

### Outline

- Autoregressive Processes
- ☐ The Mixed Autoregressive Moving Average Model
- □ Stationarity of AR(1) and AR(2)
- Invertibility

## Autoregressive Processes

Autoregressive processes are as their name suggests—regressions on themselves. Specifically, a pth-order autoregressive process  $\{Y_t\}$  satisfies the equation

$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + e_t$$

The current value of the series  $Y_t$  is a linear combination of the p most recent past values of itself plus an "innovation" term  $e_t$  that incorporates everything new in the series at time t that is not explained by the past values. Thus, for every t, we assume that  $e_t$  is independent of  $Y_{t-1}$ ,  $Y_{t-2}$ ,  $Y_{t-3}$ , ... Yule (1926) carried out the original work on autoregressive processes

### The First-Order Autoregressive Processes

$$\mathsf{AR}(1) | Y_t = \phi Y_{t-1} + e_t$$

 $e_t \sim iid(0, \sigma_e^2)$ 

we are assuming that  $Y_t$  has zero mean.

$$E(Y_t) = 0$$

we also assume that  $e_t$  is independent of  $Y_{t-1}, Y_{t-2}, Y_{t-3}, \dots$ 

$$E(e_t, Y_{t-1}) = 0, E(e_t, Y_{t-2}) = 0, ..., dst$$

$$E(e_t Y_{t-k}) = E(e_t) E(Y_{t-k}) = 0$$

$$Var(Y_t) = Var(\phi Y_{t-1} + e_t)$$

$$\gamma_0 = \phi^2 \gamma_0 + \sigma_e^2$$

Solving for  $\gamma_0$  yields

$$\gamma_0 = \frac{\sigma_e^2}{1 - \phi^2}$$

 $\phi^2 < 1$  or that  $|\phi| < 1$ .

### The First-Order Autoregressive Processes

$$\mathsf{AR}(1) \mid Y_t = \phi Y_{t-1} + e_t$$

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$$E(Y_t) = 0$$

we also assume that  $e_t$  is independent of  $Y_{t-1}, Y_{t-2}, Y_{t-3}, \dots$ 

$$E(e_t, Y_{t-1}) = 0, E(e_t, Y_{t-2}) = 0, ..., dst$$

$$E(e_t Y_{t-k}) = E(e_t) E(Y_{t-k}) = 0$$

$$E(Y_{t-k}Y_t) = \phi E(Y_{t-k}Y_{t-1}) + E(e_tY_{t-k})$$

or

$$\gamma_k = \phi \gamma_{k-1} + E(e_t Y_{t-k})$$

$$\gamma_k = \phi \gamma_{k-1}$$
 for  $k = 1, 2, 3,...$ 

Setting k = 1, we get  $\gamma_1 = \phi \gamma_0 = \phi \sigma_e^2 / (1 - \phi^2)$ . With k = 2, we obtain  $\gamma_2 = \phi^2 \sigma_e^2 / (1 - \phi^2)$ . Now it is easy to see that in general

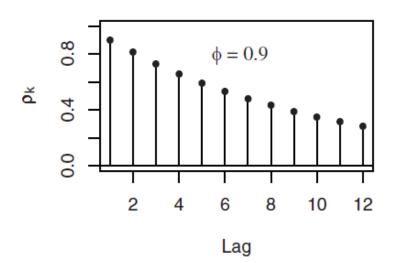
$$\gamma_k = \phi^k \frac{\sigma_e^2}{1 - \phi^2}$$

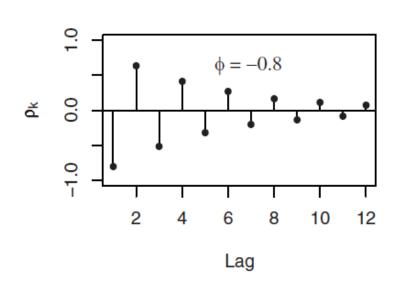
and thus

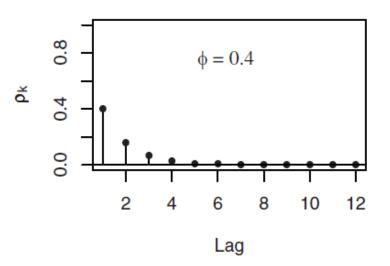
$$\rho_k = \frac{\gamma_k}{\gamma_0} = \phi^k$$
 for  $k = 1, 2, 3,...$ 

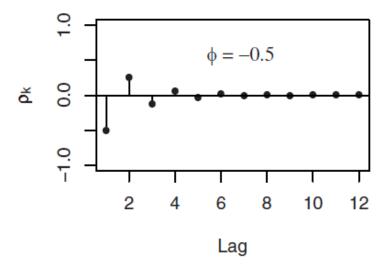
### The First-Order Autoregressive Processes

- Since  $|\phi| < 1$ , the magnitude of the autocorrelation function decreases exponentially as the number of lags, k, increases.
- □ If  $0 < \phi < 1$ , all correlations are positive;
- If  $-1 < \phi < 0$ , the lag 1 autocorrelation is negative  $(\rho_1 = \phi)$  and the signs of successive autocorrelations alternate from positive to negative, with their magnitudes decreasing exponentially.

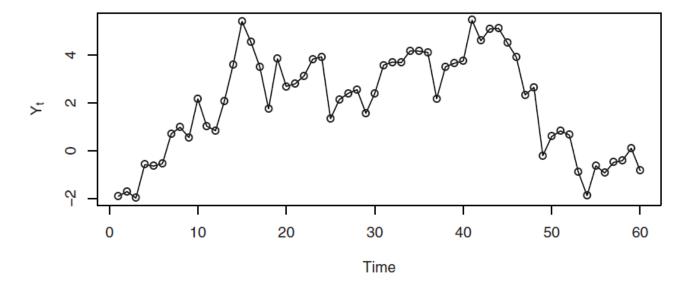








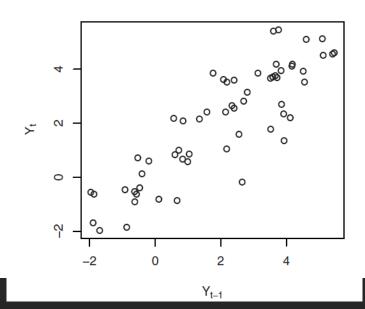
#### Time Plot of an AR(1) Series with $\phi = 0.9$



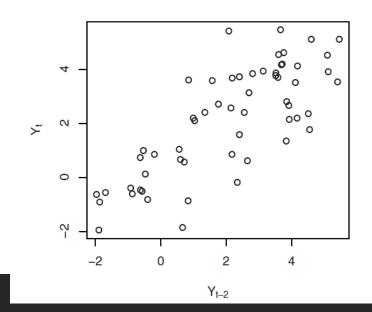
# AR(1)

$$Y_t = \phi Y_{t-1} + e_t$$

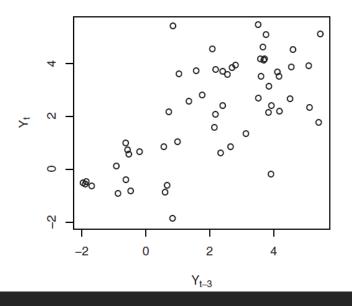
#### Plot of Yt versus Yt - 1 for AR(1)



#### Plot of Yt versus Yt - 2 for AR(1)



#### Plot of Yt versus Yt - 3 for AR(1)



#### The General Linear Process Version of the AR(1) Model

The General Linear Process Version of the AR(1) Model can be written as follows:

$$Y_t = e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \dots + \phi^{k-1} e_{t-k+1} + \phi^k Y_{t-k}$$

Assuming  $|\phi| < 1$  and letting k increase without bound, it seems reasonable (this is almost a rigorous proof) that we should obtain the infinite series representation

$$Y_t = e_t + \phi e_{t-1} + \phi^2 e_{t-2} + \phi^3 e_{t-3} + \cdots$$

### The Second-Order Autoregressive Processes

AR(2) 
$$Y_t = \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + e_t$$

$$e_t \sim iid(0, \sigma_e^2)$$

we also assume that  $e_t$  is independent of  $Y_{t-1}$ ,  $Y_{t-2}$ ,  $Y_{t-3}$ , ...

$$E(e_t, Y_{t-1}) = 0, E(e_t, Y_{t-2}) = 0, ..., dst$$

How to get the variance for AR(2) model?

$$\gamma_k = \phi_1 \gamma_{k-1} + \phi_2 \gamma_{k-2}$$
 for  $k = 1, 2, 3, ...$ 

or, dividing through by  $\gamma_0$ ,

$$\rho_k = \phi_1 \rho_{k-1} + \phi_2 \rho_{k-2}$$
 for  $k = 1, 2, 3, ...$ 

### These Equations are usually called the **Yule-Walker equations**

Setting k = 1 and using  $\rho_0 = 1$  and  $\rho_{-1} = \rho_1$ , we get and so

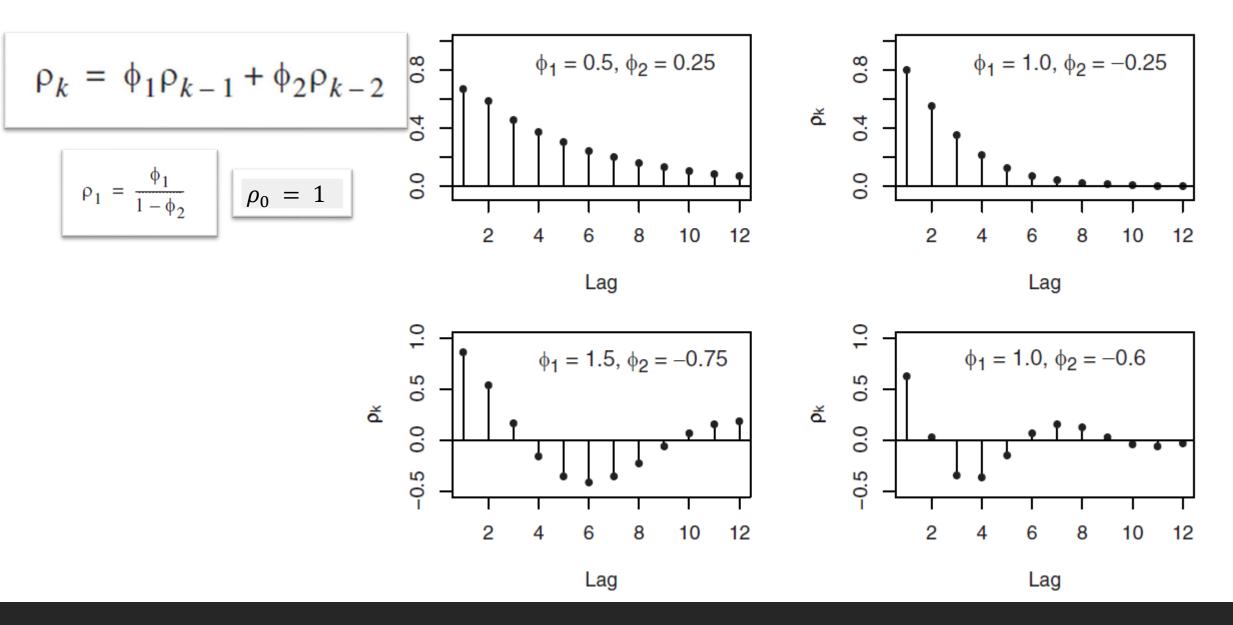
$$\rho_1 = \frac{\phi_1}{1 - \phi_2}$$

Using the now known values for  $\rho_1$  (and  $\rho_0$ ) and  $\rho_k$ , for k=2 we can obtain

$$\rho_2 = \phi_1 \rho_1 + \phi_2 \rho_0$$

$$= \frac{\phi_2 (1 - \phi_2) + \phi_2}{1 - \phi_2}$$

#### Autocorrelation Functions for Several AR(2) Models



### The Mixed Autoregressive Moving Average Model

If we assume that the series is partly autoregressive and partly moving average, we obtain a quite general time series model. In general, if

$$\begin{split} Y_t &= \phi_1 Y_{t-1} + \phi_2 Y_{t-2} + \dots + \phi_p Y_{t-p} + e_t - \theta_1 e_{t-1} - \theta_2 e_{t-2} \\ &- \dots - \theta_q e_{t-q} \end{split}$$

we say that  $\{Y_t\}$  is a mixed autoregressive moving average process of orders p and q, respectively; we abbreviate the name to ARMA(p,q)

#### The ARMA(1,1) Model

The defining equation can be written

$$Y_t = \phi Y_{t-1} + e_t - \theta e_{t-1}$$

To derive Yule-Walker type equations, we first note that

$$\begin{split} E(e_t Y_t) &= E[e_t (\phi Y_{t-1} + e_t - \theta e_{t-1})] \\ &= \sigma_e^2 \end{split}$$

$$\begin{split} E(e_{t-1}Y_t) &= E[e_{t-1}(\phi Y_{t-1} + e_t - \theta e_{t-1})] \\ &= \phi \sigma_e^2 - \theta \sigma_e^2 \\ &= (\phi - \theta)\sigma_e^2 \end{split}$$

$$\gamma_0 = \phi \gamma_1 + [1 - \theta(\phi - \theta)] \sigma_e^2 \gamma_1$$

$$\gamma_1 = \phi \gamma_0 - \theta \sigma_e^2$$

$$\gamma_k = \phi \gamma_{k-1} \quad \text{for } k \ge 2$$

Solving the first two equations yields

$$\gamma_0 = \frac{(1 - 2\phi\theta + \theta^2)}{1 - \phi^2} \sigma_e^2$$

and solving the simple recursion gives

$$\rho_k = \frac{(1 - \theta\phi)(\phi - \theta)}{1 - 2\theta\phi + \theta^2} \phi^{k - 1} \quad \text{for } k \ge 1$$

## Stationarity of AR(1)

- □ It can be shown that, subject to the restriction that  $e_t$  be independent of  $Y_{t-1}$ ,  $Y_{t-2}$ ,  $Y_{t-3}$ , ... and that  $\sigma_e^2 > 0$ ,
- □ The solution of the AR(1) defining recursion  $Y_t = \phi Y_{t-1} + e_t$  will be stationary if and only  $|\phi| < 1$ .
- □ The requirement  $|\phi|$  < 1 is usually called the stationary condition for the AR(1) process.

# Stationarity of AR(2)

The AR(2) process is called the stationary if and only if three conditions are satisfied:

$$\phi_1 + \phi_2 < 1$$
,  $\phi_2 - \phi_1 < 1$ , and  $\left| \phi_2 \right| < 1$ 

As with the AR(1) model, we call these the stationarity conditions for the AR(2) model.

# Invertibility $(MA(q) \rightarrow AR(p))$

An MA(1) model:

$$Y_t = e_t - \theta e_{t-1}$$

First rewriting this as  $e_t = Y_t + \theta e_{t-1}$  and then replacing t by t-1 and substituting for  $e_{t-1}$  above, we get

$$e_{t} = Y_{t} + \theta(Y_{t-1} + \theta e_{t-2})$$
$$= Y_{t} + \theta Y_{t-1} + \theta^{2} e_{t-2}$$

If  $|\theta| < 1$ , we may continue this substitution "infinitely" into the past and obtain the expression

$$e_t = Y_t + \theta Y_{t-1} + \theta^2 Y_{t-2} + \cdots$$

or

$$Y_t = (-\theta Y_{t-1} - \theta^2 Y_{t-2} - \theta^3 Y_{t-3} - \cdots) + e_t$$

- □ If  $|\theta|$  < 1, we see that the MA(1) model can be inverted into an infinite-order autoregressive model.
- □ We say that the MA(1) model is invertible if and only if  $|\theta| < 1$ .

## Invertibility

- □ Invertibility refers to the fact that the moving average (MA) models can be written as an autoregressive (AR) model
- Or more generally, if ARMA models can be written as AR models, we say that the time series model is invertible.

# Assignment

- 1. Find the autocovariance function and the autocorrelation function for AR (3)
- 2. Sketch the autocorrelation function for AR (2) model as follows:
  - $\phi_1 = 0.6 \text{ and } \phi_2 = 0.3$
  - $\phi_1 = -0.4$  and  $\phi_2 = 0.5$
  - $\phi_1 = 1.2 \text{ and } \phi_2 = -0.7$
- 3. Find the autocovariance function and the autocorrelation function for ARMA(1,2)
- 4. Find the autocovariance function and the autocorrelation function for ARMA(2,1)
- 5. Find the invertibility condition for MA(2) process

# Thanks