

Lecture Notes #6  
(Chapter 8)  
Economics 120B  
Econometrics

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# Nonlinear Regression Functions

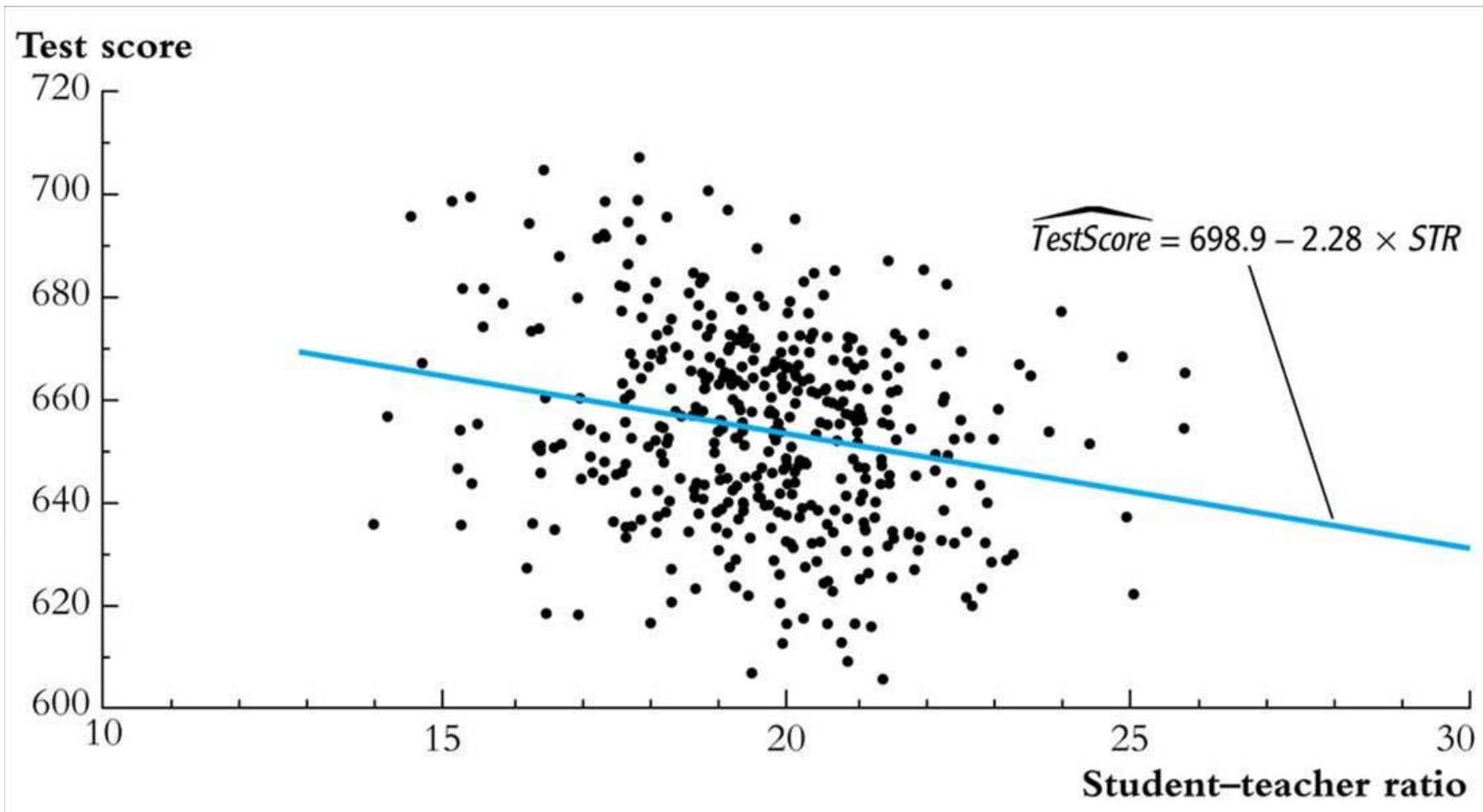
## (SW Chapter 8)

- Everything so far has been linear in the  $X$ 's
- But the linear approximation is not always a good one
- The multiple regression framework can be extended to handle regression functions that are nonlinear in one or more  $X$ .

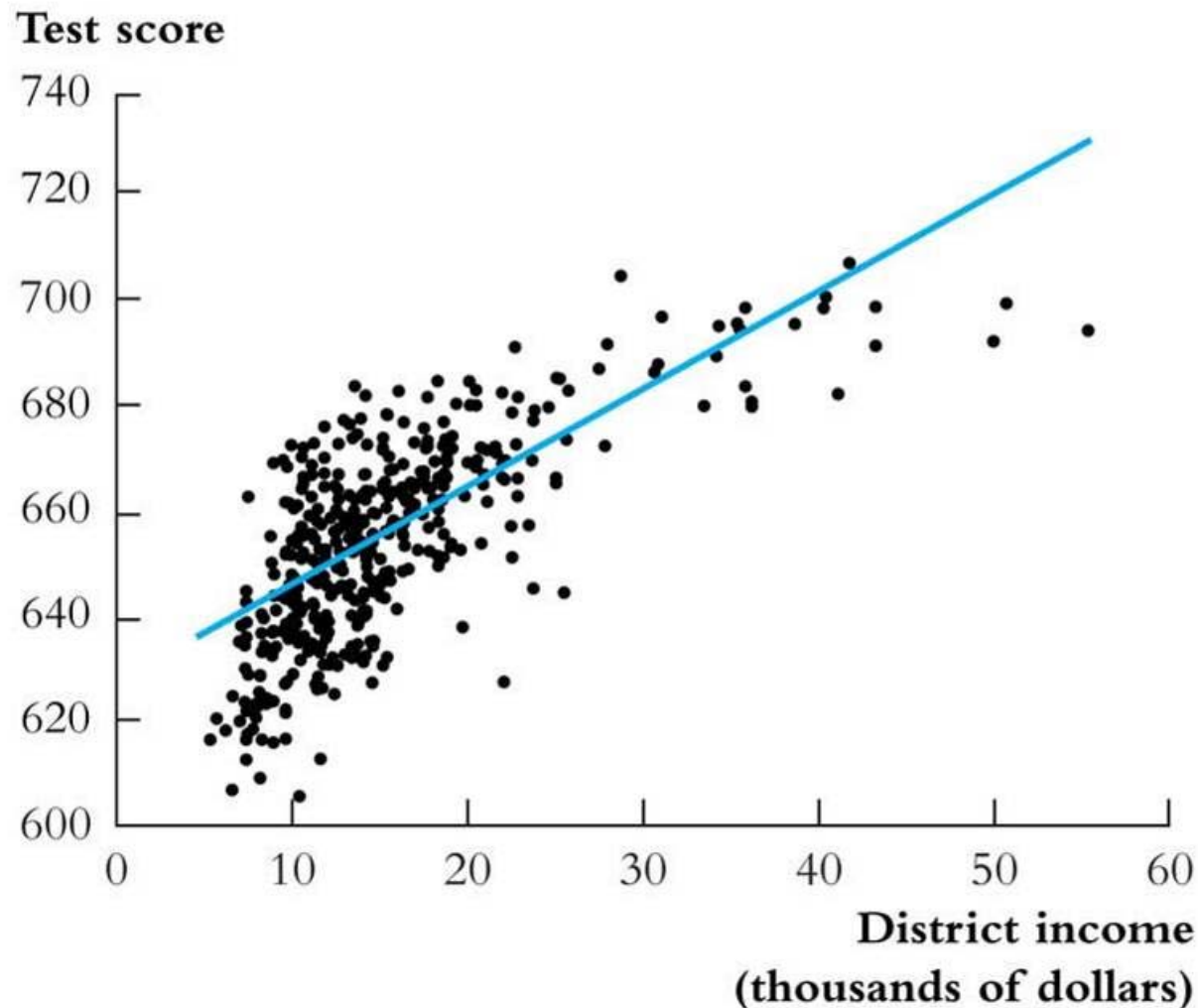
### Outline

1. Nonlinear regression functions – general comments
2. Nonlinear functions of one variable
3. Nonlinear functions of two variables: interactions

# The *TestScore* – *STR* relation looks linear (maybe)...



But the *TestScore* – *Income* relation looks nonlinear...



# Nonlinear Regression Population Regression Functions – General Ideas (SW Section 8.1)

If a relation between  $Y$  and  $X$  is **nonlinear**:

- The effect on  $Y$  of a change in  $X$  depends on the value of  $X$  – that is, the marginal effect of  $X$  is not constant
- A linear regression is mis-specified – the functional form is wrong
- The estimator of the effect on  $Y$  of  $X$  is biased – it needn't even be right on average.
- The solution to this is to estimate a regression function that is nonlinear in  $X$

# Nonlinear Functions of a Single Independent Variable (SW Section 8.2)

We'll look at two complementary approaches:

## 1. Polynomials in $X$

The population regression function is approximated by a quadratic, cubic, or higher-degree polynomial

## 2. Logarithmic transformations

- $Y$  and/or  $X$  is transformed by taking its logarithm
- this gives a “percentages” interpretation that makes sense in many applications

# 1. Polynomials in $X$

Approximate the population regression function by a polynomial:

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \dots + \beta_r X_i^r + u_i$$

- This is just the linear multiple regression model – except that the regressors are powers of  $X$ !
- Estimation, hypothesis testing, etc. proceeds as in the multiple regression model using OLS
- The coefficients are difficult to interpret, but the regression function itself is interpretable

# ***Example: the TestScore – Income relation***

$Income_i$  = average district income in the  $i^{th}$  district  
(thousands of dollars per capita)

Quadratic specification:

$$TestScore_i = \beta_0 + \beta_1 Income_i + \beta_2 (Income_i)^2 + u_i$$

Cubic specification:

$$TestScore_i = \beta_0 + \beta_1 Income_i + \beta_2 (Income_i)^2 + \beta_3 (Income_i)^3 + u_i$$



# Estimation of the quadratic specification in STATA

```
generate avginc2 = avginc*avginc;  
reg testscr avginc avginc2, r;
```

Create a new regressor

Regression with robust standard errors

```
Number of obs =      420  
F(   2,   417) =   428.52  
Prob > F       =    0.0000  
R-squared      =    0.5562  
Root MSE     =   12.724
```

testscr	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
avginc	3.850995	.2680941	14.36	0.000	3.32401	4.377979
avginc2	-.0423085	.0047803	-8.85	0.000	-.051705	-.0329119
_cons	607.3017	2.901754	209.29	0.000	601.5978	613.0056

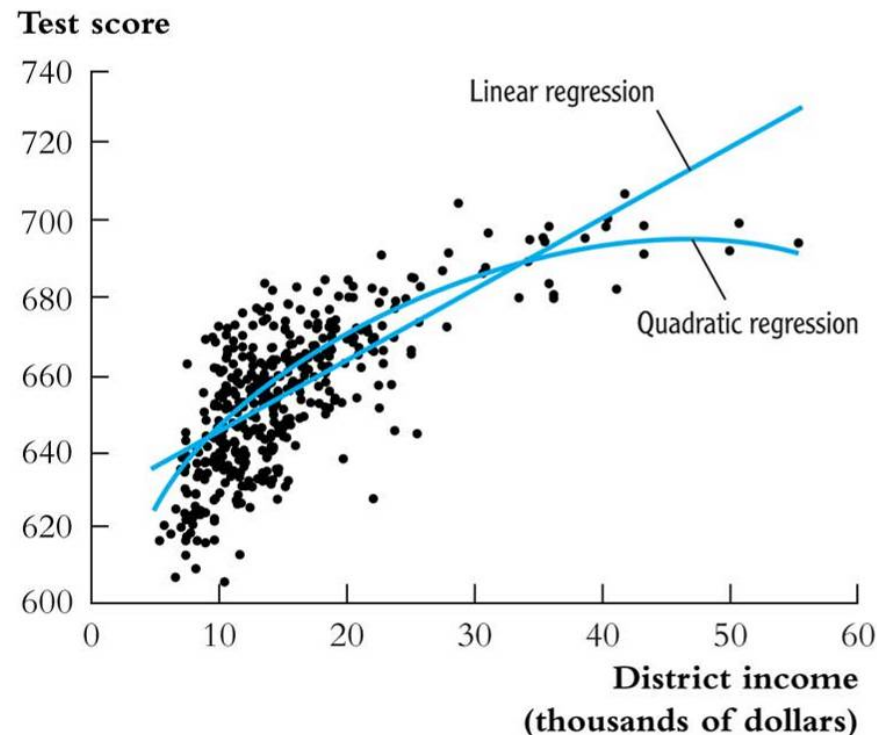
Test the null hypothesis of linearity against the alternative that the regression function is a quadratic....

# Interpreting the estimated regression function:

(a) Plot the predicted values

$$\widehat{TestScore} = 607.3 + 3.85 * Income_i - 0.0423(Income_i)^2$$

(2.9)      (0.27)                      (0.0048)



## ***Interpreting the estimated regression function, ctd:***

(b) Compute “effects” for different values of  $X$

$$\widehat{TestScore} = 607.3 + 3.85 * Income_i - 0.0423(Income_i)^2$$

(2.9)      (0.27)                      (0.0048)

Predicted change in  $\widehat{TestScore}$  for a change in income from \$5,000 per capita to \$6,000 per capita:

$$\begin{aligned}\Delta \widehat{TestScore} &= 607.3 + 3.85 * 6 - 0.0423 * 6^2 \\ &\quad - (607.3 + 3.85 * 5 - 0.0423 * 5^2) \\ &= 3.4\end{aligned}$$

# Estimation of a cubic specification in STATA

```
gen avginc3 = avginc*avginc2;  
reg testscr avginc avginc2 avginc3, r;
```

Create the cubic regressor

Regression with robust standard errors

Number of obs = 420  
F( 3, 416) = 270.18  
Prob > F = 0.0000  
R-squared = 0.5584  
Root MSE = 12.707

	Robust					
testscr	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
avginc	5.018677	.7073505	7.10	0.000	3.628251	6.409104
avginc2	-.0958052	.0289537	-3.31	0.001	-.1527191	-.0388913
avginc3	.0006855	.0003471	1.98	0.049	3.27e-06	.0013677
_cons	600.079	5.102062	117.61	0.000	590.0499	610.108

Testing the null hypothesis of linearity, against the alternative that the population regression is quadratic and/or cubic, that is, it is a polynomial of degree up to 3:

$H_0$ : pop'n coefficients on  $Income^2$  and  $Income^3 = 0$

$H_1$ : at least one of these coefficients is nonzero.

```
test avginc2 avginc3; Execute the test command after running the regression
```

```
( 1)  avginc2 = 0.0
```

```
( 2)  avginc3 = 0.0
```

```
      F( 2, 416) = 37.69
```

```
      Prob > F = 0.0000
```

The hypothesis that the population regression is linear is rejected at the 1% significance level against the alternative that it is a polynomial of degree up to 3.

# Summary: polynomial regression functions

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + \dots + \beta_r X_i^r + u_i$$

- Estimation: by OLS after defining new regressors
- Coefficients have complicated interpretations
- To interpret the estimated regression function:
  - plot predicted values as a function of  $x$
  - compute predicted  $\Delta Y / \Delta X$  at different values of  $x$
- Hypotheses concerning degree  $r$  can be tested by  $t$ - and  $F$ -tests on the appropriate (blocks of) variable(s).
- Choice of degree  $r$ 
  - plot the data;  $t$ - and  $F$ -tests, check sensitivity of estimated effects; judgment.
  - *Or use model selection criteria (later)*

## 2. Logarithmic functions of $Y$ and/or $X$

- $\ln(X)$  = the natural logarithm of  $X$
- Logarithmic transforms permit modeling relations in “percentage” terms (like elasticities), rather than linearly.

*Here's why:*  $\ln(x+\Delta x) - \ln(x) = \ln\left(1 + \frac{\Delta x}{x}\right) \cong \frac{\Delta x}{x}$

(calculus:  $\frac{d \ln(x)}{dx} = \frac{1}{x}$ )

*Numerically:*

$$\ln(1.01) = .00995 \cong .01;$$

$$\ln(1.10) = .0953 \cong .10 \text{ (sort of)}$$

# ***The three log regression specifications:***

Case	Population regression function
I. linear-log	$Y_i = \beta_0 + \beta_1 \ln(X_i) + u_i$
II. log-linear	$\ln(Y_i) = \beta_0 + \beta_1 X_i + u_i$
III. log-log	$\ln(Y_i) = \beta_0 + \beta_1 \ln(X_i) + u_i$

- The interpretation of the slope coefficient differs in each case.
- The interpretation is found by applying the general “before and after” rule: “figure out the change in  $Y$  for a given change in  $X$ .”



# ***I. Linear-log case***

$$Y_i = \beta_0 + \beta_1 \ln(X_i) + u_i$$

for small  $\Delta X$ ,

$$\beta_1 \cong \frac{\Delta Y}{\Delta X / X}$$

Now  $100 \times \frac{\Delta X}{X}$  = percentage change in  $X$ , so ***a 1% increase in  $X$***

***(multiplying  $X$  by 1.01) is associated with a  $.01\beta_1$  change in  $Y$ .***

(1% increase in  $X \Rightarrow .01$  increase in  $\ln(X)$ )

$\Rightarrow .01\beta_1$  increase in  $Y$ )

## ***Example: TestScore vs. ln(Income)***

- First defining the new regressor,  $\ln(\text{Income})$
- The model is now linear in  $\ln(\text{Income})$ , so the linear-log model can be estimated by OLS:

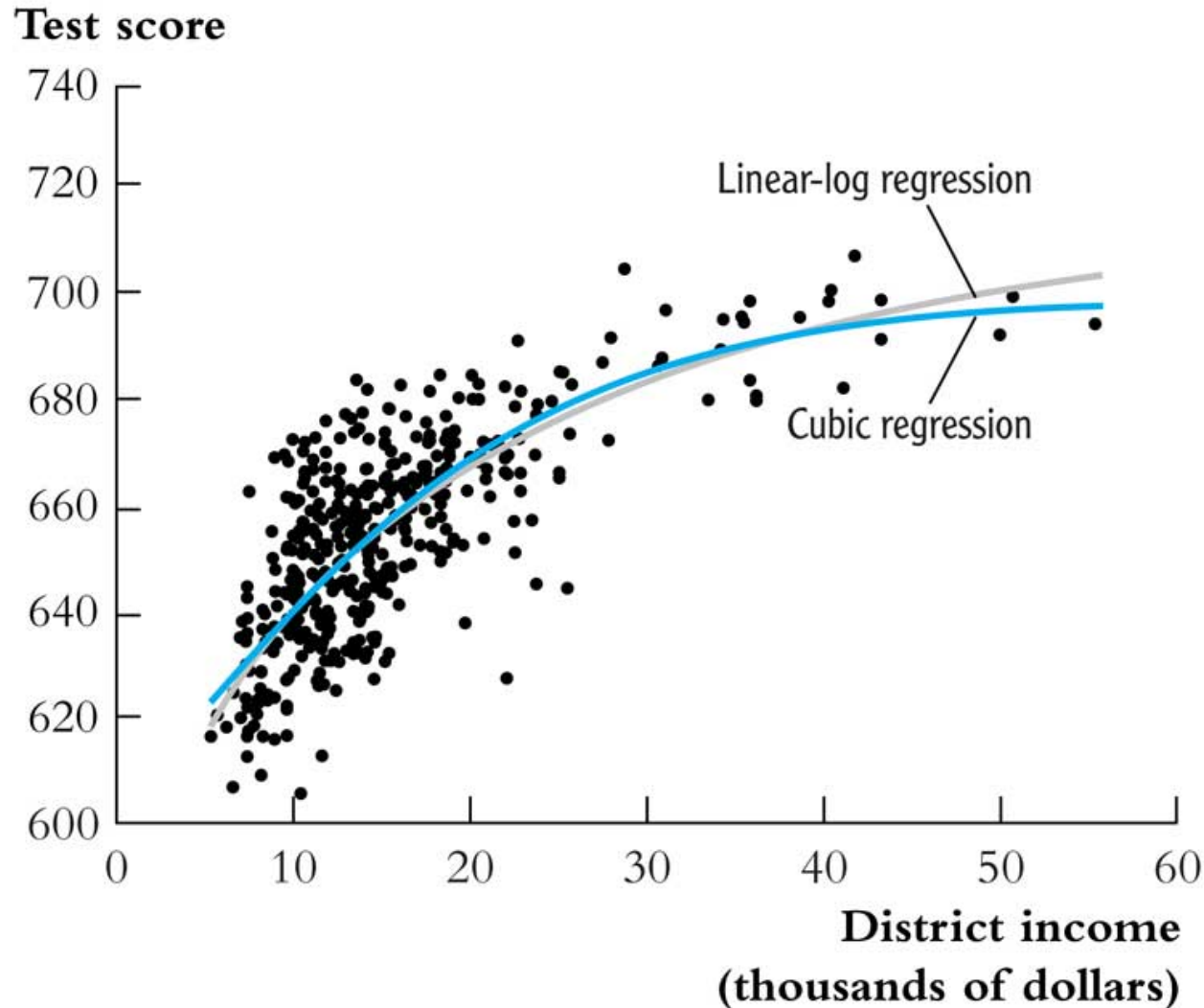
$$\widehat{\text{TestScore}} = 557.8 + 36.42 * \ln(\text{Income}_i)$$

(3.8)      (1.40)

so a 1% increase in *Income* is associated with an increase in *TestScore* of 0.36 points on the test.

- Standard errors, confidence intervals,  $R^2$  – all the usual tools of regression apply here.
- How does this compare to the cubic model?

# The linear-log and cubic regression functions



## II. Log-linear case

$$\ln(Y_i) = \beta_0 + \beta_1 X_i + u_i$$

for small  $\Delta X$ ,  $\beta_1 \cong \frac{\Delta Y / Y}{\Delta X}$

- Now  $100 \times \frac{\Delta Y}{Y}$  = percentage change in  $Y$ , so *a change in  $X$  by one unit ( $\Delta X = 1$ ) is associated with a  $100\beta_1\%$  change in  $Y$ .*
- 1 unit increase in  $X \Rightarrow \beta_1$  increase in  $\ln(Y)$   
 $\Rightarrow 100\beta_1\%$  increase in  $Y$

### III. Log-log case

$$\ln(Y_i) = \beta_0 + \beta_1 \ln(X_i) + u_i$$

for small  $\Delta X$ ,

$$\beta_1 \cong \frac{\Delta Y / Y}{\Delta X / X}$$

Now  $100 \times \frac{\Delta Y}{Y}$  = percentage change in  $Y$ , and  $100 \times \frac{\Delta X}{X}$  =

percentage change in  $X$ , so *a 1% change in  $X$  is associated with a  $\beta_1\%$  change in  $Y$ .*

- *In the log-log specification,  $\beta_1$  has the interpretation of an elasticity.*

# Summary: Logarithmic transformations

- Three cases, differing in whether  $Y$  and/or  $X$  is transformed by taking logarithms.
- The regression is linear in the new variable(s)  $\ln(Y)$  and/or  $\ln(X)$ , and the coefficients can be estimated by OLS.
- Hypothesis tests and confidence intervals are now implemented and interpreted “as usual.”
- The interpretation of  $\beta_1$  differs from case to case.
- Choice of specification should be guided by judgment (which interpretation makes the most sense in your application?), tests, and plotting predicted values

# Interactions Between Independent Variables (SW Section 8.3)

- Perhaps a class size reduction is more effective in some circumstances than in others...
- Perhaps smaller classes help more if there are many English learners, who need individual attention
- That is,  $\frac{\Delta TestScore}{\Delta STR}$  might depend on  $PctEL$
- More generally,  $\frac{\Delta Y}{\Delta X_1}$  might depend on  $X_2$
- How to model such “interactions” between  $X_1$  and  $X_2$ ?
- We first consider binary  $X$ ’s, then continuous  $X$ ’s

## (a) Interactions between two binary variables

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + u_i$$

- $D_{1i}$ ,  $D_{2i}$  are binary
- $\beta_1$  is the effect of changing  $D_1=0$  to  $D_1=1$ . In this specification, *this effect doesn't depend on the value of  $D_2$ .*
- To allow the effect of changing  $D_1$  to depend on  $D_2$ , include the “interaction term”  $D_{1i} \times D_{2i}$  as a regressor:

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 (D_{1i} \times D_{2i}) + u_i$$



# Interpreting the coefficients

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 (D_{1i} \times D_{2i}) + u_i$$

General rule: compare the various cases

$$E(Y_i | D_{1i}=0, D_{2i}=d_2) = \beta_0 + \beta_2 d_2 \quad (b)$$

$$E(Y_i | D_{1i}=1, D_{2i}=d_2) = \beta_0 + \beta_1 + \beta_2 d_2 + \beta_3 d_2 \quad (a)$$

subtract (a) – (b):

$$E(Y_i | D_{1i}=1, D_{2i}=d_2) - E(Y_i | D_{1i}=0, D_{2i}=d_2) = \beta_1 + \beta_3 d_2$$

- The effect of  $D_1$  depends on  $d_2$  (what we wanted)
- $\beta_3$  = increment to the effect of  $D_1$ , when  $D_2 = 1$

# Example: TestScore, STR, English learners

Let

$$HiSTR = \begin{cases} 1 & \text{if } STR \geq 20 \\ 0 & \text{if } STR < 20 \end{cases} \quad \text{and} \quad HiEL = \begin{cases} 1 & \text{if } PctEL \geq 10 \\ 0 & \text{if } PctEL < 10 \end{cases}$$

$$\boxed{TestScore} = 664.1 - 18.2HiEL - 1.9HiSTR - 3.5(HiSTR \times HiEL)$$

(1.4)   (2.3)            (1.9)            (3.1)

- “Effect” of  $HiSTR$  when  $HiEL = 0$  is  $-1.9$
- “Effect” of  $HiSTR$  when  $HiEL = 1$  is  $-1.9 - 3.5 = -5.4$
- Class size reduction is estimated to have a bigger effect when the percent of English learners is large
- This interaction isn’t statistically significant:  $t = 3.5/3.1$

## (b) Interactions between continuous and binary variables

$$Y_i = \beta_0 + \beta_1 D_i + \beta_2 X_i + u_i$$

- $D_i$  is binary,  $X$  is continuous
- As specified above, the effect on  $Y$  of  $X$  (holding constant  $D$ ) =  $\beta_2$ , which does not depend on  $D$
- To allow the effect of  $X$  to depend on  $D$ , include the “interaction term”  $D_i \times X_i$  as a regressor:

$$Y_i = \beta_0 + \beta_1 D_i + \beta_2 X_i + \beta_3 (D_i \times X_i) + u_i$$

# ***Binary-continuous interactions: the two regression lines***

$$Y_i = \beta_0 + \beta_1 D_i + \beta_2 X_i + \beta_3 (D_i \times X_i) + u_i$$

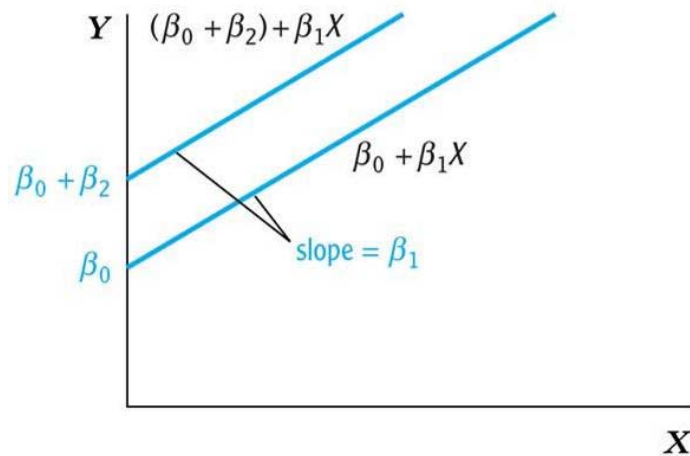
Observations with  $D_i = 0$  (the “ $D = 0$ ” group):

$$Y_i = \beta_0 + \beta_2 X_i + u_i \quad \textit{The } D=0 \text{ regression line}$$

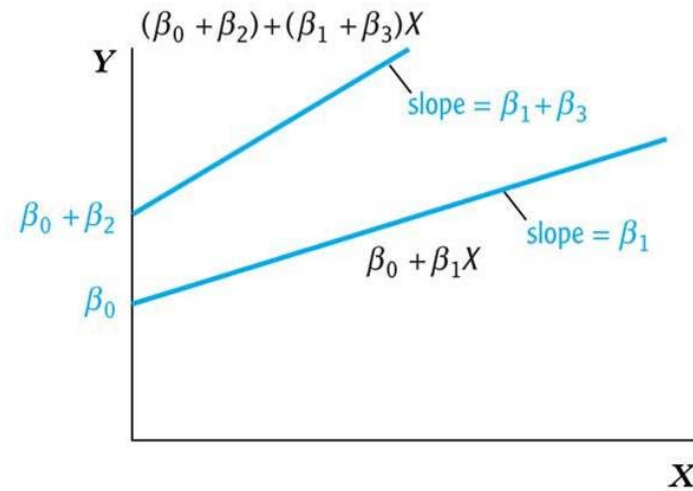
Observations with  $D_i = 1$  (the “ $D = 1$ ” group):

$$\begin{aligned} Y_i &= \beta_0 + \beta_1 + \beta_2 X_i + \beta_3 X_i + u_i \\ &= (\beta_0 + \beta_1) + (\beta_2 + \beta_3) X_i + u_i \quad \textit{The } D=1 \text{ regression line} \end{aligned}$$

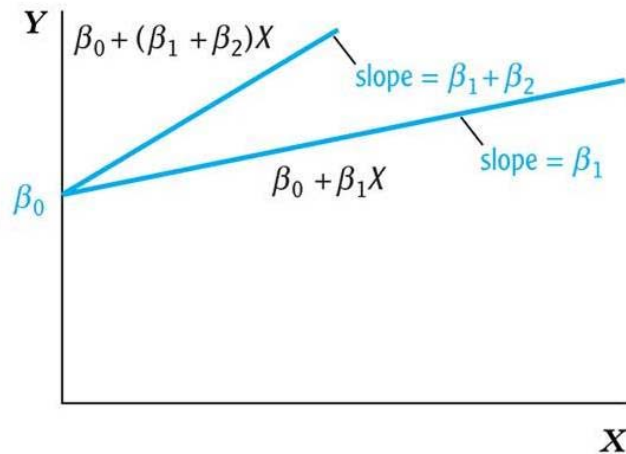
# Binary-continuous interactions, ctd.



(a) Different intercepts, same slope



(b) Different intercepts, different slopes



(c) Same intercept, different slopes

# Interpreting the coefficients

$$Y_i = \beta_0 + \beta_1 D_i + \beta_2 X_i + \beta_3 (D_i \times X_i) + u_i$$

General rule: compare the various cases

$$Y = \beta_0 + \beta_1 D + \beta_2 X + \beta_3 (D \times X) \quad (b)$$

Now change  $X$ :

$$Y + \Delta Y = \beta_0 + \beta_1 D + \beta_2 (X + \Delta X) + \beta_3 [D \times (X + \Delta X)] \quad (a)$$

subtract (a) – (b):

$$\Delta Y = \beta_2 \Delta X + \beta_3 D \Delta X \quad \text{or} \quad \frac{\Delta Y}{\Delta X} = \beta_2 + \beta_3 D$$

- The effect of  $X$  depends on  $D$  (what we wanted)
- $\beta_3$  = increment to the effect of  $X$ , when  $D = 1$

## ***Example: TestScore, STR, HiEL*** ***(=1 if PctEL > 10, = 0 otherwise)***

$$\widehat{TestScore} = 682.2 - 0.97 * STR + 5.6 * HiEL - 1.28(STR * HiEL)$$

(11.9)    (0.59)                    (19.5)                    (0.97)

- When  $HiEL = 0$ :

$$\widehat{TestScore} = 682.2 - 0.97 * STR$$

- When  $HiEL = 1$ ,

$$\begin{aligned}\widehat{TestScore} &= 682.2 - 0.97 * STR + 5.6 - 1.28 * STR \\ &= 687.8 - 2.25 * STR\end{aligned}$$

- Two regression lines: one for each  $HiSTR$  group.
- Class size reduction is estimated to have a larger effect when the percent of English learners is large.

## Example, ctd: Testing hypotheses

$$\widehat{TestScore} = 682.2 - 0.97 * STR + 5.6 * HiEL - 1.28(STR * HiEL)$$

(11.9)    (0.59)                    (19.5)                    (0.97)

- The two regression lines have the same **slope**  $\Leftrightarrow$  the coefficient on  $STR \times HiEL$  is zero:  $t = -1.28/0.97 = -1.32$
- The two regression lines have the same **intercept**  $\Leftrightarrow$  the coefficient on  $HiEL$  is zero:  $t = -5.6/19.5 = 0.29$
- The two regression **lines** are the same  $\Leftrightarrow$  population coefficient on  $HiEL = 0$  **and** population coefficient on  $STR \times HiEL = 0$ :  $F = 89.94$  ( $p$ -value  $< .001$ ) !!
- We reject the joint hypothesis but neither individual hypothesis (*how can this be?*)



## (c) Interactions between two continuous variables

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + u_i$$

- $X_1, X_2$  are continuous
- As specified, the effect of  $X_1$  doesn't depend on  $X_2$
- As specified, the effect of  $X_2$  doesn't depend on  $X_1$
- To allow the effect of  $X_1$  to depend on  $X_2$ , include the “interaction term”  $X_{1i} \times X_{2i}$  as a regressor:

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} \times X_{2i}) + u_i$$

# *Interpreting the coefficients:*

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} \times X_{2i}) + u_i$$

General rule: compare the various cases

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 (X_1 \times X_2) \quad (b)$$

Now change  $X_1$ :

$$Y + \Delta Y = \beta_0 + \beta_1 (X_1 + \Delta X_1) + \beta_2 X_2 + \beta_3 [(X_1 + \Delta X_1) \times X_2] \quad (a)$$

subtract (a) – (b):

$$\Delta Y = \beta_1 \Delta X_1 + \beta_3 X_2 \Delta X_1 \quad \text{or} \quad \frac{\Delta Y}{\Delta X_1} = \beta_1 + \beta_3 X_2$$

- The effect of  $X_1$  depends on  $X_2$  (what we wanted)
- $\beta_3$  = increment to the effect of  $X_1$  from a unit change in  $X_2$

# Application: Nonlinear Effects on Test Scores of the Student-Teacher Ratio (SW Section 8.4)

Nonlinear specifications let us examine more nuanced questions about the Test score – *STR* relation, such as:

1. Are there nonlinear effects of class size reduction on test scores? (Does a reduction from 35 to 30 have same effect as a reduction from 20 to 15?)
2. Are there nonlinear interactions between *PctEL* and *STR*? (Are small classes more effective when there are many English learners?)

# Strategy for Question #1 (different effects for different *STR*?)

- Estimate linear and nonlinear functions of *STR*, holding constant relevant demographic variables
  - *PctEL*
  - *Income* (remember the nonlinear *TestScore-Income* relation!)
  - *LunchPCT* (fraction on free/subsidized lunch)
- See whether adding the nonlinear terms makes an “economically important” quantitative difference (“economic” or “real-world” importance is different than statistically significant)
- Test for whether the nonlinear terms are significant

## Strategy for Question #2 (interactions between *PctEL* and *STR*?)

- Estimate linear and nonlinear functions of *STR*, interacted with *PctEL*.
- If the specification is nonlinear (with *STR*,  $STR^2$ ,  $STR^3$ ), then you need to add interactions with all the terms so that the entire functional form can be different, depending on the level of *PctEL*.
- We will use a binary-continuous interaction specification by adding  $HiEL \times STR$ ,  $HiEL \times STR^2$ , and  $HiEL \times STR^3$ .

**TABLE 8.3** Nonlinear Regression Models of Test Scores

Dependent variable: average test score in district; 420 observations.

Regressor	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Student–teacher ratio ( <i>STR</i> )	−1.00** (0.27)	−0.73** (0.26)	−0.97 (0.59)	−0.53 (0.34)	64.33** (24.86)	83.70** (28.50)	65.29** (25.26)
<i>STR</i> <sup>2</sup>					−3.42** (1.25)	−4.38** (1.44)	−3.47** (1.27)
<i>STR</i> <sup>3</sup>					0.059** (0.021)	0.075** (0.024)	0.060** (0.021)
% English learners	−0.122** (0.033)	−0.176** (0.034)					−0.166** (0.034)
% English learners ≥ 10%? (Binary, <i>HiEL</i> )			5.64 (19.51)	5.50 (9.80)	−5.47** (1.03)	816.1* (327.7)	
<i>HiEL</i> × <i>STR</i>			−1.28 (0.97)	−0.58 (0.50)		−123.3* (50.2)	
<i>HiEL</i> × <i>STR</i> <sup>2</sup>						6.12* (2.54)	
<i>HiEL</i> × <i>STR</i> <sup>3</sup>						−0.101* (0.043)	
% Eligible for subsidized lunch	−0.547** (0.024)	−0.398** (0.033)		−0.411** (0.029)	−0.420** (0.029)	−0.418** (0.029)	−0.402** (0.033)
Average district income (logarithm)		11.57** (1.81)		12.12** (1.80)	11.75** (1.78)	11.80** (1.78)	11.51** (1.81)
Intercept	700.2** (5.6)	658.6** (8.6)	682.2** (11.9)	653.6** (9.9)	252.0 (163.6)	122.3 (185.5)	244.8 (165.7)

# Tests of joint hypotheses:

F-Statistics and p-Values on Joint Hypotheses							
(a) All <i>STR</i> variables and interactions = 0	5.64 (0.004)	5.92 (0.003)	6.31 ( $< 0.001$ )	4.96 ( $< 0.001$ )	5.91 (0.001)		
(b) $STR^2, STR^3 = 0$			6.17 ( $< 0.001$ )	5.81 (0.003)	5.96 (0.003)		
(c) $HiEL \times STR, HiEL \times STR^2, HiEL \times STR^3 = 0$				2.69 (0.046)			
<i>SER</i>	9.08	8.64	15.88	8.63	8.56	8.55	8.57
$\overline{R}^2$	0.773	0.794	0.305	0.795	0.798	0.799	0.798

These regressions were estimated using the data on K-8 school districts in California, described in Appendix 4.1. Standard errors are given in parentheses under coefficients, and *p*-values are given in parentheses under *F*-statistics. Individual coefficients are statistically significant at the \*5% or \*\*1% significance level.

*What can you conclude about question #1?*  
*About question #2?*

# Summary: Nonlinear Regression Functions

- Using functions of the independent variables such as  $\ln(X)$  or  $X_1 \times X_2$ , allows recasting a large family of nonlinear regression functions as multiple regression.
- Estimation and inference proceed in the same way as in the linear multiple regression model.
- Interpretation of the coefficients is model-specific, but the general rule is to compute effects by comparing different cases (different value of the original  $X$ 's)
- Many nonlinear specifications are possible, so you must use judgment:
  - What nonlinear effect you want to analyze?
  - What makes sense in your application?