Expectation-Maximization from ml_stanford_cs229/CS229_Lecture_Notes.pdf

Mixture of Gaussians model is an example of EM for missing data

$$\{x^{(1)}, \dots, x^{(n)}\}$$
 training set $p(x^{(i)}, z^{(i)}) = p(x^{(i)}|z^{(i)})p(z^{(i)}).$ $z^{(i)} \sim \text{Multinomial}(\phi)$ $\phi_j \geq 0, \ \sum_{j=1}^k \phi_j = 1,$

Observed, incomplete-data

the model is the joint probability that each $x^{(i)}$ was generated by randomly choosing $z^{(i)}$ from $\{1...k\}$.

Latent, unobserved cluster assignments, needed for complete-data

ϕ_j gives $p(z^{(i)} = j)$

Cluster (gaussian) weights

$$p(x^{(i)} \mid z^{(i)} = j) \sim N(\mu_j, \Sigma_j)$$

we adopt the gaussian distribution to represent x

 ϕ , μ and Σ are the model parameters. k is a hyper-parameter

Number of clusters (number of gaussians)

we need the likelihood of x, $p(x; \phi, \mu, \Sigma)$ to estimate the parameters:

$$\begin{split} \ell(\phi,\mu,\Sigma) &= \sum_{i=1}^n \log p(x^{(i)};\phi,\mu,\Sigma) \\ &= \sum_{i=1}^n \log \sum_{z^{(i)}=1}^k p(x^{(i)}|z^{(i)};\mu,\Sigma) p(z^{(i)};\phi) = \sum_{i=1}^n \log \sum_{z^{(i)}=1}^k p(x^{(i)},z^{(i)}) - \sum_{i=1}^n \log \sum_{z^{(i)}=1}^k p(z^{(i)},z^{(i)}) - \sum_{z^{(i)}=1}^n \log \sum_{z^{(i)}=1}^n p(z^{(i)},z^{(i)}) - \sum_{z^{(i)}=1}^n p(z^{(i)},z^{(i)})$$

If we knew z^(i)'s, the problem would be solvable using:

$$\ell(\phi, \mu, \Sigma) = \sum_{i=1}^{n} \log p(x^{(i)}|z^{(i)}; \mu, \Sigma) + \log p(z^{(i)}; \phi).$$

then from setting derivatives to 0 to find the maxima, we would have:

$$\phi_{j} = \frac{1}{n} \sum_{i=1}^{n} 1\{z^{(i)} = j\},$$

$$\mu_{j} = \frac{\sum_{i=1}^{n} 1\{z^{(i)} = j\}x^{(i)}}{\sum_{i=1}^{n} 1\{z^{(i)} = j\}},$$

$$\Sigma_{j} = \frac{\sum_{i=1}^{n} 1\{z^{(i)} = j\}(x^{(i)} - \mu_{j})(x^{(i)} - \mu_{j})^{T}}{\sum_{i=1}^{n} 1\{z^{(i)} = j\}}.$$

since we do NOT know the z^(i)'s, we use expectation-maximization:

in Expectation step: guess the z^(i)'s

break

in Maximization step: update the model parameters, based upon the z^(i) estimates

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step 0: init parameters k, and \phi j, \mu j, \Sigma j) for k=1 to k
prev \log I = 1E-11; tol=1E-4
while (true):
   #E-step:
      for each j=1,n, for each i=1,k: #estimate the posterior
         #set Q i(z^{(i)} = i) := p(z^{(i)} | x^{(i)}; \theta) so that ELBO(x;Q,\theta) = log p(x; \theta) for x and the current \theta
         Q_i(z^{(i)} == i) = w^{(i)} = p(z^{(i)} = i \mid x^{(i)}; \phi, \mu, \Sigma) from Bayes Rule = p(x|z)^*p(z)/p(x)
                           = N(x^{(i)} | \mu j, \Sigma j) * \phi j / sum over ell = 1 to k( \phi ell * N(x^{(i)} | \mu ell, \Sigma ell) )
         #note that these should sum to 1 for a single data point x^{(i)}
      #calculate log likelihood too as it is used to define convergence.
      \log I = I(\theta) = \text{sum over } i=1,n \text{ (log } p(x; \theta) \text{ )} = \text{sum over } i=1,n \text{ (log } (\text{sum over } i=1,k \text{ (w}^i) \text{ i)))}
      if (log I - prev log I <= tol):
         converged = true
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#M-step:

estimate ϕ , μ , Σ by maximizing ELBO:

EQN 11.12 $\theta = \operatorname{argmax} ELBO(x^{(i)}; Q_i, \theta)$ w.r.t θ while fixing the choice of Q_i

$$\begin{split} &\sum_{i=1}^{n} \sum_{z^{(i)}} Q_{i}(z^{(i)}) \log \frac{p(x^{(i)}, z^{(i)}; \phi, \mu, \Sigma)}{Q_{i}(z^{(i)})} \\ &= \sum_{i=1}^{n} \sum_{j=1}^{k} Q_{i}(z^{(i)} = j) \log \frac{p(x^{(i)}|z^{(i)} = j; \mu, \Sigma)p(z^{(i)} = j; \phi)}{Q_{i}(z^{(i)} = j)} \\ &= \sum_{i=1}^{n} \sum_{j=1}^{k} w_{j}^{(i)} \log \frac{\frac{1}{(2\pi)^{d/2}|\Sigma_{j}|^{1/2}} \exp\left(-\frac{1}{2}(x^{(i)} - \mu_{j})^{T}\Sigma_{j}^{-1}(x^{(i)} - \mu_{j})\right) \cdot \phi_{j}}{w_{j}^{(i)}} \\ &\phi_{j} &:= \frac{1}{n} \sum_{i=1}^{n} w_{j}^{(i)}, \\ &\mu_{j} &:= \frac{\sum_{i=1}^{n} w_{j}^{(i)} x^{(i)}}{\sum_{i=1}^{n} w_{j}^{(i)}}, \\ &\Sigma_{j} &:= \frac{\sum_{i=1}^{n} w_{j}^{(i)} (x^{(i)} - \mu_{j})(x^{(i)} - \mu_{j})^{T}}{\sum_{i=1}^{n} w_{j}^{(i)}} \end{split}$$

hold w^(i)_j fixed while taking the gradient of the rest of the ELBO. set gradient to zero and solve for parameters. note that: w^(i) j =Q i(j) Generalization of details of **p(z|x)**, **choosing** Q as estimate and the ELBO:

choose Q(z) proportional to the joint distribution as the model. the posterior follows.

$$Q(z) \propto p(x, z; \theta) = p(z|x; \theta)$$
 posterior of z's (11.8)

$$Q_i(z^{(i)}) = p(z^{(i)}|x^{(i)};\theta)$$

$$\log p(x^{(i)}; \theta) \ge \text{ELBO}(x^{(i)}; Q_i, \theta) = \sum_{z^{(i)}} Q_i(z^{(i)}) \log \frac{p(x^{(i)}, z^{(i)}; \theta)}{Q_i(z^{(i)})}$$

$$\ell(\theta) \ge \sum_{i} \text{ELBO}(x^{(i)}; Q_{i}, \theta)$$

$$= \sum_{i} \sum_{z^{(i)}} Q_{i}(z^{(i)}) \log \frac{p(x^{(i)}, z^{(i)}; \theta)}{Q_{i}(z^{(i)})}$$
(11.11)