Probability

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \qquad P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

If A and B are independent $P(A \cap B) = P(A)P(B)$

Expected Value

$$E[x] = \int_{-\infty}^{\infty} xp(x) dx \qquad E[g(x)] = \int_{-\infty}^{\infty} g(x)p(x) dx$$

$$E[a] = a; a \text{ is a constant} \qquad E[aX + b] = aE[x] + b$$

$$E[X + Y] = E[X] + E[y]$$

Variance

$$Var[x] = E[(x - E[x])^2] = \sigma^2 = \int_{-\infty}^{\infty} (x - E[x])^2 p(x) dx$$

$$E((x - E[x])^{2}) = E[x^{2}] - (E[x])^{2}$$

$$Var[a] = 0$$
; a is a constant $Var[aX + b] = a^2 Var[X]$

Covariance

$$cov(X_1, X_2) = E[(X_1 - m_1)(X_2 - m_2)]$$

= $E[(X_1)(X_2)] - m_1 m_2$

Correlation

$$\begin{split} \rho &= \frac{cov(X_1, X_2)}{\sqrt{V(X_1)V(X_2)}} \\ r_{xy} &= \frac{\sum\limits_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum\limits_{i=1}^{n} (x_i - \bar{x})^2} \sqrt{\sum\limits_{i=1}^{n} (y_i - \bar{y})^2}} \end{split}$$

Algebra of Random Variable

$$Z = Y + X$$

$$P_Z(Z_0) = P_Y(y) * P_X(x)$$

Probability Distribution

Cumulative Distribution Function (CDF)

Normal

$$\frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} \qquad E[X] = \mu \qquad Var[X] = \sigma^2$$

Exponential

$$\lambda e^{-\lambda x}$$
 $E[X] = \frac{1}{\lambda}$ $Var[X] = \frac{1}{\lambda^2}$

Uniform

$$\begin{cases} \frac{1}{b-a} & \text{for } x \in [a,b] \\ 0 & \text{otherwise} \end{cases}$$

$$E[X] = \frac{1}{2}(a+b) \qquad Var[X] = \frac{1}{12}(b-a)^2$$

Bernoulli

Success (p), Fail (1-p)

$$\begin{cases} q = 1 - p & \text{if } k = 0 \\ p & \text{if } k = 1 \end{cases} \quad E[X] = p \quad Var[X] = p(1 - p)$$

Binomial

n Bernoulli trials

$$\binom{n}{k}p^k(1-p)^{n-k} \qquad E[X] = np \qquad Var[X] = np(1-p)$$

Poisson

$$\frac{\lambda^k e^{-\lambda}}{k!} \qquad E[X] = \lambda \qquad Var[X] = \lambda$$

Pareto

$$\frac{\alpha x_m^a}{x^{\alpha+1}} \qquad E[x] = \begin{cases} \infty & \text{for } \alpha \leq 1 \\ \frac{\alpha x_m}{\alpha-1} & \text{for } \alpha > 1 \end{cases}$$

$$Var[X] = \begin{cases} \infty & \text{for } \alpha \le 1\\ \frac{x_m^2 \alpha}{(\alpha - 1)^2 (\alpha - 2)} & \text{for } \alpha > 1 \end{cases}$$

MLE

To find the MLE given data

- 1. The likelihood function $P(data|\lambda)$, λ is the parameter
- 2. $\frac{d}{d\lambda}(\log \text{ likelihood}) = 0$, Find λ

LLN & CLT

LLN: As n grows, the probability that X_n is close to $\mu \to 1$. CLT: As n grows, the distribution of X_n converges to the normal distribution $N(\mu, \sigma^2/n)$.

Confidence Interval (Polling)

95% Confidence Interval $\bar{x} \pm \frac{1}{\sqrt{n}}$

Null Hypothesis Significance Testing

Errors

| | | True State of Nature | | | | |
|--------------|--------------|----------------------|------------------|--|--|--|
| | | H_0 | H_A | | | |
| Our Decision | Reject H_0 | Type-I Error | correct decision | | | |
| | Reject H_0 | correct decision | Type-II Error | | | |

P-value

We usually do testing by specifying significance level and do testing using p-values. If p-value is less than the significance level we reject H_0

P-value - Probability assuming Null of seeing data at least as extreme as the experiment data.

Problems with P-value

- **P-hacking** Do experiment multiple times until the results is what we want
- Base rate fallacy Low base rate →More chance for false positive
- Low power experiments If the test has low power (underpowered study), no significant is likely to be due to not enough samples t detect small differences.

Significance level and power

Significance level =
$$P(rejectH_0|H_0)$$

= probability we incorrectly reject H_0
= $P(type I error)$

Power = probability we correctly reject
$$H_0$$

= $P(rejectH_0|H_A)$
= $1 - P(type II error)$

One Sample z-test

Use when the **Variance** (σ^2) of the data is known

$$z = \frac{\bar{x} - \mu}{\sigma / \sqrt{n}}$$

One Sample t-test

$$t = \frac{\bar{x} - \mu}{s/\sqrt{n}}$$
 where $s^2 = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})^2$ $df = n-1$

Two Sample z-test

$$z = \frac{(\bar{x_1} - \bar{x_2}) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$$

Two Sample t-test with Equal Variance

Assume that the data have $\sigma_1 = \sigma_2$

$$t = \frac{\bar{x} - \bar{y}}{s_p} \qquad s_p^2 = \frac{(n-1)s_x^2 + (m-1)s_y^2}{n + m - 2} \left(\frac{1}{n} + \frac{1}{m}\right)$$

Two Sample t-test with Unequal Variance

Assume that the data have $\sigma_1 \neq \sigma_2$

$$t = \frac{\bar{x} - \bar{y} - \mu_0}{s_P} \qquad s_P^2 = \frac{s_x^2}{n} + \frac{s_y^2}{n}$$

$$df = \frac{(s_x^2/n + s_y^2/m)^2}{(s_x^2/n)^2/(n-1) + (s_y^2/m)^2/(m-1)}$$

Paired two-sample t-test

$$t = \frac{\bar{w} - \mu_0}{s/\sqrt{n}}$$
 $w_i = x_i - y_i$ $s^2 = \frac{1}{n-1} \sum_{i=1}^n (w_i - \bar{w})^2$

One-way ANOVA (F-test for equal means)

Test if the population means from n group are all the same Data for each group is an independent normal sample drawn from distributions with (possibly) different means but the same variance.

$$w=\frac{MS_B}{MS_w} \qquad \bar{x}_i=\text{ mean of group i} \qquad \bar{x}=\text{ mean of all data}$$

$$s_i^2=\frac{1}{m-1}\sum_{j=1}m(x_{i,j}-\bar{x}_i)^2$$

 $MS_B =$ between group variance $= m \times$ sample variance of group means $= \frac{m}{n-1} \sum_{i=1}^{n} (\bar{x}_i - \bar{x})^2$

 $MS_w = \text{average within group variance}$ $= m \times \text{sample means } s_1^2, \dots, s_n^2$ $= \frac{s_1^2 + s_2^2 + \dots + s_n^2}{n}$

A/B Testing

Steps

- 1. Define relevant metrics
- 2. Split samples into comparable groups
- 3. Choose statistical tests and validate their assumptions
- 4. Decide on stopping criteria
- 5. Run and monitor the experiment
- 6. Analyze results and suggest actions

Possible event probabilities

- 1. # checkout events / # hits double-count on page refreshes
- 2. # checkout events/# sessions double-count on inactive visits; good to see which products get bought within fewer visits
- 3. # checkout events / # cookies on product page "users" as denominator; includes both logins and non-logins; different browsers/devices double-counts
- 4. # payment events / # cookies on product page captures successful purchases
- 5. # payment events / # user ids non-logins count as one user id
- 6. # payment events / # people who are people?

Attribution Period

Conversion rate of August

conversions within August / number of users that visited in August

Conversion rate of August cohort

conversions within X days / number of users that visited in August

What Frequentist Hypothesis Tests Are NOT Saying

- 1. The p-value is not the probability that the null hypothesis is true, or the probability that the alternative hypothesis is false.
- 2. The p-value is not the probability that the observed effects were produced by random chance alone.
- 3. The 0.05 significance level is merely a convention.
- 4. The p-value does not indicate the size or importance of the observed effect.
- 5. The p-value is not the observed false positive rate; that depends on the prevalence of the data.

MDE

Sample Size

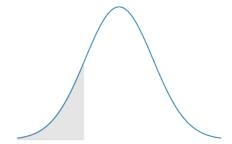
m is the split rate (50:50 = 1, 80:20 = 4)

$$n = \frac{m+1}{m} \left(\frac{(Z_{\alpha} + Z_{\beta})\sigma}{MDE} \right)^{2}$$

When not to do an A/B test

- Things that cannot be summarized into one or a few metrics are our users happy?
- Totally new things not apple-to-apple comparison; do a survey, focus group, design thinking
- **Delayed results** for example, does a change makes customer repurchase more? what if repurchase period is 3 months?
- One-off events no two big sales events such as 11.11 or Black Friday are the same
- Cannot split group independently offline store layout of the same store

Z-score table (Area to the left of the z score)



| _ | 1 | | | | | | | | | |
|-----------------|--------|--------|--------|--------|--------|--------|--------|--------|--------|--------|
| \underline{Z} | .00 | .01 | .02 | .03 | .04 | .05 | .06 | .07 | .08 | .09 |
| -3.9 | .00005 | .00005 | .00004 | .00004 | .00004 | .00004 | .00004 | .00004 | .00003 | .00003 |
| -3.8 | .00007 | .00007 | .00007 | .00006 | .00006 | .00006 | .00006 | .00005 | .00005 | .00005 |
| -3.7 | .00011 | .00010 | .00010 | .00010 | .00009 | .00009 | .00008 | .00008 | .00008 | .00008 |
| -3.6 | .00016 | .00015 | .00015 | .00014 | .00014 | .00013 | .00013 | .00012 | .00012 | .00011 |
| -3.5 | .00023 | .00022 | .00022 | .00021 | .00020 | .00019 | .00019 | .00018 | .00017 | .00017 |
| -3.4 | .00034 | .00032 | .00031 | .00030 | .00029 | .00028 | .00027 | .00026 | .00025 | .00024 |
| -3.3 | .00048 | .00047 | .00045 | .00043 | .00042 | .00040 | .00039 | .00038 | .00036 | .00035 |
| -3.2 | .00069 | .00066 | .00064 | .00062 | .00060 | .00058 | .00056 | .00054 | .00052 | .00050 |
| -3.1 | .00097 | .00094 | .00090 | .00087 | .00084 | .00082 | .00079 | .00076 | .00074 | .00071 |
| -3.0 | .00135 | .00131 | .00126 | .00122 | .00118 | .00114 | .00111 | .00107 | .00104 | .00100 |
| -2.9 | .00187 | .00181 | .00175 | .00169 | .00164 | .00159 | .00154 | .00149 | .00144 | .00139 |
| -2.8 | .00256 | .00248 | .00240 | .00233 | .00226 | .00219 | .00212 | .00205 | .00199 | .00193 |
| -2.7 | .00347 | .00336 | .00326 | .00317 | .00307 | .00298 | .00289 | .00280 | .00272 | .00264 |
| -2.6 | .00466 | .00453 | .00440 | .00427 | .00415 | .00402 | .00391 | .00379 | .00368 | .00357 |
| -2.5 | .00621 | .00604 | .00587 | .00570 | .00554 | .00539 | .00523 | .00508 | .00494 | .00480 |
| -2.4 | .00820 | .00798 | .00776 | .00755 | .00734 | .00714 | .00695 | .00676 | .00657 | .00639 |
| -2.3 | .01072 | .01044 | .01017 | .00990 | .00964 | .00939 | .00914 | .00889 | .00866 | .00842 |
| -2.2 | .01390 | .01355 | .01321 | .01287 | .01255 | .01222 | .01191 | .01160 | .01130 | .01101 |
| -2.1 | .01786 | .01743 | .01700 | .01659 | .01618 | .01578 | .01539 | .01500 | .01463 | .01426 |
| -2.0 | .02275 | .02222 | .02169 | .02118 | .02068 | .02018 | .01970 | .01923 | .01876 | .01831 |
| -1.9 | .02872 | .02807 | .02743 | .02680 | .02619 | .02559 | .02500 | .02442 | .02385 | .02330 |
| -1.8 | .03593 | .03515 | .03438 | .03362 | .03288 | .03216 | .03144 | .03074 | .03005 | .02938 |
| -1.7 | .04457 | .04363 | .04272 | .04182 | .04093 | .04006 | .03920 | .03836 | .03754 | .03673 |
| -1.6 | .05480 | .05370 | .05262 | .05155 | .05050 | .04947 | .04846 | .04746 | .04648 | .04551 |
| -1.5 | .06681 | .06552 | .06426 | .06301 | .06178 | .06057 | .05938 | .05821 | .05705 | .05592 |
| -1.4 | .08076 | .07927 | .07780 | .07636 | .07493 | .07353 | .07215 | .07078 | .06944 | .06811 |
| -1.3 | .09680 | .09510 | .09342 | .09176 | .09012 | .08851 | .08691 | .08534 | .08379 | .08226 |
| -1.2 | .11507 | .11314 | .11123 | .10935 | .10749 | .10565 | .10383 | .10204 | .10027 | .09853 |
| -1.1 | .13567 | .13350 | .13136 | .12924 | .12714 | .12507 | .12302 | .12100 | .11900 | .11702 |
| -1.0 | .15866 | .15625 | .15386 | .15151 | .14917 | .14686 | .14457 | .14231 | .14007 | .13786 |
| -0.9 | .18406 | .18141 | .17879 | .17619 | .17361 | .17106 | .16853 | .16602 | .16354 | .16109 |
| -0.8 | .21186 | .20897 | .20611 | .20327 | .20045 | .19766 | .19489 | .19215 | .18943 | .18673 |
| -0.7 | .24196 | .23885 | .23576 | .23270 | .22965 | .22663 | .22363 | .22065 | .21770 | .21476 |
| -0.6 | .27425 | .27093 | .26763 | .26435 | .26109 | .25785 | .25463 | .25143 | .24825 | .24510 |
| -0.5 | .30854 | .30503 | .30153 | .29806 | .29460 | .29116 | .28774 | .28434 | .28096 | .27760 |
| -0.4 | .34458 | .34090 | .33724 | .33360 | .32997 | .32636 | .32276 | .31918 | .31561 | .31207 |
| -0.3 | .38209 | .37828 | .37448 | .37070 | .36693 | .36317 | .35942 | .35569 | .35197 | .34827 |
| -0.2 | .42074 | .41683 | .41294 | .40905 | .40517 | .40129 | .39743 | .39358 | .38974 | .38591 |
| -0.1 | .46017 | .45620 | .45224 | .44828 | .44433 | .44038 | .43644 | .43251 | .42858 | .42465 |
| -0.0 | .50000 | .49601 | .49202 | .48803 | .48405 | .48006 | .47608 | .47210 | .46812 | .46414 |