PROBABILITY OCA-1

In this activity, we will explore how to compute the *probability density function* for the **product** of two continuous random variables using the **distribution method**.

To be concrete, we will consider two **independent** uniformly distributed random variables,  $X_1$  and  $X_2$ .

• The density for  $X_1$  we'll call g(x),

$$g(x) = \begin{cases} 1 & 0 \le x \le 1, \\ 0 & \text{otherwise.} \end{cases}$$

• And the density for  $X_2$  we'll call h(x),

$$h(x) = \begin{cases} \frac{1}{2} & 0 \le x \le 2, \\ 0 & \text{otherwise.} \end{cases}$$

Our goal is to compute the density for the random variable X formed by multiplying  $X_1$  and  $X_2$ ,

$$X = X_1 \cdot X_2.$$

You could imagine, for example,  $X_1$  being a random variable for current and  $X_2$  being a random variable for resistance. According to Georg Ohm, X would then measure voltage.

1. In the space below, sketch the densities for  $X_1$  and  $X_2$ 

2. What range of values will X obtain? Do not multiply the density values!!

$$\leq X \leq$$

3. If X = x and  $X_1 = a$ , then what value must  $X_2$  have?

$$X_2 =$$

4. Given that the RVs  $X_1$  and  $X_2$  are **independent**, what is the value of their **joint density** (which we will temporarily call  $f_{X_1,X_2}$ )? Hint: See Equation 4.7' on page 40 of the text.

$$f_{X_1, X_2}(x_1, x_2) =$$

5. At this point, we will take a departure from the **convolution** method used for sums of RVs. The reason for this is that the expression  $\frac{x}{a}$  in (3) is not one-to-one when x = 0 and is undefined when a = 0.

To side-step this issue, we will make use of *cumulative distribution functions*. Recall, that if the density for X is f(x), then the cdf for X is

$$F(x) := P(X \le x) = \int_{-\infty}^{x} f(y) \ dy.$$

(a) Compute the cdf for  $X_1$ , which we will call G(x). Sketch the graph of G(x) in the space below. Check that G'(x) = g(x) everywhere that G is differentiable.

(b) Compute the cdf for  $X_2$ , which we will call H(x). Sketch the graph of H(x) in the space below. Check that H'(x) = h(x) everywhere that H is differentiable.

6. Next, for  $x \in [0, 2]$ , we will express the cdf for X as follows

$$F(x) := P(X \le x) = \int_{y=0}^{y=1} P(X_2 \le \frac{x}{y}) g(y) \ dy.$$

(a) Convince yourself that the expression above makes sense. Write a few words explaining how the integral on the right measures the probability on the left.

(b) To compute the integral, we need to carefully substitute H(x) into the integrand. If you have computed H(x) correctly above, you should have arrived at

$$H(x) = \begin{cases} 0 & x \ge 0, \\ \frac{x}{2} & 0 < x \le 2, \\ 1 & \text{otherwise.} \end{cases}$$

In the integral, however, we have  $H(\frac{x}{y})$ . When will  $H(\frac{x}{y}) = 1$ ? What values of y will guarantee this happens?

(c) In the previous question, you should have determined that  $H(\frac{x}{y}) = 1$  whenever  $y \leq \frac{x}{2}$ . Keeping in mind that g(y) = 1 for all  $y \in [0, 1]$ , we arrive at

$$F(x) = \int_{y=0}^{y=1} P(X_2 \le \frac{x}{y}) g(y) \ dy$$
$$= \int_0^1 H(\frac{x}{y}) \ dy$$
$$= \int_0^{\frac{x}{2}} dy + \int_{\frac{x}{2}}^1 \frac{x}{2y} \ dy.$$

In the analysis above, we split the integral according to how  $H(\frac{x}{y})$  would either equal 1 or  $\frac{x}{2y}$ . Implicitly, we are assuming that  $x \in [0, 2]$ .

(d) Evaluate the integrals above to arrive at an expression for the cdf of  $X = X_1 \cdot X_2$ . Since you are integrating with respect to y, you can treat x as a fixed value for the purpose of evaluating the integral. Also, remember that  $\ln(1) = 0$ .

If you have done everything correctly so far, you should have arrived at

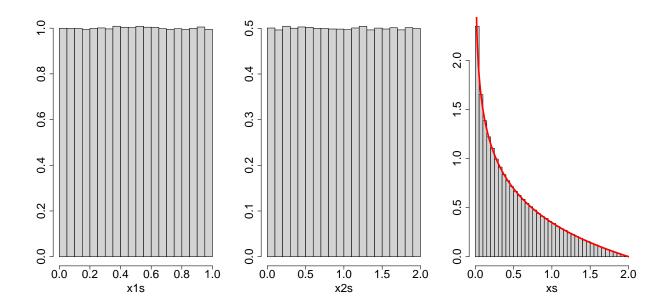
$$F(x) = \begin{cases} 0 & x \ge 0, \\ \frac{x}{2}(1 - \ln(\frac{x}{2})) & 0 < x \le 2, \\ 1 & \text{otherwise.} \end{cases}$$

7. Differentiate the expressions above to arrive at a formula for the density f(x) of the random variable X.

$$f(x) = \begin{cases} x \ge 0, \\ 0 < x \le 2, \end{cases}$$
 otherwise.

8. Excellent! In the space below sketch the graphs of both f(x) and F(x).

9. I often find it helpful to "check" analytic results by running simulations in R or Python. The following histograms were created using R – sample code is available on Canvas. The red-line on the rightmost plot corresponds to the density I computed for X using the method I presented in the preceding exercises. It's a good match!



10. Does the plot on the right look like your sketch of f(x)? Does the solution match your intuition? Why is 0 so popular?

11. BONUS ROUND!! Some of you might be wondering why we couldn't simply log-transform the product (to turn it into a sum) and then use convolutions. Perhaps you already have a workflow in mind.

1. 
$$X = X_1 \cdot X_2$$

2. 
$$\Longrightarrow \log(X) = \log(X_1) + \log(X_2)$$

- 3. Assume independence is preserved?
- 4. Convolve the densities for  $\log(X_1)$  and  $\log(X_2)$  to arrive at a density for  $\log(X)$ .
- 5. Somehow recover the density for X from the density for  $\log(X)$ .

Of course, we can't take the logarithm of 0, so there's that problem, bit it sounds promising for products in general. To make this plan actionable, however, we need a gadget that extract densities from transformed random variables. Here it is!

**TRANSFORMATION THEOREM** – Suppose X has pdf  $f_X(x)$  and form a new random variable Y by applying a monotonic function to X, that is

$$Y = g(X)$$

with g either strictly increasing or strictly decreasing on the set of all possible values of X (what we called D in class on Wednesday). One consequence of the monotonicity will be that g is invertible on D. Let's call this inverse function h, so that

$$h(g(x)) = x$$
 and  $g(h(y)) = y$ .

If h is differentiable (chances are good), then we can easily compute the density for the transformed variable Y.

$$f_Y(y) = f_X(h_Y)) \cdot |h'(y)|$$

12. For practice, suppose X is uniformly distributed on [1, 2]. What would the density be for  $X^2$ ?