

A Study on the Benefit of a User-Controlled Radial Tour for Variable Importance for Structure in High-Dimensional Data

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Abstract—Principal component analysis is a long-standing go-to method for exploring multivariate data. The principal components are linear combinations of the original variables, ordered by descending variance. The first few components typically provide a good visual summary of the data. *Tours* also make linear projections of the original variables but offer many different views, like examining the data from different directions. The grand tour shows a smooth sequence of projections as an animation following interpolations between random target bases. The manual radial tour rotates the selected variable’s contribution into and out of a projection. This allows the importance of the variable to structure in the projection to be assessed. This work describes a within-participants user study evaluating the radial tour’s efficacy compared with principal component analysis and the grand tour. A supervised classification task is assigned to participants who evaluate variable attribution of the separation between two classes. Their accuracy in assigning the variable importance is measured across various factors. Data were collected from 108 crowdsourced participants, who performed two trials with each visual for 648 trials in total. Mixed model regression finds strong evidence that the radial tour results in a large increase in accuracy over the alternatives. Participants also reported a preference for the radial tour in comparison to the other two methods.

Index Terms—Linear Dimension Reduction; Visual Analytics; Grand Tour; Data Science; Machine Learning; Data Clustering; Explainable Artificial Intelligence

I. INTRODUCTION

Despite decades of research, multivariate data continues to provide fascinating challenges for visualization. Data visualization is important because it is a key element of exploratory data analysis (EDA, Tukey 1977) for assessing model assumptions and as a cross-check on numerical summarization (Anscombe 1973; Matejka

and Fitzmaurice 2017; Yanai and Lercher 2020). One of the challenges is determining whether a new technique yields a better perception of information than current practices for multivariate data.

Dimension reduction is commonly used with visualization to provide informative low-dimensional summaries of quantitative multivariate data. Principal component analysis (PCA) (Pearson 1901) is one of the first methods ever developed, and it remains very popular. Visualization of PCA is typically in the form of static scatterplots of a few leading components. When accompanied by a representation of the linear combination of the original variables (magnitude and angles of the variable contributions are inscribed on a unit circle), they are called biplots (Gabriel 1971).

Dynamic visualizations called *tours* (Asimov 1985), animate through a sequence of linear projections (orthonormal bases or frames). Instead of a static view, tours provide a smoothly changing view by interpolating between frames. There are various types of tours distinguished by how the paths are generated. Asimov originally animated between randomly selected bases in the *grand* tour. The *manual* tour (Cook and Buja 1997) allows for user control over the basis changes. A selected variable (or component) can be rotated into or out of view or to a particular value. The *radial tour* (Spyrison and Cook 2020) is a variant of the manual tour that fixes the contribution angle and changes the magnitude along the radius. The permanence of the data points from frame to frame holds information between intermediate interpolated frames, and the user control of the basis could plausibly lead to more information being perceived than a static display. This is a hypothesis that a user study could assess.

Empirical studies have rarely assessed tours. An exception is Nelson, Cook, and Cruz-Neira (1999), who compares scatterplots of grand tours on 2D monitors with 3D (stereoscopic, not head-mounted) over $n = 15$ participants. Participants perform cluster detection, dimensionality estimation, and radial sparseness tasks on six-dimensional data. They find that stereoscopic 3D leads to more accuracy in the cluster identification, though the time to interact with the display was much higher in the 3D environment. In this work, we extend the evaluation of tours which compares the radial tour as benchmarked against the grand tour and discrete pairs of principal components.

The contribution of this paper is an empirical user study comparing the radial tour against PCA and the grand tour for assessing variable attribution on clustered data. This is the first empirical evaluation of the radial or manual tour. We discuss how this fits with other multivariate data visualization techniques and coordinated views of linear projections.

We are particularly interested in assessing the effectiveness of the new radial tour relative to common practice with PCA and grand tour. The user influence over a basis, uniquely available in the radial tour, is crucial to testing variable sensitivity to the structure visible in projection. If the contribution of a variable is reduced and the feature disappears, then we say that the variable was sensitive to that structure. For example, Figure 1 shows two frames of simulated data. Panel (a) has identified separation between the two clusters. The contributions in panel (b) show no such cluster separation. The former has a large contribution of V2 in the direction of separation, while it is negligible in the right frame. Because of this, we say that V2 is sensitive to the separation of the clusters.

appendix.rmd ch4_tab1_model_comp_ls.rds data_study data_survey figures R_spyrison-c

Knowing which variables to use is also important for statistical modeling and their interpretations. Models are becoming increasingly complex, and the nonlinear interactions of the terms cause opaqueness to the model's interpretability. Exploratory Artificial Intelligence (XAI, Adadi and Berrada 2018; Arrieta et al. 2020) is an emerging field that extends the interpretability of such black-box models. Multivariate data visualization is essential for exploring features spaces and communicating interpretations of models (Biecek 2018; Biecek and Burzykowski 2021; Wickham, Cook, and Hofmann 2015).

The paper is structured as follows. Section II provides

background on standard visualization methods and linear dimension reduction techniques. Section III describes the experimental factors, task, and accuracy measure used. The results of the study are discussed in Section IV. Conclusions and potential future directions are discussed in Section VI. More results, participant demographics, and analysis of the response time are available in the appendix, Section IX.

II. RELATED WORK

Consider the data to be a matrix of n observations (rows) and p variables (columns), denoted as $X_{n \times p}$.

A. Orthogonal multivariate visualization

Grinstein, Trutschl, and Cvek (2002) illustrate many multivariate visualization methods. In particular, this work shows examples of actual visuals. Liu et al. (2017) give a good classification and taxonomy of such methods. The content below focuses on the most common visuals that use the full data space before discussing linear combinations of those variables in projections.

1) *Scatterplot matrix*: One could consider looking at p histograms or univariate densities. Doing so will miss features in two or more dimensions. A scatterplot matrix (Chambers et al. 1983) is a $p \times p$ matrix typically with univariate densities on the diagonal and all combinations of pairs of variables in off-diagonal elements. Figure 2 shows a scatterplot matrix of the four principal components of simulated data. Such displays do not scale well with dimensions because each plot would get less and less space. Scatterplot matrices can only display information in two orthogonal dimensions, so features in three dimensions may not be fully resolved.

2) *Parallel coordinates plot*: Another common way to display multivariate data is with a parallel coordinates plot (Ocagne 1885), which shows observations by quantile or normalized values for each variable connected by lines to the quantile value in subsequent variables. Parallel coordinates plots scale well with dimensions but poorly with observations as the lines overcrowd the display.

Parallel coordinate plots are asymmetric across variable ordering. In that, shuffling the order of the variable can lead to different conclusions. Another shortcoming is the graphical channel used to convey information. Munzner (2014) suggests that position is the visual channel that is most perceptible to humans. In the case of parallel coordinates plots, the horizontal axes span variables rather than the values of one variable, causing the loss of a display dimension to be used by our most perceptible visual channel.

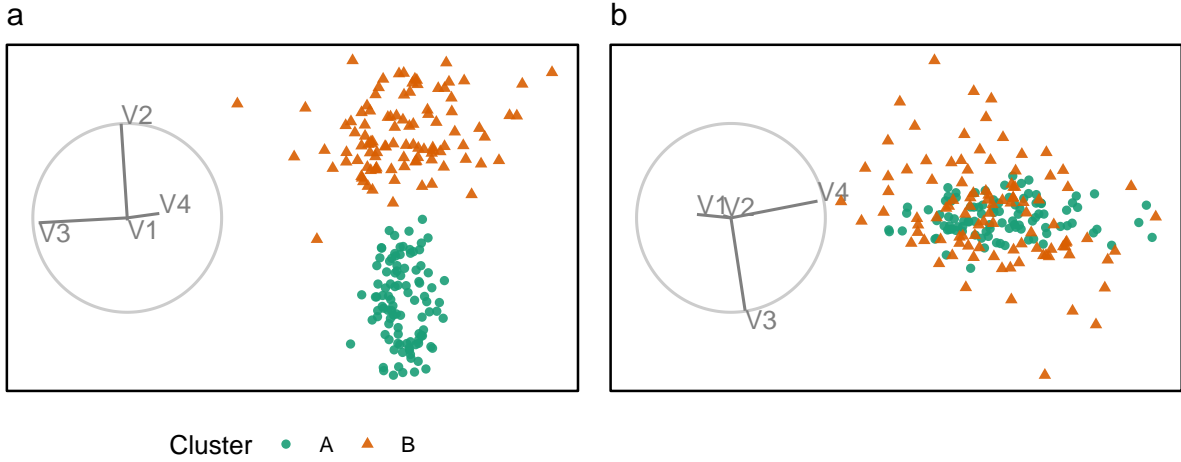


Fig. 1. Illustration of cluster separation affected by variable importance. Panel (a) is a projection mostly of V2 and V3, and the separation between clusters is in the direction of V2, not V3. This suggests V2 is important for clustering, but V3 is not. Panel (b) shows a projection of mostly V3 and V4, with no contribution from V2 and little from V3. That there is no separation between the clusters indicates that V3 and V4 are not important.

B. Multivariate projections

At some point, visualization will be forced to turn to dimension reduction to scale better with the dimensionality of the data. Below we introduce linear projections and the common principal component analysis. Then we touch on nonlinear projections and exclude them from consideration.

1) *Linear*: Let data, X , contain n observations of p variables. A linear projection maps a higher p -dimensional space onto a smaller d -space with an affine mapping (where parallel lines stay parallel). A projection, Y , is the resulting space of the data multiplied by a *basis*, A , such that $Y_{n \times d} = X_{n \times p} \times A_{p \times d}$. This is essentially a reorientation of the original variable. This intuition is conveyed by thinking of a shadow as a 2D projection of a 3D object. Rotating the object changes the shadow it casts and, correspondingly, the basis that maps the reorientation of the object.

2) *Principal component analysis*: PCA is a good baseline of comparison for linear projections because of its frequent and broad use across disciplines. PCA (Pearson 1901) defines new components, linear combinations of the original variables, ordered by decreasing variation through the help of eigenvalue matrix decomposition. While the resulting dimensionality is the same size, the benefit comes from the ordered nature of the components. The data can be said to be approximated by the first several components. The exact number is subjectively selected given the variance contained in each

component, typically guided from a scree plot (Cattell 1966). Features with sizable signal regularly appear in the leading components that commonly approximate data. However, this is not always the case and component spaces should be fully explored to look for signal in components with less variation. This is especially true for cluster structure (Donnell, Buja, and Stuetzle 1994).

3) *Nonlinear*: Nonlinear transformations bend and distort spaces that are not entirely accurate or faithful to the original variable space. There are various quality metrics, such as Trustworthiness, Continuity, Normalized stress, and Average local error, have been introduced to describe the distortion of the space (Espadoto et al. 2021; Gracia et al. 2016). Unfortunately, these distortions are hard to visualize and comprehend, effectively breaking the variable interpretability of the resulting space. The intuition of this can be demonstrated with map projections. Snyder (1987) lists over 200 different projections that distort the surface of the earth to display as a 2D map, each with unique properties and use cases.

Because of the difficulty of interpreting the distortions of nonlinear spaces and the added subjectivity of hyperparameter selection, we exclude nonlinear techniques and instead, decide to compare three linear techniques.

C. Tours, animated linear projections

A *tour* animates through many linear projections. One of the insightful features of the tour is the permanence of the data points; one can track the relative changes of observations as the basis changes, as opposed to

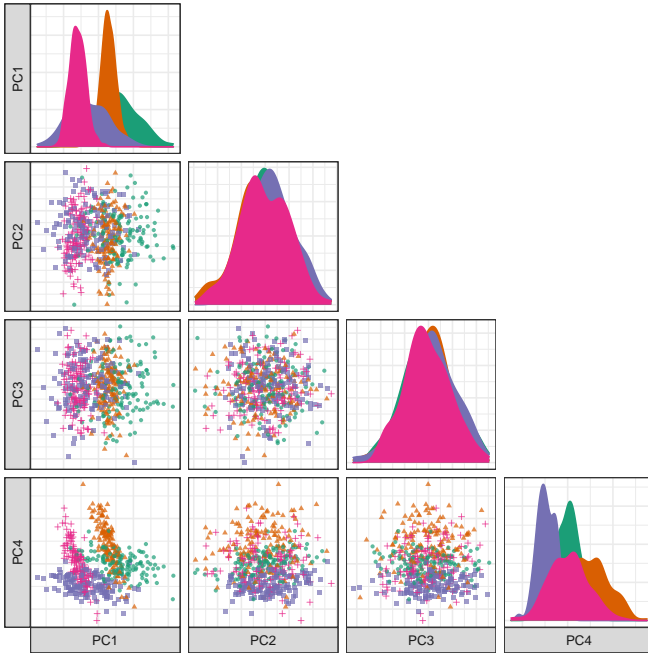


Fig. 2. Scatterplot matrix of the first four principal components of 6D simulated data containing four classes. The separation between classes is primarily in PC1 and PC4. This is not uncommon because PCA is summarizing variance, not cluster structure.

discretely jumping to an orthogonal view angle with no intermediate information. Types of tours are distinguished by the generation of their basis paths (Lee et al. 2021; Cook et al. 2008). To contrast with the discrete orientations of PCA, we compare continuous linear projection changes with grand and radial tours.

1) *Grand tours*: Target bases are selected randomly in a grand tour (Asimov 1985). These target bases are then geodesically interpolated for a smooth, continuous path. The grand tour is the first and most widely known tour. The random selection of target bases makes it a general unguided exploratory tool. The grand tour will make a good comparison that has continuity of data points similar to the radial tour but lacks the user control enjoyed by PCA and radial tours.

2) *Manual and radial tours*: Whether an analyst uses PCA or the grand tour, cannot influence the basis. They cannot explore the structure identified or change the contribution of the variables. User-controlled steering is a key aspect of *manual* tours that helps to test variable attribution.

The manual tour (Cook and Buja 1997) defines its basis path by manipulating the basis contribution of a selected variable. A manipulation dimension is appended onto the projection plane, giving a full contribution to

the selected variable. The target bases are then chosen to rotate this newly created manipulation space. This manipulation space is similarly orthogonally restrained. The data is projected through its interpolated frames and rendered into an animation. When the contribution of one variable changes, the contributions of the other variables must also change, to maintain the orthonormality of the basis. A key feature of the manual tour is that it allows users to control the variable contributions to the basis. Such manipulations can be queued in advance or selected in real-time for human-in-the-loop analysis (Karwowski 2006). Manual navigation is relatively time-consuming due to the vast volume of resulting view space and the abstract method of steering the projection basis. First, it is advisable to identify a basis of particular interest and then use the manual tour as a more directed, local exploration tool to explore the sensitivity of a variable’s contribution to the feature of interest.

To simplify the task and keep its duration realistic, we consider a variant of the manual tour called a *radial* tour. In a radial tour, the magnitude of along the radius with a fixed angle of contribution to the frame, as seen in Figure 3. The radial tour benefits from both continuity of the data alongside grand tours and user-steering via choosing the variable to rotate.

Manual tours have been recently made available in the **R** package **spinifex** (Spyrison and Cook 2020), which facilitates manual tour (and radial variant). It also provides an interface for a layered composition of tours and exporting to gif and mp4 with **gganimate** (Pedersen and Robinson 2020) or html widget with **plotly** (Sievert 2020). It is also compatible with tours made by **tourr** (Wickham et al. 2011). Now that we have a readily available means to produce various tours, we want to see how they fare against traditional discrete displays commonly used with PCA.

D. Other animated linear projections

The work of Elmqvist, Dragicevic, and Fekete (2008) allows users to interactively change the face of a local display by navigating to adjacent faces on a global overview scatterplot matrix. This offers analysts a way to geometrically explore the transition between adjacent faces of a scatterplot matrix as though rotating the face of dice at right angles. The interpolated frames between the orthogonal faces display linear combinations of three variables at varying degrees. This is what McDonald (1982) called a *little tour* with the addition of user control. It is a particular type of manual tour where only horizontal or vertical rotation is allowed.

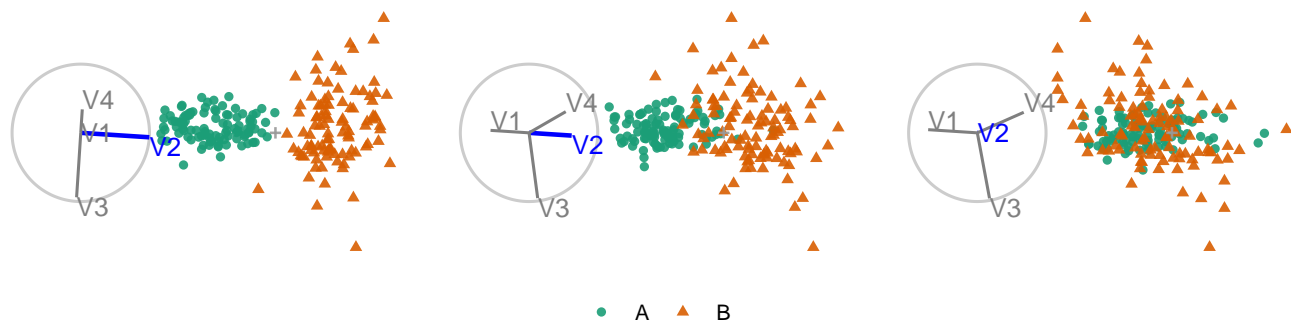


Fig. 3. A radial tour changing the contribution of V2. The contribution is in the direction of cluster separation. When its contribution is removed, the clusters overlap (right). Because of this, we say that V2 is sensitive to the separation of these two species.

Star Coordinates (Kandogan 2000) also arrive at the biplot scatterplot displays starting from the perspective of radial parallel coordinates. Lehmann and Theisel (2013) extend this idea, mapping it back to orthogonal projections. They provide a means to interpolate through PCA components, the orthogonal contributions of scatterplot matrix, and the grand tour. It also defines user-controlled interaction similar to manual and radial tours.

TripAdvisor (Nam and Mueller 2012) is an interactive application that plans sequential interpolation between distant target frames. It also provides an additional global context of a subset of possible frames with glyph representation and an overview of variable attribution by summarizing the top ten principal components. It allows for use steering by using a “touchpad polygon.” This touchpad allows for contribution magnitudes to be changed. This is similar to an incremental change with the manual tour.

E. Empirical evaluation

Some studies compare visualizations across complete contributions of variables. Chang, Dwyer, and Marriott (2018) conducted an $n = 51$ participant study comparing parallel coordinate plots and scatterplot matrix either in isolation, sequentially, or as a coordinated view. Accuracy, completion time, and eye focus were measured for six tasks. Three tasks were more accurate with scatterplot matrix and three with parallel coordinates, while the coordinated view was usually marginally more accurate than the max of the separate visuals. Cao et al. (2018) compare nonstandardized line-glyph and star-glyphs with

standardized variants (with and without fill under the curve). Each of the $n = 18$ participants performed 72 trials across the six visuals, two levels of dimensions, and two levels of observations. Visuals with variable standardization outperformed the nonstandardized variants, and the radial star-glyph reportedly outperformed the line variant.

Other studies have investigated the relative benefits of projecting to 2- or 3D scatterplots in PCA-reduced spaces. Gracia et al. (2016) conducted an $n = 40$ user study comparing 2- and 3D scatterplots on traditional 2D monitors. Participants perform point classification, distance perception, and outlier identification tasks. The results are mixed and primarily have small differences. There is some evidence to suggest a lower error in distance perception from a 3D scatterplot. Wagner Filho et al. (2018) performed an $n = 30$ within-participants study on PCA reduced space using scatterplot displays between 2D on monitors, 3D on monitors, and 3D display with a head-mounted display. None of the tasks on any dataset lead to a significant difference in accuracy. However, the immersive display reduced effort and navigation, resulting in higher perceived accuracy and engagement. Sedlmair, Munzner, and Tory (2013) instead, use two expert coders to evaluate 75 datasets and four dimension reduction techniques across the displays of 2D scatterplots, interactive 3D scatterplots, and 2D scatterplot matrices. They suggest a tiered guidance approach finding that 2D scatterplots are often sufficient to resolve a feature. If not, try 2D scatterplots on a different dimension reduction technique before going to

scatterplot matrix display or conclude a true negative. They find that interactive 3D scatterplots help in very few cases.

F. Conclusion

Orthogonal visualizations either scale poorly with dimensionality or introduce an asymmetry of the variable ordering. Projections visualize the full p -data as fewer dimensions, typically 1-3 at a time. In linear projections, the resulting space is composed of a linear combination of the original variables that maintain variable interpretability. While nonlinear techniques distort and bend space in different ways that are hard to visualize and communicate.

Tours are linear projections that are animated over changes in the basis. There are several more-recent other methods that animate or facilitate manual change the basis that are related to tour techniques. Some quality metrics and empirical studies compare techniques but scarcely with animated methods. Below we conduct a user study to compare the radial tour with PCA and the grand tour on a variable attribution task on clustered data.

III. USER STUDY

The experiment was designed to assess the performance of the radial tour relative to the grand tour and PCA for interpreting the variable attribution to the separation between two clusters. Data were simulated across three experimental factors: location of the cluster separation, cluster shape, and data dimensionality. Participant responses were collected using a web application and crowdsourced through *prolific.co*, (Palan and Schitter 2018) an alternative to MTurk.

A. Objective

PCA will be used as a baseline for comparison as it is the most commonly used linear embedding. It will use static, discrete jumps between orthogonal components. The grand tour will act as a secondary control that will help evaluate the benefit of observation trackability between nearby animation frames but without user-control of its path. Lastly, the radial tour will be compared, which benefits from the continuity of animation and user control of the basis.

Then for some subset of tasks, we expect to find that the radial tour performs most accurately. Conversely, there is less to be certain about the accuracy of such limited grand tours as there is no objective function in selecting the bases; it is possible that the random selection of the target bases altogether avoids the bases

showing cluster separation. However, given that the data dimensionality is modest, and it is probable that the grand tour coincidentally regularly crossed bases with the correct information for the task.

Experimental factors and the definition of an accuracy measure are given below. The null hypothesis can be stated as:

H_0 : accuracy does not change across the visual methods

H_a : accuracy does change across the visual methods

B. Visual factors

The visual methods are tested within-participants, with each visual being evaluated twice by each participant. The order in which experimental factors are experienced is randomized with the assignment, as illustrated in Figure 6. Below discusses the design standardization and unique input associated with each visual.

The visualization methods were standardized wherever possible. Data were displayed as 2D scatterplots with biplots. All aesthetic values (color-blind safe colors, shapes, sizes, absence of legend, and axis titles) were constant. The variable contribution biplot was always shown left of the scatterplot embeddings with their aesthetic values consistent. What did vary between visuals were their inputs.

PCA allowed users to select between the top four principal components for each axis regardless of the data dimensionality (four or six). Upon changing an axis, the visual would change to the new view of orthogonal components without displaying intermediate bases. There was no user input for the grand tour; users were instead shown a 15-second animation of the same randomly selected path (variables containing cluster separation were shuffled after simulation). Participants could view the same clip up to four times within the time limit. Radial tours allowed participants to select the manipulation variable. The starting basis was initialized to a half-clock design, where the variables were evenly distributed in half of the circle. This design was created to be variable agnostic while maximizing the independence of the variables. Selecting a new variable resets the animation where the new variable is manipulated to a complete contribution, zeroed contribution, and then back to its initial contribution. Animation and interpolation parameters were held constant across grand and radial tour (five frames per second with a step size of 0.1 radians between interpolated frames).

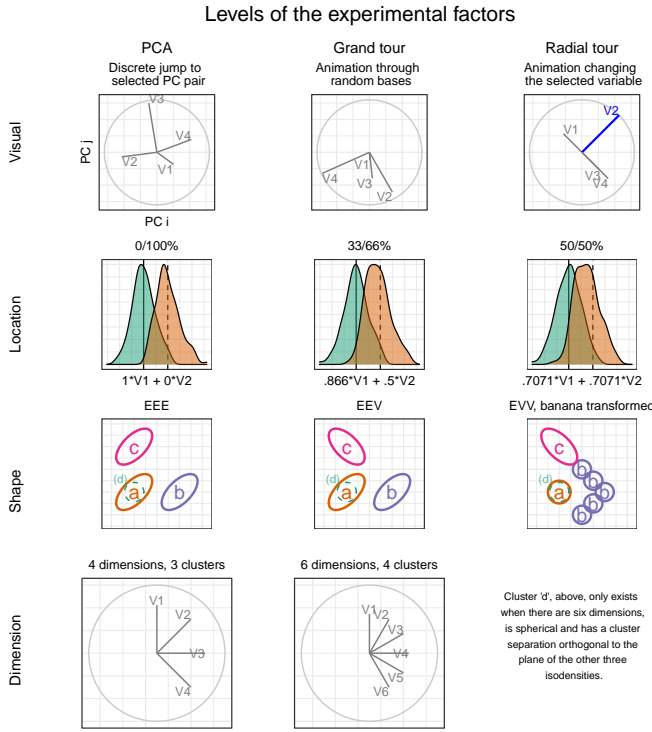


Fig. 4. Levels of the visuals and three experimental factors: location of cluster separation, the shape of clusters, and dimensionality of the sampled data.

C. Experimental factors

In addition to the visual method, data are simulated across three experimental factors. First, the *location* of the separation between clusters is controlled by mixing a signal and a noise variable at different ratios. Secondly, the *shape* of the clusters reflects varying distributions of the data. And third, the *dimension*-ality of the data is also tested. The levels within each factor are described below, and Figure 4 gives a visual representation.

The *location* of the separation between the clusters is at the heart of the measure. It would be good to test a few varying levels. To test the sensitivity, a noise and signal variable are mixed at different ratios. The separation between clusters is mixed at the following percentages: 0/100% (not mixed), 33/66%, 50/50% (evenly mixed).

In selecting the *shape* of the clusters, the convention given by Scrucca et al. (2016) is followed. They describe 14 variants of model families containing three clusters. The model family name is the abbreviation of the clusters' respective volume, shape, and orientation. The levels are either *Equal* or *Vary*. The models EEE, EEV, and EVV are used. For instance, in the EEV model, the volume and shape of clusters are constant, while the

shape's orientation varies. The EVV model is modified by moving four-fifths of the data out in a ">" or banana-like shape.

Dimension-ality is tested at two modest levels: four dimensions containing three clusters and six with four clusters. Such modest dimensionality is required to limit the difficulty and search space to make the task realistic for crowdsourcing.

D. Task and evaluation

With our hypothesis formulated, let us turn our attention to the task and how it is evaluated. Participants were asked to "check any/all variables that contribute more than average to the cluster separation green circles and orange triangles." This was further explained in the explanatory video as "mark any and all variable that carries more than their fair share of the weight, or one quarter in the case of four variables."

The instructions iterated several times in the video were: 1) use the input controls to find a frame that contains separation between the clusters of green circles and orange triangles, 2) look at the orientation of the variable contributions in the grey circle (biplot axes orientation), and 3) select all variables that contribute more than uniform distributed cluster separation in the scatterplot. Independent of the experimental level, participants were limited to 60 seconds for each evaluation of this task. This restriction did not impact many participants as the 25th, 50th, and 75th quantiles of the response time were about 7, 21, and 30 seconds, respectively.

The accuracy measure of this task was designed with a couple of features in mind. 1) symmetric about the expected value, without preference for under- or over-guessing. 2) heavier than linear weight with an increasing difference from the expected value. The following measure is defined for evaluating the task.

Let the data $\mathbf{X}_{n, p, k}$ be a simulation containing clusters of observations of different distributions. Where n is the number of observations, p is the number of variables, and k indicates the observation's cluster. Cluster membership is exclusive; an observation cannot belong to more than one cluster.

The weights, w , is a vector of the variable-wise difference between the mean of two clusters of less $1/p$, the expected cluster separation if it were uniformly distributed. Accuracy, A is defined as the signed square of these weights if selected by the participant. Participant responses are a logical value for each variable — whether

or not the participant thinks each variable separates the two clusters more than uniformly distributed separation.

$$w_j = \frac{(\bar{X}_{\cdot,j=1,k=1} - \bar{X}_{\cdot,1,2}, \dots, (\bar{X}_{\cdot,p,1} - \bar{X}_{\cdot,p,2}))}{\sum_{j=1}^p (|\bar{X}_{\cdot,j,k=1} - \bar{X}_{\cdot,j,2}|)} - \frac{1}{p}$$

$$A = \sum_{j=1}^p I(j) \cdot \text{sign}(w_j) \cdot w_j^2$$

Where $I(j)$ is the indicator function, the binary response for variable j . Figure 5 shows one projection of a simulation with its observed variable separation (wide bars), expected uniform separation (dashed line), and accuracy if selected (thin vertical lines).

E. Randomized factor assignment

Now, with simulation and their artifacts in hand, this section covers how the experimental factors are assigned and demonstrate how this is experienced from the participant's perspective.

The study is sectioned into three periods. Each period is linked to a randomized level of visual and location. The order of dimension and shape are of secondary interest and are held constant in increasing order of difficulty; four then six dimensions and EEE, EEV, then EVV-banana, respectively.

Each period starts with an untimed training task at the simplest remaining experimental levels; location = 0/100%, shape = EEE, and four dimensions with three clusters. This serves to introduce and familiarize participants with input and visual differences. After the training, the participant performs two trials with the same visual and location level across the increasing difficulty of dimension and shape. The plot was removed after 60 seconds, though participants rarely reached this limit.

The order of the visual and location levels is randomized with a nested Latin square where all levels of the visuals are exhausted before advancing to the next level of location. This requires $(3!)^2 = 36$ participants to evaluate all permutations of the experimental factors once. This randomization controls for potential learning effects the participant may receive. Figure 6 illustrates how an arbitrary participant experiences the experimental factors.

Through pilot studies sampled by convenience (information technology and statistics Ph.D. students attending Monash University), it was estimated that three complete evaluations are needed to power the study properly, a total of $N = 3 \times 3!^2 = 108$ participants.

F. Participants

$N = 108$ participants were recruited via prolific.co (Palan and Schitter 2018). Participants are restricted based on their claimed education requiring that they have completed at least an undergraduate degree (some 58,700 of the 150,400 users at the time). This restriction is used on the premise that linear projections and biplot displays will not be regularly used for consumption by general audiences. There is also the implicit filter that Prolific participants must be at least 18 years of age and have implicit biases of timezone, location, and language biases. Participants were compensated for their time at £7.50 per hour, whereas the mean duration of the survey was about 16 minutes. Previous knowledge or familiarity was minimal, as validated in the follow-up survey. The appendix, Section IX-C, contains a heatmap distribution of age and education paneled across preferred pronouns of the participants that completed the survey, who are relatively young, well educated, and slightly more likely to identify as males.

G. Data collection

Data were recorded in a **shiny** application and written to a Google Sheet after each third of the study. Especially at the start of the study, participants experienced adverse network conditions due to the volume of participants hitting the application with modest allocated resources. In addition to this, API read/write limitations further hindered data collection. To mitigate this, the thru-put of participants was throttled, and over-collect survey trials until three evaluations were collected for all permutation levels.

The processing steps were minimal. The data were formatted and then filtered to the latest three complete studies of each experimental factor, which should have experienced the least adverse network conditions. The bulk of the studies removed were partial data and a few over-sampled permutations. This brings us to the 108 studies described in the paper, from which models and aggregation tables were built. The post-study surveys were similarly decoded to human-readable format and then filtered to include only those 84 associated with the final 108 studies.

IV. RESULTS

To recap, the primary response variable is accuracy, as defined in Section III-D. Two primary data sets were collected; the user study evaluations and the post-study survey. The former is the 108 participants with the experimental factors: visual, location of the cluster

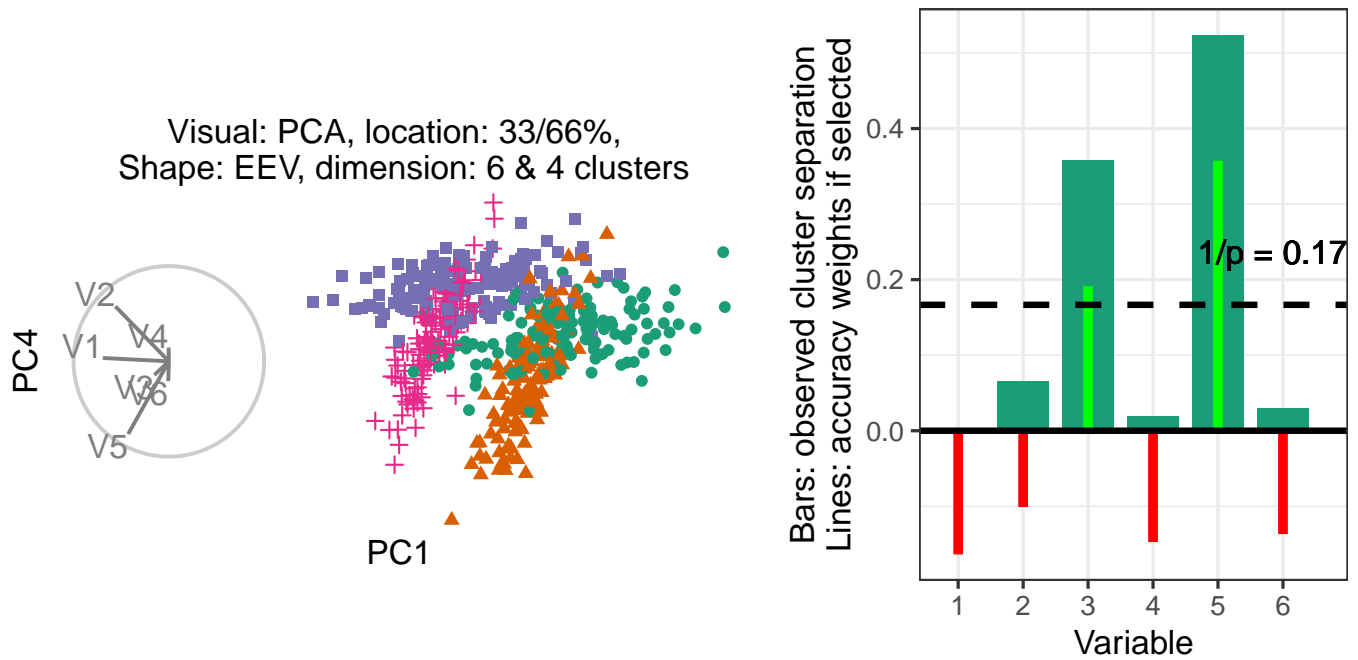


Fig. 5. Illustration of how accuracy is measured. (L), Scatterplot and biplot of PC1 by PC4 of a simulated data set (R) illustrate cluster separation between the green circles and orange triangles. Bars indicate observed cluster separation, and (red/green) lines show the accuracy weight of the variable if selected. The horizontal dashed line has a height $1/p$, the expected value of cluster separation. The accuracy weights equal the signed square of the difference between each variable value and the dashed line.

Randomization for the 63rd participant

- 1) Set the visual order:
 $1 + (63 - 1) \bmod 6 =$
permutation 4;
grand, PCA, & radial

- 2) Set location order:
 $1 + \text{floor}((63 - 1) / 6) \bmod 36 =$
permutation 3; 33/67, 50/50,
& 0/100 % noise/signal mix

Fixed factors:

- 3) Variance-covariance shape increments with each period: EEE, EEV, EVV-banana
- 4) Data dimension is fixed within each period: 4 then 6

visual order permutations	Period 1	Period 2	Period 3	location order permutations	Period 1	Period 2	Period 3
1	P	G	R	1	1	2	3
2	P	R	G	2	1	3	2
3	G	R	P	3	2	3	1
4	G	P	R	4	2	1	3
5	R	P	G	5	3	1	2
6	R	P	G	6	3	2	1

Increment through all permutations of visual, before incrementing 1 permutation of location

Period	Evaluation order	Visual ¹	Location ²	Shape ³	Dimensions ⁴
1	Train 1	Grand	0/100	EEE	4 (3cl)
1	1	Grand	33/67	EEE	4 (3cl)
1	2	Grand	33/67	EEE	6 (4cl)
2	Train 2	PCA	0/110	EEE	4 (3cl)
2	3	PCA	50/50	EEV	4 (3cl)
2	4	PCA	50/50	EEV	6 (4cl)
3	Train 3	Radial	0/100	EEE	4 (3cl)
3	5	Radial	0/100	Ban	4 (3cl)
3	6	Radial	0/100	Ban	6 (4cl)

Fig. 6. Illustration of how a hypothetical participant 63 is assigned experimental factors. Each of the six visual order permutations is exhausted before iterating to the next permutation of location order.

separation signal, the shape of the variance-covariance matrix, and the dimensionality of the data. Experimental factors and randomization were discussed in section III-C. A follow-up survey was completed by 84 of these 108 people. It collected demographic information (preferred pronoun, age, and education) and subjective measures for each visual (preference, familiarity, ease of use, and confidence).

Below a battery of mixed regression models is built to examine the degree of the evidence and the size of the effects from the experimental factors. Then, Likert plots and rank-sum tests to compare the subjective measures between the visuals.

A. Accuracy

To quantify the contribution of the experimental factors to the accuracy, mixed-effects models were fit. All models have a random effect term on the participant and the simulation. These terms explain the amount of error attributed to the individual participant's effect and variation due to the random sampling data.

In building a set of models to test, a base model with only the visual term being compared with the full linear model term and progressively interacting with an additional experimental factor. The models with three and four interacting variables are rank deficient; there is not enough varying information in the data to explain all interacting terms.

Fixed effects	Full model
α	$\hat{Y} = \mu + \alpha_i + \mathbf{Z} + \mathbf{W} + \epsilon$
$\alpha + \beta + \gamma + \delta$	$\hat{Y} = \mu + \alpha_i + \beta_j + \gamma_k + \delta_l + \mathbf{Z} + \mathbf{W} + \epsilon$
$\alpha \times \beta + \gamma + \delta$	$\hat{Y} = \mu + \alpha_i \times \beta_j + \gamma_k + \delta_l + \mathbf{Z} + \mathbf{W} + \epsilon$
$\alpha \times \beta \times \gamma + \delta$	$\hat{Y} = \mu + \alpha_i \times \beta_j \times \gamma_k + \delta_l + \mathbf{Z} + \mathbf{W} + \epsilon$
$\alpha \times \beta \times \gamma \times \delta$	$\hat{Y} = \mu + \alpha_i \times \beta_j \times \gamma_k \times \delta_l + \mathbf{Z} + \mathbf{W} + \epsilon$

where

- μ , the intercept of the model
- α_i , visual term | $i \in (\text{pca, grand, radial})$
- β_j , location term | $j \in (0/100, 33/66, 50/50\% \text{ mix})$
- γ_k , shape term | $k \in (\text{EEE, EEV, EVV, banana})$
- δ_l , dimension term | $l \in (4 \& 3, 6 \& 4) \text{ var \& clusters,}$
- $\mathbf{Z} \sim \mathcal{N}(0, \tau)$, the random effect of participant
- $\mathbf{W} \sim \mathcal{N}(0, \nu)$, the random effect of simulation
- $\epsilon \sim \mathcal{N}(0, \sigma)$, the remaining error of the model

Table I compares the model summaries across increasing complexity. The $\alpha \times \beta + \gamma + \delta$ model is selected to examine in more detail as it has relatively high condition R^2 and not overly complex interacting terms. Table II looks at the coefficients for this model. There is

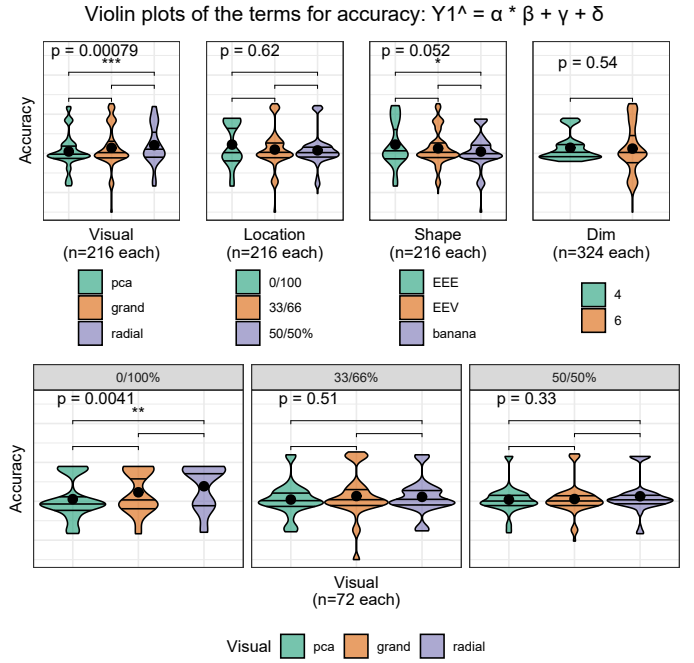


Fig. 7. Violin plots of terms of the model $\hat{Y} = \alpha \times \beta + \gamma + \delta$. Overlaid with global significance from the Kruskal-Wallis test and pairwise significance from the Wilcoxon test, both are non-parametric, ranked-sum tests. Viewing the marginal accuracy of the terms corroborates the primary findings that the use of the radial tour leads to a significant increase in accuracy, at least over PCA, and this effect is particularly well supported when no location mixing is applied.

strong evidence suggesting a relatively large increase in accuracy from the radial tour, though there is evidence that almost of increase is lost under 33/66% mixing.

We also want to visually examine the conditional variables in the model. Figure 7 illustrates the violin plots of accuracy for each model term.

B. Subjective measures

Modeling has shown us that use of the radial tour leads to a sizable improvement in the accuracy measure for this task. This is not the whole story. It is desirable to know what the users think of using the visuals. We follow the lead set by Wagner Filho et al. (2018). They observe four subjective measures. The following were used in this study: confidence, ease of use, prior familiarity, and preference. Each of these questions were asked of all for each visual as 5-point Likert items.

The 84 evaluations of the post-study survey are shown in Figure 8. The figure uses Likert plots or stacked percentage bar plots, alongside violin plots with the same non-parametric, ranked sum tests previously used.

There was strong evidence to support that participants preferred the radial tour to either alternative. There is

TABLE I

MODEL PERFORMANCE OF RANDOM EFFECT MODELS REGRESSING ACCURACY. COMPLEX MODELS PERFORM BETTER IN TERMS OF R^2 AND RMSE, YET AIC AND BIC PENALIZE THEIR LARGE NUMBER OF FIXED EFFECTS IN FAVOR OF THE MUCH SIMPLER MODEL CONTAINING ONLY THE VISUALS. CONDITIONAL R^2 INCLUDES ERROR EXPLAINED BY THE RANDOM EFFECTS, WHILE MARGINAL DOES NOT.

Model	Terms	Lvls	AIC	BIC	R2 cond.	R2 marg.	RMSE
a	1	3	-71	-44	0.303	0.018	0.194
a+b+c+d	4	8	-45	4	0.334	0.056	0.194
a*b+c+d	5	12	-26	41	0.338	0.064	0.193
a*b*c+d	8	28	28	167	0.383	0.108	0.19
a*b*c*d	15	54	105	360	0.37	0.222	0.185

TABLE II

THE TASK ACCURACY MODEL COEFFICIENTS FOR $\hat{Y} = \alpha \times \beta + \gamma + \delta$, WITH VISUAL = PCA, LOCATION = 0/100%, SHAPE = EEE, AND DIM = 4 HELD AS BASELINES. VISUAL BEING RADIAL IS THE FIXED TERM WITH THE STRONGEST EVIDENCE SUPPORTING THE HYPOTHESIS. INTERACTING WITH THE LOCATION TERM, THERE IS EVIDENCE SUGGESTING RADIAL PERFORMS WITH MINIMAL IMPROVEMENT FOR 33/66% LOCATION MIXING.

	Estimate	Std. Error	df	t value	Prob	
(Intercept)	0.10	0.06	16.1	1.54	0.143	
Factor						
Visualgrand	0.06	0.04	622.1	1.63	0.104	
Visualradial	0.14	0.04	617.0	3.77	0.000	***
Fixed effects						
Location33/66%	-0.02	0.07	19.9	-0.29	0.777	
Location50/50%	-0.04	0.07	20.0	-0.66	0.514	
ShapeEEV	-0.05	0.06	11.8	-0.82	0.427	
Shapebanana	-0.09	0.06	11.8	-1.54	0.150	
Dim6	-0.01	0.05	11.8	-0.23	0.824	
Interactions						
Visualgrand:Location33/66%	-0.02	0.06	588.9	-0.29	0.774	
Visualradial:Location33/66%	-0.12	0.06	586.5	-2.13	0.033	*
Visualgrand:Location50/50%	-0.03	0.06	591.6	-0.47	0.641	
Visualradial:Location50/50%	-0.06	0.06	576.3	-1.16	0.248	

less evidence that the radial tour led to more confidence and found easier to use than the grand tour. In confirmation of expectations, crowdsourced participants had low familiarity with all visuals, with no difference in mean supported.

V. DISCUSSION

Data visualization is an integral part of understanding relationships in data and how models are fitted. When it comes to multivariate data giving a comprehensive view quickly becomes difficult as the dimensions become sizable. Developers and analysts have the task of choosing which visualization technique to use. We have discussed orthogonal visualization, observation-link glyphs, and then focus on linear projections of data. There are many empirical evaluations comparing visualization and display dimensionality for multivariate data. However,

animated techniques are rarely compared in the literature. This user study is the first to compare the radial tour.

The radial tour is a method for the analyst to choose a variable to alter its contribution to the basis. The animation over small changes to the basis allow the sensitivity of the structure to be assessed from the variable contribution. The hypothesis that user control over the basis and the permanence of observations between intermediate frames may lead to better perception of the variable attribution causing the separation of clusters.

Mixed modeling analysis of the study provides strong support for this conclusion. That is, there is significant evidence suggests the use of the radial tour leads to a sizable increase in accuracy. One unexpected caveat is that mixing the location of the signal at 33/66% almost completely negates this gain. Perhaps this is because the “half-clock” basis used did not give enough weight to

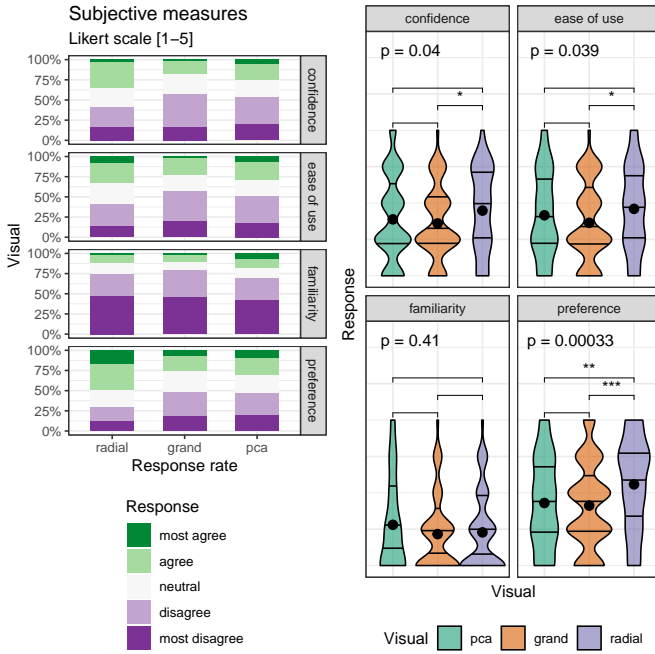


Fig. 8. The subjective measures of the 84 responses of the post-study survey with five-point Likert items levels of agreement. (L) Likert plots (stacked percent bar plots) with (R) violin plots of the same measures. Violin plots are overlaid with global significance from the Kruskal-Wallis test and pairwise significance from the Wilcoxon test. Participants are more confident using the radial tour and find it easier to use than the grand tour. The radial tour is the most preferred visual.

the variable containing the small fraction. It was also interesting to note that no level of the experimental factors alone had a significant effect on this setup. Lastly, the follow-up survey asked participants to evaluate measures of the visuals. Most notably, participants preferred the radial tour to the other visuals. Knowing that the radial tour outperforms alternatives and is the preferred choice can help inform the selection of visual methods for developers and analysts.

VI. CONCLUSION

This paper discussed a crowdsourced within-participant user study ($n = 108$) comparing the efficacy of three linear projection techniques: PCA, grand tour, and radial tour. The participants performed a supervised cluster task, explicitly identifying which variables contribute to the separation of two target clusters. This was evaluated evenly over four experimental factors. In summary, mixed model regression finds strong evidence that using the radial tour sizably increases accuracy, especially when cluster separation location is not mixed at 33/66%. The effect sizes on accuracy are large relative to the change from the other experimental

factors and the random effect of data simulation, though smaller than the random effect of the participant. The radial tour was the most preferred of the three visuals.

There are several implicit limitations to this study: the task, type of data, levels of the factors selected. There are several ways that this study could be extended. In addition to expanding the support of the experimental factors, more exciting directions include: introducing a new task, including more visualizations, and changing the experience level of the target population. It is difficult to achieve good coverage given the number of possible experimental factors.

VII. ACKNOWLEDGMENTS

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IX. APPENDIX

This section covers extended analysis. First, it illustrates of the different visuals are provided. Then, the participant demographics are covered. Lastly, a parallel modeling analysis on log response time is conducted.

A. Data simulation

Each dimension is distributed initially as $\mathcal{N}(0,1)$, given the covariance set by the shape factor. Clusters were initially separated by a distance of two before location mixing. Signal variables had a correlation of 0.9 when they had equal orientation and -0.9 when their orientations varied. Noise variables were restricted to zero correlation. Each cluster is simulated with 140 observations and is offset in a variable that did not distinguish previous variables.

Clusters of the EVV shape are transformed into the banana-chevron shape (illustrated in Figure 4, shape row). Then location mixing is applied by post-multiplying a rotation matrix to the signal variable and a noise variable for the clusters in question. All variables are then standardized by standard deviations away from the mean. The columns are then shuffled randomly.

Each of these replications is then iterated with each level of the visual. For PCA, projections were saved (to `png`) for the 12 pairs of the top four principal components. A grand tour basis path is saved for each dimensionality level. The data from each simulation is then projected through its corresponding bases path and saved to `gif` file. The radial tour starts at either the four or six-variable “half-clock” basis. A radial tour is then produced for each variable and saved as a `gif`.

B. Illustrations of visual methods

Below illustrates the three visual methods evaluated in the user study. Data was collected from a **shiny** application, and pre-rendered `gif` files were displayed based on the selected inputs.

C. Survey participant demographics

The target population is relatively well-educated people, as linear projections may prove difficult for generalized consumption. Hence, Prolific.co participants are restricted to those with an undergraduate degree (58,700 of the 150,400 users at the study time). From this cohort, 108 performed a complete study. Of these participants, 84 submitted the post-study survey, represented in the following heatmap. All participants were compensated for their time at £7.50 per hour, with a mean time of about 16 minutes. Figure 10 shows a heat map of the demographics for these 84 participants.

D. Response time

As a secondary explanatory variable, response time is considered. Response time is first log-transformed to remove its right skew. The same modeling procedure is repeated for this response. 1) Compare the performance of progressively more complex models. Table III shows the higher level performance of these models over increasing model complexity. 2) Select the model with the same effect terms, $\alpha \times \beta + \gamma + \delta$, with relatively high conditional R^2 without becoming overly complex from variable interactions. The coefficients of this model are displayed in Table IV.

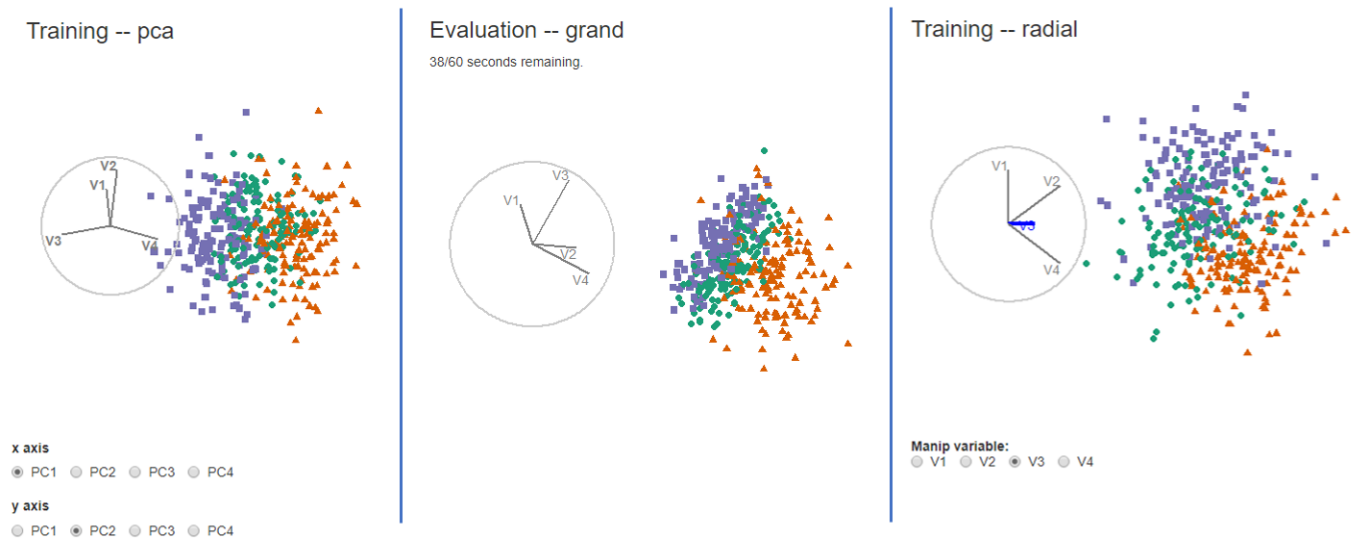


Fig. 9. Examples of the application displays for PCA, grand tour, and radial tour.

TABLE III

MODEL PERFORMANCE REGRESSING ON LOG RESPONSE TIME [SECONDS], \widehat{Y}_2 RANDOM EFFECT MODELS. CONDITIONAL R^2 INCLUDES THE RANDOM EFFECTS, WHILE MARGINAL DOES NOT. THE MODEL $\alpha \times \beta + \gamma + \delta$ MODEL IS SELECTED TO EXAMINE FURTHER AS IT HAS RELATIVELY HIGH MARGINAL R^2 WHILE HAVING MUCH LESS COMPLEXITY THAN THE COMPLETE INTERACTION MODEL.

Model	Terms	Lvls	AIC	BIC	R2 cond.	R2 marg.	RMSE
a	1	3	1448	1475	0.645	0.007	0.553
a+b+c+d	4	8	1467	1516	0.647	0.017	0.552
a*b+c+d	5	12	1474	1541	0.656	0.024	0.548
a*b*c+d	8	28	1488	1627	0.673	0.054	0.536
a*b*c*d	15	54	1537	1792	0.7	0.062	0.523

TABLE IV

MODEL COEFFICIENTS FOR LOG RESPONSE TIME [SECONDS] $\widehat{Y}_2 = \alpha \times \beta + \gamma + \delta$, WITH FACTOR = PCA, LOCATION = 0/100%, SHAPE = EEE, AND DIM = 4 HELD AS BASELINES. LOCATION = 50/50% IS THE FIXED TERM WITH THE STRONGEST EVIDENCE AND TAKES LESS TIME. IN CONTRAST, THE INTERACTION TERM LOCATION = 50/50%:SHAPE = EEV HAS THE MOST EVIDENCE AND TAKES MUCH LONGER ON AVERAGE.

	Estimate	Std. Error	df	t value	Prob	
(Intercept)	2.71	0.14	42.6	19.06	0.000	***
Factor						
Visualgrand	-0.23	0.12	567.6	-1.97	0.049	*
Visualradial	0.16	0.12	573.5	1.34	0.181	
Fixed effects						
Location33/66%	0.05	0.14	40.9	0.34	0.737	
Location50/50%	-0.05	0.14	42.1	-0.35	0.729	
ShapeEEV	-0.15	0.09	8.3	-1.61	0.145	
Shapebanana	-0.13	0.09	8.3	-1.42	0.192	
Dim6	0.14	0.08	8.3	1.90	0.093	
Interactions						
Visualgrand:Location33/66%	0.24	0.18	580.9	1.34	0.181	
Visualradial:Location33/66%	-0.24	0.18	582.4	-1.32	0.188	
Visualgrand:Location50/50%	0.12	0.18	578.6	0.69	0.491	
Visualradial:Location50/50%	0.05	0.18	584.4	0.25	0.800	

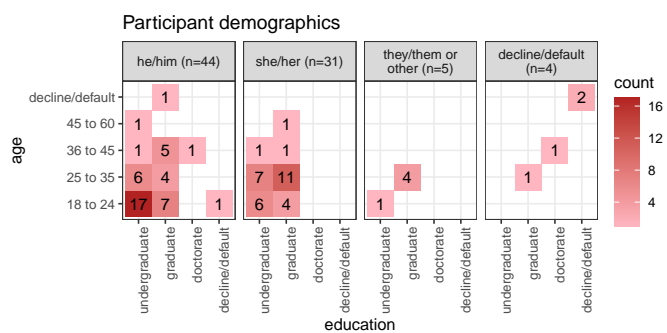


Fig. 10. Heatmaps of survey participant demographics; counts of age group by completed education as faceted across preferred pronoun. Our sample tended to be between 18 and 35 years of age with an undergraduate or graduate degree.