



Why

We have precepts in \mathbf{R}^d and want to predict postcepts in \mathbf{R} . We put a probability measure on a set of linear statistical models.¹

Setup

We work over a probability space $(\Omega, \mathcal{A}, \mathbf{P})$. We have n precepts in \mathbf{R}^d . So let $x^1, \dots, x^n \in \mathbf{R}^d$ with data matrix $X \in \mathbf{R}^{n \times d}$.

Let $\theta : \Omega \rightarrow \mathbf{R}^d$ and $e : \Omega \rightarrow \mathbf{R}^n$ be random vectors with normal density (mean zero and covariances Σ_θ and Σ_e respectively). For each $\omega \in \Omega$, define the map $f : \Omega \rightarrow (\mathbf{R}^d \rightarrow \mathbf{R})$ by $f(\omega)(x) = \theta(\omega)^\top x^i + e_i(\omega)$.

Define $y : \Omega \rightarrow \mathbf{R}^n$ by $y(\omega) = X\theta(\omega) + e$. Let $g : \mathbf{R}^d \times \mathbf{R}^n \rightarrow \mathbf{R}$ be the density of (θ, y) . Let $g_{\theta|y} : \mathbf{R}^d \times \mathbf{R}^n \rightarrow \mathbf{R}$ be the conditional density of θ given y .

Proposition 1. (θ, y) has covariance $\begin{pmatrix} \Sigma_\theta & \Sigma_\theta X^\top \\ X \Sigma_\theta & X \Sigma_\theta X^\top + \Sigma_e \end{pmatrix}$

Proposition 2. There exists $c \in \mathbf{R}$, so that for all $\alpha \in \mathbf{R}^d$ and $\gamma \in \mathbf{R}^n$, $\log g(\alpha, \gamma)$ is

$$-\frac{1}{2}(\alpha^\top \Sigma \alpha + \alpha^\top \Sigma X^\top \gamma + \gamma^\top X \Sigma \alpha + \gamma^\top X \Sigma X^\top \gamma) + c.$$

Proposition 3. A solution to maximize $g(\alpha, \gamma)$ with respect to α is $\alpha = -\Sigma^{-1} \Sigma X^\top \gamma$.

¹The name of this sheet will change in future editions. And future editions will include accounts.

Proposition 4. $g_{\theta|y}(\alpha, \gamma)$ is normal with mean

$$\tilde{\mu}(\gamma) = \Sigma X^\top (X \Sigma X^\top)^{-1} \gamma$$

and covariance

$$\tilde{\Sigma} = \Sigma - \Sigma X^\top (X \Sigma X^\top)^{-1} X \Sigma.$$

Proposition 5. A solution to maximize $g_{\theta|y}(\alpha, \gamma)$ w.r.t. α is

$$\tilde{\Sigma} \tilde{\Sigma}^{-1} \tilde{\mu}(\gamma).$$

But, of course, y also has a density. Denote the density of y by $g : \mathbf{R}^n \rightarrow \mathbf{R}$. In other words, $g \geq 0$ and $\int g = 1$.

Proposition 6.

$$\log g(\gamma) = -1/2 (\gamma^\top (X \Sigma X^\top)^{-1} \gamma) - \frac{d}{2} \log 2\pi - \frac{1}{2} \log \mathbf{det} (X \Sigma X^\top)$$

Test

This expression makes clear that y has a normal density with mean $X \mathbf{E}(x)$ and covariance $X \mathbf{E}(x) X^\top$.

Let $w : \Omega \rightarrow \mathbf{R}^d$ be a random vector with mean 0 and covariance ηI . Let $x^1, \dots, x^n \in \mathbf{R}^d$. Define $y^i : \Omega \rightarrow \mathbf{R}$ by $y_i(\omega) = w(\omega)^\top x^i$ for $i = 1, \dots, d$.

Noise setup

Let $e : \Omega \rightarrow \mathbf{R}^n$ be a normal random vector with mean 0 and covariance σI . Define $\tilde{y} : \Omega \rightarrow \mathbf{R}^n$ by $\tilde{y} = y(\omega) + e(\omega)$.

Proposition 7. \tilde{y} is a normal random vector with mean zero and covariance $X \Sigma X^\top + \sigma I$.

