

1 Data is generated from the exponential distribution with density

$$f(y) = \lambda \exp(-\lambda y), \text{ where } \lambda, y > 0.$$

- (a) Show that it belongs to the exponential family distributions by identifying θ , $b(\theta)$, ϕ , $a(\phi)$ and $c(y; \phi)$.

An exponential family distribution can be written in the form

$$\exp \left\{ \frac{y\theta - b(\theta)}{a(\phi)} + c(y; \phi) \right\}.$$

We write:

$$f(y|\lambda) = \exp(-\lambda y + \log \lambda)$$

and equate $\theta = -\lambda$, $b(\theta) = -\log \lambda = -\log(-\theta)$, (note that $\lambda > 0$, so $\theta < 0$), $\phi = 1$, $a(\phi) = 1$, $c(y; \phi) = 1$.

- (b) What is the canonical link and variance functions for a GLM with the response following the exponential distribution?

The link function connects the linear predictor μ to the parameter θ in the exponential family distribution definition above. To find the canonical link, we want $\mu = E(Y) = b'(\theta)$. we find the first moment of y :

$$EY = \int_0^\infty \lambda y e^{-\lambda y} dy = \frac{1}{\lambda}$$

$$\text{Note } b'(\theta) = \frac{-1}{\theta} = \frac{1}{\lambda} = \mu \text{ and } b''(\theta) = \frac{1}{\theta^2} = \frac{1}{\lambda^2} = \mu^2$$

We write θ as a function of μ

$$\theta(\mu) = -\frac{1}{\mu}$$

$$b'^{-1}(\cdot) = \text{negative inverse function.}$$

Since $\text{var}(Y) = b''(\theta)a(\phi)$, and $a(\phi) = 1$, $\text{var}(Y) = b''(\theta) = \mu^2$.

- (c) Is there any practical difficulty for using the canonical link in practice?

Especially in small samples, canonical links have desirable properties. However, they may not be the best fit for a model (McCullagh and Nelder pg 32).

Note in this case that the exponential mean is restricted to positive values. However our μ is a linear combination of predictors. This does not guarantee a positive restriction on our estimates of the mean.

- (d) Express the deviance as a function of y_i and fitted mean μ_i ($i = 1, \dots, n$).

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We have scaled deviance given by

$$\frac{D(y; \hat{\mu})}{\phi} = 2 \sum \frac{w_i}{\phi} \{y_i(\tilde{\theta}_i - \hat{\theta}_i) - b(\tilde{\theta}_i) + b(\hat{\theta}_i)\}$$

with $a(\phi) = \phi/w$, $\tilde{\theta} = \theta(y)$ denoting the full model (n parameter) estimate of θ , and $\hat{\theta} = \theta(\hat{\mu})$ denoting the null model (one parameter) estimate of θ .

Evaluating for $b(\theta) = -\log(-\theta)$ and $\phi = 1$, $w_i = 1$ gives

$$D(y; \hat{\mu}) = 2 \sum \left(y_i(\tilde{\theta}_i - \hat{\theta}_i) + \log\left(\frac{\tilde{\theta}_i}{\hat{\theta}_i}\right) \right)$$

From above, we have that $\theta(\mu) = -\frac{1}{\mu}$. Evaluating for $\hat{\theta}_i = 1/\hat{\mu}_i$ and $\tilde{\theta}_i = y_i$ gives

$$D(y; \hat{\mu}) = 2 \sum \left(y_i(y_i - \hat{\mu}_i) + \log\left(\frac{\hat{\mu}_i}{y_i}\right) \right)$$