

Bootstrapping.

Bootstrap in iid case

Suppose $X_t \sim_{iid} \uparrow (\mu, \sigma^2)$
distribution unknown.

Use \bar{X} as $\hat{\mu}$

Q: what is the $v(\hat{\mu})$?

① Theoretical consideration

Assume $\overset{\text{iid}}{X_i} \sim N(\mu, \sigma^2)$

then $\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$

Exact
for all n .

$$V(\hat{\mu}) = V(\bar{X}) = \frac{\sigma^2}{n}$$

→ In many cases, you can only derive asymptotic $V(\hat{\mu})$.

$$V(\hat{\mu}) \xrightarrow[\text{as } n \rightarrow \infty]{\text{formula}}$$

② Monte Carlo Simulation

Assume $X_t \sim N(\mu_0, \sigma_0^2)$ for some chosen value of μ_0, σ_0 .

In PC, generate ~~many~~ $\{x_1, \dots, x_n\} \rightarrow \bar{x}$

repeat many times. get many \bar{x}

look at Distribution of this.

to estimate $V(\bar{x})$.

→ Must assume μ_0, σ_0 as well as N .

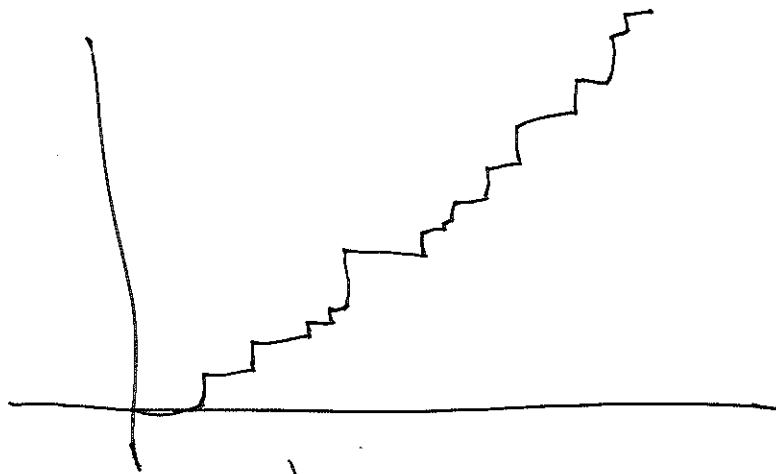
→ Can obtain result for small n .

③ Bootstrapping

Do not assume
 $X_t \sim N(\mu_0, \sigma_0^2)$.

Use the original data $\{x_1, \dots, x_n\}$ to draw

Empirical Distribution Function $\hat{F}(x) = \sum_{i=1}^n \frac{1}{n} \mathbb{I}\{x_i < x\}$



↓
Generate new set $\{x_1^*, \dots, x_n^*\}$
from EDF.

Generate new set $\{x_1^*, \dots, x_n^*\}$ from EDF
made from original data $\{x_1, \dots, x_n\}$

||

Sample with replacement from $\{x_1, \dots, x_n\}$

R code

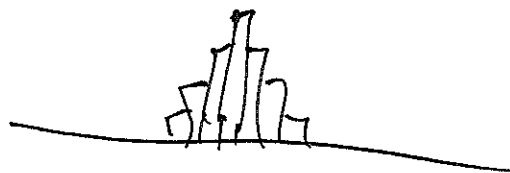
Sample (1:10 , 20 , replace = TRUE)

Original data $\{x_1, \dots, x_n\}$

$\hookrightarrow \{x_1^*, \dots, x_n^*\} \rightarrow \bar{x}^*$
another $\{x_1^*, \dots, x_n^*\} \rightarrow$ another \bar{x}^* } repeat many times.

\Downarrow
many \bar{x}^*

\downarrow
look at distribution of \bar{x}^*



$$v(\bar{x}^*) \approx v(\bar{x})$$

→ Did not assume anything about
distribution of ~~X_t~~ X_t .

→ Completely data driven.

Bootstrapping in ARMA.

ARMA(1,1)

$$Y_t - \phi_1 Y_{t-1} = e_t - \theta_1 e_{t-1}$$

unknown

$$e_t \sim_{iid} (0, \sigma^2)$$

Data $\{Y_1, \dots, Y_n\}$.

1) Fit ARMA(1,1) ~~to~~ get $\hat{\phi}_1$ $\hat{\theta}_1$.

Say, $\hat{\phi}_1 = .521$

$$\hat{\theta}_1 = .297$$

2) Get residuals from the estimation

$$\hat{e}_t = \sum_{j=0}^{n-t} \hat{\psi}_j Y_{t+j} \quad (\text{invertible representation})$$

↑
 $\hat{\psi}_j$ calculated by $\hat{\phi}_1$ and $\hat{\theta}_1$

residuals.

3) Bootstrap $\{\hat{e}_1, \dots, \hat{e}_n\}$ and generate ARMA

$$Y_t^* - \underset{\substack{\uparrow \\ .521}}{\hat{\phi}_1} Y_{t-1}^* = \overset{MA}{e_t^*} - \underset{\substack{\uparrow \\ .297}}{\hat{\theta}_1} \overset{AR}{e_{t-1}^*}$$

↑
bootstrapped errors.

4) For $\{Y_1^*, \dots, Y_n^*\}$ pretend that you don't know $\hat{\phi}_1 = .521$ and $\hat{\theta}_1 = .297$, and estimate

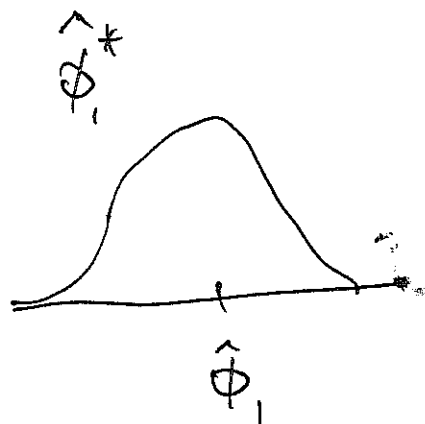
$$\hat{\phi}_1^* \quad \text{and} \quad \hat{\theta}_1^*$$

5) Repeat (3) and (4). Original set of residuals $\{\hat{e}_1, \dots, \hat{e}_n\}$ does not change. Neither do $\hat{\phi}_1 = .521$ and $\hat{\theta}_1 = .297$ in step (3).

In step (4), $\hat{\phi}_1^*$ and $\hat{\theta}_1^*$ are different each time.

Bootstrap

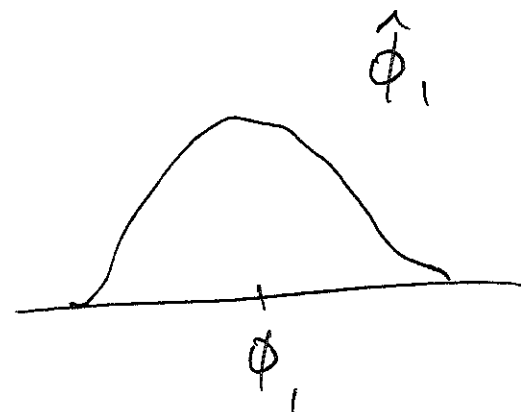
many $\hat{\phi}_i^*$ and $\hat{\theta}_i^*$
around
 $\hat{\phi}_i = .521$ and $\hat{\theta}_i = .297$



use $\sqrt{(\hat{\phi}_i^*)}$ as estimate of $\sqrt{(\hat{\phi})}$.

Monte Carlo

many $\hat{\phi}_i$ and $\hat{\theta}_i$
around
 ϕ_i and θ_i
(true values)



Order Selection of

ARMA(p, q)

Property of ACF and PACF

AR(p)

MA(q)

ARMA(p, q)

ACF

Tails off

Cuts off at q

Tails off

PACF

Cuts off at p.


Tails off

Tails off.

Akaike Information Criteria AIC

$$AIC = -2 \log(\text{maximum likelihood}) + 2(p+q+1)$$

$p+q$ if there's no μ .



Best model = lowest AIC,

AICC

$$AICC = -2 \log(\text{maximum likelihood}) + \frac{2(k+1)(k+2)}{n-k-2}$$

$$k = \begin{cases} p+q+1 & \text{with } \mu \\ p+q & \text{w/o } \mu \end{cases}$$

BIC Bayesian Information Criteria

$$BIC = -2 \log(\text{maximum likelihood}) + \frac{1}{2} \log(n)$$

$$\frac{1}{2} = \begin{cases} p+q+1 & \text{w/o } \mu \\ p+q & \text{w/ } \mu \end{cases}$$