

NHPP

intensity func.
✓

$$N(t_2) - N(t_1) \sim \text{NHPP} \left(\overset{\lambda(t)}{\cancel{g(t_1, t_2)}} \right)$$

then

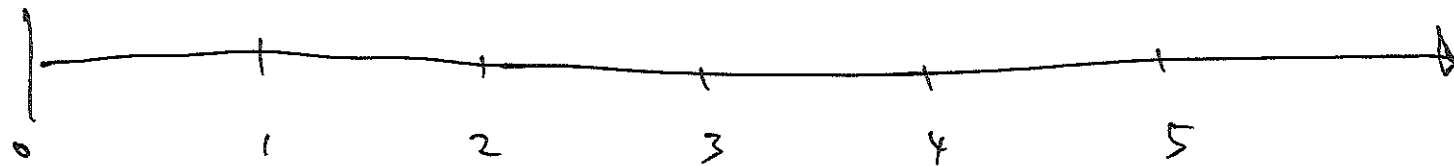
$$P(N(t_2) - N(t_1) = x) = \frac{g(t_1, t_2)^x \cdot e^{-g(t_1, t_2)}}{x!}$$

Poisson pmf

$$\frac{\lambda^x e^{-\lambda}}{x!}$$

Monte Carlo for NHPP

(Approximation)



$g(0,1)$

$g(1,2)$

$g(2,3)$

$g(3,4)$

$PP(g(0,1))$ for 1 day.

To perform kinetic MC, we need distribution of inter-arrival times.

let S_m = time of m th event.

then $S_m = \sum_{i=1}^m T_i$ T_i = i th inter-arrival times.

$$P(S_m < t) = P(\text{mth event was before } t)$$

$$= P(\text{In time } 0 \sim t, \text{ there was more than } m \text{ events})$$

$$= P(N(t) \geq m)$$

This is CDF of S_m ,

$$\begin{aligned}
 F_{S_m}(t) &= P(N(t) \geq m) \\
 &= \sum_{x=m}^{\infty} \frac{g(0,t)^x e^{-g(0,t)}}{x!}
 \end{aligned}$$

(For HPP, $S_m \sim \text{GAM}(m, \frac{1}{\lambda})$ and $T_i \sim \text{Exp}(\lambda)$.)

Take $\frac{d}{dt}$ to obtain pdf of S_m .

~~$$f_{S_m}(t) = \sum_{x=m}^{\infty} \frac{g(0,t)^x e^{-g(0,t)}}{x!}$$~~

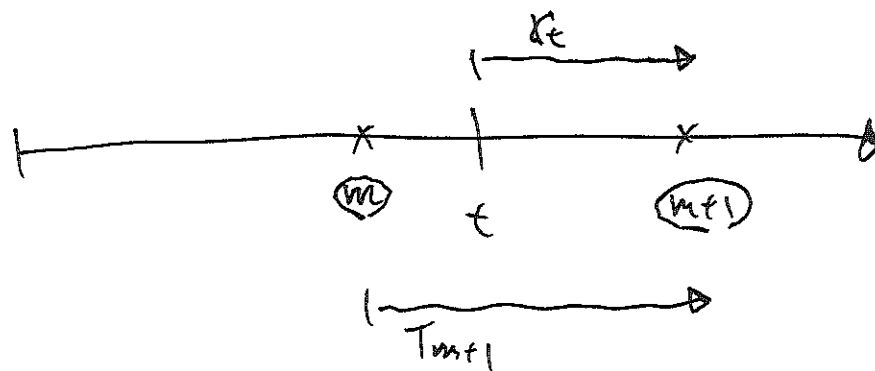
$$\begin{aligned}
 f_{S_m}(t) &= \sum_{x=m}^{\infty} \frac{x g(0,t)^{x-1} e^{-g(0,t)}}{x!} g'(0,t) \\
 &\quad + \sum_{x=m}^{\infty} \frac{g(0,t)^x e^{-g(0,t)}}{x!} (-g'(0,t))
 \end{aligned}$$

All term cancels except the 1st term on the left.

$$f_{S_m}(t) = \frac{g(0,t) e^{-(m-1)g(0,t)}}{(m-1)!} g'(0,t)$$

Now we turn to

δ_t = waiting time from t to next event.

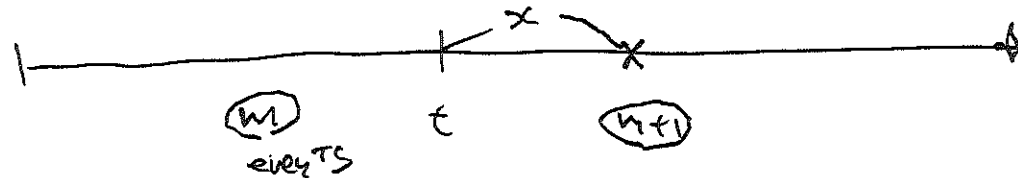


$$m = 0, 1, 2, 3, \dots$$

For HPP, $T_{m+1} \sim \text{Exp}(\lambda)$ and $\delta_t \sim \text{Exp}(\lambda)$ by memoryless.

$$P(X_t > x) = P(T_{n+1} > t+x)$$

$$+ \sum_{n=1}^{\infty} P(S_n \leq t, S_{n+1} > t+x)$$



First term,

$$N(t+x) = 0$$

$$P(T_1 > t+x) = P(\text{no events})$$

$$= e^{-g(0, t+x)}$$

Second term.

$$\sum_{m=1}^{\infty} P(S_m \leq t, S_{m+1} > t+x)$$

$$= \sum_{m=1}^{\infty} \int_0^t P(S_{m+1} > t+x \mid S_m = y) \cdot f_{S_m}(y) dy$$

$$= \sum_{m=1}^{\infty} \int_0^t P(N(t+x) - N(y) = 0) \cdot f_{S_m}(y) dy$$

$$= \sum_{m=1}^{\infty} \int_0^t e^{-g(y, t-x)} \left(\frac{g(0, y)^{m-1} e^{-g(0, y)} g'(0, y)}{(m-1)!} \right) dy$$

$$\sum_{x=0}^{\infty} \frac{a^x}{x!} = e^a$$

$$\begin{aligned}
 & \text{Since} \\
 & = \int_0^t e^{-g(y, t-x)} \cdot e^{g(0, y)} \cdot e^{-g(0, y)} \cdot g'(0, y) dy
 \end{aligned}$$

$$= \int_0^t g'(0, y) e^{-g(y, t-x)} dy$$

Combining, we get

$$P(X_t > x) = e^{-g(0, t+x)} + \int_0^t g'(0, y) e^{-g(y, t-x)} dy$$

Then we have CDF for Δ_t ,

$$\begin{aligned} F_{\Delta_t}(x) &= \cancel{\text{P}} P(\Delta_t \leq x) = 1 - P(\Delta_t > x) \\ &= 1 - e^{-g(0, t+x)} - \int_0^t g'(0, y) e^{-g(y, t+x)} dy \end{aligned}$$

Take $\frac{d}{dx}$ to get pdf,

$$\begin{aligned} f_{\Delta_t}(x) &= g'(0, t+x) e^{-g(0, t+x)} \\ &\quad + \int_0^t g'(0, y) \cdot g'(y, t+x) \cdot e^{-g(y, t+x)} dy \end{aligned}$$

$$g'(a, b) = \frac{d}{dt} g(a, t) \Big|_{t=b}$$

Use Inverse method with $F_{R_2}(t)$ or

Accept/Reject method with $f_{R_2}(t)$ to

perform Kinetic Monte Carlo.

$$\underline{F_k}.$$

$$\lambda(t) = e^{-bt} + a$$

intensity
time.

$$g(t_1, t_2) = \int_{t_1}^{t_2} \lambda(u) du$$

$$= \frac{1}{b} \left[e^{-bt_1} - e^{-bt_2} \right] + a(t_2 - t_1)$$

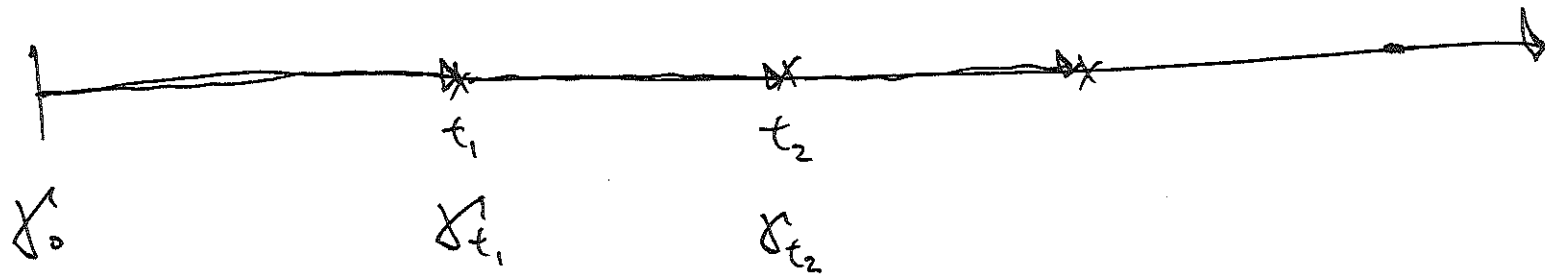
$$g'(0, t+x) = \frac{d}{dx} g(0, t+x)$$

$$= \cancel{\frac{1}{b}} e^{-b(t+x)} + a$$

$$g'(0, y) = \frac{d}{dy} g(0, y) = e^{-b(y)} + a$$

$$g'(y, t+x) = \frac{d}{dx} g(y, t+x) = e^{-b(t+x)} + a$$

Kinetic Monte Carlo for NHPP



Each time δ_t must be generated with
new value of τ .

Simulating
Non homogeneous PP

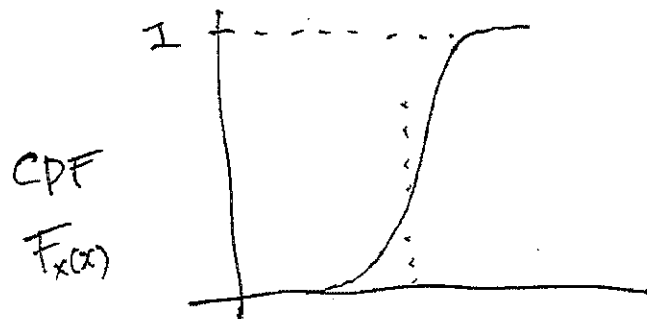
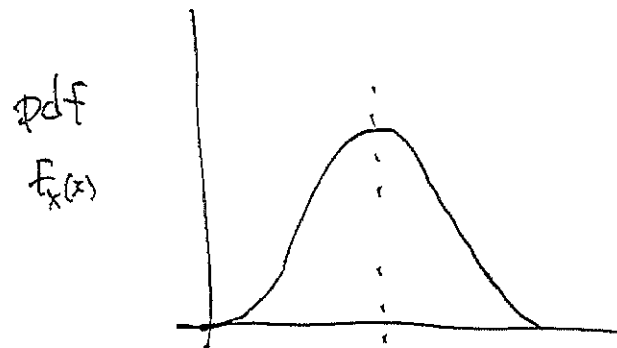
- Monte Carlo
- Kinetic Monte Carlo.

Inverse Method

$$X \sim F_X(x)$$

(CDF)

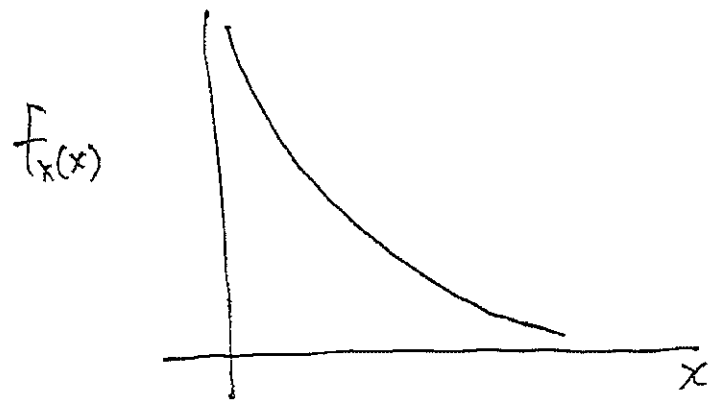
Normal



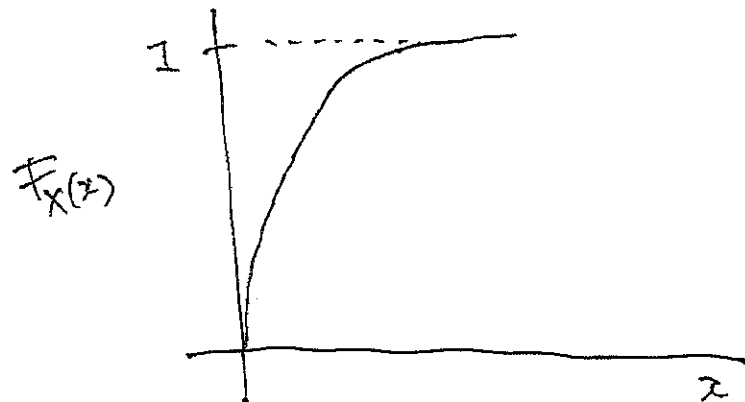
$$f_X(x) = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$F_X(x) = \int_{-\infty}^x f_X(t) dt$$

Exp



$$f_X(x) = \frac{1}{\lambda} e^{-x/\lambda} \quad x \geq 0$$



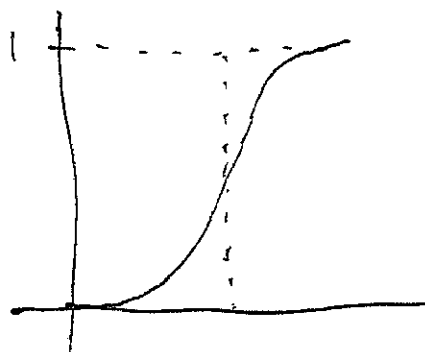
$$\begin{aligned} F_X(x) &= \int_0^x \frac{1}{\lambda} e^{-t/\lambda} dt \quad x \geq 0 \\ &= 1 - e^{-x/\lambda} \end{aligned}$$

Thm: what ever the distribution is ;

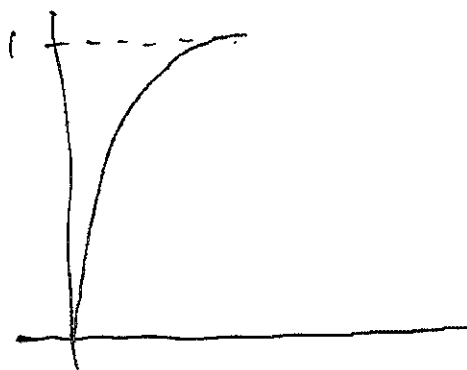
If $X \sim F_X$, then

$$F_X(X) \sim \text{UNIF}(0, 1)$$

Normal



Exp



Inverse Method for Generating R.V.

want to generate $X \sim F_X(x)$.

~~The idea is to know what~~

We know that

$$F_X(X) \stackrel{\text{in distribution}}{=} U$$

That means,

$$X \stackrel{\text{in dist.}}{=} F_X^{-1}(U)$$

where $U \sim \text{UNIF}(0,1)$.

We know how to generate U .

in R: `runif(n)`

Ex

$$X \sim \text{Exp}(5)$$

$$F_X(x) = 1 - e^{-x/5}$$

let

$$U = 1 - e^{-x/5}$$

solving for x ,

$$U - 1 = -e^{-x/5}$$

$$-\frac{x}{5} = \ln(1 - U)$$

$$X = -5 \cdot \ln(1 - U)$$

but, since

$$1 - U \sim \text{Unif}(0, 1)$$

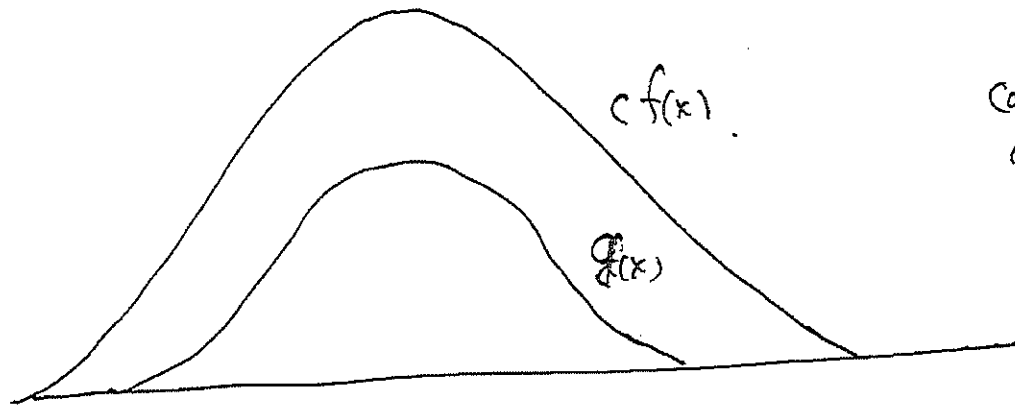
$$\boxed{X = -5 \ln(U)}$$

① generate n ~~ts~~ from $\text{Unif}(0, 1)$, call it u .

② let $X = -5 \ln(u)$,

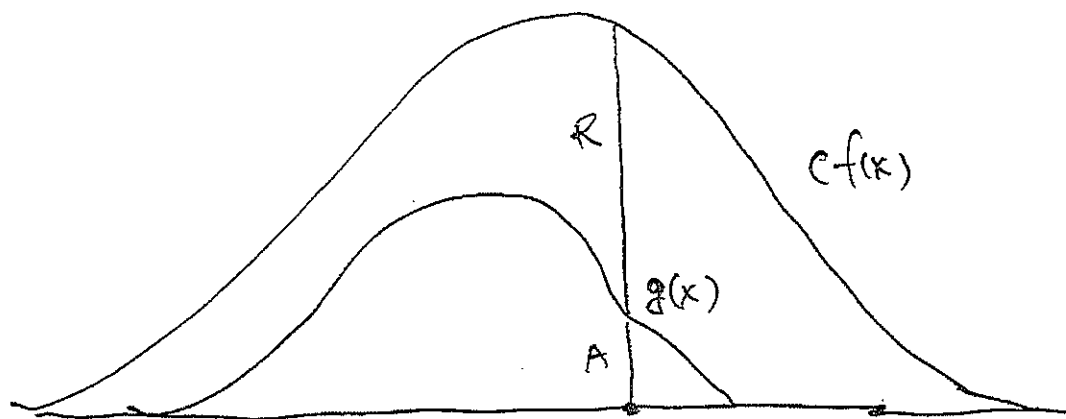
③ X is exponentially distributed with mean 5.

Acceptance - Rejection Method



$$g(x) \leq c f(x)$$

↑ ↑
can't can
generate generate



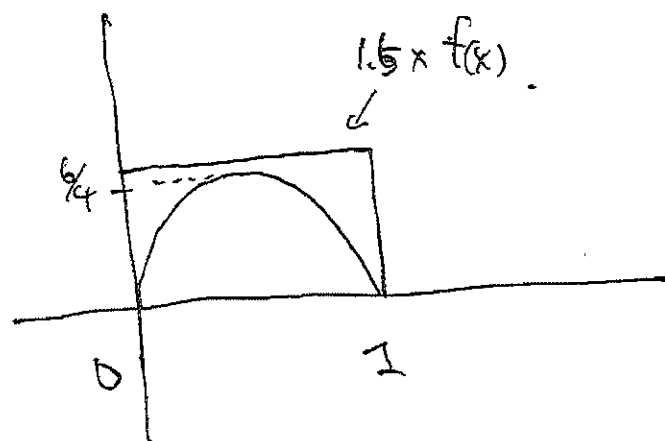
- ① generate X_1 from $f(x)$.
- ② ~~with~~ generate $U \sim \text{UNIF}(0, c f(x_1))$
- ③ if $U < g(x)$, then accept X_1 as generation from $g(x)$.
if not, disregard X_1 .
- ④ repeat.

Ex . A-R for Beta (2, 2)

Beta (2, 2)

$$f(x) = 6x(1-x)$$

pdf



① generate $X_1 \sim \text{UNIF}(0, 1)$

② generate $U \sim \text{UNIF}(0, 1.5)$

③ if $U < 6X_1(1-X_1)$

accept.

④ repeat.

$f(x)$ = pdf of $\text{UNIF}(0, 1)$.