Use of multistate models to jointly model progression-free and overall survival and improve decision-making in clinical trials

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Who

Meller et al. (2019):







Beyer et al. (2019):











Multistate model for PFS and OS

Introduction

Oncology endpoints:

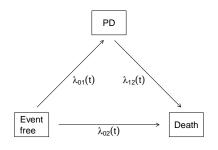
- Progression-free survival (PFS): Time from randomization to earlier of progression or death.
- Overall survival (OS): Time from randomization to death.

PFS common surrogate for OS in clinical trials.

Sophisticated methods to quantify amount of surrogacy, e.g. Buyse et al. (2016).

Correlation between PFS and OS important aspect, Li and Zhang (2015).

Multistate model for PFS and OS

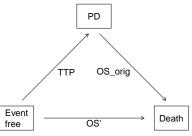


Standard illness-death model without recovery:

- Process $X(t) \in \{0, 1, 2\}, t \ge 0$ models the state occupied at time t.
- All patients in state 0 at time 0: P(X(0) = 0) = 1.
- PFS: waiting time in initial state 0, PFS = $\inf\{t : X(t) \neq 0\}$.
- OS: time until reaching state 2, OS = $\inf\{t : X(t) = 2\}$.

Alternatives

Latent failure time model



Latent failure time model (LFTM), Fleischer et al. (2009), Li and Zhang (2015):

- PFS = $min(TTP, OS_{orig})$.
- $\bullet \ \mathsf{OS} = \mathsf{PFS} \ \mathsf{if} \ \mathsf{PFS} \neq \mathsf{TTP}, \ \mathsf{TTP} + \mathsf{OS'} \ \mathsf{else}.$

Challenges:

- Impossible sampling space: TTP > OS_{orig} ⇒ progression after death ⇒ awkward idea.
- Issue with assumptions for estimation.

Copulas

Copulas, e.g. Burzykowski et al. (2001), Fu et al. (2013), Emura et al. (2017):

Model general bivariate survival data (lifetimes of twins).

Challenges:

- PFS OS structure more specific.
- PFS \leq OS \Rightarrow copulas do not place such a restriction on pair of event times.
- Reality: death without progression ⇒ P(PFS = OS) > 0. Copula model with continuous "marginal" survival functions for PFS and OS: P(PFS = OS) = 0.

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Multistate model formulation

Transition probabilities:

- Full description of multistate model by only assuming existence of intensities α_{01} , α_{02} and α_{12} .
- Formulas, even for non-Markov case: Aalen et al. (2008).

Meller et al. (2019):

- Embed PFS and OS in multistate model framework,
- formulas for P_{lm}'s assuming Weibull transition hazards for time-inhomogeneous Markov and semi-Markov (explicit),
- inference via counting process likelihood,
- $P(PFS \le u, OS \le v)$ for X non-Markov (generic).

Allows derivation of any functional of PFS and OS.

Exemplary application: Pearson correlation.

Multistate model for PFS and OS

Marginal distributions:

$$S_{PFS}(t) = P(PFS > t) = P_{00}(0, t),$$

 $S_{OS}(t) = P(OS > t) = P_{00}(0, t) + P_{01}(0, t),$

Joint distribution:

$$\begin{split} P(\text{PFS} \leq u, \text{OS} \leq v) &= P(X(u) \in \{1, 2\}, X(v) = 2) \\ &= P(X(v) = 2 | X(u) = 1) \cdot P_{01}(0, u) + P_{02}(0, u). \end{split}$$

X inhomogeneous Markov: $P(X(v) = 2|X(u) = 1) = P_{12}(u, v)$ independent of progression time $t_1 \le u$.

X non-Markov:

- Integrate P₁₂(u, v; t₁) over conditional distribution of all possible progression times t₁ < u.
- Formula tedious (see Meller et al. (2019)) ⇒ simulate in applications.

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Advantages of multistate model

Multistate model:

- Assumptions on X(t) induce properties of transition intensities, (joint) probabilities, and thus PFS and OS.
- No progression after death.
- PFS = OS easily possible.
- Estimation: No assumption about "in-/dependence" of PFS and OS.

Multistate = (most?) parsimonious model

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Correlation coefficient

Correlation coefficient

$$\mathrm{Corr}(\mathrm{PFS},\mathrm{OS}) \quad = \quad \frac{\mathrm{Cov}(\mathrm{PFS},\mathrm{OS})}{\sqrt{\mathrm{Var}(\mathrm{PFS})\,\mathrm{Var}(\mathrm{OS})}} \ = \ \frac{\mathbb{E}(\mathrm{PFS}\cdot\mathrm{OS}) - \mathbb{E}(\mathrm{PFS})\,\mathbb{E}(\mathrm{OS})}{\sqrt{\mathrm{Var}(\mathrm{PFS})\,\mathrm{Var}(\mathrm{OS})}}.$$

Mean, variance of PFS and OS: via survival functions.

 $\mathbb{E}(PFS \cdot OS)$: Use

$$P(\mathrm{PFS}\cdot\mathrm{OS}>t) \quad = \quad P(\mathrm{PFS}>\sqrt{t}) + \int_{(0,\sqrt{t}]} P_{11}(u,t/u;u) P(\mathrm{PFS}>u-)\alpha_{01}(u) \ \mathrm{d}u.$$

Proof: manipulations using law of total probability.



Estimation and inference for Markov models

Parametric:

- Plug parametric assumption in formulas for $P_{lm}(s,t)$, S_{PFS} , S_{OS} , Corr(PFS, OS).
- Estimate parameters using Counting Process Likelihood, Andersen et al. (1993).
 Product of patient-specific likelihood-contributions to each state transition.
- Inference via delta method or bootstrap (results comparable).

Nonparametric:

- Transition probabilities: Aalen-Johansen estimator, Aalen and Johansen (1978).
- Plug in estimates into formulas for PFS, OS, Corr(PFS, OS).
- Challenge: need to extrapolate tail beyond where we have data.
- Inference via bootstrap.

Estimation and inference for Markov models

LFTM in Fleischer et al. (2009) and Li and Zhang (2015):

- Group patients depending on their path from 0 to 1 or 2, or censored.
- Likelihood uses assumption of independence of TTP, OS_{orig}. Cannot tell from (even uncensored!) data! Aalen (1987): "artificial problem", as LFTM not needed, see also Beyersmann et al. (2012).

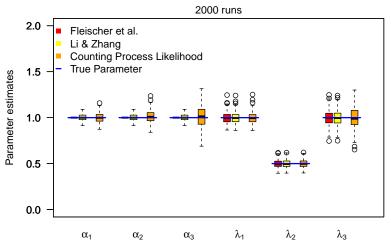
Weber and Titman (2019):

- Kendall's τ , based on multistate, nonparametric, and copula models.
- Use again LFTM for estimation.

Results

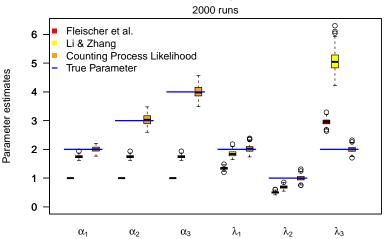
Results: estimated parameters Exponential

Estimated Parameters using different Likelihoods



Results: estimated parameters time-inhomogeneous Weibull

Estimated Parameters using different Likelihoods



Conclusions

Conclusions & outlook

Model PFS and OS within illness-death without recovery multistate model.

Advantages:

- Compared to LFTM avoids questionable and uncheckable assumptions,
- properties of PFS and OS induced through transparent assumption on X(t),
- allows for straightforward derivation of survival functions and correlation for parametric models, e.g. no need to assume common Weibull shape parameter as in Li and Zhang (2015) to get tractable formulas,
- allows for parametric and nonparametric estimation and inference (at least) in Markov models using standard multistate modelling tools,
- engine to simulate PFS and OS times.

Outlook:

- How to best extrapolate tail of nonparametric survival function estimates?
- Shorten time for computation of bootstrap confidence intervals?
- R package?

Multistate models for early decision-making

How do we typically decide whether to move an oncology molecule into Phase 3?

Decision-making in early oncology development

- **1** Small single-arm trial for **experimental** drug (e.g. n = 40).
- 2 Response proportion, duration of response.
- Ompare to "corresponding" quantities from literature for control treatment.

But:

- P(wrong decision) may be high.
- Primary endpoint in Phase 3: Overall survival.

Proposal:

Decide in early phase based on OS prediction.

Decrease P(wrong decision).

Challenges and proposal

Challenges:

- Response-type endpoint?
- Surrogacy? Poor in many indications.
- Immunotherapy (CIT): no effect on response, relevant OS effect.
- **Non-randomized** comparison ⇒ confounding.

Proposal: Base decision-making on OS prediction from multistate model.

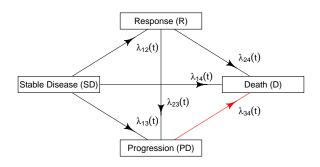
- Predicted survival function for experimental arm.
- **2** Combine S_{exp} with S_{control} to get **predicted OS HR**.
- ② Experimental drug might act on certain transitions only ⇒ not captured through simple modelling of OS. Potential efficiency gain!
- Propensity scoring.

Idealized scenario: Retrospective data from Phase 3 RCTs.

Long-term follow-up in both arms.

Randomization \Rightarrow no confounding.

Multistate model for early decision-making



- Follow-up of patient until PD or death without PD.
- Post-progression hazard λ_{34} : borrowing from historical data.
- \bullet Transitions SD \to D, R \to D rare, hazards \approx same in both arms.
- Markov assumption.

Predicted survival function in experimental arm, S_{exp}

Compute transition probabilities for each transition.

$$\begin{split} S_{\text{exp}}(t) &= 1 - \Big(P_{SD \to D}(0, t) + P_{SD \to \textbf{PD} \to \textbf{D}}(0, t) + \\ &P_{SD \to R \to D}(0, t) + P_{SD \to R \to \textbf{PD} \to \textbf{D}}(0, t)\Big). \end{split}$$

 λ_{34} corresponding to PD \rightarrow D transition borrowed from historical data.

Historical borrowing for λ_{34}

Experimental treatment expected to provide benefit beyond PD?

No:

- E.g. chemotherapy or antibody-dependent cellular cytotoxicity.
- Plug-in hazard function estimate from historical control.
- No post-PD information required for experimental arm.

Yes:

- E.g. chemoimmunotherapy.
- Estimate post-PD hazard ratio assuming proportionality.
- How much post-PD deaths needed in experimental arm to reliably estimate post-PD HR?

Benefit beyond PD: Oak

Oak

Previously treated non-small-cell lung cancer.

Rittmeyer et al. (2017).

	Atezolizumab	Chemotherapy	Hazard ratio
Effect post-PD	expected	not expected	
Objective Response	58 (13.6%)	57 (13.4%)	
Duration of Response	26.3 (10 - ∞)	6.2 (4.9 - 7.6)	
Overall Survival			0.73 (0.62, 0.87)

If this were early phase data - would you initiate Phase 3?

Competitors used this mechanism of action.

OS prediction when post-PD hazards assumed proportional

Random variable:

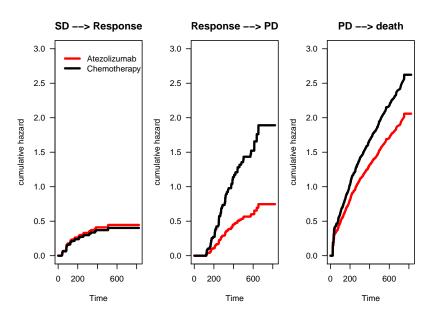
$$Z = \begin{cases} 0 & \text{if patient in control,} \\ 1 & \text{if in experimental group.} \end{cases}$$

$$\lambda_{34}(t | Z) = \lambda_{34,0}(t) \exp(\beta_{34} Z)$$

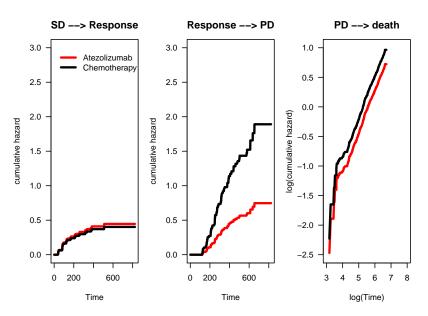
Baseline hazard $\lambda_{34,0}$ estimated from both arms combined.

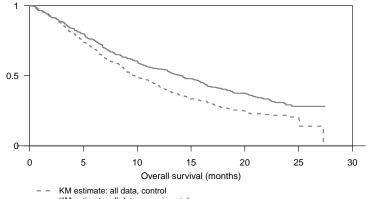
Post-progression hazard ratio β_{34} ?

Oak: raw cumulative hazard estimates (of interest)

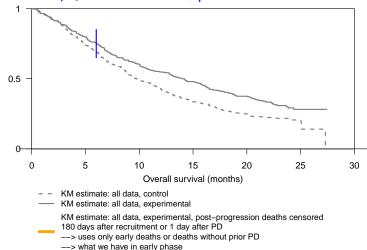


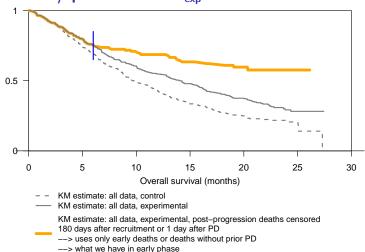
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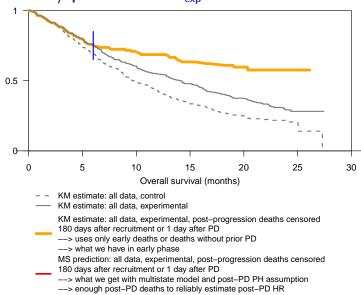


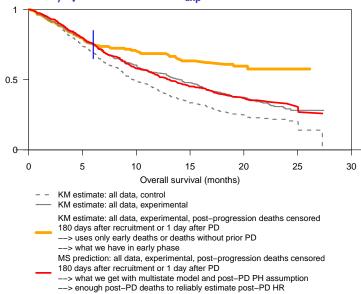


— KM estimate: all data, experimental









Early phase decision based on multistate prediction:

P(wrong decision)?

OS prediction from mimicked early phase data

Historical control: Oak control arm data.

False-positive decision: Sample early phase trial from Oak control arm.

False-negative decision: Sample early phase trial from Oak experimental arm.

Sample early phase trial:

- 40 patients,
- 6 months uniform recruitment,
- analysis 15 months after first patient entered,
- censor post-PD follow-up one day after PD,
- estimate $\lambda_{12}, \lambda_{13}, \lambda_{14}, \lambda_{23}, \lambda_{24}$ from this data.

Cox regression for post-PD transition $\Rightarrow \widehat{\lambda}_{34}(t|Z)$.

Compute prediction of S_{exp} .

OS HR prediction based on early phase trial

Approximate HR by fitting exponential distribution to both arms $\Rightarrow \widehat{HR}$.

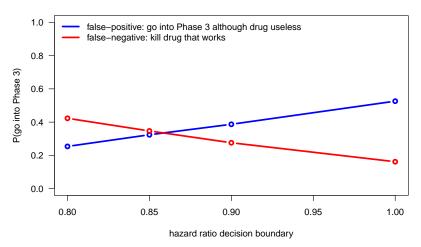
Decision to move to Phase 3: $\widehat{\textit{HR}} \leq \text{boundary} \in \{0.80, 0.85, 0.90, 1.00\}.$

Repeat 1000 times.

 ${\sf Resampling} \Rightarrow \textbf{quantification of uncertainty}.$

Oak: P(wrong decision)

P(go into Phase 3) = P(approximated HR <= boundary)



How many post-PD deaths to estimate HR of PD → death transition?

How many post-PD deaths to estimate HR of PD → death transition?

Ask during Q&A.



Conclusions

Early phase decision-making based on multistate OS prediction:

- Assumption on $\lambda_{34} \Rightarrow$ need to understand disease and treatment.
- Avoids difficulty in interpretation of response-type endpoints.
- Feasibility assessed in idealized scenario.
- Recommendation how much post-PD follow-up needed to estimate β_{34} .
- Needs long-term individual-patient data in control arm!

What about confounding?

Real-world data as historical control.

Combine proposal with propensity scoring.

Conclusions

Multistate models

Multistate models useful:

- Canonical extension of survival analysis.
- Get more insight in how disease and drug work.
- Prediction in well-specified, as opposed to black-box, model.
- Jointly model three key oncology endpoints: response, PFS, OS.
- Applications by no means restricted to oncology!

Many potential applications:

- Improved early stage decision-making ⇒ Beyer et al. (2019).
- Improved communication of effect and optimized sample size computation.
- Bivariate modelling of PFS and OS to help inform surrogacy questions ⇒ Meller et al. (2019).

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Conclusions

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Thank you for your attention.

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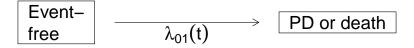
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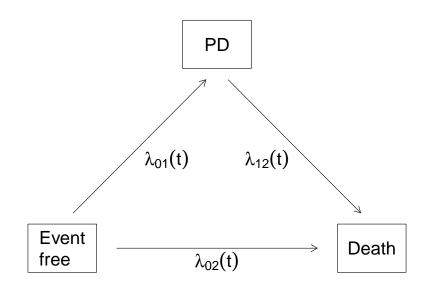
Backup

Multistate models

Canonical extension of survival analysis



Canonical extension of survival analysis



Multistate models

Multistate model:

- 1-1 correspondence hazard probability breaks down.
- Transition probabilities: (Markov) process $X(t)_{t\geq 0}$ with state space $\{0,1,2\}=\{$ event-free, progression, death $\}$. Then,

$$P_{lj}(s,t) := P(X_t = j | X_s = l, Past).$$

- Estimate P_{li} 's nonparametrically by Aalen-Johansen estimator.
- PFS: Kaplan-Meier of time-to-progression simply censoring death is biased!
- OS: Aalen-Johansen offers higher precision compared to simple Kaplan-Meier estimate, Andersen et al. (1993) (p. 315 and Fig. IV.4.16).
- Markov assumption stronger than what is needed for Kaplan-Meier though.

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Prediction in multistate models

Rates (hazards, intensities):

- Modelling of effects of covariates on transition hazards.
- Hazard ratios (HR) from Cox regression.

Transition probabilities look at cumulative effects:

- Effects on transition probabilities may be different from what HRs suggest.
- Intermediate events in multistate model also contribute to cumulative effects.
- How to estimate such cumulative effects?

Prediction from multistate model!

Prediction in multistate models

General problem: estimate conditional probability of some future clinical event, given

- event history,
- set of values for prognostic factors of a patient.

Derive formulas for these conditional probabilities, or simulate.

Final result: survival function for OS, as function of

- covariates and
- relevant cumulative hazards.

PFS - OS

Multistate vs. latent failure time model

Fleischer *et al.* (2009), Li and Zhang (2015): LFTM with uncheckable and questionable (unrealistic?) independence assumption.

Connections to multistate model? We are still figuring that out, work in progress.

Parametric models: formula for S_{PES} identical for all three models below, and

- Time-homogeneous Markov, Exponential: model so simple that ∄ time-inhomogeneous Markov process. Sos identical to Exponential LFTM.
- Time-homogeneous Markov, Weibull: formula for S_{OS} identical to Weibull LFTM ⇒ are model assumptions equivalent? No!
- Time-inhomogeneous Markov, Weibull: formulas for S_{OS} are different.

BUT: values of estimated parameters differ between LFTM and multistate for all three parametric models, as likelihoods differ!

Not clear (?) how to nonparametrically estimate LFTM \Rightarrow possible for (Markov) multistate.

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Assumptions for multistate model

Assumptions for multistate model

Multistate model sufficiently smooth so that following intensities exist:

$$\begin{split} \alpha_{0j}(t) &= \lim_{\Delta t \searrow 0} \frac{P(\text{PFS} \in [t, t + \Delta t), X(\text{PFS}) = j \mid \text{PFS} \ge t)}{\Delta t}, j = 1, 2, \\ \alpha_{12}(t; t_1) &= \lim_{\Delta t \searrow 0} \frac{P(X(t + \Delta t) = 2 \mid X(t -) = 1, \text{PFS} = t_1)}{\Delta t} \\ &= \lim_{\Delta t \searrow 0} \frac{P(\text{OS} - \text{PFS} \in [t - t_1, t - t_1 + \Delta t) \mid \text{OS} \ge t, \text{PFS} = t_1)}{\Delta t} \quad \text{for } \mathbf{t_1} < \mathbf{t}. \end{split}$$

t₁: observed PFS time, i.e. time when leaving state 0.

Assumptions for multistate model

X(t) Markov:

- Time-inhomogeneous: intensity of death after progression does not depend on time of progression, $\alpha_{12}(t;t_1) = \alpha_{12}(t)$ for all $t_1 < t$.
- Homogeneous: intensities are time-constant, i.e. Exponential, $\alpha_{ij}(t) = \alpha_{ij}, i, j = 0, 1, 2.$

X(t) non-Markov (= semi-Markov for illness-death model without recovery):

- Intensities depend on state patient is in at s and entire history $\leq s$, i.e. all transitions.
- ullet Relevant for 1 o 2 transition only, as 0 o 1, 2 are rooted in initial state 0.

As soon as a quantity depends on $1 \rightarrow 2$ transition we need to be specific about assumption on X(t).

Illness-death multistate model for PFS and OS

Transition probabilities to move from state I at time s to state m at time t:

$$P_{lm}(s,t) := P(X(t) = m|X(s) = l, \text{history}).$$

Illness-death model w/o recovery, P_{lm} as functions of transition intensities, Aalen *et al.* (2008):

$$P_{00}(s,t) = \exp\left(-\int_{s}^{t} \alpha_{01}(u) + \alpha_{02}(u) du\right),$$

$$P_{11}(s,t;\mathbf{t_{1}}) = \exp\left(-\int_{s}^{t} \alpha_{12}(u;t_{1}) du\right),$$

$$P_{22}(s,t) = 1,$$

$$P_{01}(s,t) = \int_{s}^{t} P_{00}(s,u_{-})\alpha_{01}(u)P_{11}(u,t;u) du,$$

$$P_{12}(s,t;\mathbf{t_{1}}) = 1 - P_{11}(s,t;\mathbf{t_{1}}),$$

$$P_{02}(s,t) = 1 - \left(P_{00}(s,t) + P_{01}(s,t)\right).$$

If X(t) non-Markov:

- P_{11} and P_{12} depend on PFS time t_1 .
- Although P_{01} , P_{02} depend on α_{12} they do not depend on t_1 .

Intuition behind transition probabilities

 $P_{00}(s,t)$, $P_{11}(s,t;t_1)$: exp of cumulative hazards \Rightarrow standard survival functions.

$$P_{01}(s,t) = \int_{s}^{t} P_{00}(s,u_{-})\alpha_{01}(u)P_{11}(u,t;u) du$$
: integral of

- $P_{00}(s, u_-)\alpha_{01}(u)$: "infinitesimal probabilities" to move from 0 to 1 at time u, $u \in (s, t]$,
- $P_{11}(u, t; u)$: subsequently stay in state 1 until at least time t, with progression happened in u.

Illness-death multistate model for PFS and OS

Marginal distributions:

$$S_{PFS}(t) = P(PFS > t) = P_{00}(0, t),$$

 $S_{OS}(t) = P(OS > t) = P_{00}(0, t) + P_{01}(0, t),$

Joint distribution:

$$P(PFS \le u, OS \le v) = P(X(u) \in \{1, 2\}, X(v) = 2)$$

$$= P(X(u) = 1, X(v) = 2) + P(X(u) = 2)$$

$$= P(X(v) = 2|X(u) = 1) \cdot P(X(u) = 1|X(0) = 0)$$

$$+ P(X(u) = 2|X(0) = 0)$$

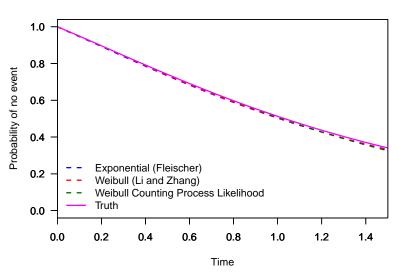
$$= P(X(v) = 2|X(u) = 1) \cdot P_{01}(0, u) + P_{02}(0, u).$$

X inhomogeneous Markov: $P(X(v) = 2|X(u) = 1) = P_{12}(u, v)$ independent of progression time $t_1 \le u$.

X non-Markov: integrate $P_{12}(u,v;t_1)$ over conditional distribution of all possible progression times $t_1 \leq u \Rightarrow$ final formula tedious.

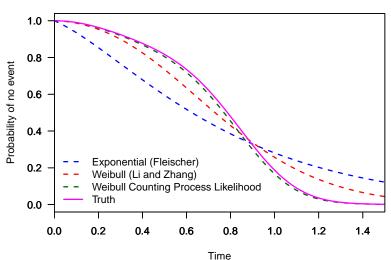
Results: S_{OS} for Exponential

Exponential

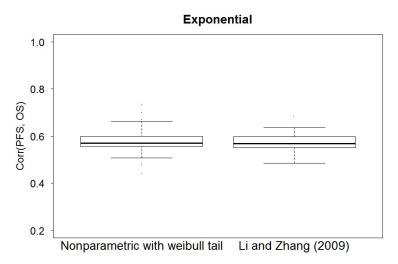


Results: S_{OS} for Weibull

Data from time-inhomogeneous Markov, Weibull with different shape, n = 500

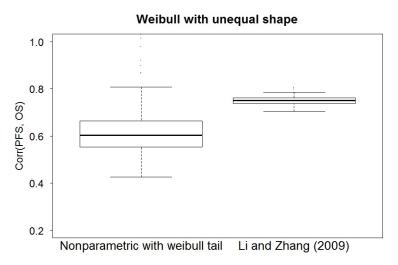


Results: correlations Exponential



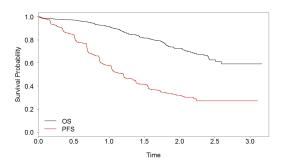
Corr(PFS, OS) for 200 simulated dataset from time-inhomogeneous Markov process.

Results: correlations Weibull



Corr(PFS, OS) for 200 simulated dataset from time-inhomogeneous Markov process.

Results: CLEOPATRA, Baselga and Cortes (2012).



			Weibull	Nonparametric
	Exponential	Weibull	Markov	Markov
Corr(PFS, OS)	0.611	0.643	0.483	0.450
95% Bootstrap CI	[0.541; 0.673]	[0.584; 0.699]	[0.342; 0.643]	[0.297; 0.655]

Table: Correlation between PFS and OS in CLEOPATRA (1000 bootstrap samples).

Early decision-making

How many post-PD deaths needed?

Assumption:

$$\lambda_{34}(t \mid \mathsf{Z}) = \lambda_{34,0}(t) \exp(\beta_{34} \mathsf{Z}).$$

How many post-PD deaths needed in experimental arm to reliably estimate λ_{34} ?

Planning stage: only data for control arm are available.

- Fit multistate model to control data.
- Assume transition-specific hazard ratios corresponding to clinically meaningful OS effect.
- Simulate.

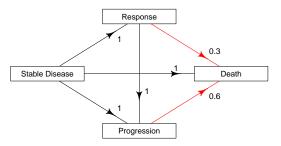
Various scenarios for post-PD follow-up time.

Simulation details – mimick Oak

NOT power computation for hypothesis test – sample size too small anyway.

Rather: find cutoff timepoint from which on OS HR estimate remains stable.

• Simulate 40 patient from experimental arm as before.

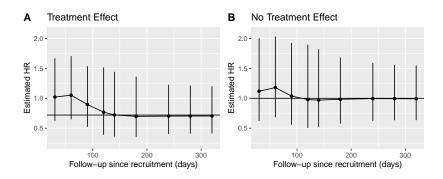


- Resulting OS HR = 0.73. Close to Oak OS HR.
- Follow-up post-PD for experimental arm truncated at 30, 60, 90, 120, 150, 180 and 240 days after recruitment.

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Repeat 1000 times.

Stability of hazard ratio estimate



180-240 days sufficient to obtain stable point estimate over time.

Typical early phase follow-up: Post-PD deaths censored 180 days after recruitment in experimental arm.

Kaspar Rufibach Multistate models in clinical trials Early decision-making

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Example 1: Cleopatra

Cleopatra

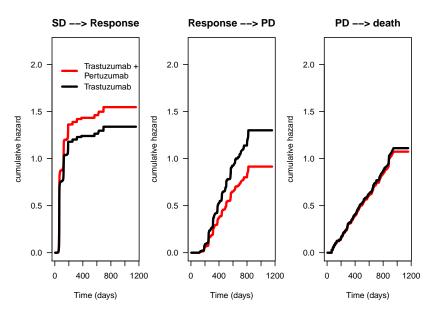
Baselga and Cortes (2012), Swain and Baselga (2015).

Previously untreated HER2-positive metastatic breast cancer patients.

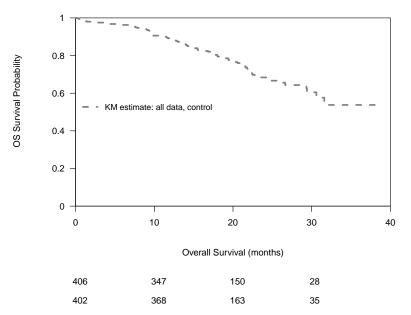
	Pertuzumab + Trastuzumab	Trastuzumab	HR (95% CI)
Survival	N=402	N=406	
Overall Survival			0.64
			(0.47, 0.88)
Progression-free Survival			0.62 (0.51,0.75)
Response	N=343	N=336	
Objective Response	275 (80.2%)	233 (69.3%)	
Stable Disease	50 (14.6%)	70 (20.8%)	
Progressive Disease	13 (3.8%)	28 (8.3%)	
Duration of Response	N=275	N=233	
Median (months, 95% CI)	20.2 (16.0,24.0)	12.5 (10.0-15.0)	

- Moderate difference in response.
- Prolonged duration of response in experimental arm.
- Clear OS benefit
- Experimental treatment induces antibody-dependent cellular cytotoxicity \Rightarrow no benefit beyond PD expected $\Rightarrow \lambda_{34}$ same in both arms.

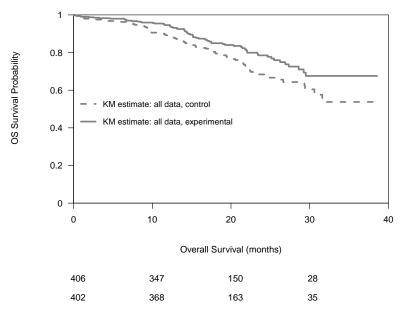
Cleopatra: raw cumulative hazard estimates (of interest)



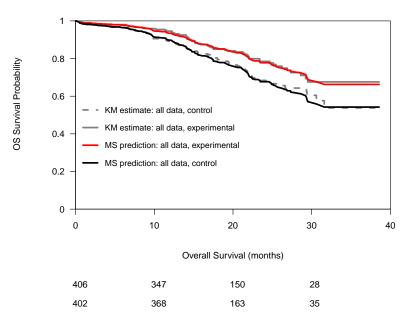
Cleopatra: estimates / predictions of S_{exp}



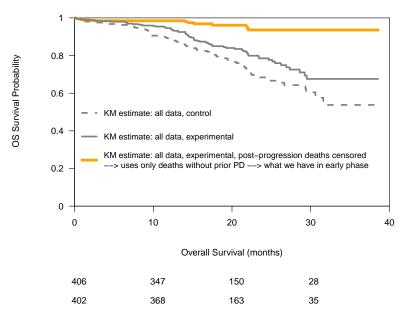
Cleopatra: estimates / predictions of S_{exp}



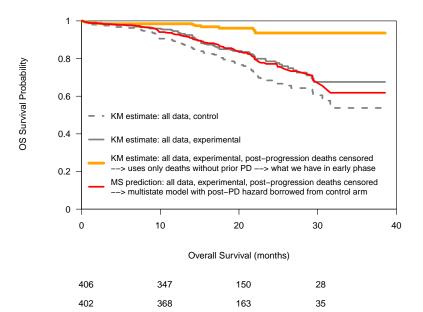
Cleopatra: estimates / predictions of S_{exp}



Cleopatra: estimates / predictions of S_{exp}



Cleopatra: estimates / predictions of S_{exp}



Conclusions for Cleopatra

For estimated / predicted survival function in experimental arm, based on all data:

- Majority of patients dies after observed PD.
- KM estimate of simply censoring post-PD deaths does not work ⇒ very few deaths observed.
- Multistate model prediction assuming post-PD hazards as in control provides good prediction.

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Early phase decision based on multistate prediction:

Operating characteristics?

OS prediction from mimicked early phase data

Sample early phase trial from Cleopatra experimental arm:

- 40 patients,
- 6 months uniform recruitment.
- analysis 15 months after first patient entered,
- censor post-PD follow-up one day after PD,
- estimate $\lambda_{12}, \lambda_{13}, \lambda_{14}, \lambda_{23}, \lambda_{24}$ from this data,
- borrow $\hat{\lambda}_{34}$ from historical data = Cleopatra control arm in idealized scenario,
- compute prediction of S_{exp} as described above.

Resampling of operating characteristics

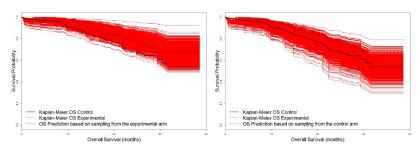
Setup:

- Use all data in control arm ⇒ corresponds to historical control.
- False-positive decision: Sample early phase trial from Cleopatra control arm.
- False-negative decision: Sample early phase trial from Cleopatra experimental arm.
- Approximate HR by fitting exponential distribution to both arms $\Rightarrow \widehat{HR}$.
- Decision to move to Phase 3: $\widehat{\mathit{HR}} \leq \mathsf{boundary} \in \{0.80, 0.85, 0.90, 1.00\}.$
- Repeat 1000 times.

Resampling easily allows for quantification of uncertainty.

Cleopatra: operating characteristics

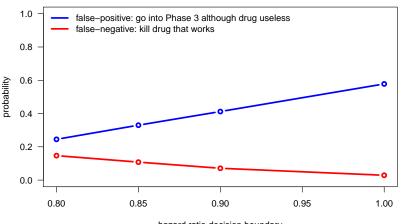
Sampled from experimental and control arm.



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Cleopatra: operating characteristics

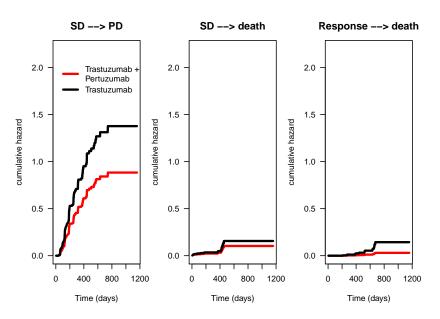
probability to go into Phase 3: P(approximated HR <= boundary)



hazard ratio decision boundary

Decision based on response: $\approx 10\%$ difference, some prolongation of DOR \Rightarrow moved to Phase 3.

Cleopatra: cumulative hazards of secondary interest



Oak

Previously treated non-small-cell lung cancer. Rittmeyer et al. (2017).

- Control: no benefit post-PD expected.
- Experimental: CIT ⇒ benefit post-PD expected.

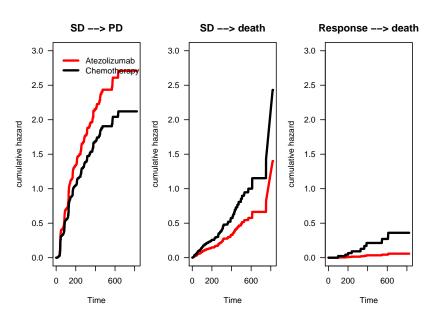
	Atezolizumab	Chemotherapy	HR (95% CI)
Survival	N=425	N=425	
Overall Survival			0.73 (0.62,0.87)
Progression-free Survival			0.95 (0.82,1.10)
Response	N=425	N=425	
Objective Response	58 (13.6%)	57 (13.4%)	
Stable Disease	150 (35%)	177 (42%)	
Progressive Disease	187 (44%)	117 (28%)	
Duration of Response	N=58	N=57	
Median (months, 95% CI)	26.3 (10,NE)	6.2 (4.9-7.6)	

No observed difference in response.

Prolonged duration of response in experimental arm.

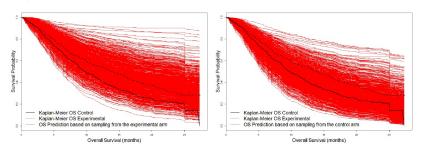
Clear survival benefit

Oak: cumulative hazards of secondary interest



Oak: operating characteristics

Sampled from experimental and control arm.



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Immunotherapy: 1) no difference in PFS, 2) non-proportional hazards for OS.

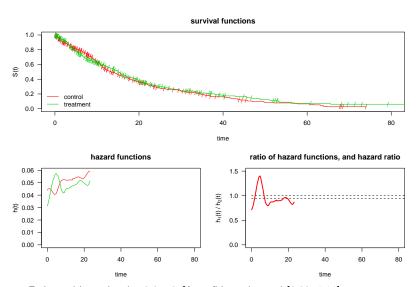
How to quantify effect?

A fictional clinical trial

Simulated clinical trial:

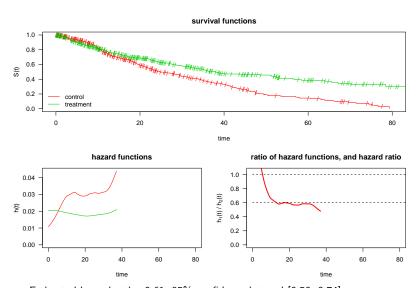
- 1:1 randomized, 400 and 400 patients per arm.
- No administrative censoring, but drop-out.

PFS for simulated clinical trial



- Estimated hazard ratio: 0.94, 95% confidence interval [0.80, 1.11].
- Hypothesis test for PH: p = 0.24.

OS for simulated clinical trial



- Estimated hazard ratio: 0.61, 95% confidence interval [0.50, 0.74].
- Hypothesis test for PH: p < 0.0001.

Summarize treatment effect

Non-proportional hazards for OS. How to summarize effect of treatment?

Data was generated according to:

Transition	Control arm	Treatment arm
0 ightarrow 1	$\lambda_{01}^c = \log(2)/25$	$\lambda_{01}^t = \lambda_{01}^c \cdot 1$
0 o 2	$\lambda_{02}^c = \log(2)/30$	$\lambda_{02}^t = \lambda_{02}^c \cdot 0.8$
1 o 2	$\lambda_{12}^c = \log(2)/15$	$\lambda_{12}^t = \lambda_{12}^c \cdot 0.4$

	coef	HR = exp(coef)	95% CI	<i>p</i> -value
transition event-free -> PD	-0.04	0.96	[0.77, 1.19]	0.72
transition event-free -> death	-0.09	0.91	[0.70, 1.18]	0.49
transition PD -> death	-1.09	0.34	[0.24, 0.46]	< 0.0001

Gaschler-Markefski et al. (2014).

Big vs. small data

Big vs. small data

Often, information removed/altered in small data:

- (Artificial) response cateogries instead of actual measurements: dichotomization,
- response proportions only: ignoring the dynamics between states,
- complicated subsets, e.g. those that respond only: selection bias,
- effect quantification in one number where biological process might suggest few numbers.
- ...

Maximize information from small data. AND look at BIG data.

Biostatisticians ideally placed to contribute to this!

Doing now what patients need next

R version and packages used to generate these slides:

R version: R version 3.6.0 (2019-04-26)

Base packages: stats / graphics / grDevices / utils / datasets / methods / base

 $Other\ packages:\ nls2\ /\ proto\ /\ diagram\ /\ shape\ /\ ggplot2\ /\ rocheBCE\ /\ muhaz\ /\ flexsurv\ /\ reporttools\ /\ xtable\ /\ mstate\ /\ etm\ /\ dplyr\ /\ nlssurv\ /\ reporttools\ /\ xtable\ /\ mstate\ /\ etm\ /\ dplyr\ /\ nlssurv\ /\ n$

mvna / prodlim / biostatKR / survival

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