{j=1}^{K} p(x|Y=j)\mathbb{P}(Y=j)},$$
$$= \frac{\pi_k p_k(x)}{\sum_{j=1}^{K} \pi_j p_j(x)}$$

- \triangleright $p_k(x) \equiv p(x|Y=k)$ is the density for X in class k
- \blacktriangleright $\pi_k \equiv p(Y = k)$ is the weight, or prior probability of class k

Bayes classifier

Definition

The Bayes classification rule f^* is defined as

$$f^*(x) = \arg\max_{k \in \mathcal{Y}} \mathbb{P}(Y = k | X = x) = \arg\max_{k \in \mathcal{Y}} \mathbb{P}(Y = k) p(x | Y = k).$$

Theorem

The Bayes classification rule f^* is optimal in the misclassification rate sense where $\mathcal{R}(f) = \mathbb{P}(f(X) \neq Y)$:

for any rule
$$f$$
, $\mathcal{R}(f) \geq \mathcal{R}(f^*)$,

Remarks

▶ In practice, the distribution of (x, y) is unknown \Rightarrow no analytical expression of $f^*(x)$. But useful reference on academic examples.

Estimation of $f^*(X)$

$$f^*(x) = \arg \max_{k \in \mathcal{Y}} \mathbb{P}(Y = k | X = x) = \arg \max_{k \in \mathcal{Y}} \mathbb{P}(Y = k) p(x | Y = k)$$

- 1. Discriminative approaches : direct learning of $\mathbb{P}(Y|X)$
 - ▶ e.g. logistic regression : $\mathbb{P}(Y = k|x) \approx \operatorname{sigmoid}_k(f(x)) = \frac{e^{f(x)}k}{\sum_{\ell} e^{f(x)}\ell}$ where f is a linear function, a neural net...
 - Very powerful, but not very interpretable
- 2. Generative models: learning of the joint distribution p(X, Y)

$$\mathbb{P}(X=x,Y=k)=p(x|Y=k)\mathbb{P}(Y=k),$$

- $ightharpoonup p_k(x) = p(x|Y=k)$ is the data distribution of class k
- $ightharpoonup \pi_k = \mathbb{P}(Y = k)$ is the weight (proportion) of class k
- ▶ linear/quadratic discriminant analysis, Naïve Bayes
- Interpretable, but requires good generative models (difficult)
- Can generate new data!

Maximum likelihood estimation (MLE)

- ► How to learn a generative model?
- MLE is a general methodology which consists in maximizing the probability of observing the training data with respect to some parametric distirbution p_{θ} . The likelihood function is :

$$\mathcal{L}(\theta) = p_{\theta}((x_1, y_1), \dots, (x_n, y_n))$$

▶ Generally (x_i, y_i) are supposed i.i.d., and

$$\mathcal{L}(\theta) = \prod_{i} p_{\theta}(x_{i}, y_{i})$$

Products are complicated, hence one always work with the log-likelihood:

$$\ell(\theta) = \log \mathcal{L}(\theta) = \sum_i \log p_{\theta}(x_i, y_i) = \sum_i (\log \pi_{y_i} + \log p_{y_i}(x_i))$$

MLE is a classic, interpretable source of loss functions! The log-likelihood is directly the empirical risk! (up to 1/n)

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Discriminant Analysis

Two kinds of Discriminant Analysis: Linear and Quadratic. In both cases, the key assumption is that, within each class, the input variables X_i are assumed to be normally distributed.

Supplementary materials

- Wikipedia page (quite complete and detailed) https://en.wikipedia.org/wiki/Linear_discriminant_analysis
- short and simple Scikit-learn documentation (with examples) https://scikit-learn.org/stable/modules/lda_qda.html

Ouadratic Discriminant Analysis (ODA)

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Quadratic Discriminant Analysis (QDA)

Supervised classification assumptions

- \triangleright $x \in \mathbb{R}^d$, $y \in \mathcal{Y} = \{1, \dots, K\}$,
- ightharpoonup sized n training set $(x_1, y_1), \ldots (x_n, y_n)$

QDA Assumptions

The input variables x, given a class y=k, are distributed according to a Gaussian distribution :

$$X|Y = k \sim \mathcal{N}(\mu_k, \Sigma_k) \Leftrightarrow p_k(x) = \frac{1}{(2\pi)^{p/2} |\Sigma_k|^{1/2}} e^{-\frac{1}{2}(x-\mu_k)^T \Sigma_k^{-1}(x-\mu_k)}$$

The Gaussian parameters are, for each class k = 1, ..., K

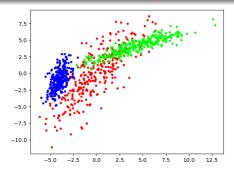
- ightharpoonup mean vectors $\mu_k \in \mathbb{R}^d$,
- \triangleright covariance matrices $\Sigma_k \in \mathbb{R}^{d \times d}$.
- set of parameters $\theta_k = \{\mu_k, \Sigma_k\}$, plus the weights π_k , for k = 1, ..., K.

- Linear/Quadratic Discriminant Analysis
 - Quadratic Discriminant Analysis (QDA)

Example

Mixture of K = 3 Gaussians

- $V = \{1, 2, 3\}$
- ightharpoonup d = 2

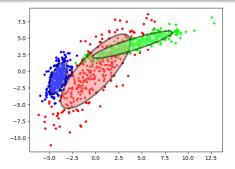


- Linear/Quadratic Discriminant Analysis
 - Quadratic Discriminant Analysis (QDA)

Example

Mixture of K = 3 Gaussians

- $V = \{1, 2, 3\}$
- $\rightarrow d=2$



True mean μ_k and covariance Σ_k parameters, for k = 1, 2, 3

Linear/Quadratic Discriminant Analysis

Quadratic Discriminant Analysis (QDA)

QDA parameter estimation

Log-likelihood

The log-likelihood with Gaussians is:

$$\begin{split} \ell\left(\theta\right) &= \sum_{i=1}^{n} \log \pi_{y_{i}} + \log p\left(x_{i}, \theta_{y_{i}}\right) \\ &= \sum_{i=1}^{n} \log \pi_{y_{i}} - \frac{1}{2} \log |\Sigma_{y_{i}}| - \frac{1}{2} (x_{i} - \mu_{y_{i}})^{T} \Sigma_{y_{i}}^{-1} (x_{i} - \mu_{y_{i}}) \end{split}$$

Remark : $\pi_K = 1 - \sum_{j=1}^{K-1} \pi_j$ so there is one less parameter.

- Linear/Quadratic Discriminant Analysis
 - Quadratic Discriminant Analysis (QDA)

QDA parameter estimation (Cont'd)

Maximizing $\ell(\theta)$ by setting its gradient to 0, we obtain

- $\hat{\pi}_k = \frac{n_k}{n}$ where $n_k = \sharp \{y_i = k\}$. Sample proportion, valid for any model p_k .
- $\hat{\mu}_k = rac{\sum_{y_i=k} x_i}{n_k}$: empirical mean, a classic quantity. Easy derivation.
- $\hat{\Sigma}_k = \frac{1}{n_k} \sum_{y_i = k} (x_i \hat{\mu}_k) (x_i \hat{\mu}_k)^T$. Empirical covariance, again classic. Zeroing the gradient is a bit harder! hint: derive wrt Σ^{-1} and not Σ , use the chain rule.
 - Unlike $\hat{\mu}$, $\hat{\Sigma}$ is biased. An unbiased version is $\frac{n_k}{n_k 1} \hat{\Sigma}_k = \frac{1}{n_k 1} \sum_{y_i = k} (x_i \hat{\mu}_k) (x_i \hat{\mu}_k)^T$

see https://people.eecs.berkeley.edu/~jordan/courses/260-spring10/other-readings/chapter13.pdf

Generative models: Discriminant Analysis, Naïve Bayes

Linear/Quadratic Discriminant Analysis

☐ Quadratic Discriminant Analysis (QDA)

QDA decision rule

Since we have estimated p(x, y), we can derive a classification rule. Starting from the expression of the Bayes estimator :

$$\begin{split} f(x) &= \arg\max_{k \in \mathcal{Y}} \mathbb{P}(Y = k) p(x|Y = k) \\ &\approx \arg\max_{k \in \mathcal{Y}} \hat{\pi}_k p(x|\hat{\theta}_k) = \arg\max_{k \in \mathcal{Y}} \hat{\pi}_k p(x|\hat{\theta}_k) \end{split}$$

Taking the logarithm, which does not change the $\operatorname{argmax}: f(x) = \operatorname{argmax}_k \delta_k(x)$ where

$$\delta_k(x) = -\frac{1}{2}\log\left|\hat{\Sigma}_k\right| - \frac{1}{2}(x - \hat{\mu}_k)^T\hat{\Sigma}_k^{-1}(x - \hat{\mu}_k) + \log\hat{\pi}_k + \text{Lest},$$

is the discriminant function

Remarks

If the Gaussian model is correct:

- 1. this is an estimation of the Bayes classifier with θ replaced by $\hat{\theta}$ (and π replaced by $\hat{\pi}$)
- 2. when $n\gg d$, $\hat{\theta}\to \theta$ (and $\hat{\pi}\to\pi$) : convergence to the Bayes classifier

QDA decision boundary

The boundary between two classes k and l is described by the equation

$$\delta_k(x) = \delta_l(x) \Leftrightarrow C_{k,l} + L_{k,l}^T x + x^T Q_{k,l}^T x = 0,$$

where

$$\blacktriangleright \ L_{k,l} = \hat{\Sigma}_k^{-1} \hat{\mu}_k - \hat{\Sigma}_l^{-1} \hat{\mu}_l, \quad \leftarrow \text{vector in } \mathbb{R}^d$$

$$\blacktriangleright \ \ Q_{k,l} = \frac{1}{2} \left(-\hat{\Sigma}_k^{-1} + \hat{\Sigma}_l^{-1} \right), \quad \leftarrow \mathsf{matrix} \ \mathsf{in} \ \mathbb{R}^{d \times d}$$

- ► This is a quadratic equation, which defines an ellipsoid
- make hence Quadratic discriminant analysis

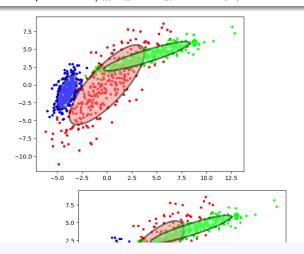
Generative models: Discriminant Analysis, Naïve Bayes

- Linear/Quadratic Discriminant Analysis
 - Quadratic Discriminant Analysis (QDA)

QDA example

Mixture of K = 3 Gaussians

Estimation of the parameters $\hat{\mu}_k$, $\hat{\Sigma}_k$ and $\hat{\pi}_k$, for k = 1, 2, 3

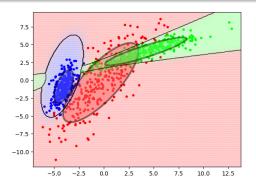


- Linear/Quadratic Discriminant Analysis
 - Quadratic Discriminant Analysis (QDA)

QDA example

Mixture of K = 3 Gaussians

- ► Classification rule : $\arg \max_{k=1,2,3} \delta_k(x)$
- ▶ Quadratic boundaries $\{x; \delta_k(x) = \delta_l(x)\}$



Linear Discriminant Analysis (LDA)

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LDA principle

LDA Assumptions

Additional simplifying assumption w.r.t. QDA : all the class covariance matrices are identical ("homoscedasticity"), i.e. $\Sigma_k = \Sigma$, for k = 1, ..., K

Maximum likelihood estimators (MLE)

- $ightharpoonup \hat{\pi}_k$ and $\hat{\mu}_k$ are unchanged,
- $\hat{\Sigma} = \frac{1}{n} \sum_{k=1}^{K} \sum_{v_i = k} (x_i \hat{\mu}_k) (x_i \hat{\mu}_k)^T$ pooled covariance
 - Again, this is a biased estimator. The unbiased version is $\frac{n}{n-K}\hat{\Sigma}$.

LDA discriminant function

$$\delta_k(x) = -\frac{1}{2}\log\left|\hat{\Sigma}\right| - \frac{1}{2}(x - \hat{\mu}_k)^T\hat{\Sigma}^{-1}(x - \hat{\mu}_k) + \log\hat{\pi}_k + \text{-est},$$

LDA decision boundary

The boundary between two classes k and l reduces to the equation

$$\delta_k(x) = \delta_l(x) \Leftrightarrow C_{k,l} + L_{k,l}^T x = 0$$

where

- ► This is a linear equation
- Linear discriminant analysis

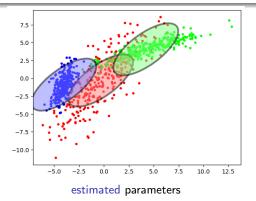
Linear/Quadratic Discriminant Analysis

Linear Discriminant Analysis (LDA)

Linear Discriminant Analysis (LDA)

Mixture of K = 3 Gaussians

Estimation of the parameters $\hat{\mu}_k$, $\hat{\pi}_k$, for k = 1, 2, 3, and $\hat{\Sigma}$



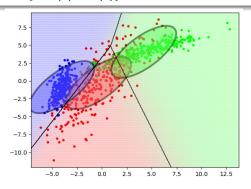
Linear/Quadratic Discriminant Analysis

Linear Discriminant Analysis (LDA)

Linear Discriminant Analysis (LDA)

Mixture of K = 3 Gaussians

- ► Classification rule : $\arg \max_{k=1,2,3} \delta_k(x)$
- linear boundaries $\{x; \delta_k(x) = \delta_l(x)\}$



Linear Discriminant Analysis (LDA)

Complexity of discriminant analysis methods

Effective number of parameters

- ▶ LDA : $K 1 + Kd + \frac{d(d+1)}{2} = O(Kd + d^2)$
- ► QDA : $K 1 + Kd + K \frac{d(d+1)}{2} = O(Kd^2)$

Remarks

- ▶ in high dimension, i.e. $d \approx n$ or d > n, LDA is more stable than QDA which is more prone to overfitting,
- both methods appear however to be robust on a large number of real-word datasets
- ▶ LDA can be viewed in some cases as a least squares regression method
- ▶ LDA performs a dimension reduction to a subspace of dimension $\leq K 1$ generated by the vectors $z_k = \hat{\Sigma}^{-1}\hat{\mu}_k \leftarrow$ dimension reduction from p to K 1! (same for QDA, but more rarely used)

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Non-parametric modelling

Non-parametric estimation of $p_k(x) = p(x|Y = k)$: density estimation. Then $f(x) = \arg\max_k \hat{\pi}_k \hat{p}_k(x)$ as usual.

Parzen kernel approach

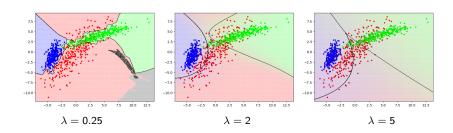
To locally estimate the density, take a weighted average of the number of points in the neighborhood of the desired location :

$$\hat{\rho}_k(x) = \frac{1}{n_k} \sum_{y_i = k} k_{\lambda}(x, x_i)$$

for a kernel function k_{λ} . Usually λ is a bandwidth, and $k_{\lambda}(x,x') = \frac{1}{\lambda^{d}} k(\frac{x-x'}{\lambda})$, with $\int k = 1$. Classic choice includes :

- ▶ 0-1 kernel : $k(x,x') = 1/V_d$ if $||x-x_i|| \le 1$, 0 otherwise, where V_d is the volume of the d-sphere. True unweighted average of the number of points in a fixed-radius neighborhood.
- ► Gaussian kernel : $k(x, x') = \frac{1}{(2\pi)^{d/2}} e^{-\frac{1}{2}||x-x'||^2}$. Classic choice.
- \triangleright Same problem than k-NN: in high-dimension, the space is mostly empty!

KDE example



Complexity parameter λ (kernel bandwidth)

- ▶ large λ w.r.t. to the dispersion of $X \rightarrow$ under-fitting
- \triangleright small λ w.r.t. to the dispersion of $X \rightarrow$ over-fitting

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NB classifiers

Family of "probabilistic classifiers" based on applying Bayes' theorem on a generative model, with strong (naïve) independence assumptions between the features. Particularly useful for high-dimensional data (avoids quadratic cost d^2). Can be coupled with

- parametric models (Gaussian, Bernoulli, Multinomial,...) with maximum likelihood estimation
- or non-parametric models with kernel density estimation

Supplementary materials

- Wikipedia page (quite detailed) https://en.wikipedia.org/wiki/Naive_Bayes_classifier
- ► short and simple Scikit-learn documentation
 https://scikit-learn.org/stable/modules/naive_bayes.html

Naïve Bayes (NB)

$$x = (x^1, ..., x^d) \in \mathbb{R}^d, y \in \mathcal{Y} = \{1, ..., K\}$$

Naive Bayes Assumption

Simplifying assumption : given Y, the components x^1, \ldots, x^d are assumed to be independent :

$$p(x|Y = k) = p_k(x) = \prod_{j=1}^d p_{k,j}(x^j).$$

Remarks

- ▶ independence reduces one estimation problem in d dimensions to d much simpler 1D estimation problems ← prevent from curse of dimensionality
- ▶ independence assumption is naïve, i.e. not realistic in practice... but yields efficient/stable/robust approaches especially in high dimension!

Naïve Bayes for parametric estimation

Gaussian model

- ▶ NB + QDA : $X|Y = k \sim \mathcal{N}(\mu_k, \Sigma_k)$, where the Σ_k are diagonal
 - $ightharpoonup \hat{\mu}_k$ don't change
 - $(\hat{\Sigma}_k)_{jj} = \frac{1}{n_k 1} \sum_{y_i = k} (x_i^j \mu_k^j)^2$
- ▶ NB + LDA : $X|Y = k \sim \mathcal{N}(\mu_k, \Sigma)$, where Σ is diagonal.
 - $\hat{\Sigma}_{jj} = \frac{1}{n-K} \sum_{k} \sum_{y_i = k} (x_i^j \mu_k^j)^2$

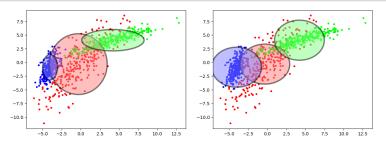
Other classical parametric models

- Bernoulli NB for binary events models (e.g., word occurrence vectors in text processing)
- Multinomial NB for multiple events models (e.g., word count vectors in text processing)
- Mixed models (e.g. Gaussian and Multinomial) for mixed quantitative/qualitative features

NB + QDA example

Mixture of K = 3 Gaussians

▶ Gaussian model : $X|Y = k \sim \mathcal{N}(\mu_k, \Sigma_k)$ with $\hat{\Sigma}_k = \begin{pmatrix} (\hat{\Sigma}_k)_{11} & 0 \\ 0 & (\hat{\Sigma}_k)_{22} \end{pmatrix}$

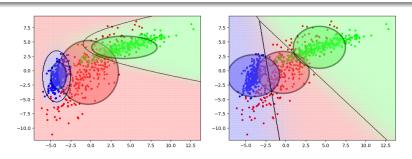


Naive Bayes QDA (left), LDA (right). The Gaussians are "axis-aligned"

Naïve Bayes (NB)

Mixture of K = 3 Gaussians

- ► Classification rule : arg max_{k=1,2,3} $\delta_k(x)$
- quadratic boundaries $\{x; \delta_k(x) = \delta_l(x)\}$



Naïve Bayes for non-parametric estimation

Non-parametric estimation of $p_{k,j}(x^j) = p(x^j|Y=k)$, where x^j is the jth component of x: univariate density estimation. Then $\hat{p}_k(x) = \prod_j p_{k,j}(x^j)$, and $\hat{f}(x) = \arg\max_k \hat{\pi}_k$.

Parzen kernel approach

Apply Parzen window to each component, with a univariate kernel (not necessarily the same for each component) :

$$\hat{\rho}_{k,j}(x^j) = \frac{1}{n_k \lambda} \sum_{y_i = k} k^j (\frac{x_j - x_{j,i}}{\lambda})$$

- Avoids the curse of dimensionality, to the price of simplification
- Note: when $n \to \infty$, regular KDE with classic Gaussian kernel is already Naive Bayes! Since $\frac{1}{(2\pi)^{d/2}}e^{-\|x-x_i\|^2}=\prod_j\frac{1}{\sqrt{2\pi}}e^{-(x^j-x_i^j)^2}$. The only difference is how we compute their empirical versions.

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Generative models

- ▶ learning/estimation of $p(X, Y) = p(X|Y) \Pr(Y)$,
- derivation of Pr(Y|X) from Bayes rule,

Different assumptions on the class densities $p_k(x) = p(X = x | Y = k)$

- ▶ QDA/LDA : Gaussian parametric model
 - performs well on many real-word datasets
 - ► LDA is especially useful when *n* is small
- Parzen window (aka KDE) : non-parametric
 - more flexible, necessitates a lot of data, poor performance in high-dimension
- ▶ Naive Bayes : independence of the feature X components given Y
 - useful when d is very large (high dimension)

Incoming...

Discriminative approaches: direct learning of Pr(Y|X)