
STAT 223: Project 2

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Applied Analytics

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Medical expenses using multiple regression

Use historical data on individuals to predict medical expenses.

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BRIEF DESCRIPTION OF THE DATA

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The data set has six potential predictor variables for a total of 1338 individuals. Three variables are numerical and three are categorical. The numerical variables are the age of the policy holder, his or her bmi (body mass index), and the number of children of the policy holder. Sex, whether the policy holder is a smoker, and geographical region are the categorical predictors. The response is calendar year total medical expenses for the covered members of the policy holder's family.

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HOW TO USE THE DATA FILE

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The data file is in csv format, under the name "insurance.csv". Each of the 1338 rows after the row containing variable names has the information on a policy holder that is described above, so there are six predictor columns followed by the response column. There are no missing values. The variable names are:

age.....integer values.

sex.....character values, "female" or "male". (R may consider this as a factor variable, with codes 1 and 2 if you use read.csv to input the data)

bmi.....numerical values.

children.....integer values.

smoker.....character values, “no” or “yes”. (R may consider this as a factor variable, with codes 1 and 2 respectively for “no” and “yes”)

region.....character values “northeast”, “northwest”, “southeast”, “southwest”. (As a factor, the codes for these will be alphabetized 1,2,3,4 respectively)

expenses.....numerical values

The data file is “insurance.csv”.

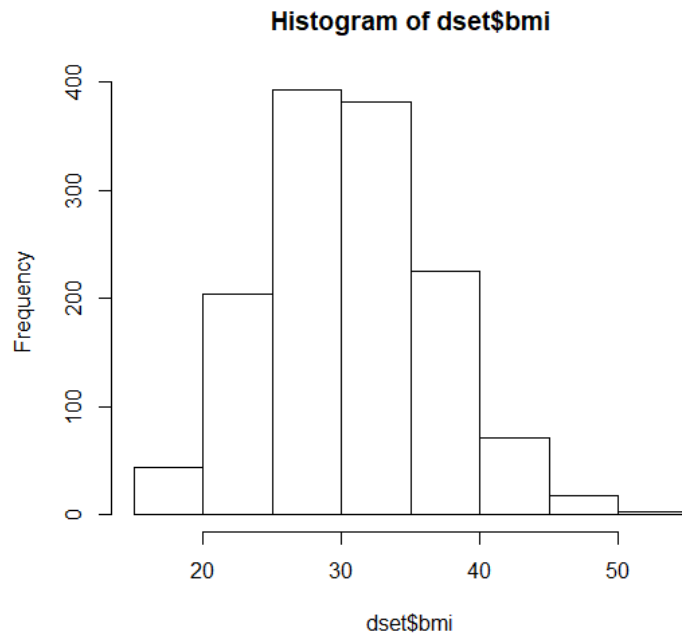
1. Understanding the variables. Use R's table() command to get a frequency distribution of each of the categorical variables and summary() to get basic statistics on the integer and numeric variables. Produce histograms to decide whether the predictor bmi and the response expenses appear to be normally distributed. If you set up a new variable for the log of expenses, does that appear to be more normal?

```
> table(dset$sex)
female  male
  662    676
> table(dset$smoker)
  no  yes
1064  274
> table(dset$region)
northeast northwest southeast southwest
      324       325       364       325
.         .         .         .
```

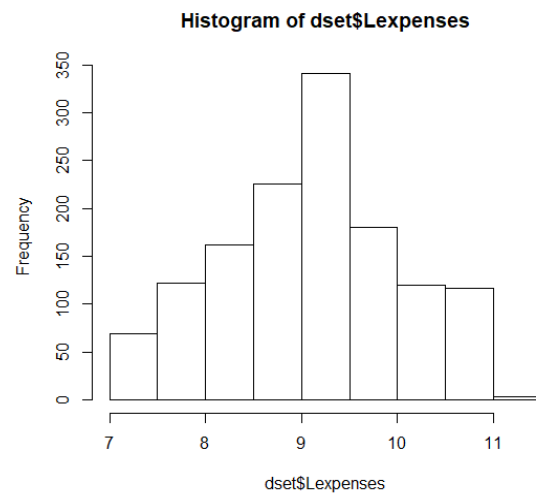
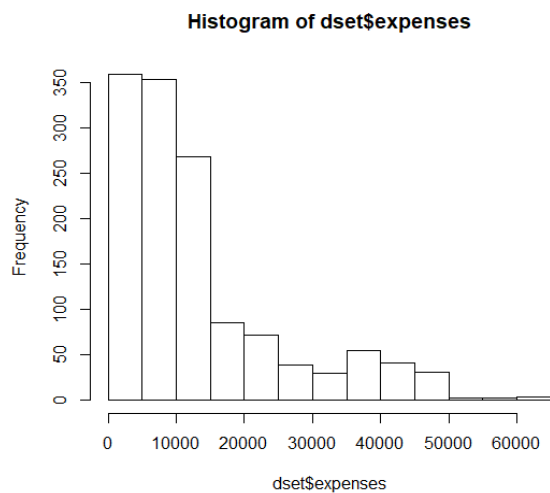
We notice from above that the dataset had 662 females and 676 males. 1064 were non-smokers and 274 were smokers. We also see about equal distribution in the regions.

```
> summary(dset)
   age      sex      bmi      children      smoker
Min.   :18.00 female:662 Min.   :16.00 Min.   :0.000 no :1064
1st Qu.:27.00 male  :676 1st Qu.:26.30 1st Qu.:0.000 yes: 274
Median :39.00
Mean   :39.21
3rd Qu.:51.00
Max.   :64.00
   region      expenses      Lexpenses      SQage      bmiSmoker
northeast:324 Min.   : 1122 Min.   : 7.023 Min.   : 324 Mode:logical
northwest:325 1st Qu.: 4740 1st Qu.: 8.464 1st Qu.: 729 NA's:1338
southeast:364 Median : 9382 Median : 9.147 Median :1521
southwest:325 Mean   :13270 Mean   : 9.099 Mean   :1734
               3rd Qu.:16640 3rd Qu.: 9.720 3rd Qu.:2601
               Max.   :63770 Max.   :11.063 Max.   :4096
```

The summary function is very useful as it tells us min/max and quartiles information of integer and numeric values in the dataset.

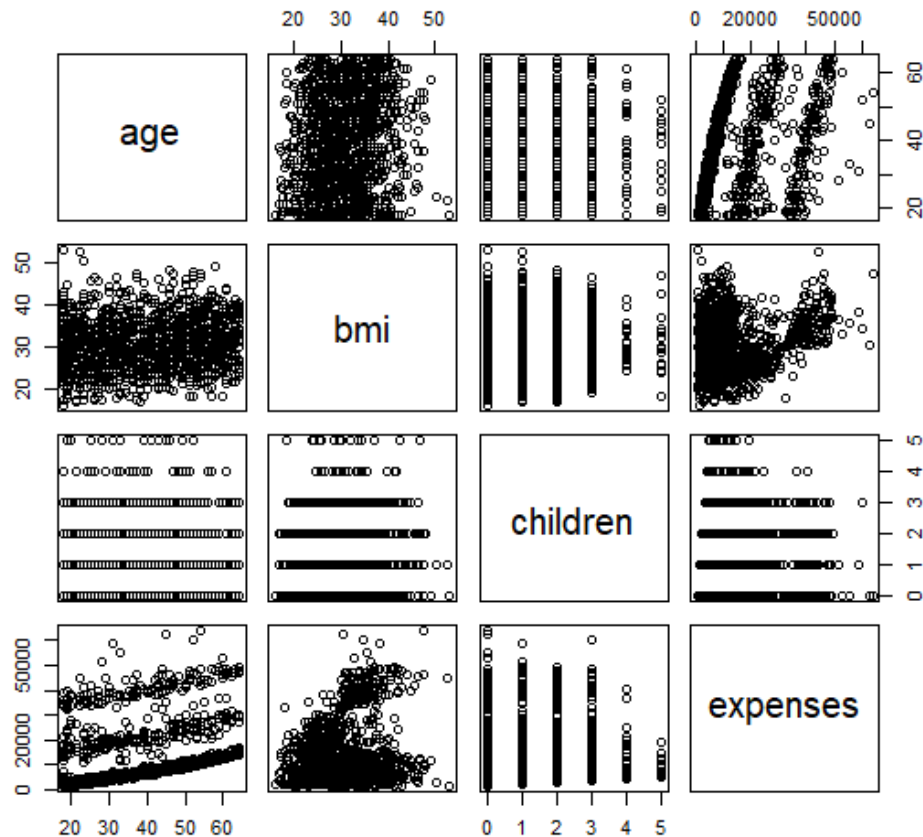


The histogram plot of predictor bmi appears normal.

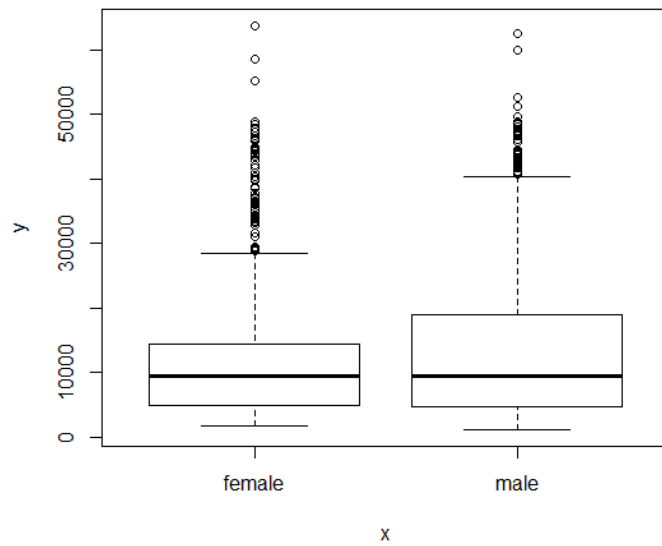


The histogram plot of response expenses alone did not appear to be normally distributed but when we set up a new variable by taking the log of expenses, the data appears to be more normal which we will further use for our analysis.

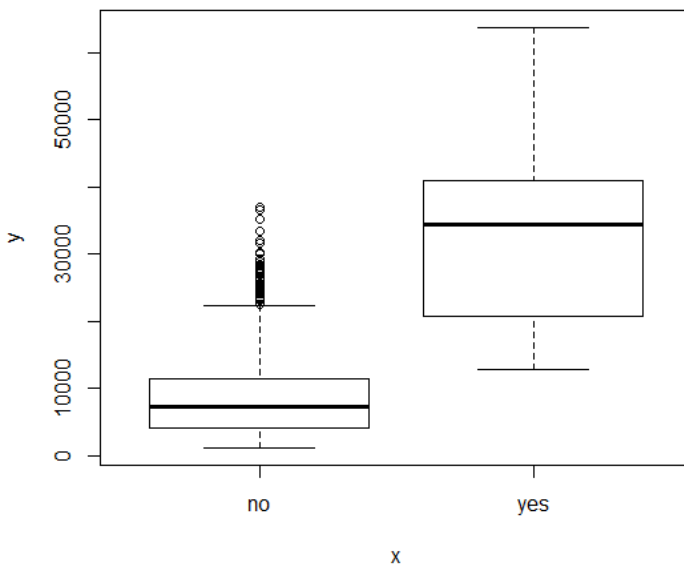
2. Understanding the relationships. Use the `pairs(insurance[c("age","bmi","children","expenses")])` command in R to produce a scatterplot matrix of the integer and numeric variables and note any features that stand out. What do you get when you apply the `plot(varname,expenses)` command to the categorical variables? Try both again with the logged expenses.



The pair plot above is for all categorical variables and expenses. Most of the plots do not give us a pattern or correlation to work with. However, the plot expenses vs age shows a linear correlation with each other. This positive relationship entails that the older the policyholder, the more they spend. One other thing that stands out about this scatterplot is that there's slight curvature. In addition to the curves, it seems as though the data is being divided into groups. This particular influence on the shape of the graph might be a result of the categorical variables in our insurance dataset.

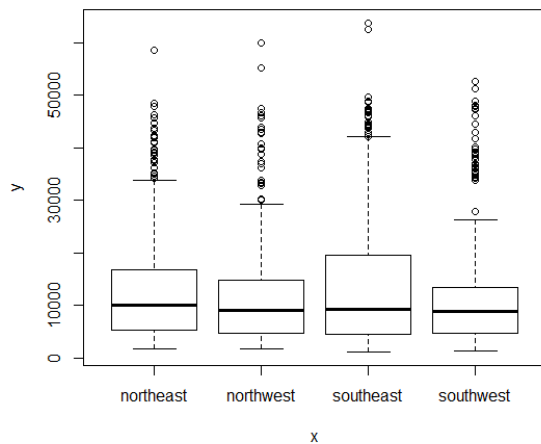


The middle values for both female and male policyholders are about \$10,000. In general, both genders seem to have the same spread, except for the fact the males have a higher 75th percentile value. The distributions of expenses for both female and male policyholder are heavily skewed to the right.

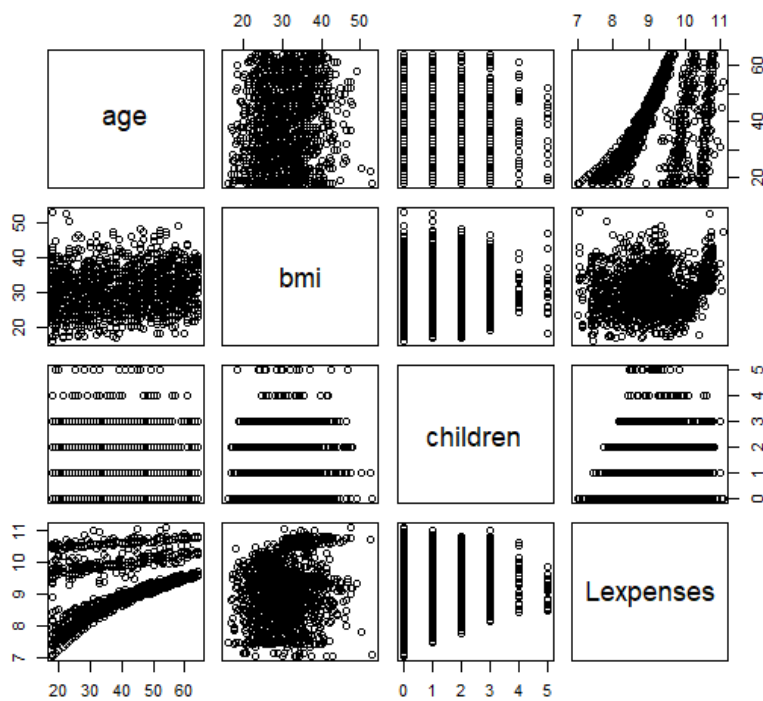


As for the box plot of the smoker variable, and the expense, we observe that the median for smokers is well above than that of non-smokers. Even the minimum value of the smokers is

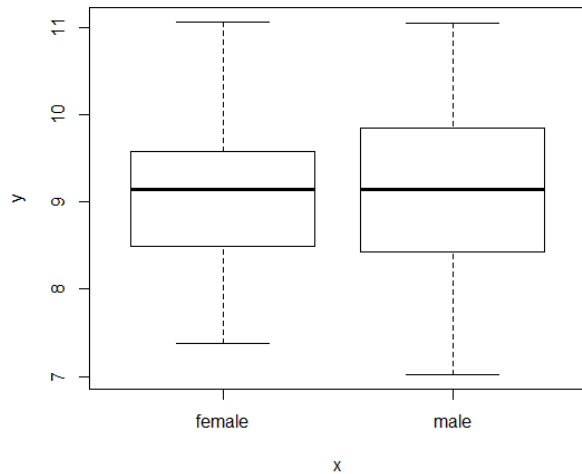
higher than the third quartile of non-smokers. In general, we would conclude that smokers have a much wider spread than non-smokers. We also notice that the distribution of the policyholder's expenses for non-smokers is affected by large outliers.



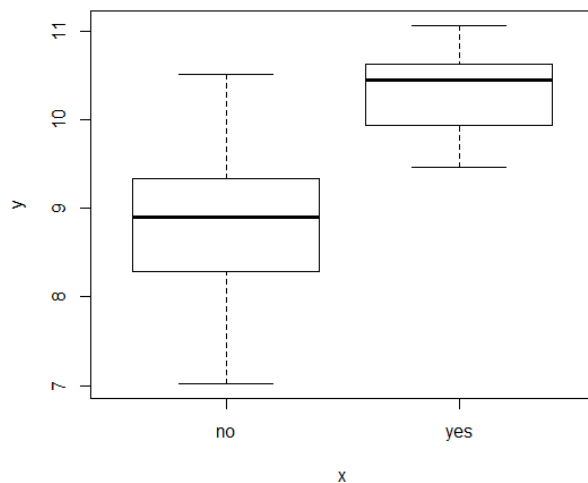
Unlike the previous two variables, it appears that all four regions have about the same spread. The distributions of expenses for all four regions, however, are once again influenced by data points that lie beyond the “whiskers” of the box plots. One way to fix this problem is to transform our response variable using a log function.



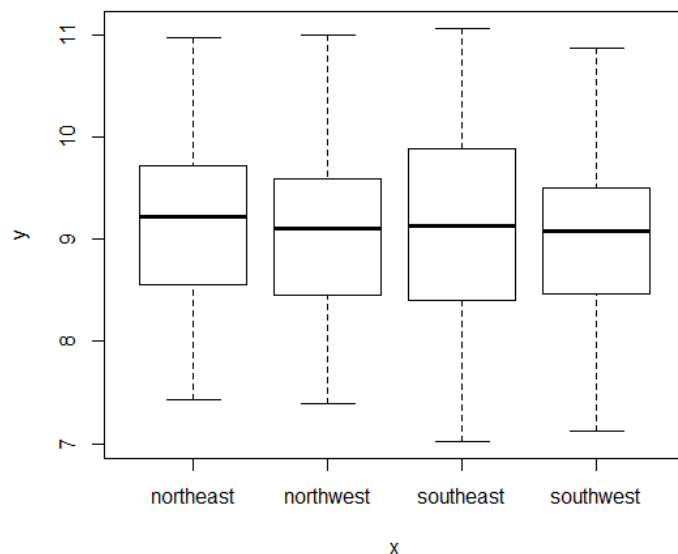
We use the pairs command again. The pair plot above is for all categorical variables and log of the expenses. Most of the plots do not give us a pattern or correlation to work with. However, the plot log expenses vs age shows a stronger linear correlation with each other.



The box and whisker plot tells us how the data is distributed among the sex categorical variables. The distribution now is now at a more suitable data range.



The box and whisker plot tells us how the data is distributed among the smoker categorical variables. 'Yes' if the subject smokes and 'no' if not. The distribution now is now at a more suitable data range.



The box and whisker plot tells us how the data is distributed among the region categorical variables. The distribution now is now at a more suitable data range.

3. The regression. Now run a regression of the expenses on the full set of variables. Interpret the results. Why do you see more than 6 predictors? (Hint: remember the concept of the indicator variable.) Compare your results to a regression of the logged expenses on the variables. Which seems better, in terms of both the significance of the predictors, the R2 value, and the distribution of the residuals?

```
call:
lm(formula = dset$expenses ~ dset$age + dset$sex + dset$bmi +
    dset$children + dset$smoker + dset$region)

Residuals:
    Min       1Q   Median       3Q      Max
-11302.7  -2850.9   -979.6   1383.9  29981.7

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)   -11941.6     987.8  -12.089  < 2e-16 ***
dset$age         256.8       11.9    21.586  < 2e-16 ***
dset$sexmale    -131.3     332.9   -0.395  0.693255
dset$bmi         339.3       28.6    11.864  < 2e-16 ***
dset$children    475.7     137.8     3.452  0.000574 ***
dset$smokeryes  23847.5    413.1   57.723  < 2e-16 ***
dset$regionnorthwest -352.8    476.3   -0.741  0.458976
dset$regionsoutheast -1035.6    478.7   -2.163  0.030685 *
dset$regionsouthwest -959.3    477.9   -2.007  0.044921 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6062 on 1329 degrees of freedom
Multiple R-squared:  0.7509,    Adjusted R-squared:  0.7494
F-statistic: 500.9 on 8 and 1329 DF,  p-value: < 2.2e-16
```

From the model summary we see that the R-squared value is 0.75 which is very close to 1 with the p-value of less than 10^{-16} which is highly significant. This means about 75% of the

variation can be explained by our model. We should notice however that a few predictors do not have significant p-values for example: sex male.

We see more than 6 predictor variables because of the concept of dummy variables. We know that if a categorical variable has n different values, you only need n-1 dummy variables to indicate which category is active. So, in the multiple regression equation, if the sex value is 0 then it means female and 1 if male. Same thing with the Smoker predictor. For the region category, if all the values are 0 then it will indicate 'Northeast'. It is just a way to account for both variables to influence the regression model.

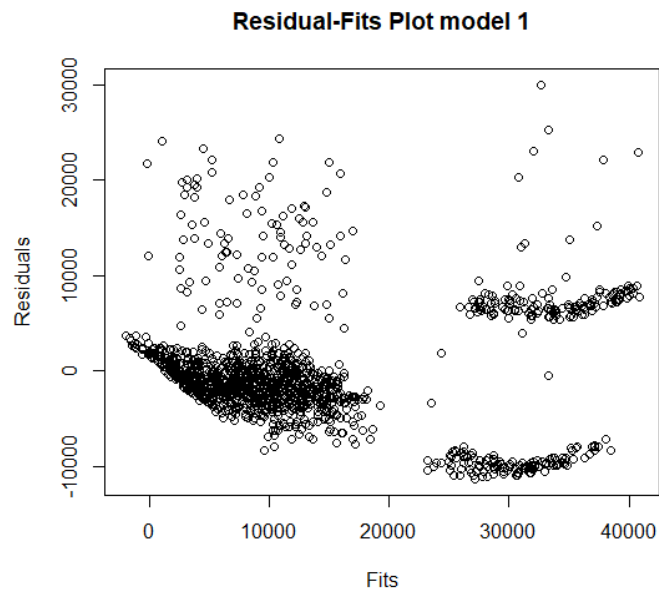
```
Call:
lm(formula = dset$lexpenses ~ dset$age + dset$sex + dset$bmi +
    dset$children + dset$smoker + dset$region)

Residuals:
    Min       1Q   Median       3Q      Max
-1.07125 -0.19783 -0.04891  0.06604  2.16655

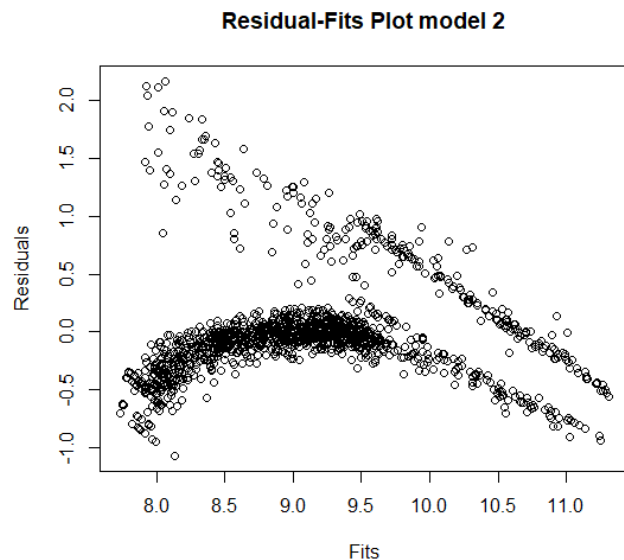
Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)    7.0307859   0.0723992   97.111 < 2e-16 ***
dset$age        0.0345816   0.0008721   39.654 < 2e-16 ***
dset$sexmale    -0.0754109   0.0244017   -3.090 0.002040 **
dset$bmi        0.0133658   0.0020960    6.377 2.49e-10 ***
dset$children   0.1018651   0.0100997   10.086 < 2e-16 ***
dset$smokeryes  1.5542783   0.0302800   51.330 < 2e-16 ***
dset$regionnorthwest -0.0637805  0.0349064   -1.827 0.067896 .
dset$regionsoutheast -0.1571654  0.0350837   -4.480 8.12e-06 ***
dset$regionsouthwest -0.1289048  0.0350274   -3.680 0.000242 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.4443 on 1329 degrees of freedom
Multiple R-squared:  0.7679,    Adjusted R-squared:  0.7665
F-statistic: 549.7 on 8 and 1329 DF,  p-value: < 2.2e-16
```

The model above uses the logged expenses on the variables. This model does better for a various reason. Firstly, we notice that all the p-values for the predictors are less than 0.05 which says that the model is highly significant considering all predictors. The R-squared value is 0.77 which is better than our previous model and is close to 1 so we can say that the model is doing a good job with the linear regression. This means about 77% of the variation can be explained by our model.



We see here that when residuals are plotted with the fitted y values in the expenses plot, the linearity assumption is not met because we are seeing a pattern of a group of points and two lines that are formed. We also notice that the variation is not constant as it varies from the start of the graph and to the end of the graph.



We see here that when residuals are plotted with the fitted y values in the log expenses plot, the linearity assumption is not met because we are seeing a curved pattern of a group of points and a linear straight line. This means that one of our predictors could be improved by squaring each

values of the predictor (Question 4). We also notice that the variation is not constant as it varies from the start of the graph and to the end of the graph.

4. Improving the model. Returning to the original response variable expenses (not the logs), you may have noticed a slight curvature in the scatterplot of expenses against age. Create a new variable that is the square of the age and rerun the regression with it included. If the quadratic term is significant, keep it in the model and then try adding an interaction term between the bmi and smoker variables. (In R this would be a term `bmi*smoker` in the model equation). Do either of these changes or both seem to result in better prediction?

After we create a new variable of square of age we rerun the regression.

```
Call:
lm(formula = dset$expenses ~ dset$age + dset$SQage + dset$sex +
    dset$bmi + dset$children + dset$smoker + dset$region)

Residuals:
    Min       1Q   Median       3Q      Max
-11665.6  -2854.7   -942.7   1300.8  30814.6

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept)   -6602.064    1689.528   -3.908  9.79e-05 ***
dset$age        -54.423      80.989   -0.672  0.501716
dset$SQage       3.925       1.010    3.885  0.000107 ***
dset$sexmale    -138.451    331.189   -0.418  0.675983
dset$bmi        335.291     28.467   11.778  < 2e-16 ***
dset$children   642.121    143.613    4.471  8.44e-06 ***
dset$smokeryes  23858.690   410.976   58.054  < 2e-16 ***
dset$regionnorthwest -367.632   473.771   -0.776  0.437905
dset$regionsoutheast -1031.998   476.164   -2.167  0.030388 *
dset$regionsouthwest -956.787   475.398   -2.013  0.044358 *
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 6030 on 1328 degrees of freedom
Multiple R-squared:  0.7537,    Adjusted R-squared:  0.7521
F-statistic: 451.6 on 9 and 1328 DF,  p-value: < 2.2e-16
```

We see here that the p-value is not significant for just the age predictor as it has a value of 0.502. However, we notice that the square age gives us a p-value of 0.0001 which is highly significant. The R-square is doing just as well like the previous models with 75% of the variations explained. Due to this we decide to keep the new age-squared variable in our model.

```
lm(formula = dset$expenses ~ dset$age + dset$Sqage + dset$sex +
    dset$bmi + dset$children + dset$smoker + dset$region + dset$bmi *
    dset$smoker)
```

Residuals:

| Min | 1Q | Median | 3Q | Max |
|--------|-------|--------|-------|-------|
| -14905 | -1592 | -1282 | -1024 | 31294 |

Coefficients:

| | Estimate | Std. Error | t value | Pr(> t) |
|---------------------------|------------|------------|---------|--------------|
| (Intercept) | 2.854e+03 | 1.390e+03 | 2.053 | 0.04028 * |
| dset\$age | -3.299e+01 | 6.459e+01 | -0.511 | 0.60957 |
| dset\$Sqage | 3.740e+00 | 8.057e-01 | 4.642 | 3.79e-06 *** |
| dset\$sexmale | -5.054e+02 | 2.644e+02 | -1.911 | 0.05617 . |
| dset\$bmi | 2.005e+01 | 2.541e+01 | 0.789 | 0.43037 |
| dset\$children | 6.750e+02 | 1.145e+02 | 5.894 | 4.77e-09 *** |
| dset\$smokeryes | -2.035e+04 | 1.635e+03 | -12.445 | < 2e-16 *** |
| dset\$regionnorthwest | -5.987e+02 | 3.779e+02 | -1.584 | 0.11336 |
| dset\$regionsoutheast | -1.206e+03 | 3.798e+02 | -3.176 | 0.00153 ** |
| dset\$regionsouthwest | -1.226e+03 | 3.792e+02 | -3.234 | 0.00125 ** |
| dset\$bmi:dset\$smokeryes | 1.441e+03 | 5.222e+01 | 27.595 | < 2e-16 *** |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

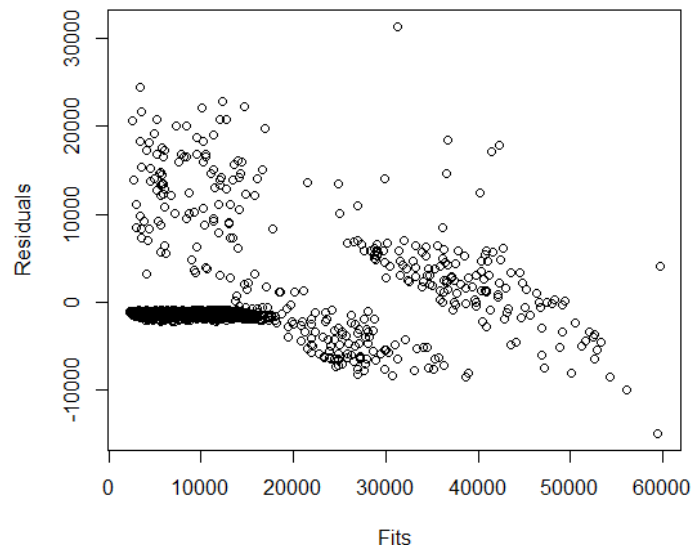
Residual standard error: 4808 on 1327 degrees of freedom

Multiple R-squared: 0.8435, Adjusted R-squared: 0.8423

F-statistic: 715.3 on 10 and 1327 DF, p-value: < 2.2e-16

After we add our interaction term, we notice the following: The age-squared predictor is still highly significant, and our interaction term is also highly significant with a p-value of 10^{-16} . This means that we cannot exclude both bmi and smoker from our model. The R-squared value is very good as we can explain up to 84% of the variation from the model.

Residual-Fits Plot model 4



When we plot a residual-fits for this model we notice that the pattern has disappeared which tells us that the linearity assumptions are met. The points are scattered, and the variation is fairly equal.

Code:

```
#-----  
# Importing data  
  
dset <- read.csv("insurance.csv", header = TRUE);  
head(dset)  
  
#-----  
# Part 1  
  
# Frequency distribution of each of the categorical variables  
  
table(dset$sex)  
table(dset$smoker)  
table(dset$region)  
  
# Summary to get basic statistics  
summary(dset)  
  
# Histogram to check normal distribution  
hist(dset$bmi)  
hist(dset$expenses)  
  
# Log of expences as a new variable  
dset$Lexpenses <- log(dset$expenses); View(dset)  
hist(dset$Lexpenses)  
  
#-----  
# Part 2  
  
# scatter plot matric of numeric variables  
pairs(dset[c("age", "bmi", "children", "expenses")])  
  
# Plot of categorical variables with expences
```

```
plot(dset$sex,dset$expenses)
plot(dset$smoker,dset$expenses)
plot(dset$region,dset$expenses)
```

```
# Repeat for log expences
pairs(dset[c("age","bmi","children","Lexpenses")])
```

```
plot(dset$sex,dset$Lexpenses)
plot(dset$smoker,dset$Lexpenses)
plot(dset$region,dset$Lexpenses)
```

```
#-----
```

```
# Part 3
```

```
#Multiple regression
```

```
model1 <-
lm(dset$expenses~dset$age+dset$sex+dset$bmi+dset$children+dset$smoker+dset$region);su
mmary(model1)
model2 <-
lm(dset$Lexpenses~dset$age+dset$sex+dset$bmi+dset$children+dset$smoker+dset$region);s
ummary(model2)
anova(model1)
anova(model2)
```

```
fits1 <- fitted(model1)
resids1 <- residuals(model1)
plot(fits1, resids1, main = "Residual-Fits Plot model 1", xlab = "Fits", ylab = "Residuals")
```

```
fits2 <- fitted(model2)
resids2 <- residuals(model2)
plot(fits2, resids2, main = "Residual-Fits Plot model 2", xlab = "Fits", ylab = "Residuals")
```

```
#-----
```

```
# Part 4
```

```
dset$SQage <- dset$age^2; View(dset)
```

```
model3 <-
```

```
lm(dset$expenses~dset$age+dset$SQage+dset$sex+dset$bmi+dset$children+dset$smoker+d  
set$region);summary(model3)
```

```
summary(model3)
```

```
model4 <-
```

```
lm(dset$expenses~dset$age+dset$SQage+dset$sex+dset$bmi+dset$children+dset$smoker+d  
set$region+dset$bmi*dset$smoker);summary(model4)
```

```
summary(model4)
```

```
fits3 <- fitted(model3)
```

```
resids3 <- residuals(model3)
```

```
plot(fits3, resids3, main = "Residual-Fits Plot model 3", xlab = "Fits", ylab = "Residuals")
```

```
fits4 <- fitted(model4)
```

```
resids4 <- residuals(model4)
```

```
plot(fits4, resids4, main = "Residual-Fits Plot model 4", xlab = "Fits", ylab = "Residuals")
```