University of Akron, Dept. of Statistics

3470:451/551 Theoretical Statistics I

Ch 4: Continuous Random Variables

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Textbook: Wackerly, Mendenhall, and Scheaffer 7e (2008)

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2.1 Continuous RV

Continuous Random Variable is a r.v. whose range is a interval on a real line or a disjoint union of such intervals. It also must satisfy that for any constant c, P(X = c) = 0.

Probability density function (pdf) of continuous r.v. X is a function f(x) such that for any two numbers a and b with $a \le b$,

$$P(a \le X \le b) = \int_{a}^{b} f(x)dx$$

Pdf must satisfy:

- 1. $f(x) \ge 0$ for all x.
- $2. \int_{-\infty}^{\infty} f(x)dx = 1.$

Cumulative Distribution Function (CDF) of r.v. X is a function F(x) defined for every number x by

$$F(x) = P(X \le x) = \int_{-\infty}^{x} f(x)dx$$

If X is a continuous r.v. with pdf f(x) and cdf F(x) then at every x at which the derivative F'(x) exists,

$$F'(x) = f(x).$$

Cdf must satisfy:

- 1. $F(-\infty) = 0$ and $F(\infty) = 1$.
- 2. non-decreasing.
- 3. right continuous.

For any number a and b with a < b,

$$P(X > a) = 1 - F(a)$$

$$P(a \le X \le b) = F(b) - F(a)$$

Percentiles

Let p a number between 0 and 1. The $(100 \times p)$ th percentile of the distribution of a continuous r.v. X, denoted η_p , is a number such that

$$F(\eta_p) = p$$

Example

Let rv X have pdf

$$f(x) = \begin{cases} Kx^2 & \text{if } 1 \le x \le 3\\ 0 & \text{otherwise} \end{cases}$$

- 1. What is K?
- 2. What is $P(1.5 \le X \le 2)$
- 3. What is F(x)
- 4. What is 70th percentile of X?

Example: K 1-24

An insurance company insures a large number of homes. The insured value, X, of a randomly selected home is assumed to follow a distribution with density function $f(x) = 3x^{-4}$ for 1 < x, and 0 otherwise. Given that a randomly selected home is insured for at least 1.5, what is the probability that it is insured for less than 2?

Example: K 1-13

The waiting time for the first claim from a good driver and the waiting time for the first claim from a bad driver are independent and follow exponential distributions with means 6 years and 3 years, respectively. What is the probability that the first claim from a good driver will be filed within 3 years and the first claim from a bad driver will be filed within 2 years?

Expected Values

Expected or mean value of a continuous r.v. X with pdf f(x) is

$$E[X] = \int_{-\infty}^{\infty} x f(x) dx.$$

To get expectation of a function of X, g(X),

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x)dx.$$

If $g(\cdot)$ is a linear function,

$$E(aX + b) = aE(X) + b.$$

Variance

Variance of a continuous r.v. X with pdf f(x) is

$$V(X) = \sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx = E\left[(x - \mu)^2\right]$$

and standard deviation (SD) of X is

$$\sigma = \sqrt{\sigma^2}$$
.

Or we can still use alternative formula

$$V(X) = E(X^2) - [E(X)]^2.$$

Example Let $\operatorname{rv} X$ have pdf

$$f(x) = \begin{cases} Kx^2 & \text{if } 1 \le x \le 3\\ 0 & \text{otherwise} \end{cases}$$

- 1. What is E(X)?
- 2. What is V(X)?

Example: K 1-10 modified

An auto insurance company insures an automobile for one year under a policy with a 1000 deductible and upper limit of 15,000. If there is a damage, the amount can be modeled by (in thousands) pdf $f(x) = 3e^{-x/3}$ for 0 < x. What is the expected claim payment?

Example: K 1-12

A device that continuously measures and records seismic activity is placed in a remote region. The time, T, to failure of this device is exponentially distributed with mean 3 years. Since the device will not be monitored during its first two years of service, the time to discovery of its failure is $X = \max(T, 2)$. Determine E(X).

2.2 Popular Distributions

2.2.1 Uniform

```
X \sim Unif(a,b)
                            pmf: f(x) = \frac{1}{b-a} for x \in [a,b]
                          CDF: F(x) = \frac{x-a}{b-a} for x \in [a,b]
                                                                V(X) = \frac{(b-a)^2}{12}
               mean and var : E(X) = \frac{b+a}{2}
                          MGF: M(t) = \frac{e^{bt} - e^{at}}{(b-a)t}
dunif(2, a, b) #pmf at x=2
punif(2, a, b) #CDF at x=2
qunif(.5, a, b) #Inv CDF at q=.5
runif(1000, a, b) # random sample of size 1000
x=seq(-1,4,.01); plot(x,dunif(x,1,3), type="l", ylim=c(0,1)) #plot pdf
```

x=seq(-1,4,.01); plot(x,punif(x,1,3), type="l", ylim=c(0,1)) #plot CDF

[top]

2.2.2 Normal

[top]

$$X \sim N(\mu, \sigma^2)$$

$$\operatorname{pdf}: f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad x \in \mathbb{R}$$

$$\operatorname{CDF}: F(x) = \int_{-\infty}^x f(t)dt$$

$$\operatorname{mean}: E(X) = \mu$$

$$\operatorname{var}: V(X) = \sigma^2$$

$$\operatorname{MGF}: M(t) = e^{\mu t + \frac{\sigma^2}{2}t^2}$$

 μ is location parameter, and σ is scale parameter.

```
dnorm(2, mu, sigma)  #pmf at x=2
pnorm(2, mu, sigma)  #CDF at x=2
qnorm(.5, mu, sigma)  #Inv CDF at q=.5
rnorm(1000, mu, sigma)  # random sample of size 1000.
x=seq(-4,4,.01); plot(x,dnorm(x,0,1), type="l", ylim=c(0,1))  #plot pdf
x=seq(-4,4,.01); plot(x,pnorm(x,0,1), type="l", ylim=c(0,1))  #plot CDF
```

2.2.3 Log-normal

[top]

 $X \sim \text{Log-normal}(\mu, \sigma)$

$$pdf: f(x) = \frac{1}{\sqrt{2\pi}\sigma x} e^{-[\ln(x)-\mu]^2/(2\sigma^2)}$$

$$CDF: F(x) = \Phi\left(\frac{\ln(x)-\mu}{\sigma}\right)$$

$$mean: E(X) = e^{\mu+\sigma^2/2}$$

$$var: V(X) = e^{2\mu+\sigma^2} \times (e^{\sigma^2}-1)$$

$$MGF: M(t) =$$

 $ln(X) \sim N(\mu, \sigma^2)$

2.2.4 Exponential

[top]

$$X \sim \text{Exp}(\beta)$$

$$pdf: p(x) = \frac{1}{\beta} e^{-\frac{x}{\beta}} \text{ for } x > 0$$

$$CDF: F(x) = 1 - e^{-x/\beta} \text{ for } x > 0$$

$$mean: E(X) = \beta$$

$$var: V(X) = \beta^2$$

$$MGF: M(t) = \left[\frac{1}{1 - t\beta}\right]$$

 β is a scale parameter

```
dexp(2, 1/b)  #pmf at x=2
pexp(2, 1/b)  #CDF at x=2
pexp(.5, 1/b)  #Inv CDF at q=.5
rexp(1000, 1/b)  # random sample of size 1000. mean should be b
x=seq(-1,5,.01); plot(x,dexp(x,1/2), type="l", ylim=c(0,1)) #plot pdf
x=seq(-1,5,.01); plot(x,pexp(x,1/2), type="l", ylim=c(0,1)) #plot CDF
```

2.2.5 Gamma

$$X \sim \text{Gam}(\alpha, \beta)$$
 [top]

pdf:
$$f(x) = \frac{1}{\Gamma(\alpha)\beta^{\alpha}} x^{\alpha-1} e^{-x/\beta} \quad x > 0$$

CDF:
$$F(x) = \frac{\Gamma(x/\beta; \alpha)}{\Gamma(\alpha)}$$
 $x > 0$

mean and var :
$$E(X) = \alpha \beta$$
 $V(X) = \alpha \beta^2$ $MGF : M(t) = \left[\frac{1}{1 - t\beta}\right]^{\alpha}$

 α is a shape parameter, β is a scale parameter

$$\text{Gamma Func:} \quad \Gamma(\alpha) = \int_0^1 x^{\alpha-1} \, e^{-x} dx, \qquad \quad \text{Incomplete Gamma Func:} \quad \Gamma(x,\alpha) = \int_0^x t^{\alpha-1} \, e^{-t} dt$$

2.2.6 Chi-square

 $[top] X \sim \chi^2(\nu)$

pdf:
$$f(x) = \frac{1}{\Gamma(\nu/2) 2^{\nu/2}} x^{\nu/2-1} e^{-x/2}$$

$$\mathrm{CDF}: \ F(x) \ = \ \Gamma(x/2;\nu/2)/\Gamma(\nu/2)$$

mean:
$$E(X) = \nu$$

$$\operatorname{var}: V(X) = 2\nu$$

MGF:
$$M(t) = \left[\frac{1}{1-2t}\right]^{\nu/2} \quad t < 1/2$$

same as $Gam(\nu/2,2)$

```
dchisq(2, a, scale=b)  #pmf at x=2
pchisq(2, a, scale=b)  #CDF at x=2
pchisq(.5, a, scale=b)  #Inv CDF at q=.5
rchisq(1000, a, scale=b)  # random sample of size 1000. mean should be a*b
x=seq(-1,4,.01); plot(x,dchisq(x,3), type="l", ylim=c(0,.5))  #plot pdf
x=seq(-1,4,.01); plot(x,pchisq(x,3), type="l", ylim=c(0,.5))  #plot CDF
```

2.2.7 Beta

$$\text{pdf}: \quad f(x) = \frac{1}{\beta(\alpha,\beta)} x^{\alpha-1} (1-x)^{\beta-1} \qquad 0 < x < 1$$

$$\text{CDF}: \quad F(x) = \frac{\beta(x;\alpha,\beta)}{\beta(\alpha,\beta)}$$

$$\text{mean and var}: \quad E(X) = \frac{\alpha}{\alpha+\beta}, \qquad V(X) = \frac{\alpha\beta}{(\alpha+\beta+1)(\alpha+\beta)^2}$$

$$\text{MGF}: \quad M(t) = 1 + \sum_{k=1}^{\infty} \Big(\prod_{r=0}^{k-1} \frac{\alpha+r}{\alpha+\beta+r}\Big) \frac{t^k}{k!}$$

$$\text{dbeta(2, a, scale=b)} \qquad \text{#pmf at } x=2$$

$$\text{pbeta(2, a, scale=b)} \qquad \text{#cDF at } x=2$$

$$\text{pbeta(3, a, scale=b)} \qquad \text{#inv CDF at } q=.5$$

$$\text{rbeta(1000, a, scale=b)} \qquad \text{# random sample of size 1000. mean should be a*b}$$

$$x=\text{seq(-1,2,.01)}; \quad \text{plot(x,dbeta(x,2,2), type="l", ylim=c(0,2))} \qquad \text{#plot cDF}$$

$$\text{Beta function} \qquad \beta(\alpha,\beta) = \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$\text{Incomplete Beta func} \qquad \beta(x;\alpha,\beta) = \int_0^x t^{\alpha-1} (1-t)^{\beta-1} dt = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$\text{Incomplete Beta func} \qquad \beta(x;\alpha,\beta) = \int_0^x t^{\alpha-1} (1-t)^{\beta-1} dt = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

2.2.8 Cauchy

[top]

$$X \sim \mathrm{Cau}(\theta, \sigma) \qquad \theta \in \mathbb{R}, \sigma > 0$$

$$\mathrm{pdf}: \quad f(x) = \frac{1}{\pi \sigma} \frac{1}{1 + (\frac{x - \theta}{\sigma})^2} \qquad x \in \mathbb{R}$$

$$\mathrm{CDF}: \quad F(x) = \int_0^x f(t) dt \quad \text{for } t > 0$$

$$\mathrm{mean \ and \ var}: \quad E(X) = \mathrm{Does \ not \ exist} \qquad V(X) = \mathrm{Does \ not \ exist}$$

$$\mathrm{MGF}: \quad M(t) = \mathrm{Does \ not \ exsit}$$

Special case of student's t, when df=1.

2.2.9 Weibull

[top]

$$\begin{split} X \sim & \text{Wei}(\alpha,\beta) \qquad \alpha > 0, \beta > 0 \\ & \text{pdf}: \quad f(x) \quad = \quad \frac{\alpha}{\beta} \, \, x^{\alpha-1} \, e^{-x^{\alpha}/\beta} \qquad 0 \leq x \\ & \text{CDF}: \quad F(x) \quad = \quad \int_0^x f(t) dt \quad \text{for } 0 \leq x \\ & \text{mean}: \quad E(X) \quad = \quad \beta^{1/\alpha} \Gamma(1+1/\alpha) \\ & \text{var}: \quad V(X) \quad = \quad \beta^{2/\alpha} [\Gamma(1+2/\alpha) - \Gamma^2(1+1/\alpha)] \\ & \text{moments}: \quad E(X^n) \quad = \quad \beta^{n/\alpha} [\Gamma(1+n/\alpha)] \\ & \text{MGF}: \quad M(t) \quad = \quad \text{Exists only for } \alpha \geq 1. \end{split}$$

If $\alpha = 1$, it is $\text{Exp}(\beta)$

2.2.10 Student-t

```
\begin{array}{lll} X \sim & t(\nu) & \nu = 1,2,3,\dots \\ & \mathrm{pdf}: & f(x) & = & \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)} \frac{1}{\sqrt{\nu\pi}} \Big(1 + \frac{x^2}{\nu}\Big)^{-(\nu+1)/2} \\ & \mathrm{CDF}: & F(x) & = & \int_0^x f(t) dt \quad \text{for } t > 0 \\ & \mathrm{mean \ and \ var}: & E(X) & = & 0, \quad \nu > 1 \qquad V(X) = \frac{\nu}{\nu-2}, \quad \nu > 2 \\ & \mathrm{moments}: & E(X^2) & = & \frac{\Gamma\left(\frac{n+1}{2}\right) \Gamma\left(\frac{\nu-n}{2}\right)}{\sqrt{\pi} \Gamma\left(\frac{\nu}{2}\right)} \nu^{n/2} & \text{if } n < \nu \text{ and even. } 0 \text{ if odd.} \\ & \mathrm{MGF}: & M(t) & = & \mathrm{Does \ not \ exsit} \end{array}
```

[top]

2.2.11 F

 $X \sim F(\nu_1, \nu_2) \qquad \nu_1, \nu_2 = 1, 2, 3, \dots$ $\text{pdf}: \quad f(x) = \frac{\Gamma(\frac{\nu_1 + \nu_2}{2})}{\Gamma(\frac{\nu_1}{2})\Gamma(\frac{\nu_2}{2})} \left(\frac{\nu_1}{\nu_2}\right)^{\nu_1/2} \frac{x^{(\nu_1 - 2)/2}}{\left(1 + (\frac{\nu_1}{\nu_2})x\right)^{(\nu_1 + \nu_2)/2}} \qquad 0 \le x$ $\text{CDF}: \quad F(x) = \int_0^x f(t)dt \quad \text{for } x > 0$ $\text{mean}: \quad E(X) = \frac{\nu}{\nu_2 - 2} \quad \nu_2 > 2$ $\text{var}: \quad V(X) = 2(\frac{\nu_2}{\nu_2 - 2})^2 \frac{\nu_1 + \nu_2 - 2}{\nu_1(\nu_2 - 4)} \quad \nu_2 > 4$ $\text{moments} \quad E(X^n) = \frac{\Gamma(\frac{\nu_1 + 2n}{2})\Gamma(\frac{\nu_2 + 2n}{2})}{\Gamma(\frac{\nu_1}{2})\Gamma(\frac{\nu_2}{2})} \left(\frac{\nu_2}{\nu_1}\right)^n, \qquad n < \frac{\nu_2}{2}$

MGF: M(t) = Does Not Exist

top

2.2.12 Overlay plots in R

[top]

```
#-- Overlay with N(0,1) pdf --
plot(x, dt(x,5), type='l')
lines(x, dnorm(x,0,1), col='red')

#-- Overlay with N(0,1) pdf (method 2 - have to specify plot range) --
plot(x, dt(x,5), type='l', xlim=c(-5,5), ylim=c(0,.4))
par(new=T)
plot(x, dnorm(x,0,1), type='l', xlim=c(-5,5), ylim=c(0,.4), col='red')
```

2.2.13 Distributional Relations

[top]

- When X and Y are independent $\text{Exp}(\lambda)$, X + Y is $\text{Gam}(2, \lambda)$.
- When you have n iid Exponential r.v. with mean of λ , $\min(X_1, \ldots, X_n)$ is Exponential with mean λ/n .
- Beta(1,1) is same as Unif(0,1).
- Cauchy is same as t(1).
- When X, and Y are independent U(0,1), X/Y is Cauchy.
- When X is $U(-\pi/2, \pi/2)$, tan(X) is Cauchy.
- If we have two independent r.v. $X_1 \sim \text{Gam}(\alpha_1, \beta)$ and $X_2 \sim \text{Gam}(\alpha_2, \beta)$ and

$$Y = \frac{X_1}{X_1 + X_2} \sim \operatorname{Beta}(\alpha_1, \alpha_2)$$

• That is same as to say, if we have two independent r.v. $X_1 \sim \chi^2(\alpha_1)$ and $X_2 \sim \chi^2(\alpha_2)$

$$Y = \frac{X_1}{X_1 + X_2} \sim \text{Beta}\left(\frac{\alpha_1}{2}, \frac{\alpha_2}{2}\right)$$

- When U is $\chi^2(r_1)$ and V is $\chi^2(r_2)$, $\frac{U/r_1}{V/r_2}$ is $F(r_1, r_2)$,
- $F_{1,\nu}$ is same as $t^2(\nu)$
- Gamma function: $\frac{1}{2}! = \Gamma(\frac{1}{2}) = \sqrt{\pi}$

2.3 Scale Parameter

[top]

• If you transform r.v. X to $Y = \theta X$,

$$F_Y(y) = P(\theta X \le y) = P(X \le y/\theta) = F_X(x/\theta)$$

$$f_Y(y) = \frac{1}{\theta} f_X\left(\frac{x}{\theta}\right)$$

 θ is called the scale parameter.

• if Y has scale parameter θ , then

$$\frac{X}{\theta}$$
 has same distribution with $\theta = 1$.

Example

If $X \sim \text{Exp}(\lambda)$, then λ is a scale parameter. Thus

$$\frac{X}{\lambda} \sim \text{Exp}(1).$$

Example

If $X \sim \text{Gam}(\alpha, \beta)$, then β is a scale parameter. Thus

$$\frac{X}{\lambda} \sim \operatorname{Gam}(\alpha, 1).$$

Example

If $X \sim \text{Gam}(\alpha, \beta)$, then $X/\beta \sim \text{Gam}(\alpha, 1)$. We can write cdf of $\text{Gam}(\alpha, 1)$ as

$$F_X(x) = \frac{1}{\Gamma(\alpha)} \underbrace{\int_0^x y^{a-1} e^{-y} dy}_{\text{(lower) incomplete gamma func}} = \frac{\Gamma(x, \alpha)}{\Gamma(\alpha)}, \qquad 0 < x < \infty$$