





Time Series Analysis Seasonal Models

Seasonal Models

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Seasonal Models

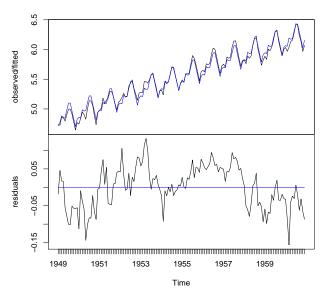
Seasonal ARIMA Models

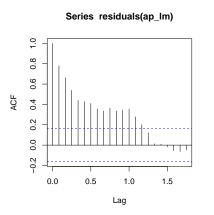
So far:

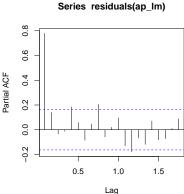
- Included seasonal patterns in deterministic models but not ARIMA models.
- In deterministic models, correlations are not exploited.
- Associated residuals are often highly autocorrelated.
- Need for stochastic seasonal models.

Illustration: Deterministic linear trend plus seasonal effects for monthly airline passenger numbers from 1949(1) to 1960(12).

```
R> data("AirPassengers", package = "datasets")
R> ap <- log(AirPassengers)
R> ap_lm <- dynlm(ap ~ trend(ap) + season(ap))</pre>
```







Residuals are not white noise and may even be nonstationary.

Idea: Introduce AR and/or MA effects at seasonal lag s, e.g., s=4 for quarterly data or s=12 for monthly data.

Example: MA(12) model with only one non-zero coefficient.

$$Y_t = e_t - \Theta e_{t-12}$$

Properties: Stationary with nonzero autocorrelation only at lag 12.

$$Cov(Y_t, Y_{t-1}) = Cov(e_t - \Theta e_{t-12}, e_{t-1} - \Theta e_{t-13}) = 0$$

 $Cov(Y_t, Y_{t-12}) = Cov(e_t - \Theta e_{t-12}, e_{t-12} - \Theta e_{t-24}) = -\Theta \sigma_e^2$

Definition: Seasonal MA(Q) model of order Q with seasonal period s.

$$Y_t = e_t - \Theta_1 e_{t-s} - \Theta_2 e_{t-2s} - \dots - \Theta_Q e_{t-Qs}$$

Characteristic polynomial:

$$\Theta(x) = 1 - \Theta_1 x^s - \Theta_2 x^{2s} - \dots - \Theta_Q x^{Qs}$$

Autocorrelation: Zero except at seasonal lags $s, 2s, \ldots, Qs$.

$$\rho_{ks} \ = \ \frac{\Theta_k + \Theta_1 \Theta_{k+1} + \dots + \Theta_{Q-k} \Theta_Q}{1 + \Theta_1^2 + \dots + \Theta_Q^2} \qquad \text{for } k = 1, \dots, Q$$

Invertibility condition: All roots of $\Theta(x) = 0$ must exceed 1 in absolute value (i.e., lie outside the complex unit circle).

Analogously: AR(12) model with only one nonzero coefficient.

$$Y_t = \Phi Y_{t-12} + e_t$$

Properties: Stationary with nonzero autocorrelation only at seasonal lags 12, 24, 36, ...

$$\rho_k = \Phi \rho_{k-12} \quad \text{for } k \ge 1$$
 $\rho_{12} = \Phi \rho_0 = \Phi$
 $\rho_{24} = \Phi \rho_{12} = \Phi^2$
 $\rho_{12k} = \Phi^k \quad \text{for } k \ge 1$
 $\rho_1 = \Phi \rho_{11}$
 $\rho_{11} = \Phi \rho_1$

and hence $\rho_1 = \rho_{11} = 0$.

Definition: Seasonal AR(P) model of order P with seasonal period s.

$$Y_t = \Phi_1 Y_{t-s} + \Phi_2 Y_{t-2s} + \ldots + \Phi_P Y_{t-Ps} + e_t$$

Characteristic polynomial:

$$\Phi(x) = 1 - \Phi_1 x^s - \Phi_2 x^{2s} - \dots - \Phi_P x^{Ps}$$

Autocorrelation: Zero except at seasonal lags $s, 2s, 3s, \ldots$ E.g., for seasonal AR(1)

$$\rho_{ks} = \Phi^k$$
 for $k = 1, 2, \dots$

Stationarity condition: All roots of $\Phi(x) = 0$ must exceed 1 in absolute value (i.e., lie outside the complex unit circle).

Seasonal Models

Multiplicative Seasonal ARMA Models

Idea: Combine autocorrelations at seasonal lags with those at neighboring lags.

Conceivable approaches: Illustrated for MA model with one seasonal and one nonseasonal coefficient.

 Subset MA: Consider subset of MA(12) with nonzero coefficients at lags 1 and 12.

$$Y_t = e_t - \theta_1 e_{t-1} - \theta_{12} e_{t-12}$$

Multiplicative seasonal MA: Consider MA characteristic polynomial

$$(1 - \theta x)(1 - \Theta x^{12}) = 1 - \theta x - \Theta x^{12} + \theta \Theta x^{13}$$

leading to the model

$$Y_t = e_t - \theta e_{t-1} - \Theta e_{t-12} + \theta \Theta e_{t-13}$$

The latter is more commonly used in practice.

Properties: Stationary with nonzero autocorrelations only at lags 1, 11, 12, 13.

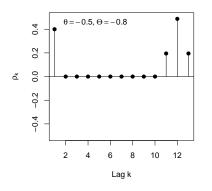
$$\gamma_0 = (1 + \theta^2)(1 + \Theta^2)\sigma_e^2$$

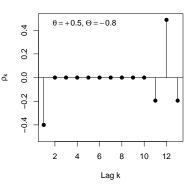
$$\rho_1 = -\frac{\theta}{1 + \theta^2}$$

$$\rho_{11} = \rho_{13} = \frac{\theta\Theta}{(1 + \theta^2)(1 + \Theta^2)}$$

$$\rho_{12} = -\frac{\Theta}{1 + \Theta^2}$$

Example: ACF for $\Theta = -0.8$ and $\theta = \pm 0.5$.





Definition: Multiplicative seasonal ARMA model with orders p and q and seasonal orders P and Q at seasonal lag s, ARMA(p, q)(P, Q) $_s$ for short.

$$\phi(B)\Phi(B)Y_t = \theta(B)\Theta(B)e_t$$

with polynomials

$$\phi(x) = 1 - \phi_1 x^1 - \phi_2 x^2 - \dots - \phi_p x^p
\Phi(x) = 1 - \Phi_1 x^s - \Phi_2 x^{2s} - \dots - \Phi_p x^{ps}
\theta(x) = 1 - \theta_1 x^1 - \theta_2 x^2 - \dots - \theta_q x^q
\Theta(x) = 1 - \Theta_1 x^s - \Theta_2 x^{2s} - \dots - \Theta_0 x^{Qs}$$

Note: A mean or intercept can be included as before.

Remarks:

- ARMA $(p, q)(P, Q)_s$ is a special case of ARMA(p + Ps, q + Qs).
- Coefficients are not completely free but determined by p + P + q + Q parameters.
- Typically, much more parsimonious, especially if s is large.
- Sometimes also called SARMA.

Special case: ARMA $(0, 1)(1, 0)_{12}$.

$$Y_t = \Phi Y_{t-12} + e_t - \theta e_{t-1}$$

which has

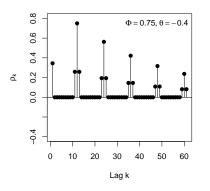
$$\gamma_1 = \Phi \gamma_{11} - \theta \sigma_e^2$$

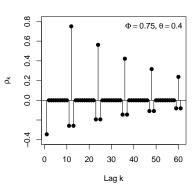
 $\gamma_k = \Phi \gamma_{k-12} \quad \text{for } k \ge 2$

Hence:

$$\begin{array}{rcl} \gamma_0 & = & \frac{1+\theta^2}{1-\Phi^2} \ \sigma_{\rm e}^2 \\ & & \\ \rho_{12k} & = & \Phi^k & {\rm for} \ k \geq 1 \\ & & \\ \rho_{12k-1} & = & \rho_{12k+1} & = & -\frac{\theta}{1+\theta^2} \Phi^k & {\rm for} \ k = 0, 1, 2, \dots \end{array}$$

Example: ACF for $\Phi = 0.75$ and $\theta = \pm 0.4$.





Seasonal Models

Nonstationary Seasonal ARIMA Models

Nonstationary seasonal ARIMA models

Idea: Analogous to differencing in nonstationary series, seasonal nonstationarity may be captured by *seasonal differences*.

$$\Delta_s = 1 - B^s$$

$$\Delta_s Y_t = Y_t - Y_{t-s}$$

The resulting series is then of length n - s.

Example: Seasonal random walk $\{S_t\}$ plus independent noise.

$$Y_t = S_t + e_t$$

 $S_t = S_{t-s} + \varepsilon_t$

If $\sigma_{\varepsilon} \ll \sigma_{e}$, $\{S_{t}\}$ captures a slowly changing seasonal component.

Nonstationary seasonal ARIMA models

Nonstationarity: Y_t inherits nonstationarity from S_t . This can be removed via seasonal differences.

$$\Delta_s Y_t = S_t + e_t - S_{t-s} - e_{t-s}$$
$$= \varepsilon_t + e_t - e_{t-s}$$

which is stationary and has the ACF of a seasonal $MA(1)_s$ model (depending on ratio of noise variances).

Analogously: Consider an additional nonseasonal stochastic trend.

$$Y_t = M_t + S_t + e_t$$

$$S_t = S_{t-s} + \varepsilon_t$$

$$M_t = M_{t-1} + \xi_t$$

with $\{e_t\}$, $\{\varepsilon_t\}$, $\{\xi_t\}$ mutually independent white noise series.

Nonstationary seasonal ARIMA models

Nonstationarity: Can be removed by taking both seasonal and nonseasonal differences.

$$\Delta \Delta_{s} Y_{t} = \Delta (M_{t} - M_{t-s} + \varepsilon_{t} + e_{t} - e_{t-s})
= (\xi_{t} + \varepsilon_{t} + e_{t}) - (\varepsilon_{t-1} + e_{t-1}) - (\xi_{t-s} + e_{t-s}) + e_{t-s-1}$$

This process is stationary and has nonzero autocorrelations only at lags 1, s-1, s, s+1. ACF is equivalent to seasonal ARMA(0, 1)(0, 1) $_s$.

Definition: Multiplicative seasonal ARIMA model with orders p, d, and q and seasonal orders P, D, and Q at seasonal lag s, denoted ARIMA(p, d, q)(P, D, Q) $_s$ for short.

$$\phi(B)\Phi(B)(1-B)^{d}(1-B^{s})^{D}Y_{t} = \theta(B)\Theta(B)e_{t}$$

Note: This can capture the structure of a wide range of empirical time series, often with just a few parameters.

Seasonal Models

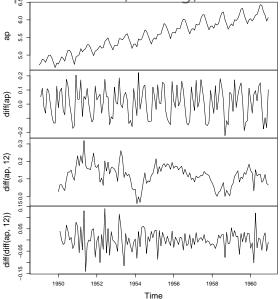
Approach: Employ the same ideas as for nonseasonal models.

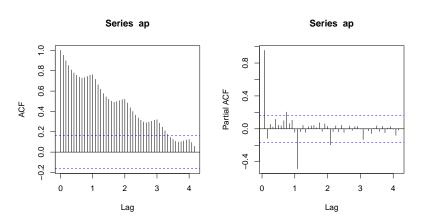
- Model specification: Inspect time series plot, ACF and PACF of suitably differenced series. Can also be accompanied by stationarity or (seasonal) unit root tests.
- Model fitting: Fit a preliminary (seasonal) ARIMA model.
- Diagnostic checking: Inspect (standardized) residuals (series, QQ plot, ACF, Ljung-Box tests, . . .).

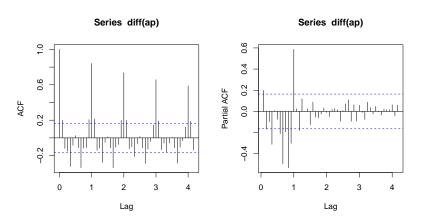
In R:

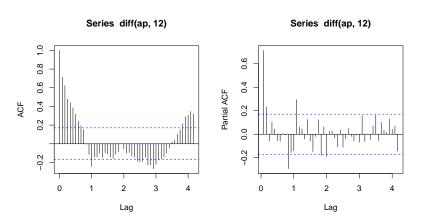
- diff(y, s) can also compute seasonal differences.
- arima() can also include seasonal orders.
- tsdiag() can be applied as before.

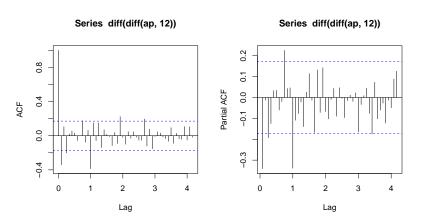
Illustration: Revisit AirPassengers data.





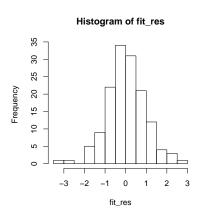


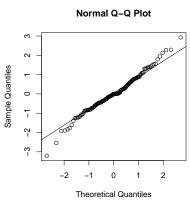


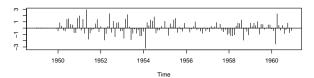


Model fitting: Employ seasonal ARIMA $(0, 1, 1)(0, 1, 1)_{12}$.

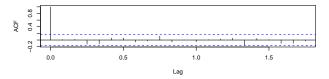
Model fitting: Employ seasonal ARIMA $(0, 1, 1)(0, 1, 1)_{12}$.



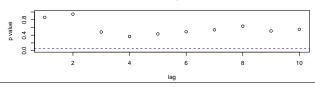




ACF of Residuals



p values for Ljung-Box statistic



Furthermore: The same model would be selected by an information-criteria-based approach, e.g., auto.arima(ap, d = 1, D = 1, approx = FALSE, stepwise = FALSE) or auto.arima(ap, d = 1, D = 1, approx = FALSE) etc.

Summary:

- ARIMA(0, 1, 1)(0, 1, 1)₁₂ model leads to satisfactory fit and has appealing simple interpretation.
- The model for this data was popularized by Box & Jenkins (1976) and is hence also called *airline model*.
- It is also found to yield satisfactory fits for many other time series.

Seasonal Models

Forecasting Seasonal Models

Again: The same strategies for computing forecasts and associated standard errors can be used as in the nonseasonal case.

In R: As before, predict() or forecast().

Example: Seasonal $AR(1)_{12}$.

$$Y_t = \Phi Y_{t-12} + e_t$$

$$\widehat{Y}_t(\ell) = \Phi \widehat{Y}_t(\ell - 12)$$

$$= \Phi^{k+1} Y_{t+r-11}$$

where k is the integer part and r/12 is the fractional part of $(\ell-1)/12$.

The $MA(\infty)$ weights are nonzero only for multiple lags of 12.

$$\psi_j \; = \; \left\{ egin{array}{ll} \Phi^{j/12} & {
m for} \, j=0,12,24, \ldots \ 0 & {
m otherwise} \end{array}
ight.$$

Hence: With *k* defined as before.

$$Var(e_t(\ell)) = \left(\frac{1 - \Phi^{2k+2}}{1 - \Phi^2}\right) \sigma_e^2$$

Example: Seasonal $MA(1)_{12}$ with intercept.

$$Y_t = \theta_0 + e_t - \Theta e_{t-12}$$
 $\widehat{Y}_t(1) = \theta_0 - \Theta e_{t-11}$
 $\widehat{Y}_t(2) = \theta_0 - \Theta e_{t-10}$
 \vdots
 $\widehat{Y}_t(12) = \theta_0 - \Theta e_t$
 $\widehat{Y}_t(\ell) = \theta_0 \text{ for } \ell > 12$

Thus: The MA(∞) weights are simply $\psi_0=1$, $\psi_{12}=\Theta$, and $\psi_j=0$ otherwise. This yields

$$\mathsf{Var}(\mathsf{e}_t(\ell)) \ = \ \left\{ egin{array}{ll} \sigma_\mathsf{e}^2 & 1 \leq \ell \leq 12 \ (1 + \Theta^2)\sigma_\mathsf{e}^2 & \ell > 12 \end{array}
ight.$$

Example: ARIMA $(0, 1, 1)(0, 1, 1)_{12}$ (airline model).

$$Y_t = Y_{t-1} + Y_{t-12} - Y_{t-13} + e_t - \theta e_{t-1} - \Theta e_{t-12} + \theta \Theta e_{t-13}$$

Then, the forecasts combine properties of AR and MA forecasts:

and

$$\widehat{Y}_t(\ell) = \widehat{Y}_t(\ell-1) + \widehat{Y}_t(\ell-12) - \widehat{Y}_t(\ell-13)$$
 for $\ell > 13$

Forecasts from ARIMA(0,1,1)(0,1,1)[12]

