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# Bayesian Inference for Least-Squares Temporal Difference Regularization

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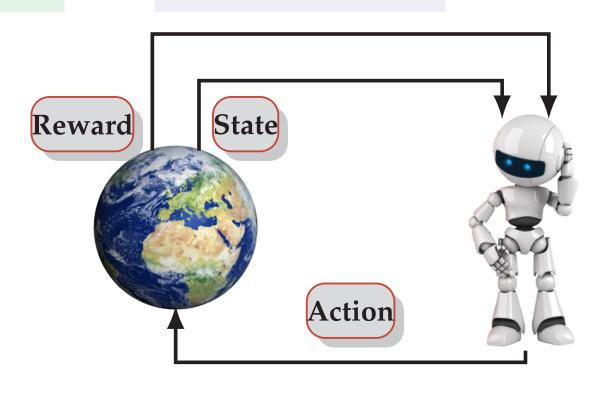
## Abstract

#### **Bayesian Least-Squares Temporal Difference:**

- ✓ Fully Bayesian approach for LSTD
- ✓ Probabilistic inference of value functions
- Quantifies our uncertainty about value function Variational Bayesian LSTD:
- ✓ Sparse model Good generalisation capabilities
- ✓ Automatically determine the model's complexity
- ✓ No need to select a regularization parameter

## Reinforcement Learning (RL)

Learning to act in an unknown environment, by interaction and reinforcement.



RL tasks folmulated as **MDPs**,  $\{S, A, P, r, \gamma\}$ **Policy**  $\pi$ :  $S \to A$  (map states to actions) Value function:  $V^{\pi}(s) \triangleq \mathbb{E}^{\pi} \left[ \sum_{t=0}^{\infty} \gamma^{t} r(s_{t}) | s_{0} = s \right]$ Bellman operator:

$$(T^{\pi}V)(s) = r(s) + \gamma \int_{\mathcal{S}} V(s')dP(s'|s, \pi(s))$$

V is the unique fixed-point of Bellman operator:

$$V^{\pi} = T^{\pi}V^{\pi} \Rightarrow V^{\pi} = (\boldsymbol{I} - \gamma P^{\pi})^{-1}\boldsymbol{r}$$

- © The value function cannot be represented in an explicit way (continuous state space)
- © The model of the MDP is unknown
- **Solution Value function approximation:**

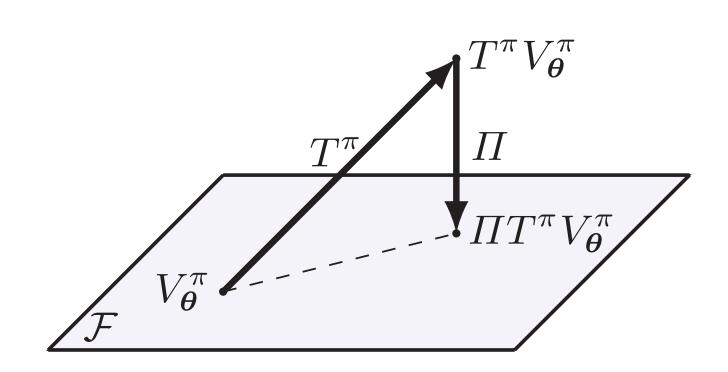
$$V_{\boldsymbol{\theta}}^{\pi}(s) = \boldsymbol{\phi}(s)^{\top} \boldsymbol{\theta} = \sum_{i=1}^{n} \phi_i(s) \theta_i,$$

where  $\mathcal{F} = \{ f_{\boldsymbol{\theta}} | f_{\boldsymbol{\theta}}(\cdot) = \boldsymbol{\phi}(\cdot)^{\top} \boldsymbol{\theta} \}.$ 

 $\odot$  Access to a set of transitions:  $\mathcal{D} = \{(s_i, r_i, s_i')\}_{i=1}^n$ , where we define:  $\tilde{\Phi} = [\phi(s_1)^\top; \dots; \phi(s_n)^\top],$   $\tilde{R} = [r_i, \dots, r_n]^\top$  and  $\tilde{\Phi}' = [\phi(s_1')^\top; \dots; \phi(s_n')^\top].$ 

## Least-Squares Temporal Difference

Minimize the mean-square projected Bellman error (MSPBE):  $\boldsymbol{\theta} = \arg\min_{\boldsymbol{\theta} \in \mathbb{R}^k} \|V_{\boldsymbol{\theta}}^{\pi} - \Pi T^{\pi} V_{\boldsymbol{\theta}}^{\pi}\|_D^2$ .



**Projection operator over**  $\mathcal{F}$ :  $\Pi = \Phi C^{-1} \Phi^{\top} D$ 

## Nested Optimization Problem

 $u^* = \arg\min \|\Phi u - T^{\pi} \Phi \theta\|_D^2$  (Projection step)  $oldsymbol{u}{\in}\mathbb{R}^k$ 

 $\boldsymbol{\theta} = \arg\min \|\boldsymbol{\Phi}\boldsymbol{\theta} - \boldsymbol{\Phi}\boldsymbol{u}^*\|_D^2 \quad (Fixed-point\ step)$  $oldsymbol{ heta} \in \mathbb{R}^k$ 

#### Solution

$$\boldsymbol{u}^* = \tilde{C}^{-1}\tilde{\Phi}(\tilde{R} + \gamma\tilde{\Phi}'\boldsymbol{\theta}),$$
  
$$\boldsymbol{\theta} = (\tilde{\Phi}^{\top}(\tilde{\Phi} - \gamma\tilde{\Phi}'))^{-1}\tilde{\Phi}^{\top}\tilde{R} = A^{-1}\boldsymbol{b},$$

where  $\tilde{C} \triangleq \tilde{\Phi}^{\top} \tilde{\Phi}$ ,  $A \triangleq \tilde{\Phi}^{\top} (\tilde{\Phi} - \gamma \tilde{\Phi}')$ , and  $\mathbf{b} \triangleq \tilde{\Phi}^{\top} \tilde{R}$ .

As the number of samples n increases, the LSTD solution  $\tilde{\Phi}\theta$  converges to the fixed-point of  $\hat{\Pi}T^{\pi}$ .

## Regularized LSTD Schemes

- $\ell_2$ -LSTD
- LARS-TD
- $\ell_{2,1}$ -LSTD
- Lasso-TD • LC-TD

#### • $\ell_1$ -PBR Dantzig-LSTD • $\ell_{2,2}$ -LSTD ODDS-TD

## Bayesian LSTD

#### **Empirical Bellman operator:**

$$\hat{T}^{\pi}V_{\boldsymbol{\theta}}^{\pi} = \boldsymbol{r} + \gamma P^{\pi}V_{\boldsymbol{\theta}}^{\pi} + N, \quad N \sim \mathcal{N}(0, \beta^{-1}\boldsymbol{I})$$

Given observations  $\mathcal{D}$  (LSTD solution):

$$V_{\boldsymbol{\theta}}^{\pi} = \hat{\Pi}\hat{T}^{\pi}V_{\boldsymbol{\theta}}^{\pi} \Leftrightarrow \tilde{\boldsymbol{\Phi}}^{\top}\tilde{R} = \tilde{\boldsymbol{\Phi}}^{\top}(\tilde{\boldsymbol{\Phi}} - \gamma\tilde{\boldsymbol{\Phi}}')\boldsymbol{\theta} + \tilde{\boldsymbol{\Phi}}^{\top}N$$

Linear regression model:  $\boldsymbol{b} = A\boldsymbol{\theta} + \tilde{\boldsymbol{\Phi}}^{\top} N$ Likelihood function:  $p(b|\theta,\beta) = \mathcal{N}(b|A\theta,\beta^{-1}\tilde{C})$ 

✓ Maximum likelihood inference corresponds to standard LSTD solution

Prior distribution over  $\theta : p(\theta | \alpha) = \mathcal{N}(\theta | 0, \alpha^{-1} I)$ . Logarithm of posterior distribution:

$$\ln p(\boldsymbol{\theta}|\mathcal{D}) \propto -\frac{\beta}{2} E_{\mathcal{D}}(\boldsymbol{\theta}) - \frac{\alpha}{2} \boldsymbol{\theta}^{\top} \boldsymbol{\theta},$$

where  $E_{\mathcal{D}}(\boldsymbol{\theta}) = (\boldsymbol{b} - A\boldsymbol{\theta})^{\top} \tilde{C}^{-1} (\boldsymbol{b} - A\boldsymbol{\theta})$  (MSPBE).

√ Maximum a posteriori inference corresponds to  $\ell_2$  regularization ( $\lambda = \alpha/\beta$ )

Posterior distribution:  $p(\theta|\mathcal{D}) = \mathcal{N}(\theta|m, S)$ 

$$S = (\alpha \mathbf{I} + \beta \underbrace{A^{\top} \tilde{C}^{-1} A})^{-1} \text{ and } \mathbf{m} = \beta S A^{\top} \tilde{C}^{-1} \mathbf{b}$$

#### Predictive distribution:

$$p(V_{\boldsymbol{\theta}}^{\pi}(s^*)|s^*, \mathcal{D}) = \int_{\boldsymbol{\theta}} p(V_{\boldsymbol{\theta}}^{\pi}(s^*)|\boldsymbol{\theta}, s^*) dp(\boldsymbol{\theta}|\boldsymbol{b}, \alpha, \beta)$$
$$= \mathcal{N}(V_{\boldsymbol{\theta}}^{\pi}(s^*)|\boldsymbol{\phi}(s^*)^{\top}\boldsymbol{m}, \boldsymbol{\phi}(s^*)^{\top}S\boldsymbol{\phi}(s^*))$$

## Variational Bayesian LSTD

Prior distribution:  $p(\boldsymbol{\theta}|\boldsymbol{\alpha}) = \prod_{i=1}^k \mathcal{N}(\theta_i|0,\alpha_i^{-1})$ **Hyperprior over \alpha:**  $p(\alpha) = \prod_{i=1}^k \mathcal{G}$ amma $(\alpha_i | h_a, h_b)$ **Hyperprior over**  $\beta$ :  $p(\beta) = Gamma(\beta|h_c, h_d)$ 

Posterior distribution over  $\mathcal{Z} = \{\theta, \alpha, \beta\}$ :

$$p(\boldsymbol{\theta}, \boldsymbol{\alpha}, \beta | \boldsymbol{b}) = \frac{p(\boldsymbol{b} | \boldsymbol{\theta}, \beta) p(\beta) p(\boldsymbol{\theta} | \boldsymbol{\alpha}) p(\boldsymbol{\alpha})}{p(\boldsymbol{b})}$$

### © Marginal likelihood is analytically intractable

#### √ We resolt to variational inference

Variational distribution:  $Q(Z) = Q_{\theta}(\theta)Q_{\alpha}(\alpha)Q_{\beta}(\beta)$ Optimal distribution for each factor:

$$Q_{\boldsymbol{\theta}}(\boldsymbol{\theta}) = \mathcal{N}(\boldsymbol{\theta}|\boldsymbol{m}, S)$$

$$\mathcal{Q}_{eta}(eta) = \mathcal{G}amma(eta| ilde{c}, ilde{d})$$

$$\mathcal{Q}_{oldsymbol{lpha}}(oldsymbol{lpha}) = \prod_{i=1}^k \mathcal{G}amma(lpha_i| ilde{a}_i, ilde{b}_i)$$

where,

$$S = (\operatorname{diag} \mathbb{E}[\boldsymbol{\alpha}] + \mathbb{E}[\boldsymbol{\beta}]\boldsymbol{\Sigma})^{-1}, \quad \boldsymbol{m} = \mathbb{E}[\boldsymbol{\beta}]SA^{\top}\tilde{C}^{-1}\boldsymbol{b},$$

$$\tilde{a}_{i} = h_{a} + \frac{1}{2}, \quad \tilde{b}_{i} = h_{b} + \frac{1}{2}\mathbb{E}[\theta_{i}^{2}],$$

$$\tilde{c} = h_{c} + \frac{k}{2}, \quad \tilde{d} = h_{d} + \frac{1}{2}\|\boldsymbol{b} - A\boldsymbol{m}\|_{\tilde{C}}^{2} + \frac{1}{2}\operatorname{tr}(\boldsymbol{\Sigma}S).$$

#### Lower bound

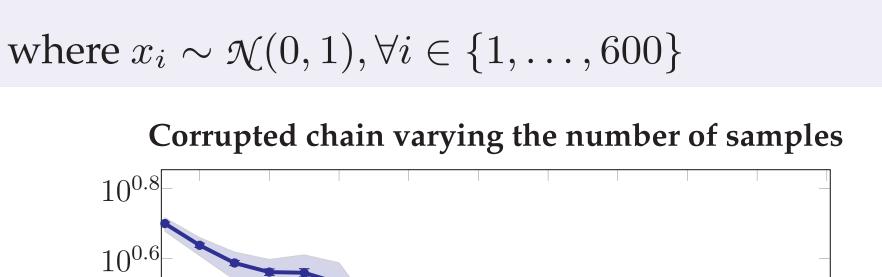
$$\mathcal{L}(\mathcal{Q}) = \frac{1}{2} \ln|S| - \frac{1}{2} |\tilde{C}| + \sum_{i=1}^{k} \{ \ln \Gamma(\tilde{a}_i) - \tilde{a}_i \ln \tilde{b}_i \}$$
$$+ \ln \Gamma(\tilde{c}) - \tilde{c} \ln \tilde{d} + \frac{k}{2} (1 - \ln 2\pi) - k \ln \Gamma(h_a)$$
$$+ kh_a \ln h_b - \ln \Gamma(h_c) + h_c \ln h_d$$

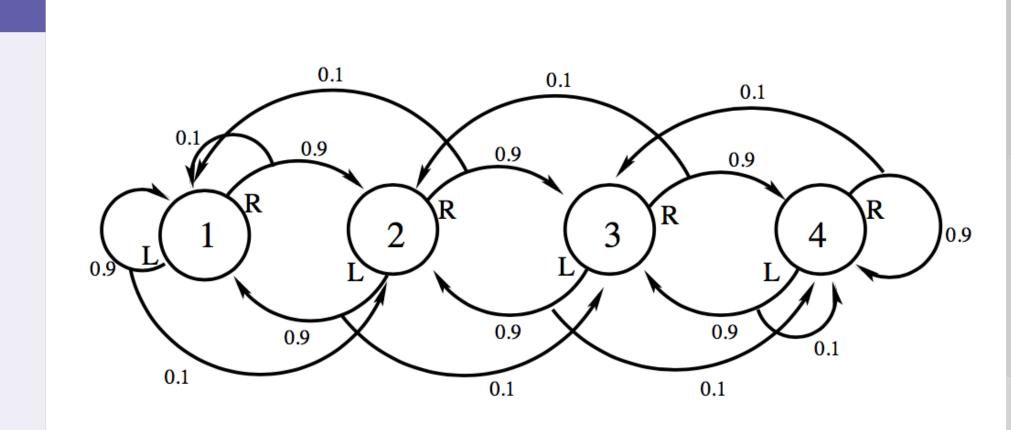
## Experimental Results

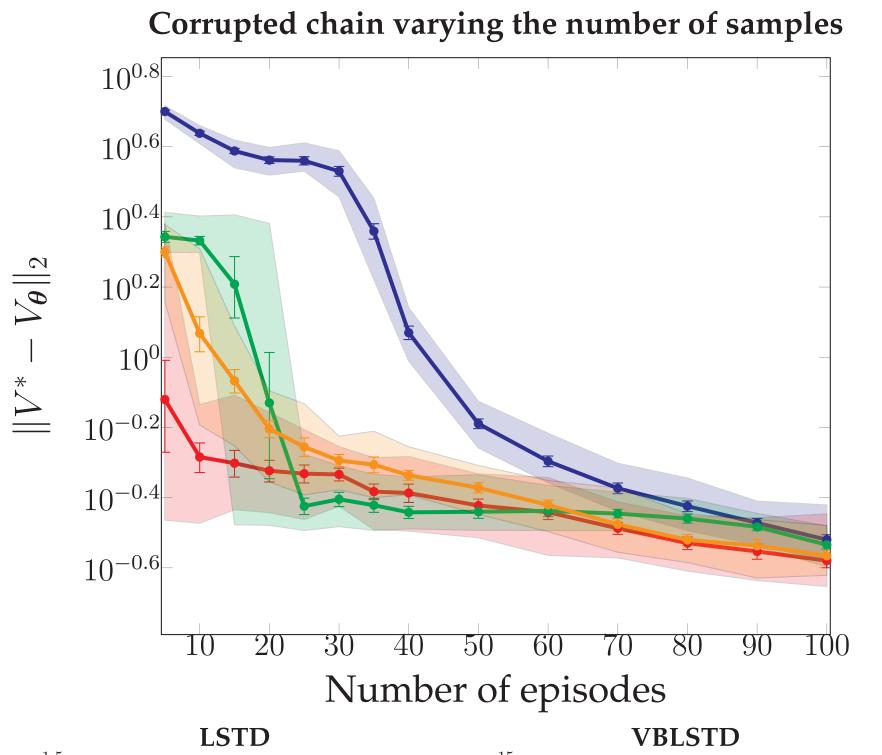
### Corrupted Chain

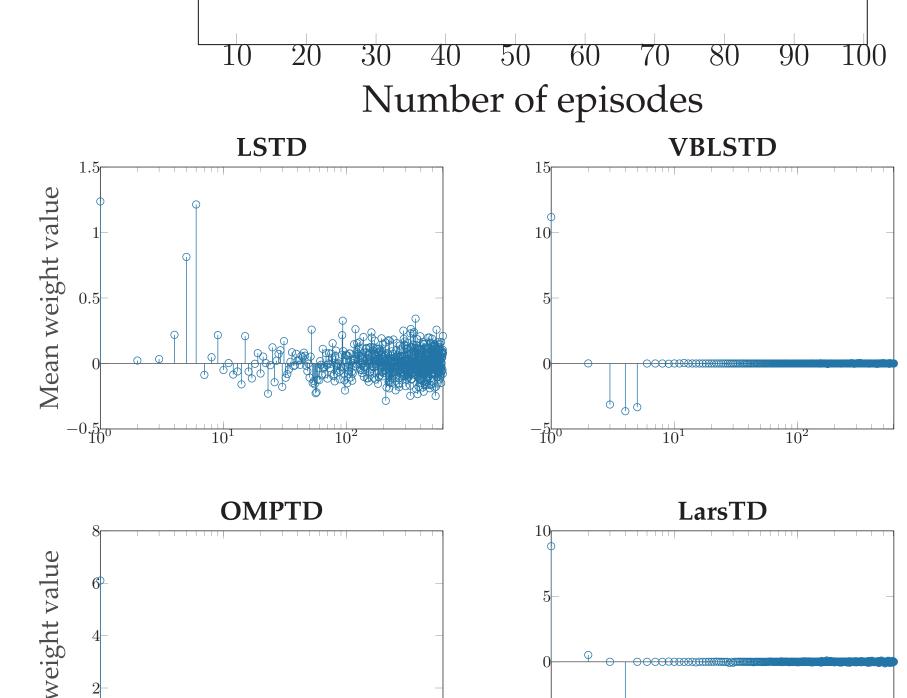
- A 20-state, 2-actions (left or right) MDP
- The probability of action's success is equal to 0.9
- A reward of 1 is given only at the ends of chain
- The horizon of each episode is set equal to 20
- Optimal policy: i)  $s \le 10$ : a = L, ii) s > 10: a = R
- Value fuction representation:

$$\phi(s) = (1, RBF_1(s), \dots, RBF_5(s), x_1, \dots, x_{600})$$



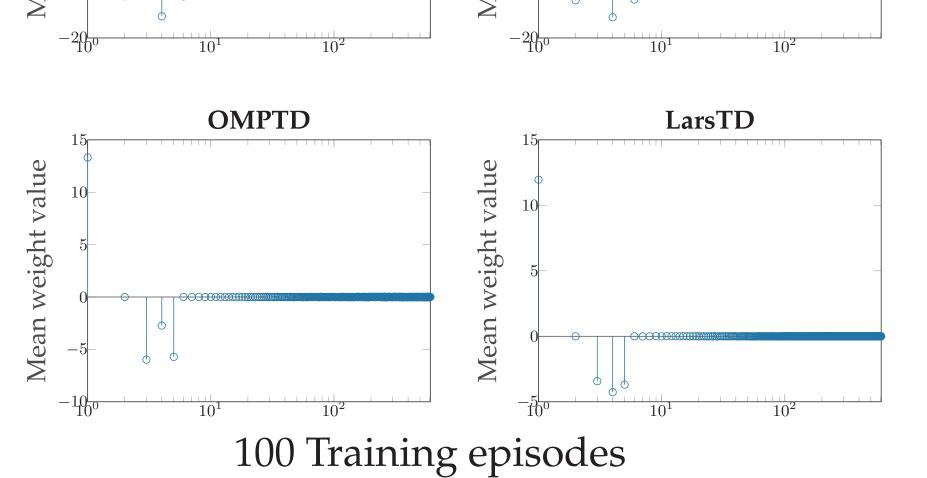






10 Training episodes

Corrupted chain varying the number of noise features  $10^{0.2}$  $--\ell_2$ -LSTD - VBLSTD - OMPTD - LarsTD  $10^{-0.2}$ Number of noisy features LSTD **VBLSTD** weight value



The 606 mean weight values. The first weight is the bias term, the next 5 correspond to the relevant features (RBFs), and the rest 600 correspond to the noise (irrelevant) features.