

# ETF3231/5231

## Business forecasting

Week 8: ARIMA models

<https://bf.numbat.space/>



# Outline

- 1 Non-seasonal ARIMA models
- 2 Estimation and order selection

# ARIMA models

**AR:** autoregressive (lagged observations as inputs)

**I:** integrated (differencing to make series stationary)

**MA:** moving average (lagged errors as inputs)

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An ARIMA model is rarely interpretable in terms of visible data structures like trend and seasonality. But it can capture a huge range of time series patterns.

Make data stationary (variance & mean), fit model, reverse, forecast.

# Outline

1 Non-seasonal ARIMA models

2 Estimation and order selection

# AR(1) model including a constant

$$y_t = c + \phi_1 y_{t-1} + \varepsilon_t, \quad |\phi_1| < 0 \quad (\text{P.T.O})$$

- When  $\phi_1 = 0$  and  $c = 0$ ,  $y_t$  is equivalent to WN;
- When  $\phi_1 = 1$  and  $c = 0$ ,  $y_t$  is equivalent to a RW;
- When  $\phi_1 = 1$  and  $c \neq 0$ ,  $y_t$  is equivalent to a RW with drift;
- When  $\phi_1 > 0$ ,  $y_t$  tends to hang below or above the mean of  $y_t$ .
- When  $\phi_1 < 0$ ,  $y_t$  tends to oscillate below and above the mean of  $y_t$ .
- If  $E(y_t) = \mu$ ,  $\mu = \frac{c}{1-\phi_1}$ ,  $c$  is related to the mean of  $y_t$ .

For stationarity we require  $-1 < \phi_1 < 1$

Let's set  $\phi_1 = 2$   $y_t = 2y_{t-1} + \varepsilon_t$  and see what happens:

$$y_1 = 2y_0 + \varepsilon_1$$

$$y_2 = 2y_1 + \varepsilon_2 = 4y_0 + 2\varepsilon_1 + \varepsilon_2$$

$$y_3 = 2y_2 + \varepsilon_3 = 8y_0 + 4\varepsilon_1 + 2\varepsilon_2 + \varepsilon_3 \dots \text{and so on}$$

⋮

- Getting longer & longer
- Hence, we only allow for the new obs. to be a fraction of the previous ones.

# Autoregressive models

A multiple regression with **lagged values** of  $y_t$  as predictors.

$$\begin{aligned}y_t &= c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \varepsilon_t \\&= c + (\phi_1 B + \phi_2 B^2 + \cdots + \phi_p B^p) y_t + \varepsilon_t\end{aligned}$$

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$$\begin{aligned}(1 - \phi_1 B - \phi_2 B^2 - \cdots - \phi_p B^p) y_t &= c + \varepsilon_t \\ \phi(B) y_t &= c + \varepsilon_t\end{aligned}$$

- $\varepsilon_t$  is white noise.
- $\phi(B) = (1 - \phi_1 B - \phi_2 B^2 - \cdots - \phi_p B^p)$

# Stationarity conditions

We normally restrict autoregressive models to stationary data, and then some constraints on the values of the parameters are required.

## General condition for stationarity

Complex roots of  $\phi(z) = 1 - \phi_1z - \phi_2z^2 - \dots - \phi_pz^p$  lie outside the unit circle on the complex plane.

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- For  $p = 1$ :  $-1 < \phi_1 < 1$ .
- For  $p = 2$ :  
$$-1 < \phi_2 < 1 \quad \phi_2 + \phi_1 < 1 \quad \phi_2 - \phi_1 < 1.$$
- More complicated conditions hold for  $p \geq 3$ .
- fable takes care of this.

# Moving Average (MA) models

A multiple regression with *past errors* as predictors.

$$\begin{aligned}y_t &= c + \varepsilon_t + \theta_1 \varepsilon_{t-1} + \theta_2 \varepsilon_{t-2} + \cdots + \theta_q \varepsilon_{t-q} \\&= c + (1 + \theta_1 B + \theta_2 B^2 + \cdots + \theta_q B^q) \varepsilon_t \\&= c + \theta(B) \varepsilon_t\end{aligned}$$

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- $\varepsilon_t$  is white noise.
- $\theta(B) = (1 + \theta_1 B + \theta_2 B^2 + \cdots + \theta_q B^q)$

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## General condition for invertibility

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- For  $q = 1$ :  $-1 < \theta_1 < 1$ .
- For  $q = 2$ :  
$$-1 < \theta_2 < 1 \quad \theta_2 + \theta_1 > -1 \quad \theta_1 - \theta_2 < 1.$$
- More complicated conditions hold for  $q \geq 3$ .
- fable takes care of this.

## PROPERTIES OF AR & MA MODELS

- AR(1)  $y_t = c + \phi y_{t-1} + \varepsilon_t ; \varepsilon_t \sim NID(0, \sigma^2); |\phi| < 1;$   $y_0$  (starting value).

$$y_t = c + \phi y_{t-1} + \varepsilon_t$$

$$= c + \phi (c + \phi y_{t-2} + \varepsilon_{t-1}) + \varepsilon_t$$

$$= c + \phi c + \phi^2 y_{t-2} + \phi \varepsilon_{t-1} + \varepsilon_t$$

$$= c + \phi c + \phi^2 c + \phi^3 y_{t-3} + \phi^2 \varepsilon_{t-2} + \phi \varepsilon_{t-1} + \varepsilon_t$$

}

$$\Rightarrow y_t = c(1 + \phi + \phi^2 + \dots) + \phi^t y_0 + \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots$$

$$\text{for } t \rightarrow \infty = \frac{c}{1-\phi} + \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots \quad MA(\infty)$$

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$$\Rightarrow y_t = c(1 + \phi + \phi^2 + \dots) + \phi^t y_0 + \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots$$

for  $t \rightarrow \infty$

$$= \frac{c}{1-\phi} + \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots \quad MA(\infty)$$

- $E(y_t) = \frac{c}{1-\phi} = \mu$

- $\text{var}(y_t) = \gamma_0 = E(y_t - \mu)^2 = \frac{\sigma^2}{1-\phi^2} \quad (\text{stat.})$

$$\begin{aligned} \gamma_0 &= E \left( \frac{c}{1-\phi} + \varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots - \frac{c}{1-\phi} \right)^2 \\ &= E(\varepsilon_t + \phi \varepsilon_{t-1} + \phi^2 \varepsilon_{t-2} + \dots)^2 \end{aligned}$$

$$\begin{aligned} &\approx E(\varepsilon_t)^2 + \phi^2 E(\varepsilon_{t-1})^2 + \phi^4 E(\varepsilon_{t-2})^2 + \dots \\ &\quad \text{cross-products} \\ &= \sigma^2 + \phi^2 \sigma^2 + (\phi^2)^2 \sigma^2 + (\phi^2)^3 \sigma^2 + \dots \end{aligned}$$

$$= \sigma^2 \left( 1 + \phi^2 + (\phi^2)^2 + (\phi^2)^3 + \dots \right) = \frac{\sigma^2}{1-\phi^2}$$

$$\text{Cov}(y_t, y_{t-1}) = \gamma_1 = \phi \frac{\sigma^2}{1-\phi^2} = \phi \gamma_0$$

$$\gamma_1 = E(y_t - \mu)(y_{t-1} - \mu)$$

$$= E\left(\varepsilon_t + \underbrace{\phi \varepsilon_{t-1}}_{\mu} + \underbrace{\phi^2 \varepsilon_{t-2}}_{\mu} + \underbrace{\phi^3 \varepsilon_{t-3}}_{\mu} + \dots\right) \left(\varepsilon_{t-1} + \underbrace{\phi \varepsilon_{t-2}}_{\mu} + \underbrace{\phi^2 \varepsilon_{t-3}}_{\mu} + \underbrace{\phi^3 \varepsilon_{t-4}}_{\mu} + \dots\right)$$

$$= \phi E(\varepsilon_{t-1})^2 + \phi^3 E(\varepsilon_{t-2})^2 + \phi^5 E(\varepsilon_{t-3})^2 + \dots$$

$$= \phi \sigma^2 + \phi \phi^2 \sigma^2 + \phi \phi^4 \sigma^2 + \dots$$

$$= \phi \sigma^2 (1 + \phi^2 + (\phi^2)^2 + (\phi^2)^3 + \dots)$$

$$= \phi \frac{\sigma^2}{1-\phi^2}$$

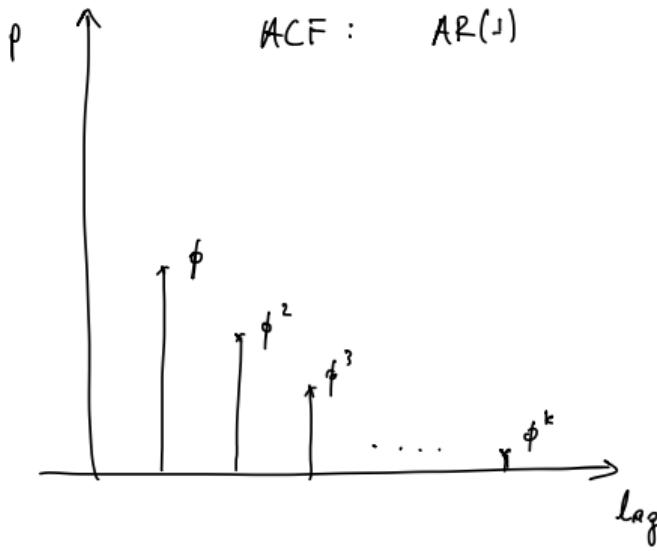
In general  $\gamma_k = \phi^k \gamma_0$  (cov stat)

$$\rho_1 = \frac{\gamma_1}{\gamma_0} = \frac{\phi \gamma_0}{\gamma_0} = \phi$$

$$\rho_2 = \frac{\gamma_2}{\gamma_0} = \frac{\phi^2 \gamma_0}{\gamma_0} = \phi^2$$

}

$$\rho_k = \frac{\gamma_k}{\gamma_0} = \frac{\phi^k \gamma_0}{\gamma_0} = \phi^k$$



Properties of an MA(1);  $y_t = c + \varepsilon_t + \theta \varepsilon_{t-1}$ ,  $\varepsilon_t \sim NID(0, \sigma^2)$   $|\theta| < 1$

$$E(y_t) = \mu = c$$

$$\text{var}(y_t) = \gamma_0 = E(y_t - \mu)^2 = E(\varepsilon_t)^2 + \theta^2 E(\varepsilon_{t-1})^2 + \cancel{\text{cov}(\varepsilon_t, \varepsilon_{t-1})} = \sigma^2 + \theta^2 \sigma^2 = (1 + \theta^2) \sigma^2$$

$$\text{cov}(y_t, y_{t-1}) = \gamma_1 = E(y_t - \mu)(y_{t-1} - \mu) = E(\varepsilon_t + \theta \varepsilon_{t-1})(\varepsilon_{t-1} + \theta \varepsilon_{t-2}) = \theta E(\varepsilon_{t-1})^2 = \theta \sigma^2$$

$$\gamma_2 = E(y_t - \mu)(y_{t-2} - \mu) = E(\varepsilon_t + \theta \varepsilon_{t-1})(\varepsilon_{t-2} - \mu) = 0$$

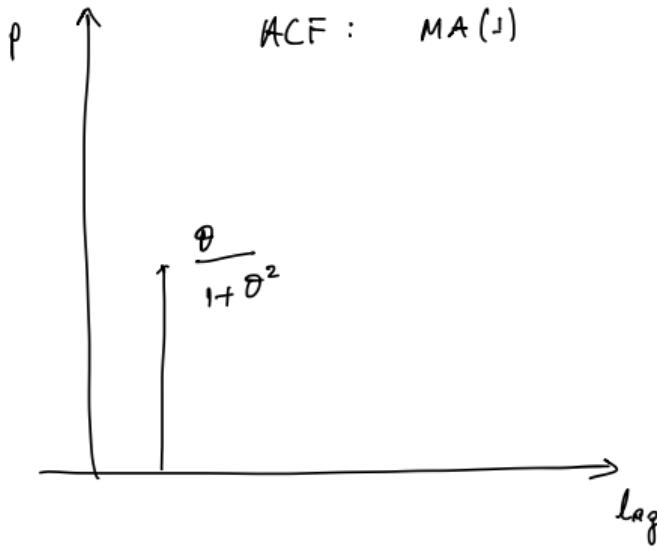
In fact for an MA(q) all  $\gamma_j = 0$   $j > q$

$$\rho_1 = \frac{\gamma_1}{\gamma_0} = \frac{\theta}{(1+\theta^2)}$$

$$\rho_2 = 0$$

{

$$\rho_k = 0$$



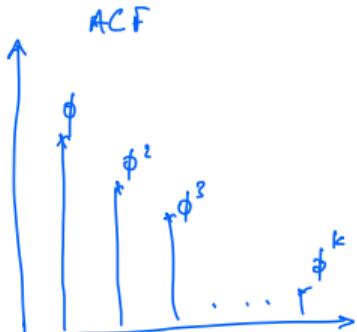
$$AR(1) \quad y_t = c + \phi y_{t-1} + \varepsilon_t \quad |\phi| < 1$$

(stationary)

$$E(y_t) = \mu = \frac{c}{1-\phi}$$

$$\text{Var}(y_t) = \gamma_0 = \frac{\sigma^2}{1-\phi^2}$$

$$\gamma_k = \phi^k \gamma_0 \quad \rho_k = \frac{\gamma_k}{\gamma_0} = \phi^k$$



$$MA(1) \quad y_t = c + \theta \varepsilon_{t-1} + \varepsilon_t \quad |\theta| < 1$$

(invertible)

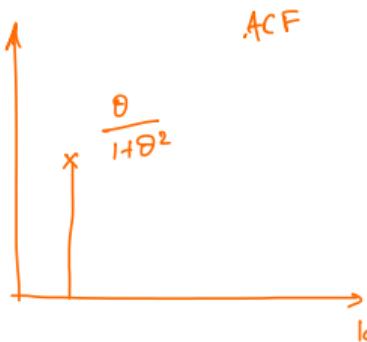
$$E(y_t) = \mu = c$$

$$\text{Var}(y_t) = \gamma_0 = (1+\theta^2)\sigma^2$$

$$\gamma_1 = \theta \sigma^2 \quad \gamma_2 = \dots = \gamma_k = 0$$

In general  $\gamma_j = 0 \quad j > q$

$$\rho_1 = \frac{\theta}{1+\theta^2} ; \quad \rho_2 = \dots = \rho_k = 0$$



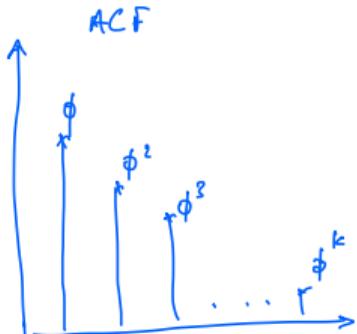
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$$\gamma_k = \phi^k \gamma_0 \quad \rho_k = \frac{\gamma_k}{\gamma_0} = \phi^k$$



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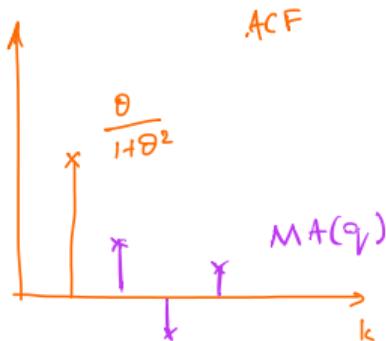
$$E(y_t) = \mu = c$$

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$$\gamma_1 = \theta \sigma^2 \quad \gamma_2 = \dots = \gamma_k = 0$$

In general  $\gamma_j = 0 \quad j > q$

$$\rho_1 = \frac{\theta}{1+\theta^2} ; \quad \rho_2 = \dots = \rho_k = 0$$



# ARIMA models

**ARIMA( $p, d, q$ ) model:**  $\phi(B)(1 - B)^d y_t = c + \theta(B)\varepsilon_t$

AR:  $p$  = order of the autoregressive part

I:  $d$  = degree of first differencing involved

MA:  $q$  = order of the moving average part.

# ARIMA models

**ARIMA( $p, d, q$ ) model:**  $\phi(B)(1 - B)^d y_t = c + \theta(B)\varepsilon_t$

- AR:  $p$  = order of the autoregressive part
- I:  $d$  = degree of first differencing involved
- MA:  $q$  = order of the moving average part.

- Conditions on AR coefficients ensure stationarity.
- Conditions on MA coefficients ensure invertibility.
- White noise model: ARIMA(0,0,0)  $y_t = \varepsilon_t$
- I(1)  
■ Random walk: ARIMA(0,1,0) with no constant  $y_t' = \varepsilon_t \Rightarrow y_t = y_{t-1} + \varepsilon_t$
- Random walk with drift: ARIMA(0,1,0) with const.  $y_t' = C + \varepsilon_t \Rightarrow y_t = C + y_{t-1} + \varepsilon_t$
- AR( $p$ ): ARIMA( $p, 0, 0$ )
- MA( $q$ ): ARIMA( $0, 0, q$ )

# R model

## Intercept form



$$(1 - \phi_1 B - \cdots - \phi_p B^p) y'_t = c + (1 + \theta_1 B + \cdots + \theta_q B^q) \varepsilon_t$$

## Mean form

$$(1 - \phi_1 B - \cdots - \phi_p B^p)(y'_t - \mu) = (1 + \theta_1 B + \cdots + \theta_q B^q) \varepsilon_t$$

\* Other books use this. Don't get confused.

- $y'_t = (1 - B)^d y_t$
- $\mu$  is the mean of  $y'_t$ .
- $c = \mu(1 - \phi_1 - \cdots - \phi_p)$ .
- ARIMA() in the fable package uses intercept form.

Recall : the mean of the (differenced) data  
is not the constant.

# Understanding ARIMA models

\* VERY IMPORTANT TO UNDERSTAND

$(P, q) \rightarrow$  short-run

$(C, d) \rightarrow$  long-run

- If  $c = 0$  and  $d = 0$ , the long-term forecasts will go to zero.
- If  $c = 0$  and  $d = 1$ , the long-term forecasts will go to a non-zero constant.
- If  $c = 0$  and  $d = 2$ , the long-term forecasts will follow a straight line.
- If  $c \neq 0$  and  $d = 0$ , the long-term forecasts will go to the mean of the data.
- If  $c \neq 0$  and  $d = 1$ , the long-term forecasts will follow a straight line.
- If  $c \neq 0$  and  $d = 2$ , the long-term forecasts will follow a quadratic trend.

## SUMMARY

constant

$c = 0$	$d = 0$	$\hat{y}_{T+\infty} \rightarrow 0$	ARMA start around $E(y_t) = 0$
$c \neq 0$	$d = 0$	$\hat{y}_{T+\infty} \rightarrow E(y_t)$	" " " $E(y_t) \neq 0$
$c = 0$	$d = 1$	$\hat{y}_{T+\infty} \rightarrow \text{const}$	RW + ARMA • diff of RW is stat and the ARMA part will converge to const.

linear trend

$c \neq 0$	$d = 1$	RW + drift + ARMA
$c = 0$	$d = 2$	$\hat{y}_{T+\infty} \rightarrow t$ Two unit roots

quadratic trend

$c \neq 0$	$d = 2$	Do IT At Your Own Risk.
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# Understanding ARIMA models

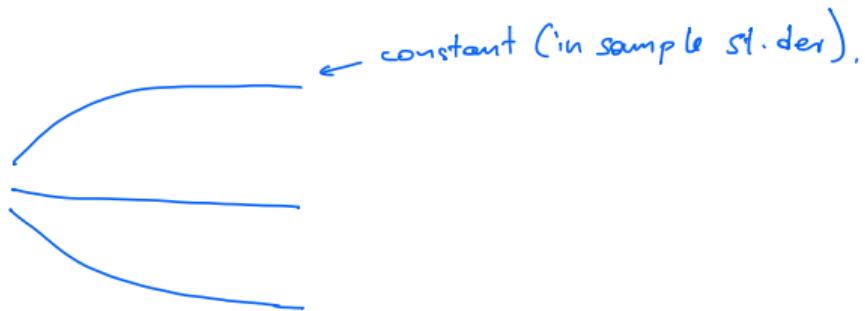
## Forecast variance and $d$

- The higher the value of  $d$ , the more rapidly the prediction intervals **increase in size**.
- For  $d = 0$ , the long-term forecast standard deviation will **go to the standard deviation** of the historical data.

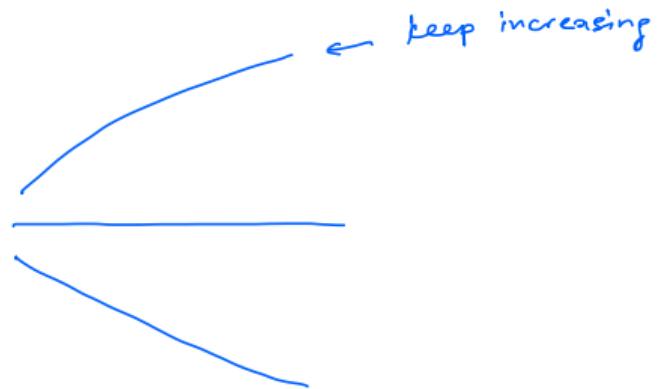
## Cyclic behaviour

- For cyclic forecasts,  $p \geq 2$  and some restrictions on coefficients are required.
- If  $p = 2$ , we need  $\phi_1^2 + 4\phi_2 < 0$ . Then average cycle of length
$$(2\pi)/[\text{arc cos}(-\phi_1(1 - \phi_2)/(4\phi_2))]$$
.

$d = 0$



$d = 1$



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# Partial autocorrelations

Partial autocorrelations measure relationship between  $y_t$  and  $y_{t-k}$ , when the effects of other time lags — 1, 2, 3, . . . ,  $k - 1$  — are removed. Why . PTO →

# Partial autocorrelations

Partial autocorrelations measure relationship between  $y_t$  and  $y_{t-k}$ , when the effects of other time lags — 1, 2, 3, …,  $k - 1$  — are removed.

$\alpha_k$  = kth partial autocorrelation coefficient  
= equal to the estimate of  $\phi_k$  in regression:

$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_k y_{t-k}. \neq \rho_1, \dots, \rho_k \text{ in ACF}$$

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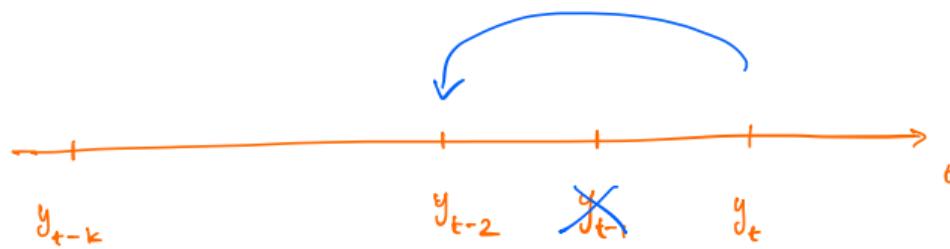
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$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_k y_{t-k}.$$

- Varying number of terms on RHS gives  $\alpha_k$  for different values of  $k$ .
- $\boxed{\alpha_1 = \rho_1}$  ALWAYS
- same critical values of  $\pm 1.96/\sqrt{T}$  as for ACF.
- \* ■ Last significant  $\alpha_k$  indicates the order of an AR model. \*

$$AR(1) \quad y_t = \phi y_{t-1} + \varepsilon_t \quad \Rightarrow \quad y_{t-1} = \phi y_{t-2} + \varepsilon_{t-1}$$

$$\Rightarrow y_t = \phi(\phi y_{t-2} + \varepsilon_{t-1}) + \varepsilon_t = \phi^2 y_{t-2} + \phi \varepsilon_{t-1} + \varepsilon_t$$



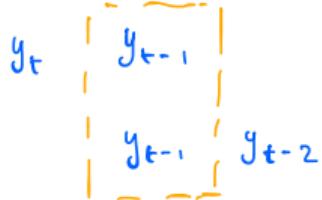
- Correlation between  $y_t$  &  $y_{t-k}$  conditioned on all other regressors
- this is what regression does

ACF

$$y_t = c + \rho_1 y_{t-1} + \varepsilon_t$$

$$y_t = c + \rho_2 y_{t-2} + \varepsilon_t$$

PACF



$$y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \varepsilon_t$$

$$\alpha_1 = \phi_1 = \rho_1 \quad (\text{Nothing between } y_t \text{ & } y_{t-1}) \quad \text{ALWAYS}$$

$$\alpha_2 = \phi_2 - \rho_2 \quad (\text{takes out the chaining/link of } y_{t-1})$$

$$\alpha_k = \phi_k = 0 \quad \text{for } k > p, \text{ hence } \max \phi_k \neq 0 \text{ gives you the AR order}$$

Note: the purpose of the PACF is largely to select the AR order. If ACF shows WN, PACF will also show WN ( $\alpha_1 = \phi_1 = 0$ ).

# ACF and PACF interpretation

## AR(1)

$$\rho_k = \phi_1^k \quad \text{for } k = 1, 2, \dots;$$

$$\alpha_1 = \phi_1 \quad \alpha_k = 0 \quad \text{for } k = 2, 3, \dots$$

So we have an AR(1) model when

- autocorrelations exponentially decay
- there is a single significant partial autocorrelation.

# ACF and PACF interpretation

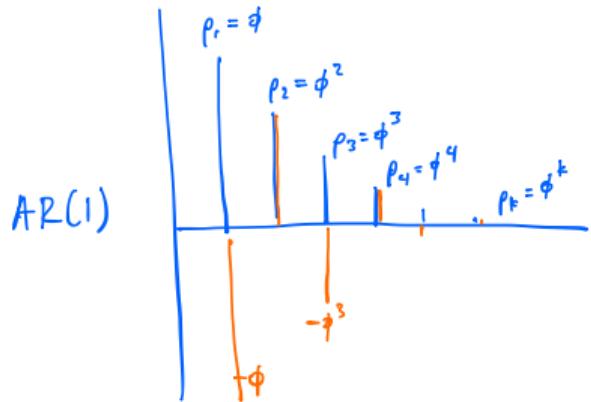
AR( $p$ )

- ACF dies out in an exponential or damped sine-wave manner
- PACF has all zero spikes beyond the  $p$ th spike

So we have an AR( $p$ ) model when

- the ACF is exponentially decaying or sinusoidal
- there is a significant spike at lag  $p$  in PACF, but none beyond  $p$

ACF

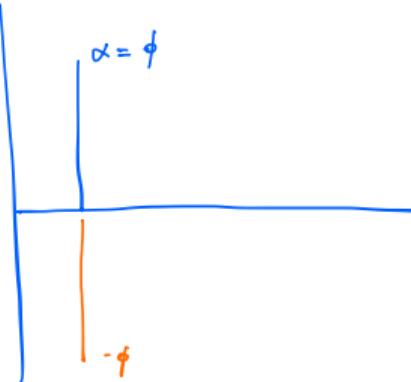


AR(1)

PACF \*

$0 < \phi < 1$

$-1 < \phi < 0$



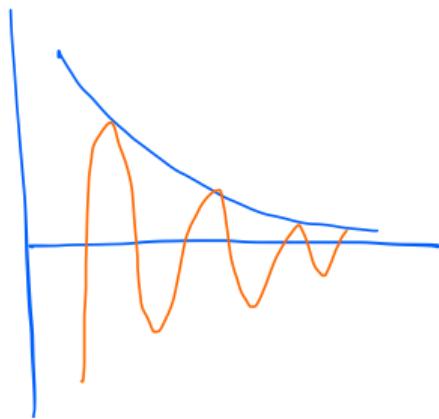
$\alpha_1 = \phi_1$

$\alpha_2 = \phi_2$

$\alpha_3 = \phi_3$

$\dots$

$\alpha_p = \phi_p$



AR( $p$ )

# ACF and PACF interpretation

## MA(1)

$$\begin{aligned}\rho_1 &= \theta_1 & \rho_k &= 0 & \text{for } k = 2, 3, \dots; \\ \alpha_k &= -(-\theta_1)^k\end{aligned}$$

So we have an **MA(1)** model when

- the PACF is exponentially decaying and
- there is a single significant spike in ACF

# ACF and PACF interpretation

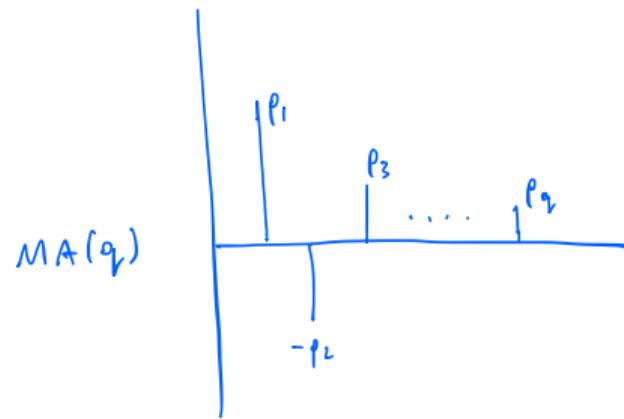
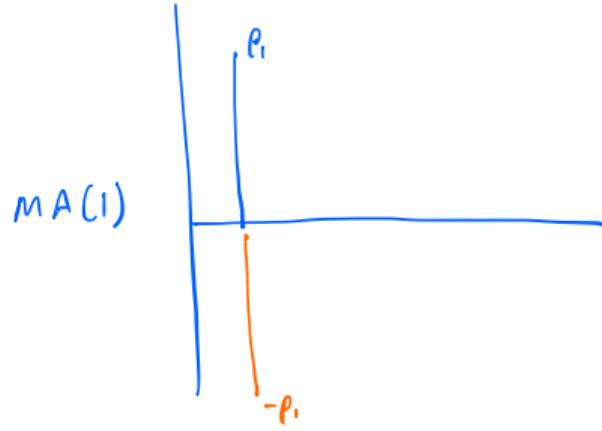
MA( $q$ )

- PACF dies out in an exponential or damped sine-wave manner
- ACF has all zero spikes beyond the  $q$ th spike

So we have an MA( $q$ ) model when

- the PACF is exponentially decaying or sinusoidal
- there is a significant spike at lag  $q$  in ACF, but none beyond  $q$

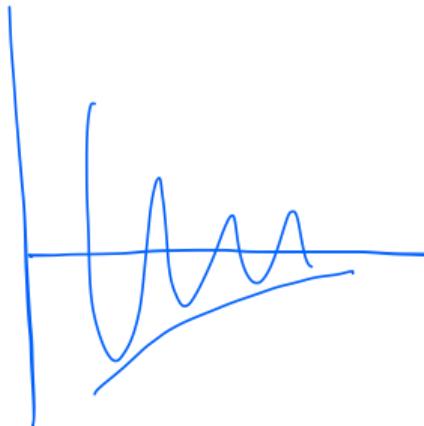
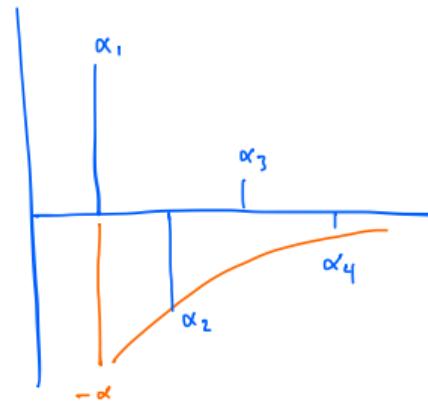
ACF \*



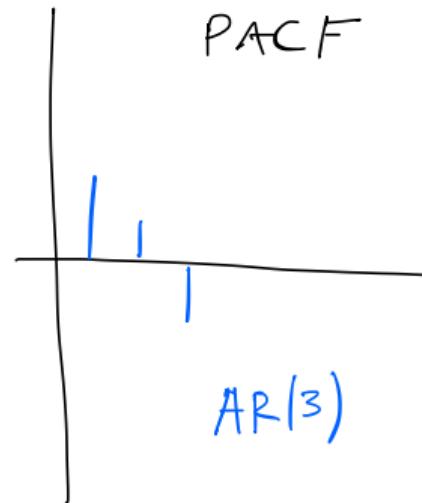
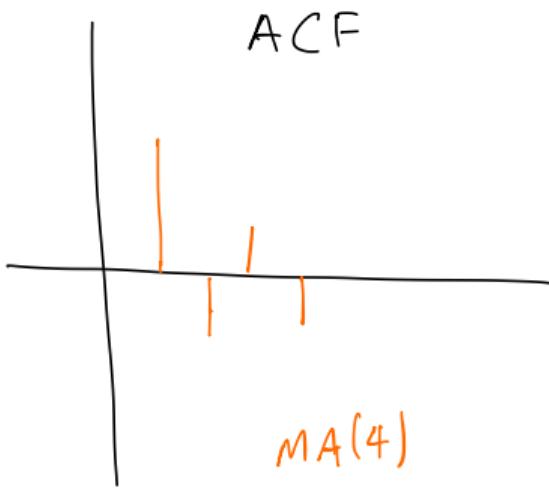
PACF

$$0 < \theta < 1$$

$$-1 < \theta < 0$$



## EXTRA JUST TO RECAP



- \* from theory / reading ACF / PACF we can tell either AR or MA orders

EGYPTIAN  
EXPORTS

ARIMA(2,0,1) with constant.

$$(1 - \phi_1 B - \phi_2 B^2) y_t = c + (1 + \theta_1 B) \varepsilon_t$$

EGYPTIAN  
EXPORTS

ARIMA(2,0,1) with constant.

$$(1 - \phi_1 B - \phi_2 B^2) y_t = c + (1 + \theta_1 B) \varepsilon_t$$

$$y_t - \phi_1 y_{t-1} - \phi_2 y_{t-2} = c + \varepsilon_t + \theta_1 \varepsilon_{t-1}$$

$$\Rightarrow y_t = c + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \theta_1 \varepsilon_{t-1} + \varepsilon_t$$

$$\Rightarrow y_t = 2.562 + 1.676 y_{t-1} - 0.803 y_{t-2} - 0.69 \varepsilon_{t-1} + \varepsilon_t$$

where  $\varepsilon_t \sim NID(0, 8.046)$

$$\text{or } \hat{\sigma} = \sqrt{8.046} = 2.837$$