Building a Prognostic Biomarker

Noah Simon and Richard Simon

July 28, 2016

Prognostic Biomarker for a Continuous Measure

On each of n patients measure

```
y_i - single continuous outcome (eg. blood pressure, tumor growth) 
 \mathbf{x}_i - p-vector of features (eg. SNPs, gene expression values)
```

Want to model y_i by \mathbf{x}_i to

- ► Predict high risk patients (give them additional care)
- Learn the underlying biology

Linear Regression:

We assume that

$$y_i = \beta_0 + x_i^{\top} \beta + \epsilon_i$$

Generally fit by solving:

$$\min_{\beta_0,\beta} \frac{1}{2} \sum \left(y_i - \beta_0 - x_i^{\top} \beta \right)^2$$

Diabetes data example

$$n = 442, p = 10$$

 y_i is quantitative measure of disease progression (one year after baseline)

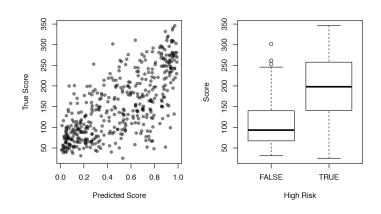
 \mathbf{x}_i includes age, sex, BMI, avg BP, and six serum measurements

Can use
$$\eta_i = \hat{\beta}_0 + \hat{\beta}^\top x_i$$
 to predict risk.

Or can stratify

$$\eta_i \ge \text{cutoff} \to \text{high risk}$$
 $\eta_i < \text{cutoff} \to \text{low risk}$

Choosing the *cutoff* can be tricky



Two Issues

When $p \sim n$ (or p > n), estimate is highly variable

When true model is far from linear, estimate is inflexible

Working in High Dimensions

For $p \sim n$ we need an often reasonable assumption:

Most of the features are [conditionally] unrelated to response

More formally: In the model

$$y_i = \beta_0 + x_i^{\top} \beta + \epsilon_i$$

Most of the β_j are 0 (or very nearly 0).

Bet on Sparsity

Is this assumption reasonable?

Often

If not, statistical trickery generally will not help.

Either need more samples

Or more benchwork

How do we fit a sparse model?

Most obvious approach is:

minimize_{$$\beta_0,\beta$$}
$$\sum_{i} \left(y_i - \beta_0 - x_i^{\top} \beta \right)^2$$
 subject to
$$\# \left\{ j \mid \beta_j \neq 0 \right\} \leq d$$

Unfortunately, this is computationally intractable.

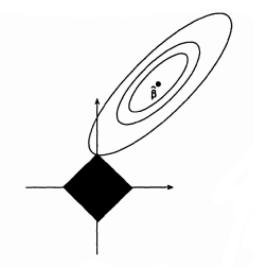
How do we fit a sparse model?

Instead we use the Lasso

$$\begin{aligned} & \text{minimize}_{\beta_0,\beta} & & \sum_i \left(y_i - \beta_0 - x_i^\top \beta \right)^2 \\ & \text{subject to} & & \sum_i |\beta_i| \leq c \end{aligned}$$

This can be solved as fast as (or faster than) least squares.

How does this give us sparsity?



Decreasing c increases sparsity.

Equivalent Penalized Form

$$\mathsf{minimize}_{\beta_0,\beta} \sum_i \left(y_i - \beta_0 - x_i^\top \beta \right)^2 + \lambda \sum_i |\beta_i|$$

c in constrained form $\longleftrightarrow \lambda$ in penalized form

In practice everyone uses

Training/test validation

OR

Cross-validation

Training/Test Validation

- 1. Choose candidate λ -values: $\lambda_1, \ldots, \lambda_M$
- 2. Split observations into two groups: training and test
- 3. For each candidate $m \leq M$
 - 3.1 Training Data \rightarrow Build model with λ_m
 - 3.2 Test Data \rightarrow Apply model to get "predictions"
- 4. Evaluate the predictions for each model choose the best.

Training/Test Validation

Evaluating each model

For each λ_m we have $\hat{y}_i^{(m)}$ $i = 1, \ldots, n_{test}$.

Simplest evaluation via mean-square-error:

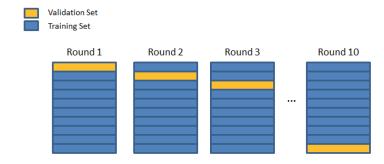
$$performance_m = \sum_{i \in test data} \left(y_i - \hat{y}_i^{(m)} \right)^2$$

Training/Test Validation





Cross-Validation



Cross-Validation

- 1. Choose candidate λ -values: $\lambda_1, \ldots, \lambda_M$
- 2. Split observations into K folds:
- 3. For each candidate $m \leq M$, and each fold $k \leq K$
 - 3.1 Data (minus fold k) \rightarrow Build model with λ_m
 - 3.2 Data (fold k) \rightarrow Apply model to get "predictions"
- 4. Evaluate models (now on all data)

What if the relationship isn't linear?

$$y = 3\sin(x) + \epsilon$$
$$y = 2e^{x} + \epsilon$$
$$y = 3x^{2} + 2x + 1 + \epsilon$$

If we know the functional form we can still use "linear regression"

$$y = 3\sin(x) + \epsilon:$$

$$\begin{pmatrix} x \end{pmatrix} \to \begin{pmatrix} \sin(x) \end{pmatrix}$$

$$y = 3x^2 + 2x + 1 + \epsilon:$$

$$\begin{pmatrix} x \end{pmatrix} \to \begin{pmatrix} x & x^2 \end{pmatrix}$$

What if we don't know the right functional form?

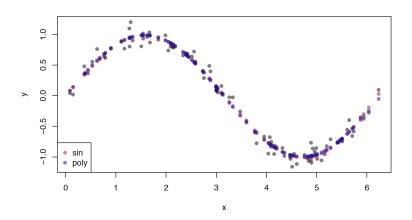
Use a flexible basis expansion:

▶ polynomial basis

$$\left(x\right) \to \left(x \mid x^2 \mid \cdots \mid x^k\right)$$

► hockey-stick (/spline) basis

$$\begin{pmatrix} x \end{pmatrix} \rightarrow \begin{pmatrix} x \mid (x-t_1)_+ \mid \cdots \mid (x-t_k)_+ \end{pmatrix}$$



For high dimensional problems, expand each variable

$$\left(\begin{array}{c|cccc} x_1 & x_2 & \cdots & x_p \end{array}\right) \rightarrow \left(\begin{array}{cccccc} x_1 & \cdots & x_1^k & x_2 & \cdots & x_2^k & \cdots & x_p \end{array}\right)$$

and use the Lasso on this expanded problem.

k must be small (\sim 5ish)

Spline basis generally outperforms polynomial

Prognostic Biomarker for a Binary Measure

On each of *n* patients measure

```
y_i - single binary outcome (eg. Progression after a year, pCR) 
 \mathbf{x}_i - p-vector of features (eg. SNPs, gene expression values)
```

More common than continuous response

Logistic Regression

Relate it back to continuous methods:

For continuous response solve:

$$\mathsf{minimize}_{\beta,\beta_0} \sum_i \left(y_i - \beta_0 - x_i^\top \beta \right)^2$$

One interpretation:

lf

$$y_i|x_i \sim N\left(x_i^{\top}\beta, \sigma^2\right)$$

then maximizing likelihood is equivalent to least squares.

Logistic Regression

Relate it back to continuous methods:

For Binary Response, consider

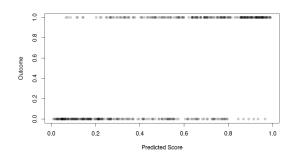
$$y_i|x_i \sim \operatorname{ber}\left(p_i = \frac{e^{\beta_0 + x_i^{\top}\beta}}{1 + e^{\beta_0 + x_i^{\top}\beta}}\right)$$

Maximizing likelihood is equivalent to minimizing

$$\ell\left(\beta,\beta_{0}\right) = -\sum_{i}\left[y_{i}\left(\beta_{0} + x_{i}^{\top}\beta\right) - \log\left(1 + e^{\beta_{0} + x_{i}^{\top}\beta}\right)\right]$$

This is just logistic regression

Diabetes Example



Additionally:

166/221 test-positive vs 55/221 test-negative patients with above median progression

Penalized Logistic Regression

As in least squares, we can induce sparsity:

$$\mathsf{minimize}_{\beta,\beta_0}\,\ell\left(\beta,\beta_0\right) + \lambda \sum |\beta_j|$$

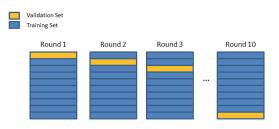
Penalized Logistic Regression

As in least squares, we can induce sparsity:

$$\mathsf{minimize}_{\beta,\beta_0}\,\ell\left(\beta,\beta_0\right) + \lambda \sum |\beta_j|$$

Choosing λ is a bit tricky.

We need a measure of the performance of our model



Using CV, we get

$$\hat{\eta}_i = \hat{\beta_0}^{train} + x_i^\top \hat{\beta}^{train}$$

For each patient we have a CV score

$$\hat{\eta}_i = \hat{eta}_0^{\; train} + x_i^{ op} \hat{eta}^{train}$$

Now plug-in

$$\hat{
ho}_i = rac{e^{\hat{\eta}_i}}{1+e^{\hat{\eta}_i}}$$

And use Cross-Validated Predictive Likelihood as our measure

$$\prod_i \hat{\rho}_i^{y_i} \left(1 - \hat{\rho}_i\right)^{1 - y_i}$$

(Some software equivalently uses the negative log-likelihood)

Can also classify patients using \hat{p}_i :

$$\widehat{class}_i = \begin{cases} 1 & : \hat{p}_i \ge 0.5 \\ 0 & : \hat{p}_i < 0.5 \end{cases}$$

And use Cross-Validated Misclassification Rate as our measure

proportion
$$(y_i \neq \widehat{class}_i)$$

Example

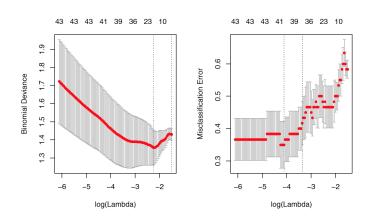
Prognostic Biomarker for pCR of HER2+ breast cancer patients on Herceptin + CT

n = 60 patients, with 28 pCR

Expression from p = 5349 genes with non-negligable variance

GEO: GSE50948

Example



Some Other Classifiers

Many other classification choices. Noteworthy high dimensional options:

Diagonal Linear Discriminant Analysis (DLDA)

Nearest Shrunken Centroid (PAM)

Support Vector Machine (SVM)

Find a *score* based on features: $x_i \to x_i^\top \beta$ indicating likelihood of each class

DLDA

Assumes:

Gaussian features within class With Pooled Diagonal Covariance:

$$(\mathbf{x_i}|y_i=0) \sim \mathsf{N}(\mu_0, D) \qquad (\mathbf{x_i}|y_i=1) \sim \mathsf{N}(\mu_1, D)$$

Estimate μ_1, μ_2, D by maximum likelihood.

Calculate Probability of each class using Bayes and plug-in.

Score is:

$$\eta_i = \hat{D}^{-1} \left(\hat{\mu}_1 - \hat{\mu}_0 \right)^\top x_i$$

PAM

Assumes:

Gaussian features within class With Pooled Diagonal Covariance:

$$(\mathbf{x_i}|y_i = 0) \sim N(\mu_0, D)$$
 $(\mathbf{x_i}|y_i = 1) \sim N(\mu_1, D)$

Estimate μ_1, μ_2 with shrinkage!

$$\tilde{\mu}_{1j} = \tilde{\mu}_{2j}$$
 for most j

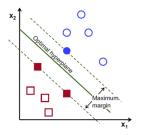
Score is:

$$\eta_i = \hat{D}^{-1} (\tilde{\mu}_1 - \tilde{\mu}_0)^\top x_i$$

SVM

No statistical model: just discriminant method.

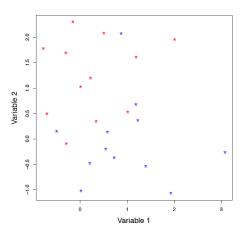
Finds the maximum-margin separating hyperplane



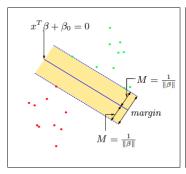
Score is (signed) distance from hyperplane

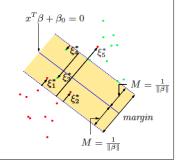
$$\eta_i = \beta^\top x_i + \beta_0$$

What if There is No Separating Hyperplane?



Support Vector Classifier: Allow for Violations



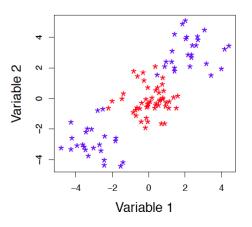


► The support vector machine is just like the support vector classifier, but it elegantly allows for non-linear expansions of the variables: "non-linear kernels".

- ► The support vector machine is just like the support vector classifier, but it elegantly allows for non-linear expansions of the variables: "non-linear kernels".
- However, linear regression, logistic regression, and other classical statistical approaches can also be applied to non-linear functions of the variables.

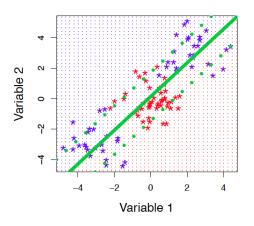
- ► The support vector machine is just like the support vector classifier, but it elegantly allows for non-linear expansions of the variables: "non-linear kernels".
- However, linear regression, logistic regression, and other classical statistical approaches can also be applied to non-linear functions of the variables.
- For historical reasons, SVMs are more frequently used with non-linear expansions as compared to other statistical approaches.

Non-Linear Class Structure

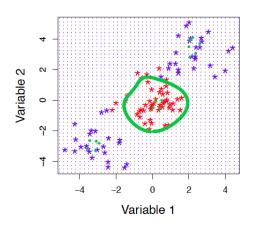


This will be hard for a linear classifier!

Try a Support Vector Classifier



Uh-oh!!



Much Better.

On each of n patients measure

```
(t_i, s_i) - time, censoring-status (eg. Disease free survival)
```

 x_i - p-vector of features

(eg. SNPs, gene expression values)

Hazard is

probability density of failure at time t given survival up to t.

Want to model the hazard as a function of covariates

We will use Cox Proportional Hazards Model

Assumes hazard is

$$\lambda(t) = h(t)e^{x_i^{\top}\beta}$$

where

h(t) is covariate-independent baseline hazard

 $e^{x_i^{\top}\beta}$ is covariate-based tilt

Considering Partial Likelihood — likelihood at only event times h(t) falls out.

$$L(\beta) = \prod_{i \in D} \frac{e^{x_i^{\top} \beta}}{\sum_{j \in R_i} e^{x_j^{\top} \beta}}$$

Maximizing this is equivalent to minimizing

$$\ell(\beta) = -\sum_{i \in D} \left[x_i^{\top} \beta - \log \left(\sum_{j \in R_i} e^{x_j^{\top} \beta} \right) \right]$$

Bells and Whistles

Similar additions as continuous/binary response

$$\mathsf{minimize}_{\beta}\,\ell\left(\beta\right) + \lambda \sum |\beta_j|$$

Trickiest for choosing λ yet.

Straightforward CV approach

Find CV estimate for each patient

$$\hat{\eta}_i = x_i^{\top} \hat{\beta}^{train}$$

Calculate CV predictive partial likelihood

$$\prod_{i \in D} \frac{e^{\hat{\eta}_i}}{\sum_{j \in R_i} e^{\hat{\eta}_j}}$$

Not necessarily a great measure