

The effect of user interaction for understanding variable contributions to structure in linear projections

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Abstract

Viewing data in its original variable space is fundamental to the exploratory data analysis. For multivariate data this is an complex task. We perform a between-participant user study to evaluate 3 types of linear embeddings, namely, biplots of principal components, grand tours, and radial tours. Crowdsourced participants ($N = 108$) were asked to identify which variable(s) explain the difference between a pair of clusters within the data as the factor of visualization is changed within participant. Principal components and grand tour are the discrete and animated benchmarks to beat. We find radial tours score higher and respond slightly faster than alternatives. Visual factor is significantly more important than the other block parameterizations (location, shape, dimension) or learning effects from the order of evaluation. Demographics, prior experience, and parameter values vary in significance, while the size of their coefficients all relatively small as compared with visual and the random effect of the participants.

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1 Introduction

Multivariate data is ubiquitous. Yet exploratory data analysis (EDA) (Tukey 1977) of such spaces becomes difficult, increasingly so as dimension increases. Numeric statistic summarization of data often doesn't explain the full complexity of the data or worse, can lead to missing obvious visual patterns (Anscombe 1973; Matejka and Fitzmaurice 2017; Goodman 2008; Coleman 1986). Data should be visually inspected in it's original variable-space before applying models or summarizations. This allows users to validate assumptions, identify outliers, and facilitates the identification of visual peculiarities.

For these reasons, it is important to use visualizations of data spaces and extend the diversity of its application. However, visualizing data containing more than a handful of variables is not trivial. Scatterplot matrices or small multiples (Chambers et al. 1983) looks at all permutation pairs of variables, but quickly becomes to vast a number of images to consider. On the other extreme, parallel coordinates plot (Ocagne 1885) and its radial variants, plot observations as lines varying across scaled variables as displayed in a line or circle. This scales well with dimensionality, while suffering from couple issues. The larger issue, being the loss of mapping multiple variables to graphic position, which is perhaps the most important visual cue for human perception (Munzner 2014). The lesser being that they suffer from asymmetry, as their interpretation is dependent on variable ordering.

Using a linear combinations of variables will allow us to keep position in 2 display axes while peering into information not contained in any one dimension. The idea of using a combination of variables may appear daunting at first, however we do it almost exclusively in the spatial dimensions. That is to say we are rarely completely aligned with rectangular objects at any one point in time. Consider a book or a filing cabinet any orientation that isn't fully a 2D rectangle, you are seeing as a linear combination of its variables, height, width, and depth. Generalizing this to arbitrary dimensionality we can project or embed a 2D profile of p -dimensional data. Its worth noting that the number of these embedded profiles, and consequently the time it takes to explore them, increase exponentially with the dimensionality of the data.

Non-linear embeddings, the compliment of the linear embedding, have also been well received recently especially with the emergence of t-Distributed stochastic neighbor embedding (Maaten and Hinton 2008). Such techniques distort the fully dimensionality on to a low, typically 2D plane. The issue with doing so is that unit of distance is not consistent with location in the embedded space, which severely hinders the interoperability of these embeddings. Additionally they often have hyperparameters that need tuning. Doing so results in completely different or contradicting embeddings. Suffice it to say we exclude their consideration for such broad application for multivariate EDA.

Additionally there are many methods for supervised data, that is, with observation classes known a priori. Linear discriminant analysis (Fisher 1936), for instance, is an example of linear projection, which creates it's components based on the mean separation and covariance shape between known classes. Because we which to extrapolate to generalized EDA of multivariate space we will not use any such supervised techniques to initialize components in the study.

In multivariate spaces, performance measures, and computational complexity are regularly compare to analogous algorithms and models. Human perception and inference from visuals is notably missing. We perform a within-participant, crowd sourced user study exploring the efficacy of 3 methods of linear embedding visualizations.

Section 2 discusses the visualization methods. Section 3 goes into the user study. The subsection 3.2 digs into the task and its evaluation. The results of the study are in section 4. Discussion is covered in section 5. An accompanying tool is discussed in section 6.

2 Background, visual methods

2.1 Linear projection notation

Consider a numeric data matrix with n observations of p variables,

$$\mathbf{X}_{[n,p]} = (\mathbf{x}_1, \dots \mathbf{x}_p)$$

$$\mathbf{x}_i = (x_{1i}, \dots x_{ni}) \mid i \in [1, p]$$

Let $\mathbf{Y}_{[n,d]}$ be the d -dimensional projection or embedding of $\mathbf{X}_{[n,p]}$ via matrix multiplication of a particular orthonormal basis matrix $\mathbf{B}_{[p,d]}$.

$$\mathbf{Y}_{[n,d]} = \mathbf{X}_{[n,p]} \mathbf{B}_{[p,d]} \mid \mathbf{B} \text{ is orthonormal}$$

$$\mathbf{y}_j = (y_{1j}, \dots y_{nj}) \mid j \in [1, d]$$

A matrix is said to be orthonormal if and only if they are 1) orthogonal, that is all column pairs are independent, having a cross product of 0, and 2) normal, each columns has a norm - Euclidian distance - of 1.

2.2 Principal Component Analysis

Considering that we want to explore multivariate data space, while maintaining position mapping of points. Linear combinations of variables becomes an ideal candidate. Principal component analysis (PCA) (Pearson 1901) creates new components that are linear combinations of the original variables. The creation of these variables is ordered by decreasing variation which is orthogonally constrained to all previous components. while the full dimensionality is in tact the benefit comes from the ordered nature of the components. For instance if nearly all of the variation in a data-space can be explained in the first half of its components than the complexity of viewing such a space is exponentially simplified.

2.3 Grand Tours

Later, Asimov (Asimov 1985), coined data visualization *tour*, an animation of many linear projections across local changes in the basis. One of key features of the tour is the object permanence of the data points. That is to say by watching near by, orthogonally-interpolated frames one can track the relative changes of observations as variable contributions change.

Asimov originally purposed the *grand* tour. To start, several target bases are randomly selected. These target bases are then orthogonally-interpolated between with a fixed target distance between interpolation frames. The data matrix is premultiplied to the array of interpolated bases and rendered into an animation. There is no user interaction in a grand tour and the target.

2.4 Manual tours

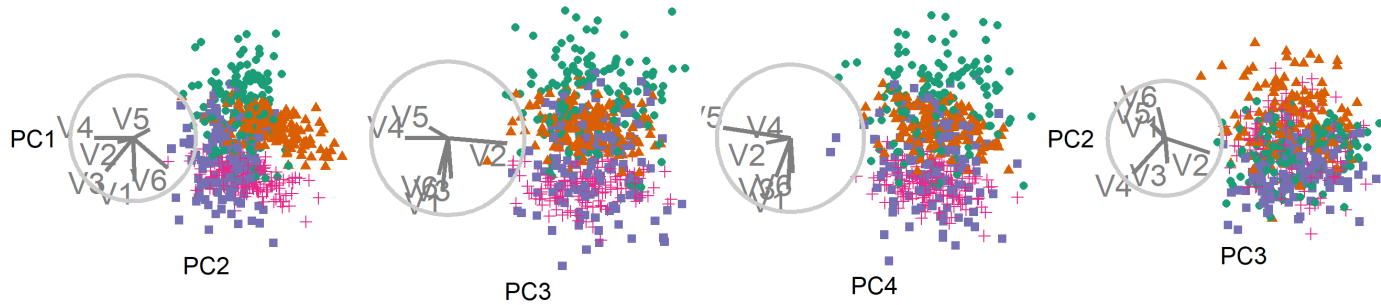
The *manual* tour (Cook and Buja 1997; Spyrisson and Cook 2020) defines its basis path by manipulating the basis contribution of a selected variable. A manipulation dimension is appended onto the projection plane, with a full contribution given to the selected variable. The target bases are then selected based on rotating this newly created manipulation space. The target bases are then similarly orthogonally-interpolated, data projected, and rendered into an animation. In order for variables to remain independent of each other the contributions of the other variables must also change, *ie.* dimension space should maintain its orthonormal structure. A key feature of the manual tour is that it affords users a way to control the variable contributions of the next target basis. This means that such manipulations can be selected and queued in advance or select on the spot for human-in-the-loop analysis (Karwowski 2006). However, this navigation is relatively time-consuming due to the huge volume of p -space (an aspect of the curse of dimensionality (Bellman 1957)) and the abstract method of steering the projection basis. It is advisable to first identify a basis of particular

interest and then use a manual tour as a finer, local exploration tool to observe how the contributions of the selected variable does or does not contribute to the feature of interest.

In order to simplify the task and keep its duration realistic we consider a variant of the manual tour, called a *radial* tour. In a radial tour the selected variable is allowed to change its magnitude of contribution, but not its angle; it must move along the direction of its original contribution radius.

PCA

- Inputs: x, y axes in [PC1, ... PC4]
- Not animated, discrete change
- 4 of the 12 unique PC combinations



Grand

- Inputs: none
- Animated through randomly selected target bases
- First 4 such target bases



Radial

- Inputs: manipulation variable in [1, ... 6]
- Animates selected variable to norm=1, norm=0, then back to start
- Target bases rotating variable 6

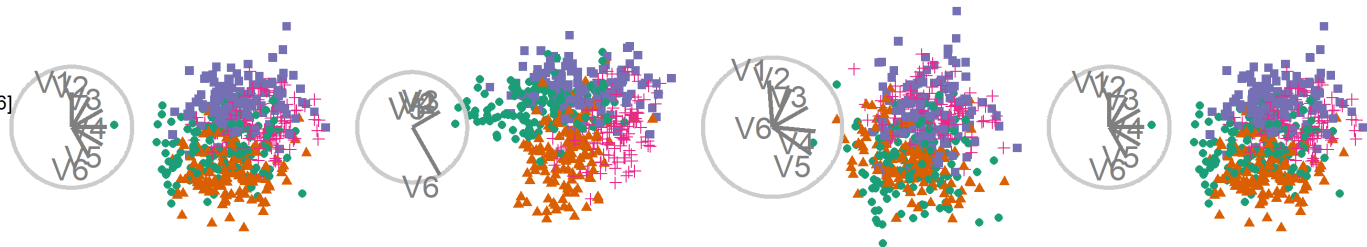


Figure 1: Example of the different visual factors. All use the same sort of biplot display to view linear projections of multivariate data. They differ in which bases are viewed which is influenced by the different factor inputs and whether or not are animated to convey the continuity of data points from 1 frame to the next.

3 User study

3.1 Hypothesis

Does the animated removal of single variables via the radial tour improve the ability of the analyst to understand the importance of variables contribution to the separation of clusters?

PCA will be used as a baseline for compared as it is the most common linear embedding. The grand tour will act as a secondary control that will help evaluate the benefit of animation, where the object permanence of the data points, but without the ability influence it's path. Lastly the radial tour should perform best as it benefits both from animation and being able to select an individual variable to change the contribution of.

3.2 Task and evaluation

The display was a 2D scatterplot with observations supervised. Cluster membership was mapped to shape and color. There were either 3 or 4 clusters each with the number of observations within each cluster. Participants were asked to 'check any/all variables that contribute more than average to the cluster separation green circles and orange triangles,' which was further explained in the explanatory video as 'mark and and all variable that carry more than their fair share of the weight, or 1 quarter in the case of 4 variables.'

The instructions iterated several times in the video was: 1) Use the input controls to find a frame that contains separation between the clusters of green circles and orange triangles, 2) look at the orientation of the variable contributions in the gray circle, a visual depiction of basis, and 3) select all variables that contribute more than average in the direction of the separation in the scatterplot. Regardless of factor and block values participants were limited to 60 seconds for each evaluation of this task.

The evaluation measure of this task was designed to have a few of features: 1) the sum of squares of the individual variable marks should be 1, and 2) symmetric about 0. With these in mind we define the following measure for evaluating the task:

Let a dataset \mathbf{X} be a simulation containing clusters of observations of different distributions. Let \mathbf{X}_k be the subset of observations in cluster k containing the p variables.

$$\begin{aligned}\mathbf{X}_{[n, p]} &= (x_1, \dots x_p) \\ \mathbf{X}_{[n_k, p]k} &= (x_1, \dots x_p) \mid n_k \in [1, n], \text{ is an observation subset of } \mathbf{X}\end{aligned}$$

where

$x_{i,j,k}$ is scalar; the observation $i \in (1, \dots n)$, of variable $j \in (1, \dots p)$, of cluster $k \in (1, \dots K)$

We define weights, W to be a vector explaining the variable-wise difference between 2 clusters. Namely the difference of each variable between clusters, as a proportion of the total difference, less $1/p$ the amount of different each variable would hold if it were uniformly distributed.

$$\begin{aligned}W &= \frac{(\overline{X_{j=1,k=1}} - \overline{X_{j=1,k=2}}, \dots (\overline{X_{j=p,k=1}} - \overline{X_{j=p,k=2}}))}{\sum_{j=1}^p (|\overline{X_{j,k=1}} - \overline{X_{j,k=2}}|)} - \frac{1}{p} \\ &= (w_1, \dots w_p)\end{aligned}$$

Participant responses, R are a vector of logical values, whether or not participant thinks the variable separates the two clusters more than if the difference uniformly distributed. Then M is a vector of variable marks.

$$M = I(r_j) * \text{sign}(w_j) * \sqrt{|w_j|}$$

$$= (m_1, \dots, m_p)$$

where I is the indicator function. Then the total marks for this task is the sum of this marks vector.

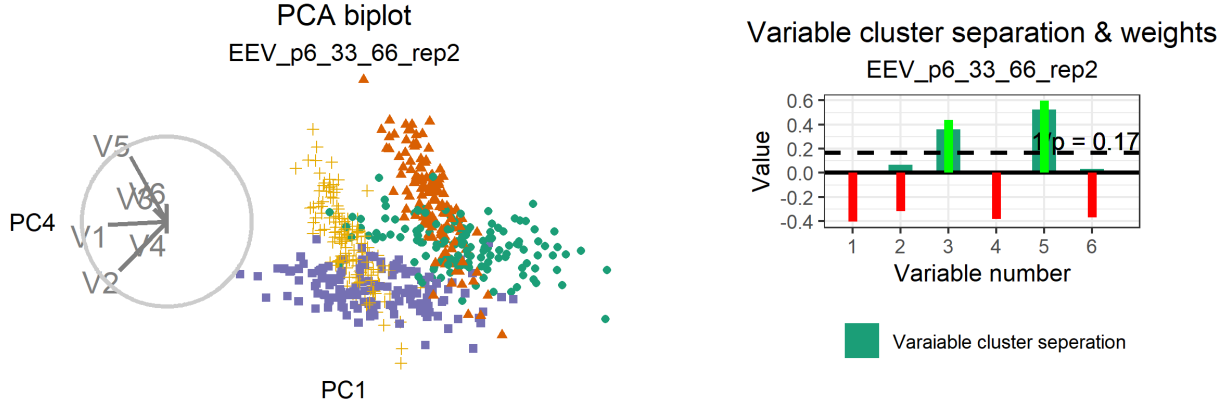


Figure 2: (L), PCA biplot of the components showing the most cluster separation with (R) Illustration of how much of the cluster separation is for each variable (wide bar) and the weight of the marks given if a variable is selected (red/green line). The horizontal dashed line is $1 / \text{dimensionality}$, the amount of separation each variable would have if evenly distributed. The weights equal the signed square of the difference between each variable value and the dashed line.

Each of the 3 periods introduced a new factor, where participants were first able to explore an untimed task with data under the simplest parameterizations of block. The training allows the participant to become familiar with the inputs and visual specific to the factor. Upon clicking a button to proceed text containing the correct answer displays with visual still intact to explore further. After the training, participant performed 2 evaluation trials. After 60 second the display was removed, though in application, few participants elapsed this time. These evaluation trials were performed under different parameterizations as explained in section 3.4.

3.3 Factor application

Section 2 gives the sources and a description of the visual factors PCA, grand tours, and radial manual tours. The factors are tested within participant, with each factor being evaluated by each participant. The order that factors are experienced in is controlled with the block assignment as illustrated below in Figure ??fig:figParmeterizationExample). Below we cover the aesthetic standardization, as well the unique input and display within each factor.

The visualization methods were standardized wherever possible. each factor was shown as a biplot, with variable contributions displayed on a unit circle. All aesthetic values (colors, shapes, sizes, absence of legend, and absence axis titles) were held consistent. Variable contributions were always shown left of the scatterplot embeddings with their aesthetic values consistent as well. What did vary between factors were their inputs which caused a discrete jump to another pair or principal components, were absent for the grand tour with target bases to animate through selected at random, or for radial tour which variable should have its contribution animated. Key frames have of each factor have been illustrated above in Figure 1.

PCA inputs allowed for users to select between the top 4 principal components for both the x and y axis regardless of the data dimensionality (either 4 or 6). There were no user input for grand tour, users were instead shown a 15 second animation of the same randomly selected path. Users were able to view the same clip up to 4 times within the time limit. Radial tours were also displayed at 5 frames per second with in interpolation step size of 0.1 radians. Users where able to swap between variables, upon which the display would change the the start of radially increasing the contribution of the selected variable till it was full, zeroed, and then back to its initial contribution. The complete animation of any variable takes about 20 seconds, and is almost fully in the projection frame at around 6 second. The starting basis of each is initialized to a half-clock design, where the variables were evenly distributed in half of the circle which is then orthonormalized. This design was created to be variable agnostic while maximizing the independence of the variables.

3.4 Blocks and parameterization

In addition to factor we vary the data across 3 aspects: 1) The dimensionality of the data. 2) the shapes of the clusters, by changing the variance-covariance of the clusters. 3) The location of the difference between clusters, by mixing a signal and a noise variable at different ratio we control which variables contain cluster separation.

Dimensionality is tested at 2 modest levels, namely, in 4 dimensions containing 3 clusters and 6 dimensions with 4 clusters. Each cluster samples 140 observations. Each dimension is originally distributed as $\mathcal{N}(2 * I(signal), 1) \mid \text{covariance } \Sigma$, before applying location mixing and standardizing by standard deviation). Signal variables have correlation 0.9 when they have equal orientation and -0.9 when their orientations vary. Noise variables were restricted to 0 correlation. Within each factor-period dimension is fixed with increasing difficulty, 4 then 6.

For choosing the shape of the clusters we follow the convention given by scrucce, (Scrucce et al. 2016) who named and categorize 14 variants of model families containing for 3 clusters. The name of the model family is the abbreviation of its respective volume, shape, and orientation, which are either equal or vary. We use the models EEE, EEV, and EVV, the latter is further modified by moving 4 fifths of the data out in an “V” or banana-like shape. Figure 3 shows the principal component biplot of the 3 model variants applied here. The training always uses 4 dimensions, while the 2 evaluations always contain 4 and 6 dimensions in order of increasing difficulty. The evaluation periods use EEE, EEV, and EVV-banana respectively in increasing order of difficulty.

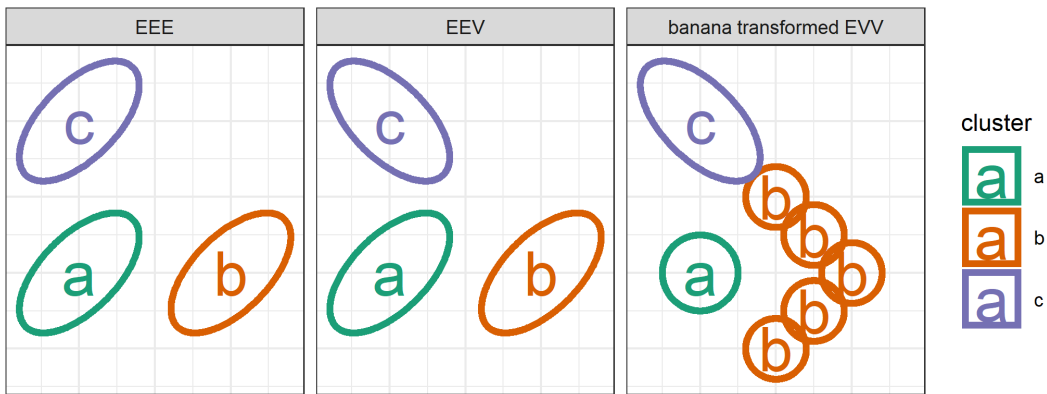


Figure 3: Ellipses of the isodensities of the model families used. Family labels are are the abbreviation for the clusters volume, shape, and orientation respectively, which are either equal or vary. We futher change the EVV model by shifting fifths of the data in banana or chevron arrow shape.

The separation of any pair of clusters is currently contained fully within a single variable at this point. In order to test the sensitivity to this we mix a noise variable with the signal-containing variable such that the

difference in the clusters is mixed at the following percentages: 0/100 (not mixed), 33/66, 50/50 (evenly mixed). The training always uses 4 dimensions, while the 2 evaluations always contain 4 and 6 dimensions in order of increasing difficulty. The training data does not mix signal Location mixing within an evaluation period is held constant and rotated through the 6 permutations of their order. Randomizing the order of the location mixing is controlled by iterating once after each of the 6 factor order permutations are evaluated. This is illustrated in Figure 4.

Consider a new participant, the 63rd participant,

- 1) Set the factor order:
 $1 + (63 - 1) \bmod 6 =$
 permutation 4;
 grand, PCA, & radial

- 2) Set location order:
 $1 + \text{floor}((63 - 1) / 6) \bmod 36 =$
 permutation 3; 33/67, 50/50, &
 0/100 percent noise/signal mix

Fixed blocks:

- 3) Variance-covariance shape increments with each period: EEE, EEV, EVV-banana
- 4) Data dimension is fixed within each period: 4, 6

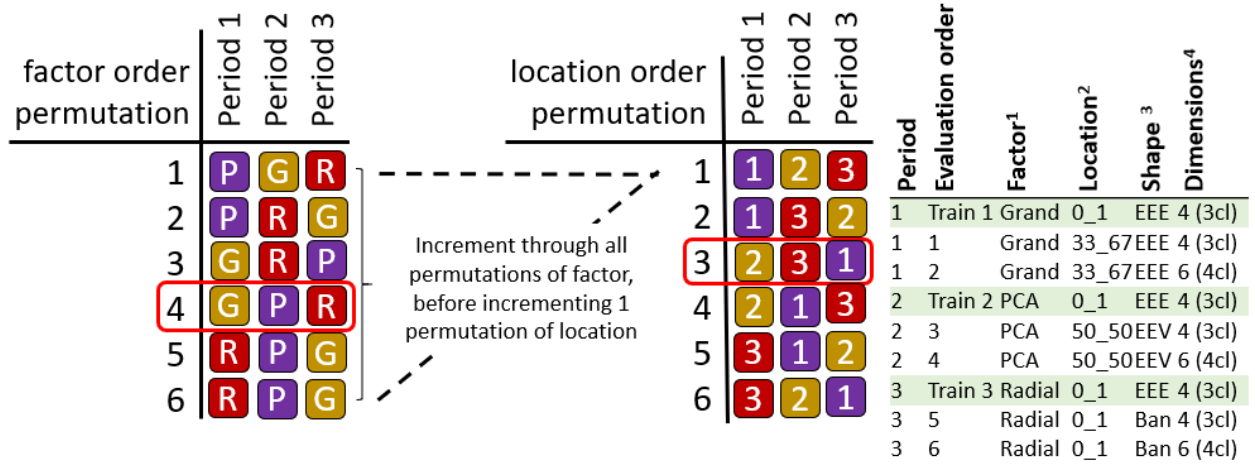


Figure 4: Illustration of how a hypothetical participant 63 is assigned factor and block parameterizations. Each of the 6 factor order permutations are exhausted before iterating to the next permutation of location order.

With this setup we test the parameter space $dimension \in (4, 6)$, $shape \in (EEE, EEV, EVV - banana)$, $location \in (0/100, 33/66, 50/50)$ percent noise/signal mixing in order to evaluate the graphic display across the $factors \in (PCA, grand, radial)$. As we iterate through the possible permutations of these factors and location we perform an even evaluation of the full parameter space every 36 participants. Via pilot studies we estimate that 3 even block evaluations should be sufficient to identify difference between the factors; we targeted for $N = 108$ participants.

In addition to the explicitly controlled block parameters we will also be discussing each participants evaluation order regardless of factor or location experiences. This will expose an learning effect from the repetition of being expose to this data or problem. Keep in my that the fixed blocks shape and location are always experienced in order of increasing difficulty.

3.5 Post study survey

After the evaluation section of the study, participants were given a short survey containing questions gauging demographics, experience, and subjective evaluation of each factor on a 5-point Likert scale. The questions and possible responses are as follows:

Demographic:

- What are your preferred pronouns? [decline to answer, he/him, she/her, they/them or other]

- Which age group do you belong to? [decline to answer, 18 to 24, 25 to 35, 36 to 45, 45 to 60, 60 and up]
- What is your highest completed education? [decline to answer, Undergraduate degree (BA/BSc/other), Graduate degree (MA/MSc/MPhil/other), Doctorate degree (PhD/other); prolific.co participants were filtered to those stating they had an least an undergraduate degree]

Within participant bias:

Likert scale [1-5], least agreement to most agreement.

- I understand the how to perform the task.
- I am experienced with data visualization.
- I am educated in multivariate statistical analysis.

Subjective by factor:

- I was already familiar with visualization.
- I found this visualization easy to use.
- I felt confident in my answers with this visualization.
- I liked using this visualization.

The code, response files, their analyses, and study application are made publicly available at on GitHub at github.com/nsprison/spinifex_study.

3.6 Sampling population

We recruited $N = 108$ via prolific.co (Palan and Schitter 2018). We make the assumption that interpretation of biplot displays used will not be commonly used for consumption by the general population and apply a single filter on education; that participants have completed at least an undergraduate degree (some 58,700 of the 150,400 users at the time). There is also the implicit filter that Prolific participants must be at least 18 years of age. Participants were compensated for their time at £7.50 per hour, whereas the mean duration of survey was about 16 minutes. We can't preclude previous knowledge or experience with the factors, but instead try to control for this in the user study. Figure 5 shows distributions of age and preferred pronouns of the participants that completed the post-study survey who are relatively young and well educated.

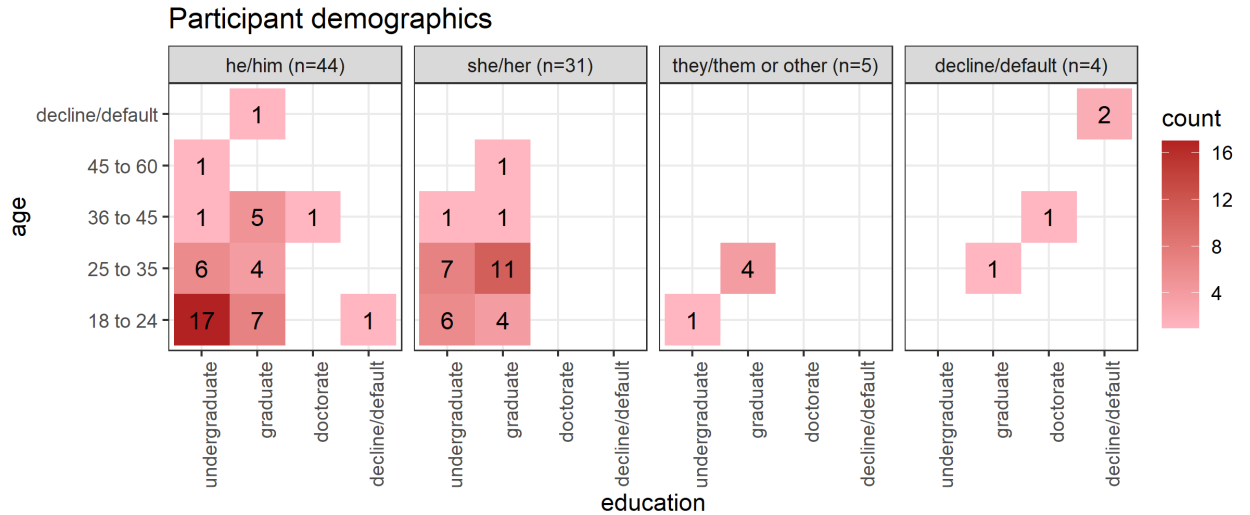


Figure 5: Heatmaps of participant demographics, counts of age group by completed education as faceted across preferred pronoun. The average participant was a young adult with an undergraduate or graduate degree.

3.7 Evenness of block evaluation

From pilot studies through a sample of convenience (primarily PhD students) we predict that we wanted 3 even block evaluations to support differences in our factor and block parameterizations. Given that factor and location each have 6 permutations we targeted $N = 108 = 3 \cdot (6 \cdot 6)$ evaluations before data was collected. In data collection we experience a number of adverse conditions, primarily: limited control of application server network configuration, throughput thresholds on data read/write API, and repeat attempts from users when experiencing disconnects. To mitigate this we over collect survey trials, exclude all partial trials, and remove the oldest attempts (mostly likely to experience adverse network conditions) from over evaluated permutations until we have our desired 3 even evaluations under each permutation.

4 Results

To recap, the primary response variable are marks as defined in section 3.2, while time till last response will be used as a secondary response variable. We have 2 primary data sets; the user study evaluations and post study survey. The former contains the 108 trials with explanatory variables: *factor*, *location* of the cluster separation signal, *shape* of variance-covariance matrix, *dim-ensionality* of the data, and *order* of evaluation within the participant. Block parameterization and randomization was discussed in section @ref(#sec:blocks). The survey was completed for 84 of these 108 trials and contains demographic information (*pronoun*, *age*, and *education*), explanatory variables (*task understanding*, *visualization experience*, and *analysis experience*), and subjective measures for each of the factors (*preference*, *familiarity*, *ease of use*, and *confidence*). The survey was covered in more detail in 3.5.

Below we look at the marginal performance of the block parameters and survey responses. After that we build a battery of regression models to explore the variables and their interactions. Lastly we look at the subjective measures between the factors.

4.1 Marginal block parameters

Consider the 108 trials and their associate block parameterizations, We apply the Kruskal-Wallis H test (or one-way anova), a non-parametric method for testing if the x-levels are from the same distribution.

TODO: XXX write up here.

Will have to figure out a more elegant solution for this. consider: size of text, global test, two-way test. $n=x$ within bars. May need to go manually run test and/or place values.

Figure 6 compares the effect of the study block parameters. Location is the inverse of what is expected, though perhaps mixing the signal diversifies the marks in that its easier to choose at least one of the correct answers of two dimensions with signal rather than the higher risk of selecting the only correct answer for all marks. EEV slightly outperforms EEE, perhaps the different shape helps distinguish the cluster or stick out from one another. Evaluation order does not have a monotonic increase. Every odd evaluation beats the next evaluation which have 4 and 6 dimensions respectively.

4.2 Marginal survey responses

We also test the demographic and experience question from the survey to see if they effect marks. Figure ?? show the variables with strongest evidence to explain marks. Keep in mind that some of the x groups have low counts any likely may not reflect the true distribution of their demographic. Perhaps unsurprisingly, participants that rated their task comprehension higher, averaged higher marks. Age group was narrowly significant, with the lowest, highest, and default groups scoring low. The explanatory power of age is more likely more of a proxy for attention and thoroughness rather than a true effect of age. Education correlates with marks, though is narrowly not significant at $\alpha = .95$, it is also be slight correlated with age and understanding which explain marks better then education alone. Decline/default groups received lower marks unanimously across survey questions; care should be taken to ensure or evaluate the level of effort of participants.

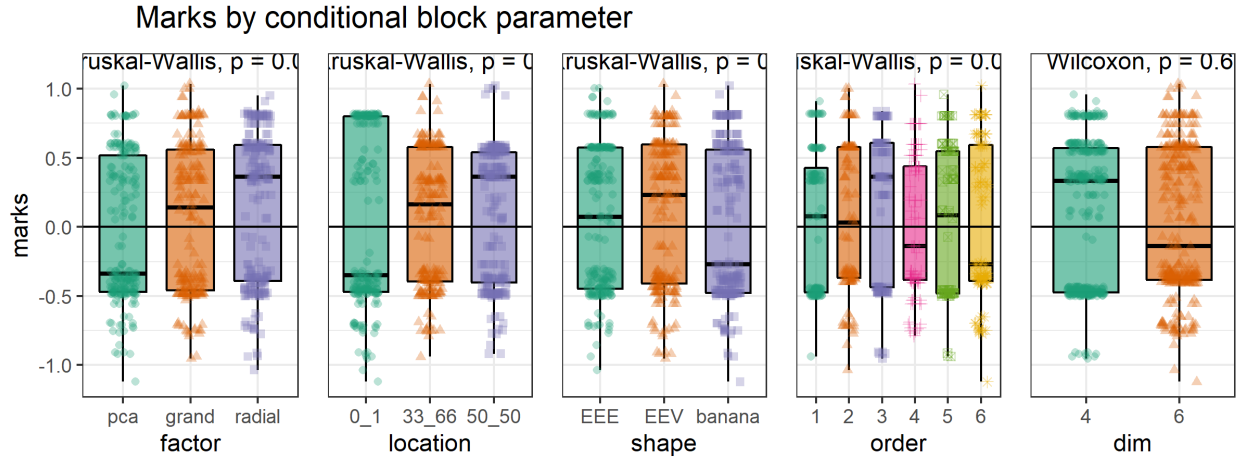


Figure 6: TODO: Marginal effect of the block parameterizations. Factor supports the hypothesis that radial tours outperform the discrete benchmark, pca, and animated benchmark, grand tour, for this task. Location interestingly shows the opposite behavior than expected, as the signal is mixed more than scores tend to be higher. There is no obvious learning effect given the increasing difficulty of the fixed blocks.

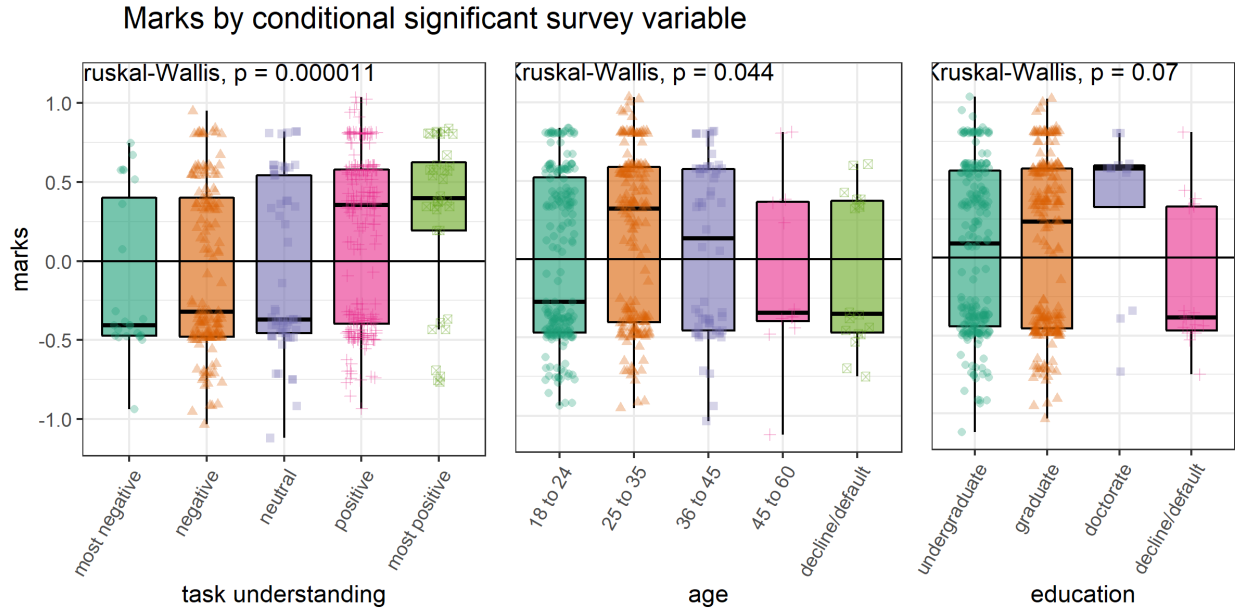


Figure 7: TODO: Marginal effect of the survey variables. TODO: XXX COMPLETE CAPTION

4.3 Random effects regression model

To more thoroughly examine explanatory variables we regress against marks. All models have a random effect term on the participant, which captures the effect of the individual participant. After we look at models of the block parameters we extend to compare against survey variables. Last, we compare how adding a random effect for data and regressing against time till last response fares against benchmark models. The matrices for models with more than a few terms quickly become rank deficient; there is not enough information in the data to explain all of the effect terms. In which case the least impactful terms are dropped.

Evaluation of the models will include Akaike's Information Criterion (AIC), Bayesian information criterion (BIC). These are comparisons measures of models based on its maximized log-likelihood function and penalize over fitting models linearly with the number of fixed terms. Root mean square error (RMSE) is the standard deviation of the residuals is also included. For the above 3 measures, the lower the value the more accurate the model. We additionally include Nakagawa's R^2 for mixed models (Nakagawa, Johnson, and Schielzeth 2017), which has conditional and marginal measures with respect to the random effect terms, and doesn't penalize overfitting.

In building a set of models to test we include all single term models, a model with all independent terms. We also include an interaction term of factor by location, allowing for the slope of each location to change across each level of the factor, which is feasible. For comparison an overly complex model with many interaction terms is included.

Abbreviation:	Model:
F:	$\widehat{marks} = \beta_0 + \sum_{f=1}^3 (\alpha_f \beta_f) + \text{effect}_{participant} + \epsilon$
L:	$\widehat{marks} = \beta_0 + \sum_{l=1}^3 (\alpha_l \beta_l) + \text{effect}_{participant} + \epsilon$
L:	$\widehat{marks} = \beta_0 + \sum_{l=1}^3 (\alpha_l \beta_l) + \text{effect}_{participant} + \epsilon$
S:	$\widehat{marks} = \beta_0 + \sum_{s=1}^3 (\alpha_s \beta_s) + \text{effect}_{participant} + \epsilon$
D:	$\widehat{marks} = \beta_0 + \sum_{d=1}^2 (\alpha_d \beta_d) \text{effect}_{participant} + \epsilon$
O:	$\widehat{marks} = \beta_0 + \sum_{o=1}^6 (\alpha_o \beta_o) + \text{effect}_{participant} + \epsilon$
F+L+S+D+O:	$\widehat{marks} = \beta_0 + \sum_{f=1}^3 (\alpha_f \beta_f) + \sum_{l=1}^3 (\alpha_l \beta_l) + \sum_{s=1}^3 (\alpha_s \beta_s) + \sum_{d=1}^2 (\alpha_d \beta_d) + \sum_{o=1}^6 (\alpha_o \beta_o) + \text{effect}_{participant} + \epsilon$
FL:	$\widehat{marks} = \beta_0 + \sum_{f=1}^3 (\sum_{l=1}^3 (\alpha_{fl} \beta_{fl})) + \text{effect}_{participant} + \epsilon$
FL+S+D+O:	$\widehat{marks} = \beta_0 + \sum_{f=1}^3 (\sum_{l=1}^3 (\alpha_{fl} \beta_{fl})) + \sum_{s=1}^3 (\alpha_s \beta_s) + \sum_{d=1}^2 (\alpha_d \beta_d) + \sum_{o=1}^6 (\alpha_o \beta_o) + \text{effect}_{participant} + \epsilon$
F(L+S+D)+O:	$\widehat{marks} = \beta_0 + \sum_{f=1}^3 (\sum_{l=1}^3 (\alpha_{fl} \beta_{fl})) + \sum_{f=1}^3 (\sum_{s=1}^3 (\alpha_{fs} \beta_{fs})) + \sum_{d=1}^2 (\sum_{s=1}^3 (\alpha_{fd} \beta_{fd})) + \sum_{o=1}^6 (\alpha_o \beta_o) + \text{effect}_{participant} + \epsilon$

where

β_0 is the sum of the intercept and the mean of the participant random effect
 $\epsilon \sim \mathcal{N}(0, \sigma)$
 $\text{effect}_{\text{participant}} \sim \mathcal{N}(0, \tau_{\text{participant}})$
 factor $f \in (\text{pca}, \text{grand}, \text{radial})$
 location $l \in (0/100, 33/66, 50/50)$ percent mixing of a noise and signal variable respectively
 shape $s \in (\text{EEE}, \text{EEV}, \text{EVV}, \text{banana})$
 dim $d \in (4, 6)$ variables, dimensions of the data with 3 & 4 clusters respectively
 order $o \in (1, \dots, 6)$, order of evaluation within participant, as a factor rather than numeric variable

We want to see how survey variables fare in comparison. Holding the factor model as a comparison benchmark we fit another battery of similarly increasing complexity. Keep in mind that the scope has changed a bit from the 108 evaluations to only the 84 studies that have available survey data.

Abbreviation:	Model:
F:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f) + \text{effect}_{\text{participant}} + \epsilon$
F + else study:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f) + \Sigma_{l=1}^3(\alpha_l \beta_l) + \Sigma_{s=1}^3(\alpha_s \beta_s) + \Sigma_{d=1}^2(\alpha_d \beta_d) + \Sigma_{o=1}^6(\alpha_o \beta_o) + \text{effect}_{\text{participant}} + \epsilon$
F + else survey:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f) + \Sigma_{p=1}^4(\alpha_p \beta_p) + \Sigma_{a=1}^5(\alpha_a \beta_a) + \Sigma_{e=1}^4(\alpha_e \beta_e) + \Sigma_{v=1}^5(\alpha_v \beta_v) + \Sigma_{ae=1}^5(\alpha_{ae} \beta_{ae}) + \text{effect}_{\text{participant}} + \epsilon$
F else study:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f)(\Sigma_{l=1}^3(\alpha_l \beta_l) + \Sigma_{s=1}^3(\alpha_s \beta_s) + \Sigma_{d=1}^2(\alpha_d \beta_d) + \Sigma_{o=1}^6(\alpha_o \beta_o)) + \text{effect}_{\text{participant}} + \epsilon$
F else survey:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f)(\Sigma_{p=1}^4(\alpha_p \beta_p) + \Sigma_{a=1}^5(\alpha_a \beta_a) + \Sigma_{e=1}^4(\alpha_e \beta_e) + \Sigma_{v=1}^5(\alpha_v \beta_v) + \Sigma_{ae=1}^5(\alpha_{ae} \beta_{ae})) + \text{effect}_{\text{participant}} + \epsilon$
F + else study&survey:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f) + \Sigma_{l=1}^3(\alpha_l \beta_l) + \Sigma_{s=1}^3(\alpha_s \beta_s) + \Sigma_{d=1}^2(\alpha_d \beta_d) + \Sigma_{o=1}^6(\alpha_o \beta_o) + (\Sigma_{p=1}^4(\alpha_p \beta_p) + \Sigma_{a=1}^5(\alpha_a \beta_a) + \Sigma_{e=1}^4(\alpha_e \beta_e) + \Sigma_{v=1}^5(\alpha_v \beta_v) + \Sigma_{ae=1}^5(\alpha_{ae} \beta_{ae})) + \text{effect}_{\text{participant}} + \epsilon$
F else study&survey:	$\widehat{\text{marks}} = \beta_0 + \Sigma_{f=1}^3(\alpha_f \beta_f)(\Sigma_{l=1}^3(\alpha_l \beta_l) + \Sigma_{s=1}^3(\alpha_s \beta_s) + \Sigma_{d=1}^2(\alpha_d \beta_d) + \Sigma_{o=1}^6(\alpha_o \beta_o) + (\Sigma_{p=1}^4(\alpha_p \beta_p) + \Sigma_{a=1}^5(\alpha_a \beta_a) + \Sigma_{e=1}^4(\alpha_e \beta_e) + \Sigma_{v=1}^5(\alpha_v \beta_v) + \Sigma_{ae=1}^5(\alpha_{ae} \beta_{ae}))) + \text{effect}_{\text{participant}} + \epsilon$

where

pronoun $p \in (\text{he/him}, \text{she/her}, \text{they/them or other}, \text{default/decline})$

textage $a \in (18 \text{ to } 24, 25 \text{ to } 35, 36 \text{ to } 45, 45 \text{ to } 60, \text{default/decline})$

texteducation $e \in (\text{Undergraduate degree (BA/BSc/other)}, \text{Graduate degree (MA/MSc/MPhil/other)}, \text{Doctorate degree (PhD)})$

texttaskUnderstanding $u \in (\text{Likert scale [1-5]: I understand the how to perform the task.})$

textdataVisualizationexperience $v \in (\text{Likert scale [1-5]: I am experienced with data visualization.})$

Analysis Experience $ae \in (\text{Likert scale [1-5]: I am educated in multivariate statistical analysis.})$

Table 1: Model comparison of our random effects models regressing marks. Each model includes a random effect term of the participant, which explains the individuals influence on their marks. Complex models perform better in terms of R2 and RMSE, yet AIC and BIC penalize their large number of fixed effect in favor of the much simpler model containing only factor.

Model name	No. factors	No. fixed effects	AIC	BIC	R2 cond. (on RE)	R2 marg. (w/o RE)	RMSE
F	1	3	**999**	**1021**	0.17	0.022	0.466
L	1	3	1014	1037	0.146	0.002	0.473
S	1	3	1013	1036	0.148	0.003	0.473
D	1	2	1007	1025	0.149	0.004	0.472
O	1	6	1020	1056	0.161	0.016	0.468
F*L	3	9	1020	1070	0.181	0.036	0.463
F+L+S+D+O	5	10	1028	1081	0.186	0.039	0.461
F*L+S+D+O	6	14	1038	1109	0.197	0.052	0.458
F*(L+S+D)+O	8	20	1055	1154	**0.207**	**0.067**	**0.455**

Among the above models, the simple factor model performs best in terms of AIC and BIC, and is thus included as a benchmark model of comparison going forward. In see the a particular simulation effect on the score an additional random effect term based on the simulation data is used. We also want to see how well the time till response can be explained by the single factor model; consider:

Abbreviation: Model:

$$\begin{aligned}
\text{marks=F+RE(participant): } \widehat{marks} &= \beta_0 + \sum_{f=1}^3 (\alpha_f \beta_f) + \text{effect}_{\text{participant}} + \epsilon \\
\text{marks=F+RE(participant+simulation): } \widehat{marks} &= \beta_0 + \sum_{f=1}^3 (\alpha_f \beta_f) + \text{effect}_{\text{participant}} + \text{effect}_{\text{simulation}} + \epsilon \\
\text{time=F+RE(participant): } \widehat{marks} &= \beta_0 + \sum_{f=1}^3 (\alpha_f \beta_f) + \text{effect}_{\text{participant}} + \epsilon \\
\text{time=F+RE(participant+simulation): } \widehat{marks} &= \beta_0 + \sum_{f=1}^3 (\alpha_f \beta_f) + \text{effect}_{\text{participant}} + \text{effect}_{\text{simulation}} + \epsilon
\end{aligned}$$

where

$$\text{effect}_{\text{simulation}} \sim \mathcal{N}(0, \pi_{\text{simulation}})$$

time is the duration of time on the page till the last response change

Residual plots had no noticeable non-linear trends and contain striped patterns as an artifact from regressing on discrete variables. Figure 8 illustrates a typically residual plot (left), while the right compares 95% confidence interval of the effect ranges for participant and simulation. The effect ranges are estimated from simulations of the posterior distribution. The effect size of participant is much larger than simulation. The most extreme participants are statistically significant at $\alpha = .95$, while none of the simulation effects significantly deviate from the null of having no effect size on the marks.

4.4 Subjective measures

TODO: XXX -Some (see <https://bookdown.org/Rmadillo/likert/>) claim that boxplot and tests are inappropriate for Likert scale data. Convert to percentile stacked bar charts? -significance testing?

Table 2: TODO: XXX CAP NEEDED

Model name	No. factors	No. fixed effects	AIC	BIC	R2 cond. (on RE)	R2 marg. (w/o RE)	RMSE
F	1	3	**803**	**824**	0.14	0.02	0.48
F+parameters	5	10	834	885	0.15	0.04	0.48
F+survey	7	25	866	980	0.18	0.12	0.48
F+parameters+survey	11	32	897	1041	0.2	0.13	0.47
F*parameters	8	20	859	952	0.18	0.08	0.47
F*survey	13	69	973	1273	0.23	0.17	0.46
F*parameters*survey	21	90	1034	1422	**0.26**	**0.22**	**0.44**
F+:t_val>1.5	5	20	845	938	0.17	0.13	0.48
F*:t_val>1.5	13	66	959	1246	0.23	0.19	0.46

Table 3: We select the simple facet model as our benchmark and compare the random effects from participants with the random effects of simulated data. The random effect of the data is small compared to the random effect of the data. The random effect of the participant is good at predicting the time users take, its conditional R2 is much larger than the marginal R2 of the factor fixed term.

Model name	No. factors	No. fixed effects	AIC	BIC	R2 cond. (on RE)	R2 marg. (w/o RE)	RMSE
marks~F+RE(parti.)	1	3	999	1021	0.17	0.02	0.47
marks~F+RE(parti.+sim)	1	3	1000	1027	0.18	0.02	0.46
time~F+RE(parti.)	1	3	5253	5275	0.50	0.02	11.29
time~F+RE(parti.+sim)	1	3	5240	5267	0.53	0.02	10.82

Figure 8: Residual plot and estimated effect ranges for factor regressing on marks with random effects for both participant and data simulation. The effect size of participant is relatively large, with several significant extrema ($\alpha = .95$). None of the simulations deviate significantly.

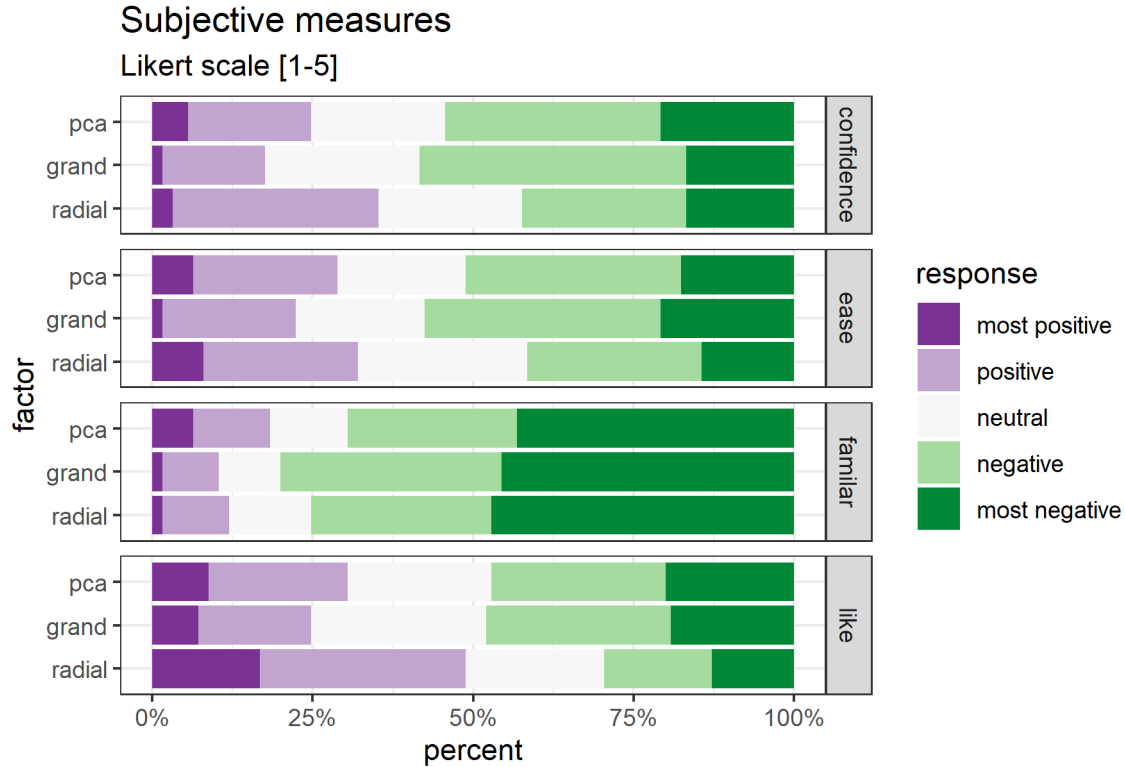


Figure 9: TODO: XXX caption needed. Convert to percentile stacked bar charts?

5 Discussion

6 Accompanying tool: spinifex application

To accompany this study we have produced a more general use tool to perform such exploratory analysis of high dimensional data. The R package, **spinifex**, (Spyrison and Cook 2020) R package contains a free, open-source **shiny** (Chang et al. 2020) application. The application allows users to explore their data with either interactive or predefined manual tours without the need for any coding. Limited implementations of grand, little, and local tours are also made available. Data can be imported in .csv and .rda format, and projections or animations can be saved as .png, .gif, and .csv formats where applicable. Run the following R code for help getting started.

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