

math 28/B HW4.

Ch4. 1.2. The pdf of Beta(a,b) is  $f(x) = \frac{x^{a-1} (1-x)^{b-1}}{B(a,b)}$

for  $x \in [0,1]$ . Therefore

$$f'(x) = \frac{x^{a-2} (1-x)^{b-2}}{B(a,b)} ((a-1)x - (b-1)(1-x)).$$

(a) increasing: Take  $a=2$ ,  $b=1$

(b) decreasing: Take  $a=1$ ,  $b=2$

(c) increasing in  $(0, p_0)$ , decreasing in  $(p_0, 1)$ .

Take  $a = 1 + p_0$ ,  $b = 2 - p_0$ .

$$\text{In this case, } f'(x) = \frac{x^{a-2} (1-x)^{b-2}}{B(a,b)} (p_0 - x).$$

(d) decreasing in  $(0, p_0)$ , increasing in  $(p_0, 1)$ .

Take  $a = 1 - p_0$ ,  $b = p_0$

$$\text{In this case } f'(x) = \frac{x^{a-2} (1-x)^{b-2}}{B(a,b)} (x - p_0).$$

1.6.  $p$  has a posterior of  $\text{Beta}(a+x, b+n-x)$ . from the conjugate prior of  $\text{Beta}(a, b)$ .

$$\text{Therefore } E(p|x) = \frac{a+x}{a+b+n}$$

$$\text{var}(p|x) = \frac{(a+x)(\cancel{a} b+n-x)}{(a+b+n)^2 (n+a+b+1)}$$

for the target  $p(1-p)$ , the Bayes estimator

$$\text{is } E(p(1-p)|x) = E(p|x) - E(p^2|x)$$

$$= E(p|x) - \text{var}(p|x) - [E(p|x)]^2$$

Some  
algebra  
here

$$\rightarrow = \frac{(x+a)(n+b-x)}{(a+b+n)(a+b+n+1)}$$

1.9. Take a look at the kernel of the posterior:

$$f(\lambda|x) \propto \pi(\lambda) \cdot f(x|\lambda) \propto \lambda^{g+1} e^{-\lambda/\alpha} \cdot \lambda^{\sum x_i} e^{-n\lambda}$$
$$= \lambda^{g+\sum x_i-1} e^{\cancel{-\lambda(\alpha+n)} - \lambda(\frac{1}{\alpha}+n)}$$

This is the kernel of  $\text{Gamma}(g+\sum x_i, (\frac{1}{\alpha}+n)^{-1})$ .

Therefore  $\lambda$  has posterior of  $\text{Gamma}(g+\sum x_i, (\frac{1}{\alpha}+n)^{-1})$ .

The Bayes estimator, which is the posterior mean, is

$$\hat{\lambda} = \frac{g + \sum x_i}{\frac{1}{\alpha} + n} = \frac{\alpha g + \alpha n \bar{X}}{\alpha n + 1} = \underbrace{\frac{1}{\alpha n + 1} \alpha g}_{\text{prior mean}} + \underbrace{\frac{\alpha n}{\alpha n + 1} \bar{X}}_{\text{sample mean}}$$

when  $n \rightarrow \infty$ , the first part shrinks, and  $\hat{\lambda} \rightarrow \bar{X}$

when  $\alpha \rightarrow \infty, g \rightarrow 0$ ,  $\hat{\lambda} \rightarrow \bar{X}$ . (prior contains less information)

If both happens.  $\hat{\lambda} \rightarrow \bar{X}$

Ch5 1.9.

$$\delta^* = \begin{cases} \frac{x}{n} & pr = 1 - \varepsilon \\ \frac{1}{2} & pr = \varepsilon. \end{cases}$$

$$R(p, \delta^*) = E((\delta^* - p)^2) = (1 - \varepsilon) \frac{p - p^2}{n} + \varepsilon \left(\frac{1}{2} - p\right)^2$$

# Set  $\varepsilon = \frac{1}{n+1}$ , one can show that

$$R(p, \delta^*) = \frac{1}{4(n+1)} \text{ which is a constant.}$$

$$\sup_p R(p, \frac{x}{n}) = \sup_p \frac{p(1-p)}{n} = \frac{1}{4n} \quad (p = \frac{1}{2})$$

$$\text{Therefore } R(p, \delta^*) < \sup_p R(p, \frac{x}{n}).$$

$$1.10 \quad \delta = \frac{X + \frac{1}{2}\sqrt{n}}{n + \sqrt{n}} \quad E(\delta) = \frac{nP + \frac{1}{2}\sqrt{n}}{n + \sqrt{n}}$$

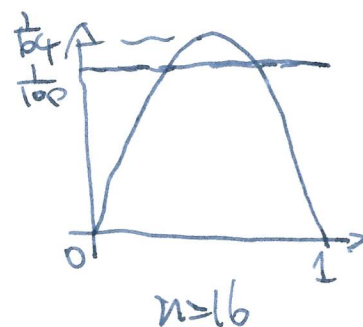
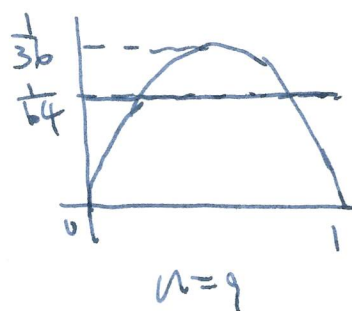
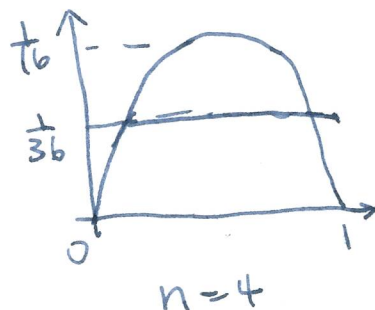
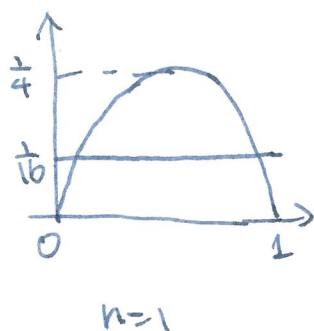
$$\text{Bias}(\delta) = E(\delta) - p = \frac{(\frac{1}{2} - p)\sqrt{n}}{n + \sqrt{n}}$$

Therefore  $p < \frac{1}{2}$  it's ~~not under~~ over-estimating

$p > \frac{1}{2}$  it's underestimating.

$$1.12 \quad R(p, \delta) = \frac{1}{4} \left( \frac{1}{1 + \sqrt{n}} \right)^2$$

$$R(p, \frac{X}{n}) = \frac{p(1-p)}{n}$$



as  $n$  increases the difference of two peaks of the estimators gets closer.