## Problem 1

 $(\Omega, \mathcal{U}, P)$  a probability space.

 $Z: \Omega \to \{0,1\}$ , measurable,  $P(Z^{-1}(\{1\})) = P(Z^{-1}(\{0\})) = 0.5$ , so let  $A = Z^{-1}(\{1\}), B = A^c$ , then  $A := \{\phi, A, B, A \cup B\}$  is a  $\sigma$ -algebra,  $\Omega = A \cup B, A \cap B = \phi$ , and Z is A-measurable, so we have  $Z = \mathbf{1}_A$ .

Let  $X_n: \Omega^n \to \{1, 2, ..., n\}$ , for  $\omega \in \Omega^n$ , write  $\omega = (\omega_k)$ , then set  $X_n((\omega_k)) = \sum_{k=1}^n Z(\omega_k)$ .

Let  $P^n$  be the product measure on  $\mathcal{A}^n$ , i.e.,  $P^n = P \times P \times ... \times P$ , n-times. Then  $E_{P^n}[f] = \int_{\Omega^n} f \, dP^n = \int$ . By Fubini's theorem,  $E_{P^n}[\mathbf{1}_A] = \int_{\Omega} \left( ... \left( \int_{\Omega} \mathbf{1}_A \, dP \right) ... \right) dP = \int_{\Omega} \left( ... \int_{\Omega} \left( P(A) \right) dP ... \right) dP = P(A) \int_{\Omega} \left( ... \int_{\Omega} \left( 1 \right) dP ... \right) dP = P(A) \cdot 1 = P(A)$ . Thus  $E_{P^n}[X_n] = \sum_{k=1}^n E_{P^n}[Z] = \sum_{k=1}^n E_{P^n}[\mathbf{1}_A] = \sum_{k=1}^n P(A) = \frac{1}{2}n$ 

So, with  $X(\omega', (\omega_k)) = \sum_{k=1}^{N(\omega')} Z(\omega_k)$ , where  $N(\omega') = \sum_{k=1}^4 k \mathbf{1}_{A_k}(\omega')$ ,  $P(A_1) = 0.5$ ,  $P(A_2) = 0.1$ ,  $P(A_3) = 0.2$ ,  $P(A_4) = 0.2$ , and  $\{A_k\}$  are independent events,  $\Omega = \bigcup_k A_k$ ,  $E[X|N] = \sum_{k=1}^4 \frac{1}{P(A_k)} E[\mathbf{1}_{A_k} X] \mathbf{1}_{A_k}$ . Now, on  $A_n$ , N = n, so  $E[\mathbf{1}_{A_n} X] = E[\sum_{k=1}^n Z(\omega_k)] = \frac{1}{2}n$ . So,  $E[X|N] = \sum_{k=1}^4 \frac{1}{P(A_k)} \frac{1}{2} k \mathbf{1}_{A_k}$ .

Then,  $E[X] = E[E[X|N]] = E\left[\sum_{k=1}^{4} \frac{1}{P(A_k)} \frac{1}{2} k \mathbf{1}_{A_k}\right] = \sum_{k=1}^{4} \frac{1}{P(A_k)} \frac{1}{2} k E[\mathbf{1}_{A_k}] = \sum_{k=1}^{4} \frac{1}{2} k = 5.$  (mistake)

 $E[x] = E[E[X|N]] = \sum_{n=1}^{4} P(N=n)\frac{1}{2}n = 0.5 * 0.5 + 0.1 * 1 + 0.2 * 1.5 + 0.2 * 2 = 1.05.$ 

 $P(X = 1|N = 2) = P({X = 1} \cap {N = 2}) \div P({N = 2}) = P({Z_1 + Z_2 = 1} \cap {N = 2}) \div P({N = 2}) = P({Z_1 = 1, Z_2 = 0} \cup {Z_1 = 0, Z_2 = 1}) \div P({N = 2}) = 2 * 0.5 * 0.5 * 0.1 = 0.05$ 

 $P(\{X=1\}) = \sum_{n=1}^{4} P(\{X=1\} \cap \{N=n\}), \text{ because } \{N=n\} \text{ are independent and } \Omega = \cup_{n=1}^{4} \{N=n\}. \text{ So } P(\{X=1\}) = P(\{Z_1=1\} \cap \{N=n\}) + P(\{Z_1+Z_2=1\} \cap \{N=n\}) + P(\{Z_1+Z_2+Z_3=1\} \cap \{N=n\}) + P(\{Z_1+Z_2+Z_3+Z_4=1\} \cap \{N=n\}) = P(\{Z_1=1\}) P(\{N=n\}) + P(\{Z_1+Z_2+Z_3=1\}) P(\{N=n\}) + P(\{Z_1+Z_2+Z_3=1\}) P(\{N=n\}) + P(\{Z_1+Z_2+Z_3=1\}) P(\{N=n\}) + P(\{Z_1+Z_2+Z_3=1\}) P(\{N=n\}) = \sum_{n=1}^{4} 0.5 * n * P(\{N=n\})$ 

# Problem 2

Consider the discrete stochastic process,  $X^x : \mathbb{N} \times \Omega \to \mathbb{Z}$ , where  $X_0^x = x > 0$ ,  $x \in \mathbb{N}$ ,  $X_{n+1}^x = 0$  if  $X_n^x = 0$ , and if  $X_n^x > 0$ , then  $X_{n+1}^x = X_n^x \pm 1$  with each half probability. x is a parameter of the process; it is not a random variable.

Define  $t^x: \Omega \to \mathbb{N}$  by  $t^x(\omega) = \min(\{k \in \mathbb{N}; X_k^x(\omega) = 0\})$ , and  $\mathcal{F}_n^x = \sigma(\{X_1^x, X_2^x, ..., X_n^x\})$ , the sigma algebra generated by the process up to time n. Then, for  $n \in \mathbb{N}$ ,  $(t^x)^{-1}(\{n\}) = \{X_n^x = 0\} \cap (\bigcap_{1 \le k < n} \{X_k^x > 0\})$ , so because  $X_k^x$  is  $\mathcal{F}_n^x$ -measurable for all  $k \le n$ ,  $t^x$ ,  $(t^x)^{-1}(\{n\}) \in \mathcal{F}_n^x$ . If we can show that  $E[X_{k+1}^x | \mathcal{F}_k^x] = X_k^x$  for all  $k \in \mathbb{N}$ , then  $X^x$  is a martingale.

Let  $A_k^x = \{\omega \in \Omega; t^x(\omega) > k\}, B_k^x = \{\omega \in \Omega; t^x(\omega) \le k\}$ , so  $A_k^x = (B_k^x)^c$ , and  $B_k^x = \bigcup_{1 \le n \le k} (t^x)^{-1}(\{n\}) \in \mathcal{F}_k^x$ , so  $A_k^x \in \mathcal{F}_k^x$ . Now  $X_{k+1}^x = X_{k+1}^x (\mathbf{1}_{A_k^x} + \mathbf{1}_{B_k^x}) = X_{k+1}^x \mathbf{1}_{A_k^x} + X_{k+1}^x \mathbf{1}_{B_k^x}$ , and  $X_{k+1}^x \mathbf{1}_{B_k^x} = X_{k+1}^x$  when  $t^x \le k$  in which case  $X_{k+1}^x = 0$ , by the definition of  $t^x$ , and when  $t^x > k$ ,  $\mathbf{1}_{B_k^x} = 0$ , so  $X_{k+1}^x \mathbf{1}_{B_k^x} = 0$  always, and thus  $X_{k+1}^x = X_{k+1}^x \mathbf{1}_{A_k^x}$ .

So  $E[X_{k+1}^x | \mathcal{F}_k^x] = E[X_{k+1}^x \mathbf{1}_{A_k^x} | \mathcal{F}_k^x] = E[(X_{k+1}^x - X_k^x + X_k^x) \mathbf{1}_{A_k^x} | \mathcal{F}_k^x] = E[(X_{k+1}^x - X_k^x) \mathbf{1}_{A_k^x} | \mathcal{F}_k^x] + E[X_k^x \mathbf{1}_{A_k^x} | \mathcal{F}_k^x].$  Now  $X_{k+1}^x - X_k^x$  is independent of  $\mathcal{F}_k^x$ , and  $\mathbf{1}_{A_k^x}$  is  $\mathcal{F}_k^x$ -measurable, so  $E[(X_{k+1}^x - X_k^x) \mathbf{1}_{A_k^x} | \mathcal{F}_k^x] = \mathbf{1}_{A_k^x} E[X_{k+1}^x - X_k^x] = \mathbf{1}_{A_k^x} E[X_{k+1}^x - X_k^x] = \mathbf{1}_{A_k^x} (-1 \cdot 0.5 + 1 \cdot 0.5) = 0.$  Then  $E[X_k^x \mathbf{1}_{A_k^x} | \mathcal{F}_k^x] = \mathbf{1}_{A_k^x} X_k^x$ , because  $X_k^x \mathbf{1}_{A_k^x} | \mathcal{F}_k^x$ -measurable.

So,  $E[X_{k+1}^x|\mathcal{F}_k^x] = X_k^x \mathbf{1}_{A_k^x}$ . If  $k < t^x$  then  $\mathbf{1}_{A_k^x} = 1$  and then  $X_k^x \mathbf{1}_{A_k^x} = X_k^x$ . If  $k \ge t^x$  then  $\mathbf{1}_{A_k^x} = 0$  and then  $X_k^x \mathbf{1}_{A_k^x} = 0$ , but in this case  $X_k^x = 0$  by the definition of  $t^x$ , so in general,  $E[X_{k+1}^x|\mathcal{F}_k^x] = X_k^x$ .

# Theory

 $X: \mathbb{N} \times \Omega \to \mathbb{N}, \text{ let } \mathbb{P}_k = (P_{k,i,j}), P_{k,i,j} = P(\{X_{k+1} = i\} | \{X_k = j\}), x_k = (x_{k,i}), x_{k,i} = P(\{X_k = i\}). \text{ Claim: } x_{k+1} = \mathbb{P}_k x_k, \text{ so } x_{k+1,i} = \sum_j P_{k,i,j} x_{k,j}. \text{ Proof: } P_{k,i,j} = P(\{X_{k+1} = i\} | \{X_k = j\}) = \frac{P(\{X_{k+1} = i\} \cap \{X_k = j\})}{P(\{X_k = i\})}, \text{ so } \sum_j P_{k,i,j} x_{k,j} = \sum_j \frac{P(\{X_{k+1} = i\} \cap \{X_k = j\})}{P(\{X_k = j\})} \cdot P(\{X_k = j\}) = \sum_j P(\{X_{k+1} = i\} \cap \{X_k = j\}), \text{ now } \{X_k = j_1\} \cap \{X_k = j_2\} = \phi \text{ when } i_1 \neq j_2, \text{ becasue inverse images commute with intersections, then } \sum_j P_{k,i,j} x_{k,i} = P(\bigcup_j (\{X_{k+1} = i\} \cap \{X_k = j\})) = P(\{X_{k+1} = i\} \cap \bigcup_j \{X_k = j\}), \text{ by countable additivity. } \text{Now, } \bigcup_{j \in \mathbb{N}} \{X_k = j\} = \Omega, \text{ so } \sum_j P_{k,i,j} x_{k,i} = P(\{X_{k+1} = i\}) = x_{k+1}, \text{ and the claim is proved.}$ 

If  $\Omega = \bigcup_k E_k$ ,  $\{E_k\}$  is disjoint,  $A, E_k \in \mathcal{U}$ , then  $\sum_k P(E_k|A) = \sum_k \frac{(P(E_k \cap A))}{P(A)} = \frac{1}{P(A)} P\left(\bigcup_k (E_k \cap A)\right) = \frac{1}{P(A)} P\left(A \cap \bigcup_k (E_k)\right) = \frac{1}{P(A)} P\left(A \cap \Omega\right) = 1$ . Then,  $X_k^{-1}(\{i\}), i \in \mathbb{N}$  generates a measurable disjoint partition of  $\Omega$ , thus  $\sum_{i \in \mathbb{N}} P_{k,i,j} = \sum_{i \in \mathbb{N}} P(\{X_{k+1} = i\} | \{X_k = j\}) = 1$ . So the columns of  $\mathbb{P}$  are are normalized, i.e. they sum to one. Then, if  $x_k$  is normalized in the same sense, so  $\sum_i x_i = 1$ , and all  $x_i \geq 0$ , then  $x_{k+1} = \mathbb{P}x_k$ ,  $\sum_i x_{k+1,i} = \sum_i \sum_j P_{k,i,j} x_{k,j} = \sum_j \sum_i P_{k,i,j} x_{k,j} = \sum_j x_{k,j} \sum_i P_{k,i,j} = \sum_j x_{k,j} \cdot 1 = 1$ , because all pobabilities are non-negative, as are the entries in  $x_k$ , and clearly  $x_{k+1,i} \geq 0$  for all i. This shows that  $\mathbb{P}$  preserves normalization, as we'd expect.

## Problem 3

Let  $\{Z_k\}: \Omega \to \mathbb{N}$  be i.i.d random variables with  $P(\{Z=0\}) = 0.1$ ,  $P(\{Z=1\}) = 0.3$ ,  $P(\{Z=2\}) = 0.2$ ,  $P(\{Z=3\}) = 0.4$ . Define  $X: \mathbb{N} \times \Omega \to \mathbb{N}$  by  $X_0 = 0$ , and  $X_k = \max(\{Z_1, Z_2, ..., Z_k\})$ . Then  $X_{k+1} = \max(\{Z_1, Z_2, ..., Z_k, Z_{k+1}\}) = \max(X_k, Z_{k+1})$ .

If  $f, g: X \to S$ , and  $\max(f, g)(x) = \max(f(x), g(x))$ , then  $\max(f, g)^{-1}(\{s\}) = \{x \in X; \max(f(x), g(x)) = s\} = \max(f, g)^{-1}(\{s\}) = \{x \in X; f(x) \ge g(x), f(x) = s\} \cup \{x \in X; f(x) < g(x), g(x) = s\} = \{f \ge g\} \cap \{f = s\} \cup \{g > f\} \cap \{g = s\}.$ 

 $P_{i,j} := P(\{X_{k+1} = i\} | \{X_k = j\}) = P(\{\max(X_k, Z_{k+1}) = i\} | \{X_k = j\}) = P(\{X_k \ge Z_{k+1}\} \cap \{X_k = i\} | \{X_k = j\}) + P(\{X_k < Z_{k+1}\} \cap \{Z_{k+1} = i\} | \{X_k = j\}), \text{ becausue } \{X_k \ge Z_{k+1}\} \text{ and } \{X_k < Z_{k+1}\} \text{ are disjoint events.}$ 

Now,  $P(\{X_k = j\} \cap \{Z_{k+1} \le X_k\}) = P(\{X_k = j \text{ and } Z_{k+1} \le X_k\}) = P(\{X_k = j\} \cap \{Z_{k+1} \le j\})$ =  $P(\{X_k = j\})P(\{Z_{k+1} \le j\})$ , because  $Z_k$  are iid. When  $i \ne j$ ,  $P(\{X_k = i\} \cap \{X_k = j\}) = 0$ . So  $P(\{X_k \ge Z_{k+1}\} \cap \{X_k = i\} | \{X_k = j\}) = \delta_{i,j}P(\{Z_{k+1} \le j\})$   $P(\{X_k < Z_{k+1}\} \cap \{Z_{k+1} = i\} | \{X_k = j\}) = P(\{X_k < Z_{k+1}\} \cap \{Z_{k+1} = i\} \cap \{X_k = j\}) \div P(\{X_k = j\})$   $= P(\{X_k < i, X_k = j\} \cap \{Z_{k+1} = i\}) \div P(\{X_k = j\}) = P(\{X_k < i, X_k = j\}) P(\{Z_{k+1} = i\}) \div P(\{X_k = j\}),$ by iid.

Now  $X: \Omega \to \mathbb{R}, a, b \in \mathbb{R}$ , then  $P(\{X < a\} | \{X = b\}) = P(X^{-1}((\infty, a)) \cap X^{-1}(\{b\})) \div P(\{X = b\}) = P(X^{-1}((\infty, a) \cap \{b\})) \div P(\{X = b\})$ . So if b > a, then  $(\infty, a) \cap \{b\}) = \phi$  and then  $P(\{X < a\} | \{X = b\}) = 0$ . if  $b \le a$ , then  $(\infty, a) \cap \{b\}) = \{b\}$ , and then  $P(\{X < a\} | \{X = b\}) = 1$ .

So, 
$$P({X_k < Z_{k+1}}) \cap {Z_{k+1} = i} | {X_k = j}) = P({Z_{k+1} = i})$$
 if  $j < i$ , and 0 else.

 $P_{i,j} = P(\{Z_{k+1} \leq j\})$  when i = j and  $P_{i,j} = P(\{Z_{k+1} = i\})$  if j < i, and  $P_{i,j} = 0$  if i < j. Computing:

$$\mathbb{P} = [P_{i,j}] = \begin{bmatrix} 0.1 & 0 & 0 & 0 \\ 0.3 & 0.4 & 0 & 0 \\ 0.2 & 0.2 & 0.6 & 0 \\ 0.4 & 0.4 & 0.4 & 1.0 \end{bmatrix}.$$

#### Problem 4

 $X: \mathbb{N} \times \Omega \to \mathbb{N}$ ,  $X_k$  is the number of infected at step k. Let  $Z = \{z_i, 1 \leq i \leq N\}$  such that  $z_i = 0$  if person i is not infected, and  $z_i = 1$  of they are. Pick two individuals evenly at random x, y, then transimmsion can occur if (x = 0 and y = 1) or (y = 0 and x = 1), and picking  $z \in Z$  at random,  $P(\{z = 1\}) = X_k/N$ , so  $P(\{(x = 0 \text{ and } y = 1) \text{ or } (y = 0 \text{ and } x = 1)\}) = P(\{x = 0\} \cap \{y = 1\}) \cup (\{y = 0\} \cap \{x = 1\}) = (X_k/N)(1 - X_k)$ . Then apply an  $\alpha = 0.1$  probability that transmission will occur. So,

$$X_0 = 1, X_{k+1} = X_k + \mathbf{1}_{E_k},$$

where  $P(E_k) = (X_k/N)(1 - X_k/N)2\alpha$ .  $P_{i,j} := P(\{X_{k+1} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\})$  $= P(\{j + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{\mathbf{1}_{E_k} = i - j\} | \{X_k = j\})$ . So  $P_{i,j} = 0$  unless  $\mathbf{1}_{E_k} = i - j$  which occurs when i - j = 0 or i - j = 1. If i = j, then  $P_{i,j} = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = j\}) = P(\{X_k + \mathbf{1}_{E_k} = i\} | \{X_k = i\} | \{X$ 

So the transition matrix is  $1 - (j/N)(1 - j/N)2\alpha$  on the diagonal, and  $(j/N)(1 - j/N)2\alpha$  on the subdiagonal, and zeros otherwise.

$$\mathbb{P} = [P_{i,j}] = \begin{bmatrix} 0.968 & 0 & 0 & 0 & 0 \\ 0.032 & 0.952 & 0 & 0 & 0 \\ 0 & 0.048 & 0.952 & 0 & 0 \\ 0 & 0 & 0.048 & 0.968 & 0 \\ 0 & 0 & 0 & 0.032 & 1 \end{bmatrix}.$$

### Problem 5

 $X: \mathbb{N} \times \Omega \to \mathbb{N}$  a markov chain such that  $\begin{bmatrix} P(\{X_{k+1}=0\}) \\ P(\{X_{k+1}=1\}) \end{bmatrix} = \mathbb{P}_X \begin{bmatrix} P(\{X_k=0\}) \\ P(\{X_k=1\}) \end{bmatrix}$  for all  $k \in \mathbb{N}$ , where  $\mathbb{P}_X := \begin{bmatrix} \alpha & 1-\beta \\ 1-\alpha & \beta \end{bmatrix}$ ,  $\alpha, \beta \in [0,1]$ . Then let  $n_0 = (0,0), n_1 = (1,0), n_2 = (0,1), n_3 = (1,1)$ , and  $n_i = (n_{i,1}, n_{i,2})$ , and define  $Z_k$  as

$$Z_k = \begin{cases} 0; & \text{if } (X_{k-1}, X_k) = (0, 0) = n_0 \\ 1; & \text{if } (X_{k-1}, X_k) = (1, 0) = n_1 \\ 2; & \text{if } (X_{k-1}, X_k) = (0, 1) = n_2 \\ 3; & \text{if } (X_{k-1}, X_k) = (1, 1) = n_3 \end{cases}$$

Then let  $\mathbb{P}_Z = [P_{i,j}], P_{i,j} = Pr(\{Z_{k+1} = i\} | \{Z_k = j\}), \text{ then}$ 

$$P_{i,j} = \frac{P(\{X_k = n_{i,1}\} \cap \{X_{k+1} = n_{i,2}\} \cap \{X_{k-1} = n_{j,1}\} \cap \{X_k = n_{j,2}\})}{P(\{X_{k-1} = n_{j,1}\} \cap \{X_k = n_{j,2}\})}.$$

If  $n_{i,1} \neq n_{j,2}$  then  $\{X_k = n_{i,1}\} \cap \{X_k = n_{j,2}\} = \phi$  and then the above numerator is zero, becasue  $P(\phi) = 0$ , so this forces  $P_{0,2} = P_{0,3} = P_{1,0} = P_{1,1} = P_{2,2} = P_{2,3} = P_{3,0} = P_{3,1} = 0$ . Assuming  $n_{i,1} = n_{j,2}$ , then  $\{X_k = n_{i,1}\} = \{X_k = n_{j,2}\}$ , and so

$$P_{i,j} = \frac{P(\{X_{k+1} = n_{i,2}\} \cap \{X_{k-1} = n_{j,1}\} \cap \{X_k = n_{j,2}\})}{P(\{X_{k-1} = n_{j,1}\} \cap \{X_k = n_{j,2}\})} = P(\{X_{k+1} = n_{i,2}\} | \{X_{k-1} = n_{j,1}\} \cap \{X_k = n_{j,2}\}).$$

Now applying the Markov property of X,

$$P_{i,j} = P(\{X_{k+1} = n_{i,2}\} | \{X_k = n_{j,2}\}).$$

So,

$$P_{0,0} = P(\{X_{k+1} = 0\} | \{X_k = 0\}) = (\mathbb{P}_X)_{0,0} = \alpha$$

$$P_{0,1} = P(\{X_{k+1} = 0\} | \{X_k = 0\}) = (\mathbb{P}_X)_{0,0} = \alpha$$

$$P_{1,2} = P(\{X_{k+1} = 0\} | \{X_k = 1\}) = (\mathbb{P}_X)_{0,1} = 1 - \beta$$

$$P_{1,3} = P(\{X_{k+1} = 0\} | \{X_k = 1\}) = (\mathbb{P}_X)_{0,1} = 1 - \beta$$

$$P_{2,0} = P(\{X_{k+1} = 1\} | \{X_k = 0\}) = (\mathbb{P}_X)_{1,0} = 1 - \alpha$$

$$P_{2,1} = P(\{X_{k+1} = 1\} | \{X_k = 0\}) = (\mathbb{P}_X)_{1,0} = 1 - \alpha$$

$$P_{3,2} = P(\{X_{k+1} = 1\} | \{X_k = 1\}) = (\mathbb{P}_X)_{1,1} = \beta$$

$$P_{3,3} = P(\{X_{k+1} = 1\} | \{X_k = 1\}) = (\mathbb{P}_X)_{1,1} = \beta,$$

and all together,

$$\mathbb{P}_Z = \begin{bmatrix} \alpha & \alpha & 0 & 0 \\ 0 & 0 & 1 - \beta & 1 - \beta \\ 1 - \alpha & 1 - \alpha & 0 & 0 \\ 0 & 0 & \beta & \beta \end{bmatrix}.$$