

Simulating Poisson Process

Monte Carlo Simulation

Kinetic Monte Carlo Simulation

Ex.

ROSS P 315

$N_1(t)$ event vs $N_2(t)$ event

$$2 \text{ Poi. Proc } \begin{cases} N_1(t) & \lambda_1 \\ N_2(t) & \lambda_2 \end{cases} \quad \underline{\text{indep.}}$$

what is $P(S_1' < S_2^2)$

$= P(\text{1st event in Proc. 1}$
 $\text{is before 1st event}$
 $\text{in Proc 2.})$

We saw before

$$P(S_1 < S_2) = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

$$Q_4 \quad S_i \sim \text{Exp}(\lambda_i)$$

$$S_2 \sim \text{Exp}(\lambda_2)$$

We could compute as.

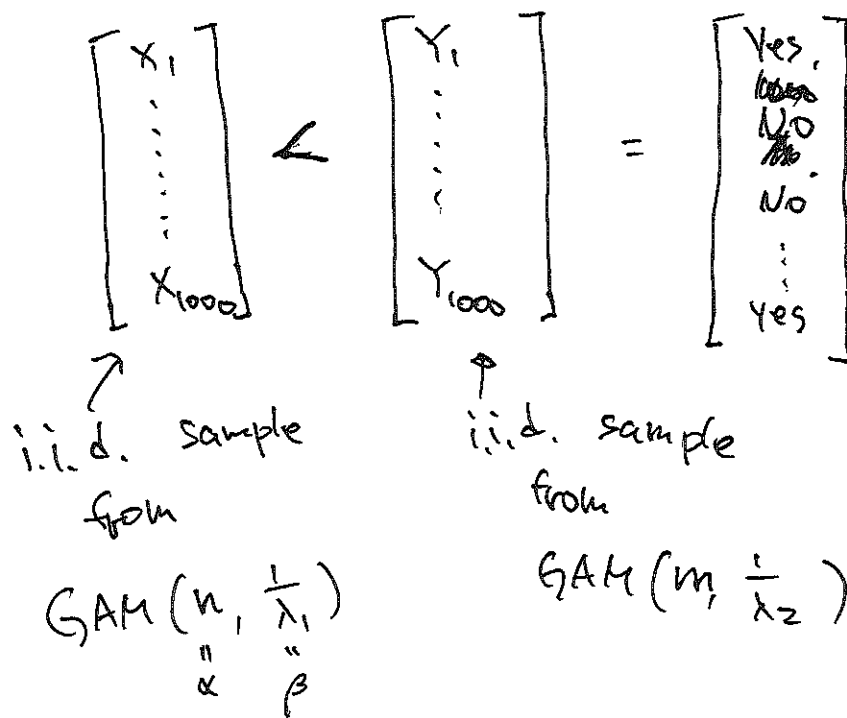
$$P(S_n' < S_m^2) = \underline{\hspace{2cm}}$$

$$S_n' \sim \text{GAM}(n, \frac{1}{\lambda_1})$$

$$S_m^2 \sim \text{GAM}(m, \frac{1}{\lambda_2})$$

$$P(S_n^1 < S_m^2) = P(S_n^1 - S_m^2 < 0)$$

Monte Carlo Simulation



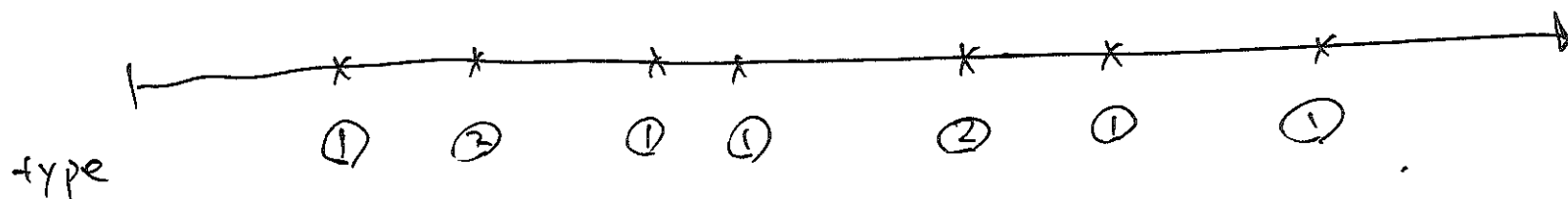
See How many % are 'Yes'

$$P(S_n^1 - S_m^2 < 0)$$

in R:

`rgamma(1000, α , $\frac{1}{\beta}$)`

Alternatively, use View 1 and see



Poi Proc. w. rate $(\lambda_1 + \lambda_2)$

events	{	①	w. prob.	$\frac{\lambda_1}{\lambda_1 + \lambda_2}$
		②	w. P.	$\frac{\lambda_2}{\lambda_1 + \lambda_2}$

$$(S_n^1 < S_m^2) = n\text{th } ① \text{ before } m\text{th } ②$$

n th Head before m th Tail

$$X \sim \text{Neg. Binomial} \left(n, \frac{\lambda_1}{\lambda_1 + \lambda_2} \right)$$

$$P(n\text{th } H \text{ before } m\text{th } T)$$

$$X = \# \text{ of throw until } n\text{th Head}$$

$$= P(X < n+m)$$

$$= P(X \leq n+m-1)$$

CDF of neg. bin.

$$P(S'_n < S_m^2)$$

$$= \text{CDF of } NB(n, \frac{\lambda_1}{\lambda_1 + \lambda_2}) \text{ at } x = n + m - 1.$$

$$= \sum_{k=n}^{n+m-1} \binom{n+m-1}{k} \left(\frac{\lambda_1}{\lambda_1 + \lambda_2} \right)^k \left(\frac{\lambda_2}{\lambda_1 + \lambda_2} \right)^{n+m-1-k}$$

$$= \text{pnbinom}(\overset{\uparrow}{\text{!}} \text{ } n+m-1, n, \frac{\lambda_1}{\lambda_1 + \lambda_2})$$

in R: $x = \# \text{ of } \underline{\text{Tails}}$ before n^{th} Head.
~~example~~ ..

Conditional Joint distribution of arrival time.

$$S_i = \sum_{k=1}^i T_k \sim \text{GAM}(i, \frac{1}{\lambda})$$

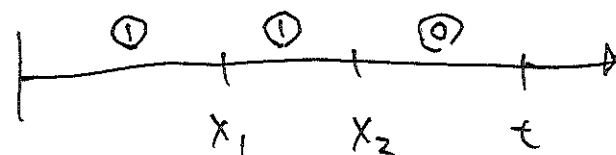
What is joint dist. of S_1, S_2 given $N(t) = 2$?

Let $x_1 \leq x_2 \leq t$

$$P(S_1 \leq x_1, S_2 \leq x_2 \mid N(t) = 2)$$

$$= \frac{P(S_1 \leq x_1, S_2 \leq x_2, N(t) = 2)}{P(N(t) = 2)}$$

$$P(S_1 \leq x_1, S_2 \leq x_2, N(t) = 2)$$

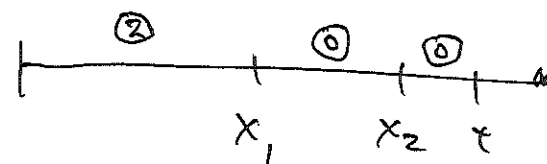


=

~~$$P(N(x_1) = 1, N(x_2) - N(x_1) = 1, N(t) - N(x_2) = 0)$$~~

$$P(N(x_1) = 1, N(x_2) - N(x_1) = 1, N(t) - N(x_2) = 0)$$

$$+ P(N(x_1) = 2, N(t) - N(x_1) = 0)$$



~~$$= e^{-\lambda x_1} \frac{\lambda x_1}{1!} e^{-\lambda(x_2 - x_1)} \frac{\lambda(x_2 - x_1)}{1!} e^{-\lambda(t - x_2)} \frac{\lambda(t - x_2)}{0!} + e^{-\lambda x_1} \frac{\lambda^2 x_1^2}{2!} e^{-\lambda(t - x_1)} \frac{\lambda(t - x_1)}{0!}$$~~

$$\begin{aligned}
& P(N(x_1) \geq 1, N(x_2) \geq 2, N(t) = 2) \\
&= P(N(x_1) = 1, N(x_2) - N(x_1) = 1, N(t) - N(x_2) = 0) \\
&\quad + P(N(x_1) = 2, N(t) - N(x_1) = 0) \\
&= \frac{e^{-\lambda x_1} \lambda x_1}{1!} \frac{e^{-\lambda(x_2-x_1)} \lambda(x_2-x_1)}{1!} e^{-\lambda(t-x_2)} \\
&\quad + \frac{e^{-\lambda x_1} (\lambda x_1)^2}{2!} e^{-\lambda(t-x_1)} \\
&= \frac{\lambda^2(x_1 x_2 - x_1^2)}{1!} e^{-\lambda t} + \frac{(\lambda x_1)^2}{2!} e^{-\lambda t} \\
&= \frac{\lambda^2(x_1 x_2)}{1!} e^{-\lambda t} - \frac{(\lambda x_1)^2}{2!} e^{-\lambda t}
\end{aligned}$$

~~Q.1~~

$$P(S_1 \leq x_1, S_2 \leq x_2 \mid N(t) = 2)$$

$$= \frac{\lambda^2 (x_1 x_2) e^{-\lambda t} - (\lambda x_1)^2 e^{-\lambda t} / 2!}{e^{-\lambda t} (\lambda t)^2 / 2!}$$

$$= \frac{1}{t^2} (2! x_1 x_2 - x_1^2)$$

Cond'l
Joint CDF

take $\frac{\partial^2}{\partial x_1 \partial x_2}$ for joint pdf.

$x_1, x_2 = t$ gives 1,

$$f_{S_1, S_2}(x_1, x_2 \mid N(t) = 2)$$

$$= \frac{\partial^2}{\partial x_1 \partial x_2} \frac{1}{t^2} (2! x_1 x_2 - x_1^2)$$

$$= \frac{2!}{t^2} \quad 0 \leq x_1 \leq x_2 \leq t,$$

\therefore Joint dist of $S_1, S_2 \mid N(t) = 2$ is

same as $U_{(1)}, U_{(2)}$

ordered iid Unif.

Remark . Thm 5.2 Ross

Joint dist. of

$$(S_1, S_2, \dots, S_k) \mid N(t) = k$$

\sim ordered k iid $\text{Unif}(0, t)$