

H_0 : All bkg, no signal

H_1 : bkg + signal

Poisson log likelihood: $P(e, o)$

$$P(e, o) = \prod_{\text{sample}} e_i - o_i + o_i \log \frac{o_i}{e_i}, o_i = \text{observed}, e_i = \text{expected}$$

dot = 2 b/c we usually want 2D graph

Critical values:

$(1 - \alpha) (\%)$	$m = 1$	$m = 2$	$m = 3$
68.27 1σ	1.00	2.30	3.53
90.	2.71	4.61	6.25
95.	3.84	5.99	7.82
95.45 2σ	4.00	6.18	8.03
99.	6.63	9.21	11.34
99.73 3σ	9.00	11.83	14.16

$P(e, o) > \text{critical value} \Rightarrow \text{reject } H_0$
(shown at left)

TABLE 1. Chi-square differences (δ) above minimum

Number of parameters = dot

Significance	α	1	2	3
0.68		1.00	2.30	3.50
0.90		2.71	4.61	6.25
0.99		6.63	9.21	11.30

$$\Delta \chi^2 = \chi^2 - \chi^2_{\min}$$

The entries are chi2.ppf

Estimator properties

Consistency: $\lim_{N \rightarrow \infty} \hat{\theta} = \theta$ \leftarrow true value of the parameter θ

Bias: $b = E[\hat{\theta}] - \theta$

Efficiency: $\hat{\theta}$ is efficient if its variance $V[\hat{\theta}]$ is small

Rao - Cramér - Fréchet minimum variance bound

Variance $\text{Var}[\hat{\theta}] \geq I^{-1}(\vec{\theta})$

← hat means estimator

θ = parameters
 x = measurement

f = unknown PDF
that we're interested in

$$\begin{aligned} \text{Information matrix } I_{jk}(\vec{\theta}) &= -E\left[\sum_{i=1}^N \frac{\partial^2 \ln f(x_i; \theta)}{\partial \theta_j \partial \theta_k}\right] \\ &= -N \int \frac{\partial^2 \ln f}{\partial \theta_j \partial \theta_k} f dx = N \int \frac{\partial \ln f}{\partial \theta_j} \frac{\partial \ln f}{\partial \theta_k} f dx \\ &= N \int \frac{1}{f} \frac{\partial f}{\partial \theta_j} \frac{\partial f}{\partial \theta_k} dx = E\left[\sum_{i=1}^N \frac{\partial \ln f}{\partial \theta_j} \frac{\partial \ln f}{\partial \theta_k}\right] \end{aligned}$$

Maximum likelihood

PDF $f(x_i; \theta, \sigma_i) = \frac{1}{\sqrt{2\pi\sigma_i^2}} \exp\left(-\frac{(x_i - \theta)^2}{2\sigma_i^2}\right)$

log-likelihood $\ln L(\theta) = -\frac{1}{2} \sum_{i=1}^N \frac{(x_i - \theta)^2}{\sigma_i^2} + \text{const.}$ Maximize $\ln L$

$$\begin{aligned} \ln L(\theta) &= \ln L(\hat{\theta}) - \frac{1}{2} (\theta - \hat{\theta})^2 \\ \Rightarrow L(\theta) &\propto \exp\left(-\frac{1}{2}(\theta - \hat{\theta})^2\right) \end{aligned}$$

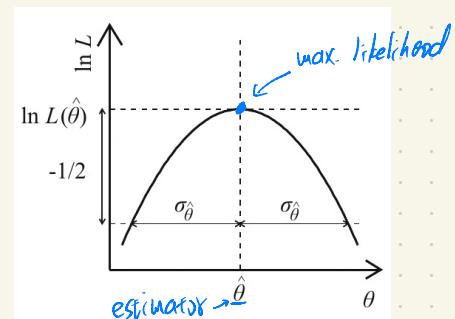
general form

$$L \propto \exp\left(-\frac{1}{2}(\vec{\theta} - \hat{\vec{\theta}})^t H(\vec{\theta} - \hat{\vec{\theta}})\right)$$

$$\text{Var}[\hat{\theta}] \rightarrow I(\hat{\theta})^{-1}$$

maximum likelihood reaches the bound

$$I_{ij}(\hat{\theta}) = -E\left[\frac{\partial^2 \ln L}{\partial \theta_j \partial \theta_k}\right] = H$$



$$\sigma_{\hat{\theta}} = h^{-\frac{1}{2}} = \left(-\frac{\partial^2 \ln L}{\partial \theta^2}\right)^{-\frac{1}{2}}$$

$$\ln L(\hat{\theta} \pm \sigma_{\hat{\theta}}) - \ln L(\hat{\theta}) = -\frac{1}{2}$$

2 ways to estimate uncertainties (or SD σ)

- $\hat{\sigma}_{\theta_j} = \sqrt{V_{jj}(\theta)}$ the diagonal element
- $\Delta \ln L = \ln L(\theta) - \ln L(\hat{\theta}) = -\frac{s^2}{2}$ for $s \cdot \sigma$ confidence interval

Least square

$$x \xrightarrow{f_{w/\theta}} y$$

$$\chi^2 = \sum_{i=1}^N \left[\frac{y_i - f(x_i; \vec{\theta})}{\sigma_i} \right]^2, \quad \sigma_i \text{ is often known } (\sim \sqrt{f(x_i; \theta)})$$

estimator $\hat{\theta}$ is found by minimizing χ^2 : $\frac{\partial \chi^2}{\partial \theta_j} = 0 \quad \forall j$

Comparing with previous, we have $\chi^2 = -2 \ln L$

general form of $\chi^2 = (\vec{y} - f(\vec{\theta}))^T V^{-1} (\vec{y} - f(\vec{\theta}))$

$$\vec{y} = \{y_i\}, \quad f(\vec{\theta}) = \{f(x_i; \vec{\theta})\}$$

V = covariance matrix of y

$$V_{ij} = E[(y_i - E[y_i])(y_j - E[y_j])]$$

$$V(\hat{\theta}) = \frac{1}{2} \left[\frac{\partial^2 \chi^2}{\partial \theta^2} \Big|_{\vec{\theta}=\hat{\vec{\theta}}} \right]^{-1} = H^{-1}$$

$$\chi^2(\vec{\theta}) = \chi^2_{\min} - s^2$$

for $s \cdot \sigma$ confidence interval

Linear least square

$$f(x; \theta) = \sum_{j=1}^N a_j(x) \theta_j$$

Suppose $f = A\theta$, where A = design matrix

$$\Rightarrow \text{chi2} = (y - A\theta)^T V^{-1} (y - A\theta)$$

$$\min \text{chi2} \Rightarrow \hat{\theta} = L y, \text{ where } L = (A^T V^{-1} A)^{-1} A^T V^{-1}$$

$$V(\hat{\theta}) = L V L^T = (A^T V^{-1} A)^{-1} = H^{-1}$$

$$\text{Hessian matrix } H = \frac{1}{2} \frac{\partial^2 \text{chi2}(L\hat{\theta})}{\partial \hat{\theta}^2} = A^T V^{-1} A$$

Jeffreys prior

The uniform prior $\pi(\theta)$ has problem if the range of θ $\leftarrow \infty$

$$1 = \int \pi(\theta) d\theta \Rightarrow \pi(\theta) \approx 0 \forall \theta$$

Robust measurement

- arbitrary choice of prior makes no great difference
- "uniform" prior is uniform is the correct variable $\pi(\theta) \sim \frac{1}{\theta^2}, \theta^2, \ln \theta, \text{etc}$

$$\pi(\theta) = \pi(\theta') \left| \frac{\partial \theta'}{\partial \theta} \right| \propto \sqrt{E\left[\left(\frac{\partial \ln L}{\partial \theta}\right)^2\right]} \left| \frac{\partial \theta'}{\partial \theta} \right|$$

parameter θ , prior π

$$\text{log-likelihood } \ln L = \sqrt{E\left[\left(\frac{\partial \ln L}{\partial \theta}\right)^2\right]} = \sqrt{I(\theta)}$$

$$\begin{aligned} \text{Fisher information } I(\theta) &= -E\left[\frac{\partial^2 \ln L}{\partial \theta^2}\right] = E\left[\left(\frac{\partial \ln L}{\partial \theta}\right)^2\right] \\ &= \int \left[\frac{\partial}{\partial \theta} \ln f(x; \theta) \right]^2 f(x; \theta) dx \end{aligned}$$

For multiple parameter, $\pi(\vec{\theta}) \propto \sqrt{\det I(\vec{\theta})}$, where I is Fisher matrix

Kolmogorov-Smirnov test

The best known unbinned goodness of fit test N is # of samples

empirical cumulative fn. $F_N(x) = N^{-1} \sum_{i=1}^N S(x_i)$, $S(x_i) = \begin{cases} 0 & x < x_i \\ 1 & x > x_i \end{cases}$

H_0 cumulative fn. $F(x)$

test stat: $D_N = \max |F_N(x) - F(x)| \quad \forall x$

p-value = $2 \sum_{i=1}^{\infty} (-1)^{i-1} \exp(-2i^2 N D_N^2) \Rightarrow$ find the D_N critical value

Smirnov-Cramér-von Mises test is alternative to this.

test stat: $W^2 = \int_{-\infty}^{\infty} [F_N(x) - F(x)]^2 dF(x)$

Feldman-Cousins confidence interval

$$\text{ordering principle } R = \frac{L(x|\mu)}{L(x|\hat{\mu})} \quad \begin{array}{l} \text{likelihood } L, \text{ observed } x \\ \text{actual mean } \mu, \text{ max likelihood mean } \hat{\mu} \end{array}$$

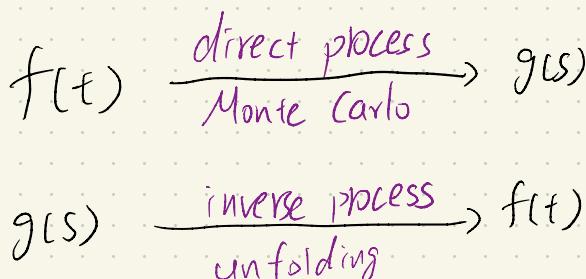
The interval $[x_1, x_2]$ is given by solving

- $R(x_1) = R(x_2)$
- $\int_{x_1}^{x_2} L(x|\mu) dx = 1 - \alpha$ for confidence level α

Unfolding

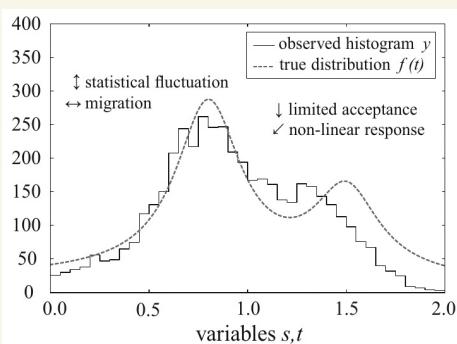
true distri. $f(t)$

measured (binned) distri. $g(s)$



Fredholm integral equation

$$\underbrace{\int K(s,t) f(t) dt}_{\text{Kernel fn., response fn.}} + b(s) = g(s)$$



↑ fluctuation is Poisson like

↔ due to finite resolution there is migration between bins (smearing)

↓ due to efficiency, some entries are missing

↙ due to non-linear detector response fn. there is a shift, on average in a certain direction

Structure of ML

Input $x \in X$ customer info

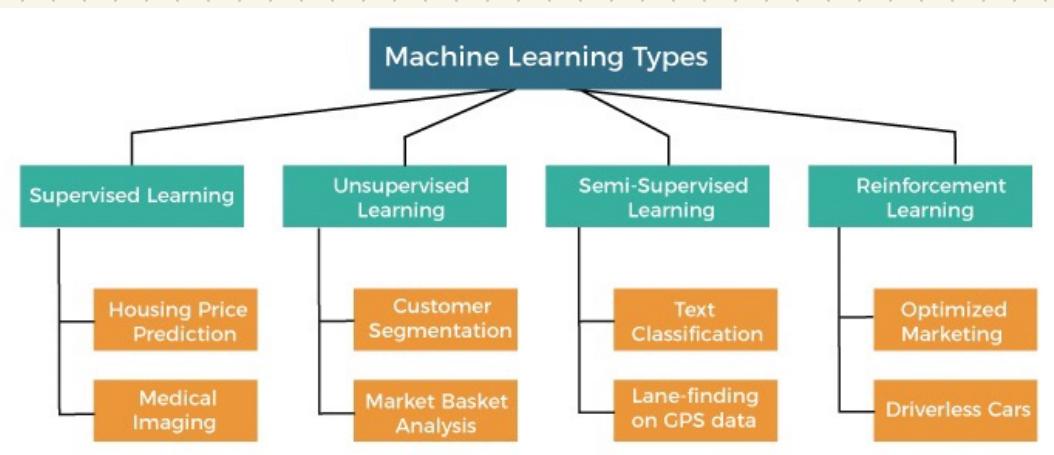
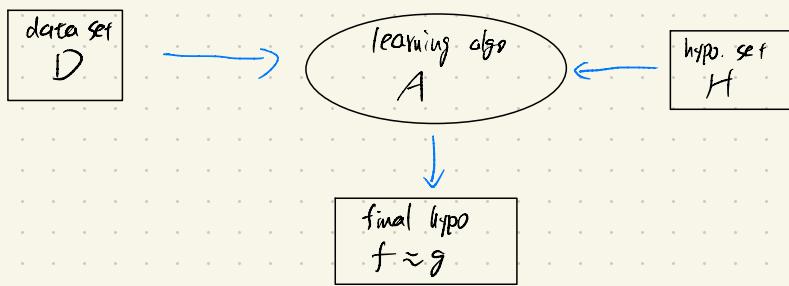
Output $y \in Y$ good or bad credit

Target fn $f: X \rightarrow Y$ ideal credit card approval formula

Hypothesis $g: X \rightarrow Y$ best credit card approval formula

Hypothesis set $G = \{ \text{all possible candidates for } g \}$

Data $D = \{(x_i, y_i) \forall i\}$

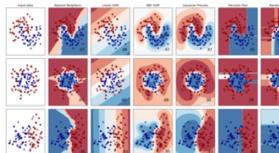


Classification

Identifying which category an object belongs to.

Applications: Spam detection, image recognition.

Algorithms: SVM, nearest neighbors, random forest, and more...

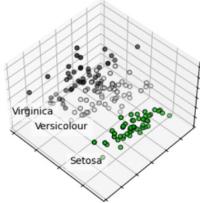


Dimensionality reduction

Reducing the number of random variables to consider.

Applications: Visualization, Increased efficiency

Algorithms: PCA, feature selection, non-negative matrix factorization, and more...

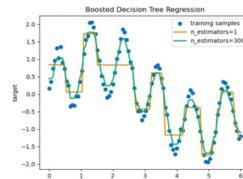


Regression

Predicting a continuous-valued attribute associated with an object.

Applications: Drug response, Stock prices.

Algorithms: SVR, nearest neighbors, random forest, and more...

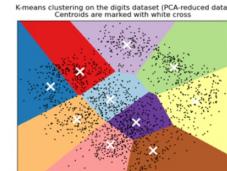


Clustering

Automatic grouping of similar objects into sets.

Applications: Customer segmentation, Grouping experiment outcomes

Algorithms: k-Means, spectral clustering, mean-shift, and more...

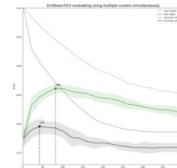


Model selection

Comparing, validating and choosing parameters and models.

Applications: Improved accuracy via parameter tuning

Algorithms: grid search, cross validation, metrics, and more...

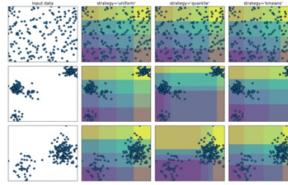


Preprocessing

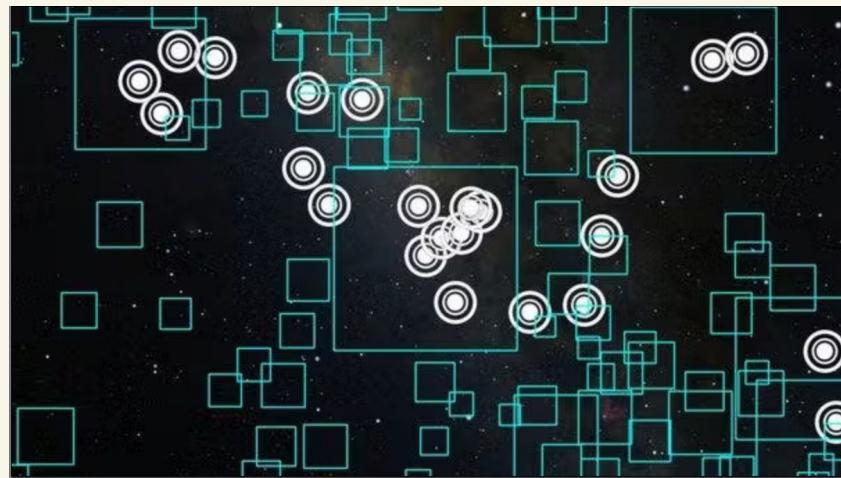
Feature extraction and normalization.

Applications: Transforming input data such as text for use with machine learning algorithms.

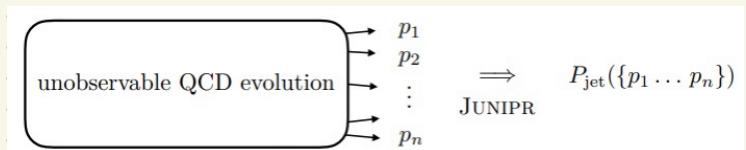
Algorithms: preprocessing, feature extraction, and more...



ML / DL in Physics

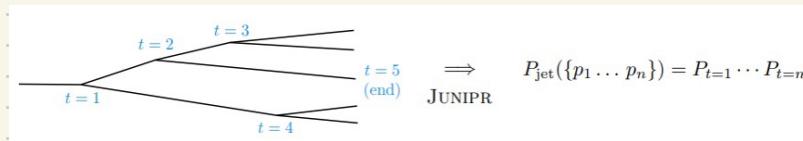


astrophysics
computer vision
to identify
stars



particle physics

neutral network
to classify high
energy particle jets



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Quantum many body

$|\Psi\rangle$ many body quantum state

neural network quantum state
to find approx. wave fn.

$$\langle s_1 \dots s_N | \Psi, w \rangle = F(s_1 \dots s_N, w)$$

w is the parameters to optimize

$$\text{energy } E(w) = \langle \Psi, w | \hat{H} | \Psi, w \rangle \text{ is minimized}$$

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