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# Use of multistate models to jointly model progression-free and overall survival and improve decision-making in clinical trials

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# Who

Meller *et al.* (2019):



Beyer *et al.* (2019):



# Multistate model for PFS and OS

# Introduction

**Oncology** endpoints:

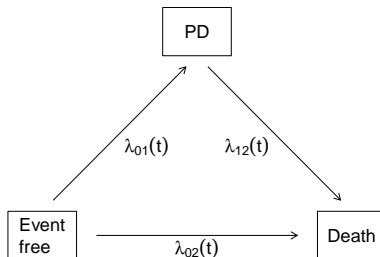
- Progression-free survival (PFS): Time from randomization to earlier of progression or death.
- Overall survival (OS): Time from randomization to death.

PFS common **surrogate** for OS in clinical trials.

Sophisticated methods to quantify amount of surrogacy, e.g. [Buyse et al. \(2016\)](#).

**Correlation** between PFS and OS important aspect, [Li and Zhang \(2015\)](#).

# Multistate model for PFS and OS

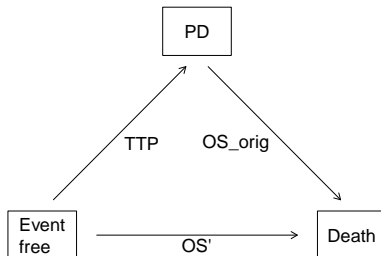


Standard **illness-death model without recovery**:

- Process  $X(t) \in \{0, 1, 2\}$ ,  $t \geq 0$  models the state occupied at time  $t$ .
- All patients in state 0 at time 0:  $P(X(0) = 0) = 1$ .
- PFS: waiting time in initial state 0,  **$\text{PFS} = \inf\{t : X(t) \neq 0\}$** .
- OS: time until reaching state 2,  **$\text{OS} = \inf\{t : X(t) = 2\}$** .

# Alternatives

# Latent failure time model



**Latent failure time model** (LFTM), [Fleischer et al. \(2009\)](#), [Li and Zhang \(2015\)](#):

- $PFS = \min(TTP, OS_{orig})$ .
- $OS = PFS$  if  $PFS \neq TTP$ ,  $TTP + OS'$  else.

Challenges:

- **Impossible sampling space:**  $TTP > OS_{orig} \Rightarrow$  progression after death  $\Rightarrow$  awkward idea.
- Issue with assumptions for estimation.

# Copulas

**Copulas**, e.g. Burzykowski *et al.* (2001), Fu *et al.* (2013), Emura *et al.* (2017):

- Model general bivariate survival data (lifetimes of twins).

Challenges:

- PFS - OS structure more specific.
- $\text{PFS} \leq \text{OS} \Rightarrow$  copulas do not place such a restriction on pair of event times.
- Reality: death without progression  $\Rightarrow P(\text{PFS} = \text{OS}) > 0$ . Copula model with **continuous** “marginal” survival functions for PFS and OS:  $P(\text{PFS} = \text{OS}) = 0$ .



# Multistate model formulation

Transition probabilities:

- **Full description** of multistate model by only assuming existence of intensities  $\alpha_{01}$ ,  $\alpha_{02}$  and  $\alpha_{12}$ .
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- Embed PFS and OS in multistate model framework,
- formulas for  $P_{lm}$ 's assuming **Weibull** transition hazards for time-inhomogeneous Markov and semi-Markov (explicit),
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Exemplary application: **Pearson correlation**.

# Multistate model for PFS and OS

**Marginal** distributions:

$$S_{PFS}(t) = P(\text{PFS} > t) = P_{00}(0, t),$$

$$S_{OS}(t) = P(\text{OS} > t) = P_{00}(0, t) + P_{01}(0, t),$$

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**Joint** distribution:

$$\begin{aligned} P(\text{PFS} \leq u, \text{OS} \leq v) &= P(X(u) \in \{1, 2\}, X(v) = 2) + P(X(u) = 2 | X(0) = 0) \\ &= P(X(v) = 2 | X(u) = 1) \cdot P_{01}(0, u; u) + P_{02}(0, u; u). \end{aligned}$$

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$X$  **non-Markov**:

- Integrate  $P_{12}(u, v; t_1)$  over conditional distribution of all possible progression times  $t_1 \leq u$ .
- Formula tedious (see [Meller et al. \(2019\)](#))  $\Rightarrow$  **simulate** in applications.



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Multistate = (most?) parsimonious model

# Correlation coefficient

## Correlation coefficient

$$\text{Corr}(\text{PFS}, \text{OS}) = \frac{\text{Cov}(\text{PFS}, \text{OS})}{\sqrt{\text{Var}(\text{PFS}) \text{Var}(\text{OS})}} = \frac{\mathbb{E}(\text{PFS} \cdot \text{OS}) - \mathbb{E}(\text{PFS}) \mathbb{E}(\text{OS})}{\sqrt{\text{Var}(\text{PFS}) \text{Var}(\text{OS})}}.$$

Mean, variance of PFS and OS: via survival functions.

$\mathbb{E}(\text{PFS} \cdot \text{OS})$ : Use

$$P(\text{PFS} \cdot \text{OS} > t) = P(\text{PFS} > \sqrt{t}) + \int_{(0, \sqrt{t}]} P_{11}(u, t/u; u) P(\text{PFS} > u-) \alpha_{01}(u) \, du.$$

Proof: manipulations using law of total probability.



# Estimation and inference for Markov models

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## Parametric:

- Plug parametric assumption in formulas for  $P_{lm}(s, t)$ ,  $S_{PFS}$ ,  $S_{OS}$ ,  $\text{Corr}(\text{PFS}, \text{OS})$ .
- Estimate parameters using **Counting Process Likelihood**, Andersen *et al.* (1993).  
Product of patient-specific likelihood-contributions to each state transition.
- Inference via delta method or bootstrap (results comparable).

## Nonparametric:

- Transition probabilities: Aalen-Johansen estimator, Aalen and Johansen (1978).
- Plug in estimates into formulas for PFS, OS,  $\text{Corr}(\text{PFS}, \text{OS})$ .
- Challenge: need to **extrapolate tail beyond where we have data**.
- Inference via bootstrap.

# Estimation and inference for Markov models

LFTM in [Fleischer et al. \(2009\)](#) and [Li and Zhang \(2015\)](#):

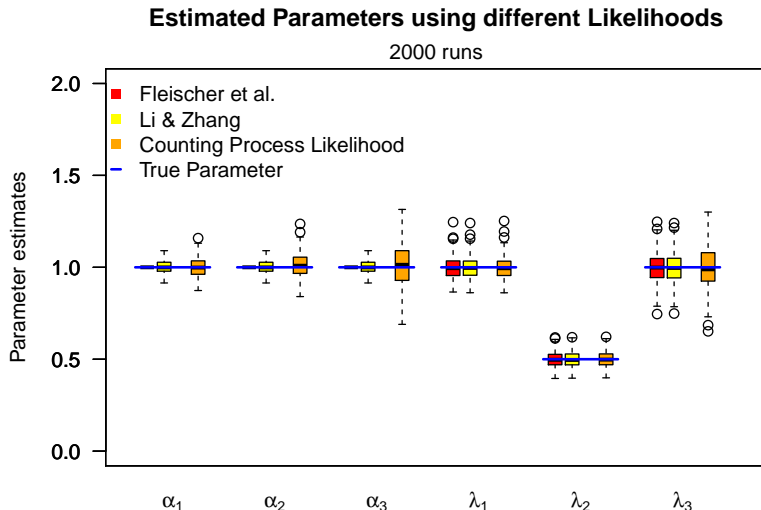
- Group patients depending on their path from 0 to 1 or 2, or censored.
- Likelihood uses **assumption of independence** of TTP,  $OS_{\text{orig}}$ . Cannot tell from (even uncensored!) data! [Aalen \(1987\)](#): “artificial problem”, as LFTM not needed, see also [Beyersmann et al. \(2012\)](#).

[Weber and Titman \(2019\)](#):

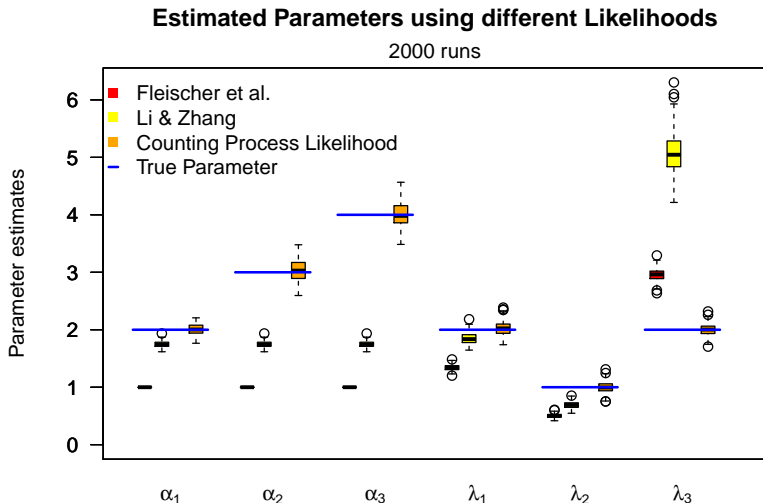
- Kendall's  $\tau$ , based on multistate, nonparametric, and copula models.
- Use again LFTM for estimation.

# Results

# Results: estimated parameters Exponential



# Results: estimated parameters time-inhomogeneous Weibull



# Conclusions

# Conclusions & outlook

Model PFS and OS within **illness-death without recovery multistate model**.

Advantages:

- Compared to LFTM avoids **questionable** and **uncheckable** assumptions,
- properties of PFS and OS induced through **transparent assumption on  $X(t)$** ,
- allows for **straightforward derivation** of survival functions and correlation for parametric models, e.g. no need to assume common Weibull shape parameter as in [Li and Zhang \(2015\)](#) to get tractable formulas,
- allows for **parametric** and **nonparametric** estimation and **inference** (at least) in Markov models using standard multistate modelling tools,
- engine to simulate PFS and OS times.

Outlook:

- How to best extrapolate tail of nonparametric survival function estimates?
- Shorten time for computation of bootstrap confidence intervals?
- R package?



# Multistate models for early decision-making

**How do we typically decide whether  
to move an oncology molecule  
into Phase 3?**

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But:

- **P(wrong decision)** may be high.
- Primary endpoint in Phase 3: **Overall survival**.

# Proposal:



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**Decrease  $P(\text{wrong decision})$ .**

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# Challenges and proposal

## Challenges:

- 1 Response-type endpoint?
- 2 Surrogacy? **Poor** in many indications.
- 3 Immunotherapy (CIT): no effect on response, relevant OS effect.
- 4 **Non-randomized** comparison  $\Rightarrow$  confounding.

Proposal: Base decision-making on **OS prediction from multistate model**.

- 1 **Predicted survival function for experimental arm**.
- 2 Combine  $S_{\text{exp}}$  with  $S_{\text{control}}$  to get **predicted OS HR**.
- 3 Experimental drug might act on certain transitions only  $\Rightarrow$  not captured through simple modelling of OS. Potential **efficiency gain!**
- 4 **Propensity scoring**.

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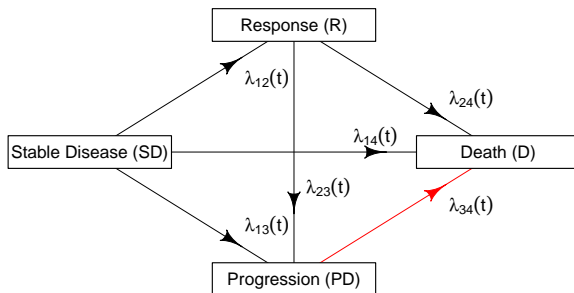
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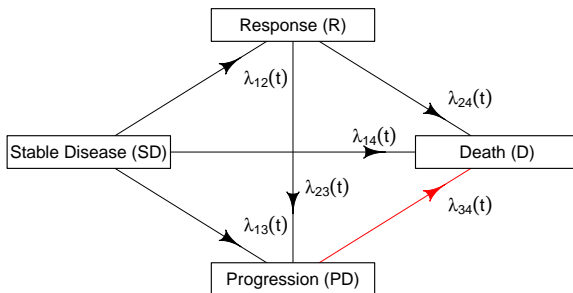
**Long-term follow-up in both arms.**

**Randomization  $\Rightarrow$  no confounding.**

# Multistate model for early decision-making

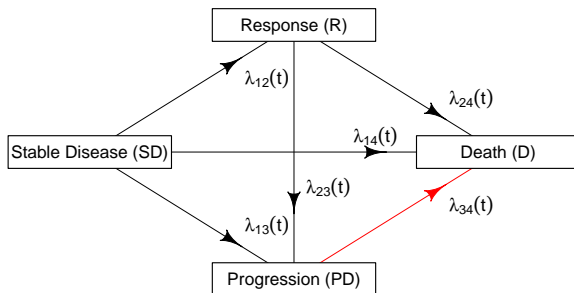


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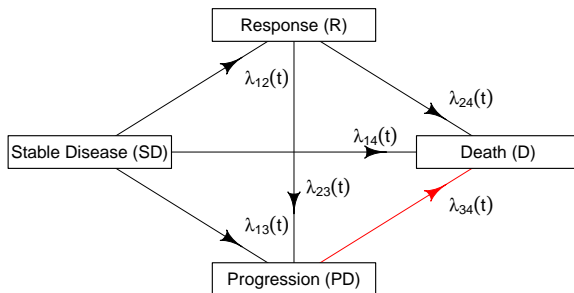
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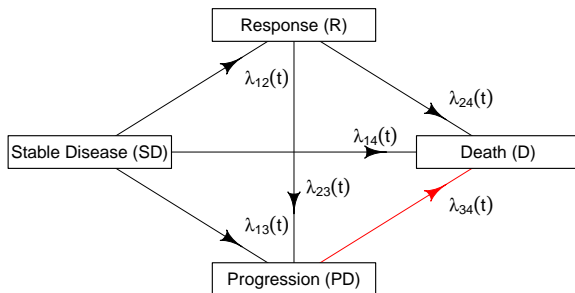
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- Transitions  $SD \rightarrow D$ ,  $R \rightarrow D$  rare, hazards  $\approx$  same in both arms.
- Markov assumption.

## Predicted survival function in experimental arm, $S_{\text{exp}}$

Compute transition probabilities for each transition.

$$S_{\text{exp}}(t) = 1 - \left( P_{SD \rightarrow D}(0, t) + P_{SD \rightarrow \text{PD} \rightarrow \text{D}}(0, t) + P_{SD \rightarrow R \rightarrow D}(0, t) + P_{SD \rightarrow R \rightarrow \text{PD} \rightarrow \text{D}}(0, t) \right).$$

$\lambda_{34}$  corresponding to  $\text{PD} \rightarrow \text{D}$  transition borrowed from historical data.



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Experimental treatment expected to provide benefit **beyond PD?**

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**No:**

- E.g. chemotherapy or antibody-dependent cellular cytotoxicity.
- **Plug-in** hazard function estimate from historical control.
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**Yes:**

- E.g. chemoimmunotherapy.
- Estimate post-PD hazard ratio assuming **proportionality**.
- How much post-PD deaths needed in experimental arm to reliably **estimate post-PD HR**?

## Benefit beyond PD: Oak

Previously treated non-small-cell lung cancer.

Rittmeyer *et al.* (2017).

	Atezolizumab	Chemotherapy	Hazard ratio
Effect post-PD	expected	not expected	
Objective Response	58 (13.6%)	57 (13.4%)	
Duration of Response	26.3 (10 - $\infty$ )	6.2 (4.9 - 7.6)	
Overall Survival			0.73 (0.62, 0.87)

**If this were early phase data -  
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**Competitors used this  
mechanism of action.**

# OS prediction when post-PD hazards assumed proportional

Random variable:

$$Z = \begin{cases} 0 & \text{if patient in control,} \\ 1 & \text{if in experimental group.} \end{cases}$$

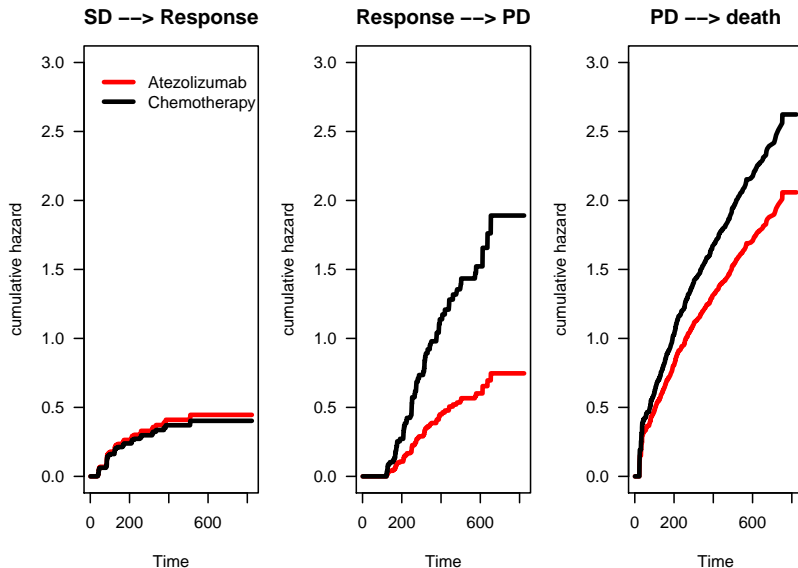
$$\lambda_{34}(t | Z) = \lambda_{34,0}(t) \exp(\beta_{34}Z)$$

Baseline hazard  $\lambda_{34,0}$  **estimated from both arms combined.**

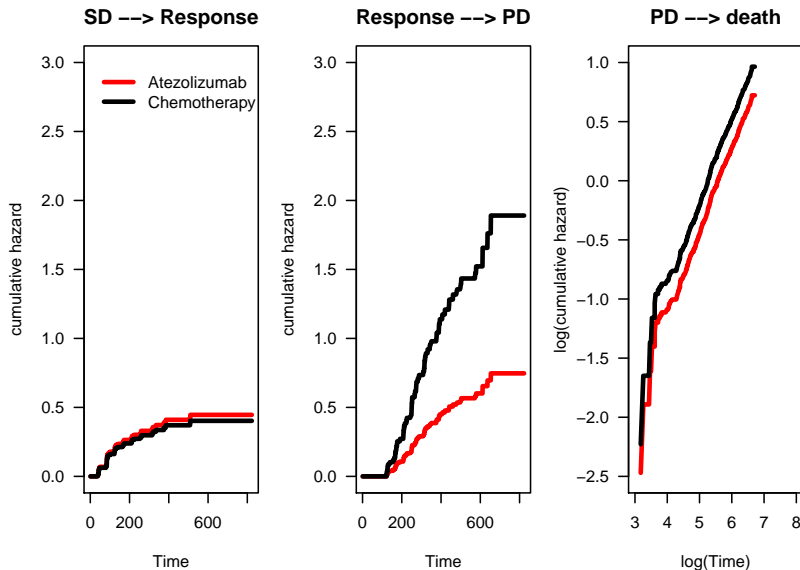
**Post-progression** hazard ratio  $\beta_{34}$ ?



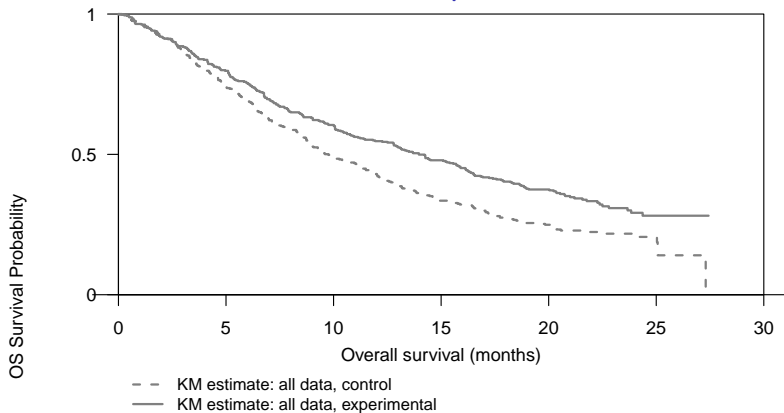
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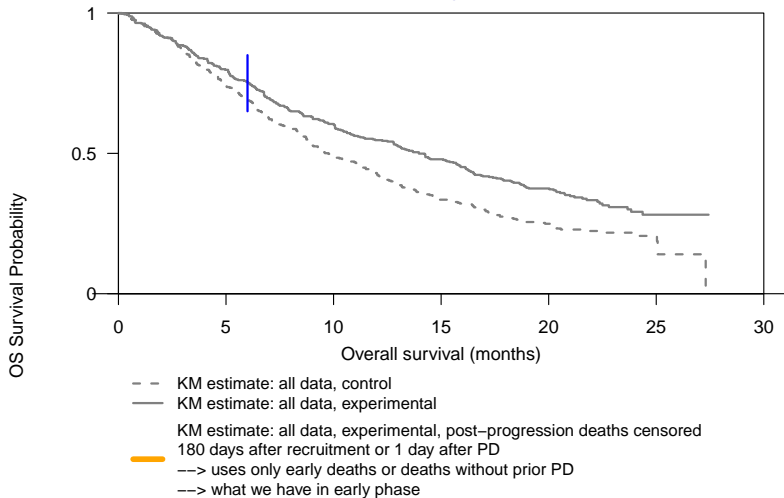
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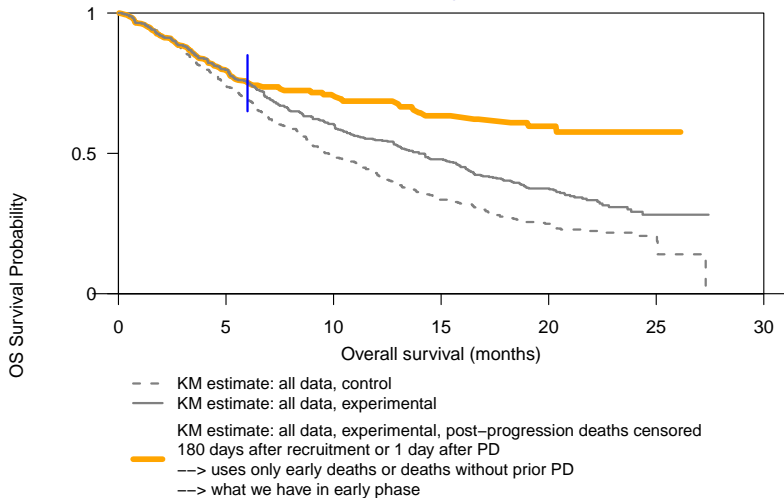
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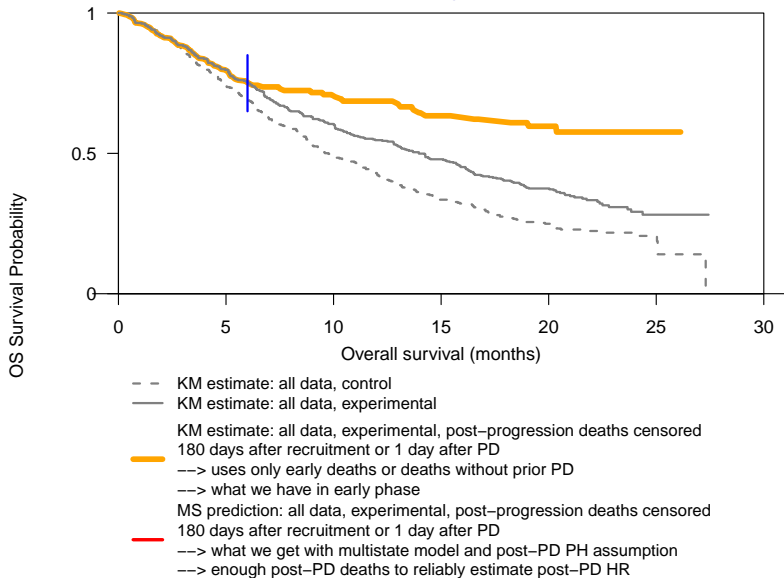
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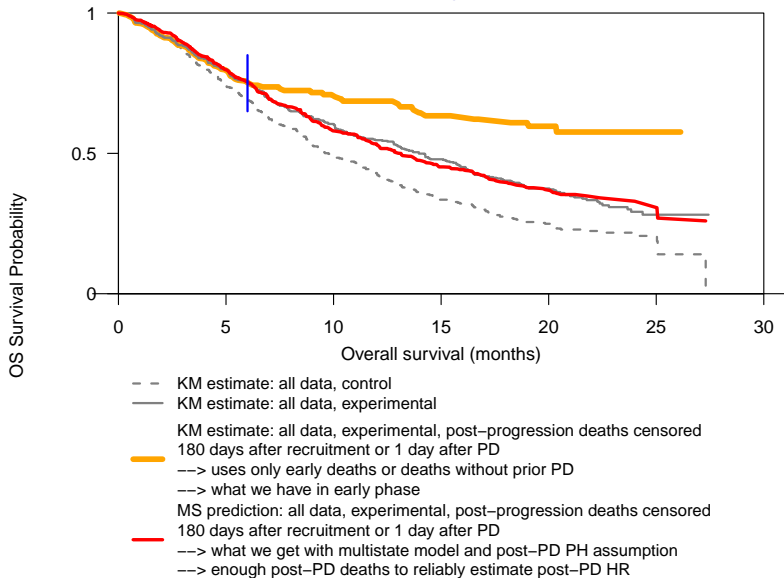
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## Early phase decision based on multistate prediction:



**Early phase decision based on  
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**$P(\text{wrong decision})?$**

# OS prediction from mimicked early phase data

Historical control: Oak control arm data.

**False-positive** decision: Sample early phase trial from Oak control arm.

**False-negative** decision: Sample early phase trial from Oak experimental arm.

Sample early phase trial:

- 40 patients,
- 6 months uniform recruitment,
- analysis 15 months after first patient entered,
- censor post-PD follow-up **one day after PD**,
- estimate  $\lambda_{12}, \lambda_{13}, \lambda_{14}, \lambda_{23}, \lambda_{24}$  from this data.

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Cox regression for post-PD transition  $\Rightarrow \hat{\lambda}_{34}(t|Z)$ .

Compute prediction of  $S_{\text{exp}}$ .

# OS HR prediction based on early phase trial

Approximate HR by fitting exponential distribution to both arms  $\Rightarrow \widehat{HR}$ .

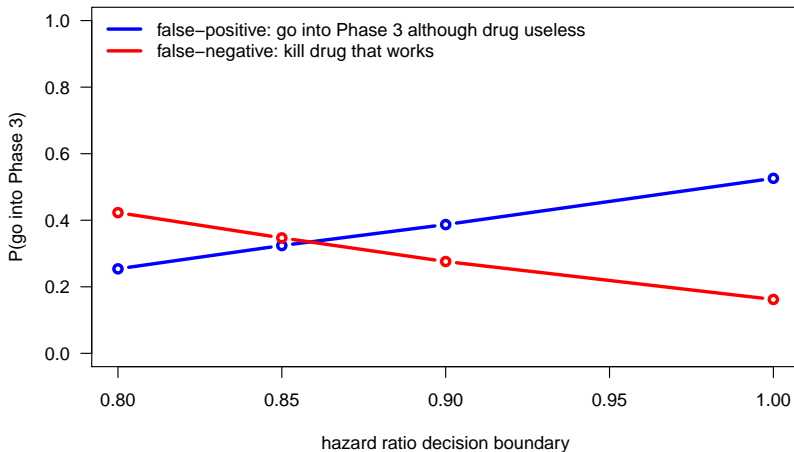
Decision to move to Phase 3:  $\widehat{HR} \leq \text{boundary} \in \{0.80, 0.85, 0.90, 1.00\}$ .

Repeat 1000 times.

Resampling  $\Rightarrow$  **quantification of uncertainty**.

## Oak: P(wrong decision)

**P(go into Phase 3) = P(approximated HR  $\leq$  boundary)**



**How many post-PD deaths to  
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**Ask during Q&A.**



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- **Avoids difficulty in interpretation of response-type endpoints**.
- Feasibility assessed in **idealized scenario**.
- Recommendation **how much post-PD follow-up** needed to estimate  $\beta_{34}$ .
- Needs **long-term individual-patient** data in control arm!

# What about confounding?



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**Combine proposal with propensity scoring.**

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- Improved **early stage decision-making**  $\Rightarrow$  [Beyer et al. \(2019\)](#).
- Improved **communication** of effect and optimized **sample size** computation.
- Bivariate modelling of PFS and OS to help inform **surrogacy** questions  $\Rightarrow$  [Meller et al. \(2019\)](#).

**Thank you for your attention.**

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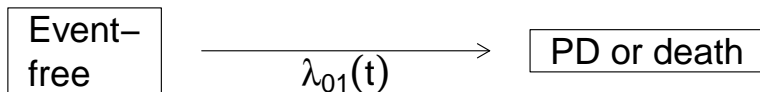
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# Backup

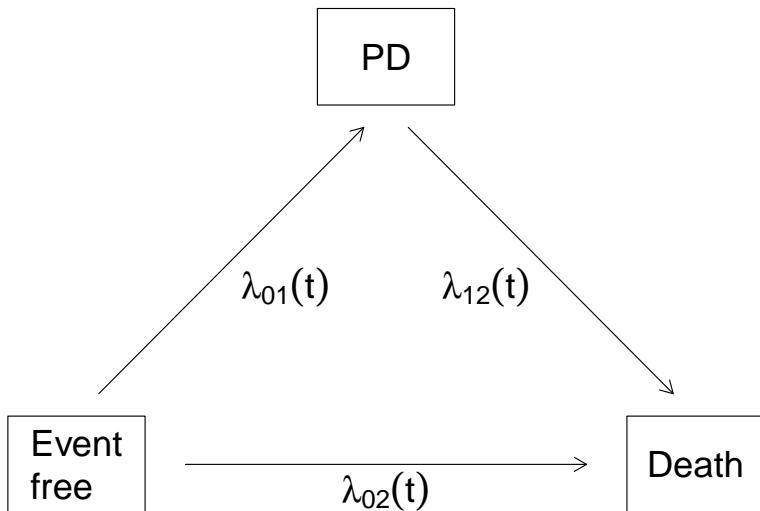
# Multistate models

# Canonical extension of survival analysis

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## Canonical extension of survival analysis



# Multistate models

Multistate model:

- 1-1 correspondence **hazard - probability** breaks down.
- Transition probabilities: (Markov) process  $X(t)_{t \geq 0}$  with state space  $\{0, 1, 2\} = \{\text{event-free, progression, death}\}$ . Then,

$$P_{lj}(s, t) := P(X_t = j | X_s = l, \text{Past}).$$

- Estimate  $P_{lj}$ 's **nonparametrically** by **Aalen-Johansen** estimator.
- PFS: Kaplan-Meier of time-to-progression simply censoring death is **biased**!
- OS: Aalen-Johansen offers **higher precision** compared to simple Kaplan-Meier estimate, [Andersen et al. \(1993\)](#) (p. 315 and Fig. IV.4.16).
- Markov assumption **stronger** than what is needed for Kaplan-Meier though.

# Prediction in multistate models

Rates (hazards, intensities):



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Derive formulas for these conditional probabilities, or simulate.

Final result: survival function for OS, as function of

- covariates and
- relevant **cumulative hazards**.

# PFS - OS

# Multistate vs. latent failure time model

Fleischer *et al.* (2009), Li and Zhang (2015): LFTM with **uncheckable** and **questionable** (**unrealistic?**) independence assumption.

Connections to multistate model? We are still figuring that out, work in progress.

Parametric models: formula for  $S_{PFS}$  identical for all three models below, and

- **Time-homogeneous Markov, Exponential:** model so simple that  $\nexists$  time-inhomogeneous Markov process.  $S_{OS}$  identical to Exponential LFTM.
- **Time-homogeneous Markov, Weibull:** formula for  $S_{OS}$  identical to Weibull LFTM  $\Rightarrow$  are model assumptions equivalent? **No!**
- **Time-inhomogeneous Markov, Weibull:** formulas for  $S_{OS}$  are **different**.

**BUT:** values of estimated parameters differ between LFTM and multistate for **all three parametric models**, as **likelihoods differ!**

Not clear (?) how to nonparametrically estimate LFTM  $\Rightarrow$  possible for (Markov) multistate.



# Assumptions for multistate model

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Multistate model **sufficiently smooth** so that following intensities exist:

$$\begin{aligned}\alpha_{0j}(t) &= \lim_{\Delta t \searrow 0} \frac{P(\text{PFS} \in [t, t + \Delta t), X(\text{PFS}) = j \mid \text{PFS} \geq t)}{\Delta t}, j = 1, 2, \\ \alpha_{12}(t; t_1) &= \lim_{\Delta t \searrow 0} \frac{P(X(t + \Delta t) = 2 \mid X(t-) = 1, \text{PFS} = t_1)}{\Delta t} \\ &= \lim_{\Delta t \searrow 0} \frac{P(\text{OS} - \text{PFS} \in [t - t_1, t - t_1 + \Delta t) \mid \text{OS} \geq t, \text{PFS} = t_1)}{\Delta t} \quad \text{for } t_1 < t.\end{aligned}$$

$t_1$ : observed PFS time, i.e. time when leaving state 0.

# Assumptions for multistate model

$X(t)$  **Markov**:

- **Time-inhomogeneous**: intensity of death after progression does not depend on time of progression,  $\alpha_{12}(t; t_1) = \alpha_{12}(t)$  for all  $t_1 < t$ .
- **Homogeneous**: intensities are time-constant, i.e. **Exponential**,  $\alpha_{ij}(t) = \alpha_{ij}$ ,  $i, j = 0, 1, 2$ .

$X(t)$  **non-Markov** (= semi-Markov for illness-death model without recovery):

- Intensities depend on state patient is in at  $s$  and entire history  $\leq s$ , i.e. all transitions.
- Relevant for  $1 \rightarrow 2$  transition only, as  $0 \rightarrow 1, 2$  are rooted in initial state 0.

As soon as a quantity depends on  **$1 \rightarrow 2$  transition** we need to be specific about assumption on  $X(t)$ .

# Illness-death multistate model for PFS and OS

Transition probabilities to move from state  $l$  at time  $s$  to state  $m$  at time  $t$ :

$$P_{lm}(s, t) := P(X(t) = m | X(s) = l, \text{history}).$$

Illness-death model w/o recovery,  $P_{lm}$  as functions of transition intensities, Aalen *et al.* (2008):

$$\begin{aligned}P_{00}(s, t) &= \exp\left(-\int_s^t \alpha_{01}(u) + \alpha_{02}(u) \, du\right), \\P_{11}(s, t; \mathbf{t}_1) &= \exp\left(-\int_s^t \alpha_{12}(u; \mathbf{t}_1) \, du\right), \\P_{22}(s, t) &= 1, \\P_{01}(s, t) &= \int_s^t P_{00}(s, u_-) \alpha_{01}(u) P_{11}(u, t; u) \, du, \\P_{12}(s, t; \mathbf{t}_1) &= 1 - P_{11}(s, t; \mathbf{t}_1), \\P_{02}(s, t) &= 1 - \left(P_{00}(s, t) + P_{01}(s, t)\right).\end{aligned}$$

If  $X(t)$  non-Markov:

- $P_{11}$  and  $P_{12}$  depend on **PFS time  $t_1$** .
- Although  $P_{01}, P_{02}$  depend on  $\alpha_{12}$  they **do not depend on  $t_1$** .

# Intuition behind transition probabilities

$P_{00}(s, t)$ ,  $P_{11}(s, t; t_1)$ : exp of cumulative hazards  $\Rightarrow$  standard survival functions.

$P_{01}(s, t) = \int_s^t P_{00}(s, u_-) \alpha_{01}(u) P_{11}(u, t; u) du$ : integral of

- $P_{00}(s, u_-) \alpha_{01}(u)$ : “infinitesimal probabilities” to move from 0 to 1 at time  $u$ ,  $u \in (s, t]$ ,
- $P_{11}(u, t; u)$ : subsequently stay in state 1 until at least time  $t$ , with progression happened in  $u$ .

# Illness-death multistate model for PFS and OS

**Marginal** distributions:

$$S_{PFS}(t) = P(\text{PFS} > t) = P_{00}(0, t),$$

$$S_{OS}(t) = P(\text{OS} > t) = P_{00}(0, t) + P_{01}(0, t),$$

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**Joint** distribution:

$$\begin{aligned} P(\text{PFS} \leq u, \text{OS} \leq v) &= P(X(u) \in \{1, 2\}, X(v) = 2) \\ &= P(X(u) = 1, X(v) = 2) + P(X(u) = 2) \\ &= P(X(v) = 2 | X(u) = 1) \cdot P(X(u) = 1 | X(0) = 0) \\ &\quad + P(X(u) = 2 | X(0) = 0) \\ &= P(X(v) = 2 | X(u) = 1) \cdot P_{01}(0, u; u) + P_{02}(0, u; u). \end{aligned}$$

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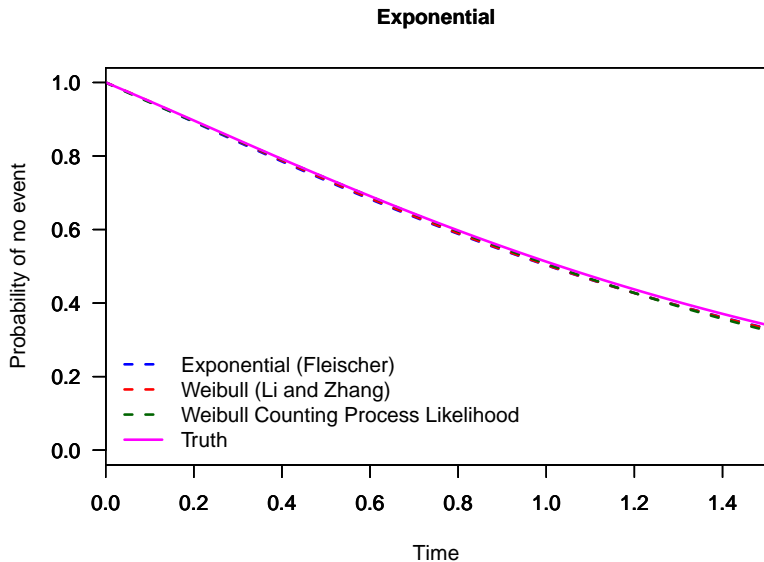
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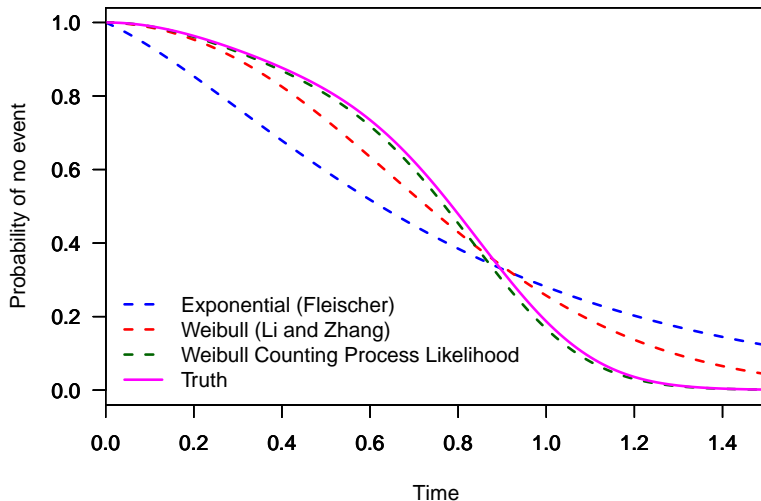
**X non-Markov:** integrate  $P_{12}(u, v; t_1)$  over conditional distribution of all possible progression times  $t_1 \leq u \Rightarrow$  final formula tedious.

## Results: $S_{OS}$ for Exponential

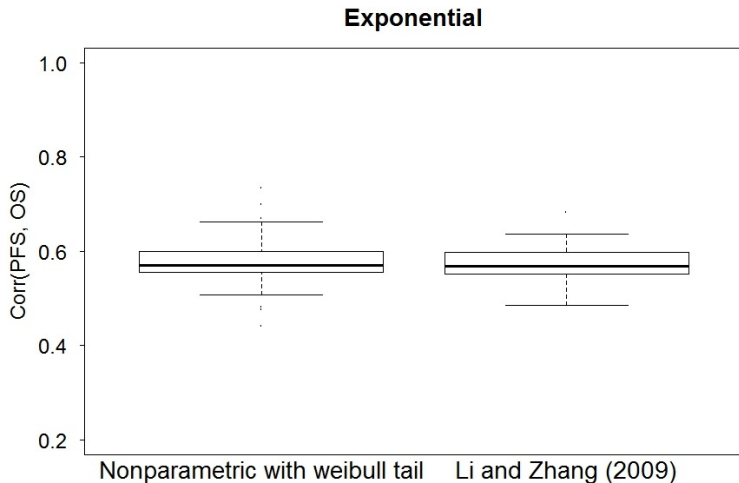


## Results: $S_{OS}$ for Weibull

Data from time-inhomogeneous Markov,  
Weibull with different shape,  $n = 500$

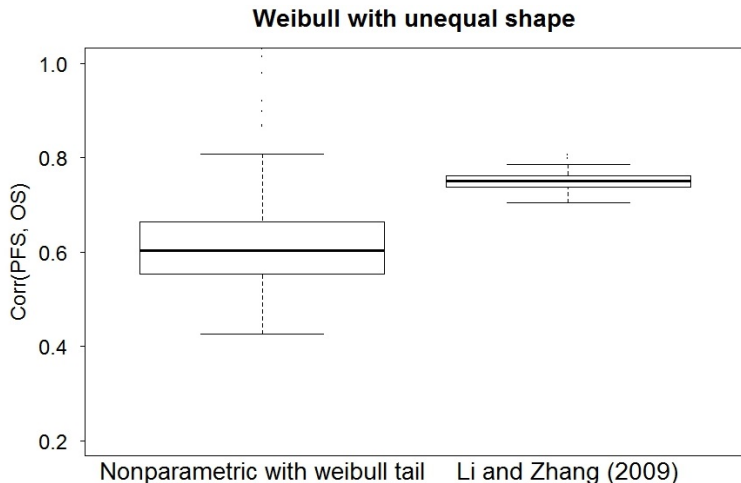


## Results: correlations Exponential



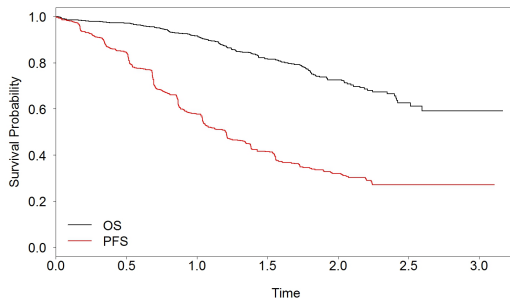
Corr(PFS, OS) for 200 simulated dataset from time-inhomogeneous Markov process.

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## Results: CLEOPATRA, Baselga and Cortes (2012).



	Exponential	Weibull	Weibull Markov	Nonparametric Markov
Corr(PFS, OS)	0.611	0.643	0.483	0.450
95% Bootstrap CI	[0.541; 0.673]	[0.584; 0.699]	[0.342; 0.643]	[0.297; 0.655]

**Table:** Correlation between PFS and OS in CLEOPATRA (1000 bootstrap samples).

# Early decision-making

# How many post-PD deaths needed?

Assumption:

$$\lambda_{34}(t | Z) = \lambda_{34,0}(t) \exp(\beta_{34}Z).$$

How many post-PD deaths needed in **experimental** arm to reliably estimate  $\lambda_{34}$ ?

Planning stage: only data for control arm are available.



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Various scenarios for post-PD follow-up time.

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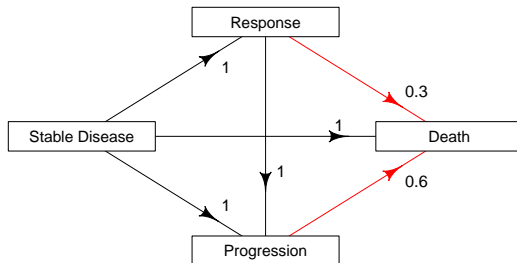
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- Simulate 40 patient from experimental arm as before.



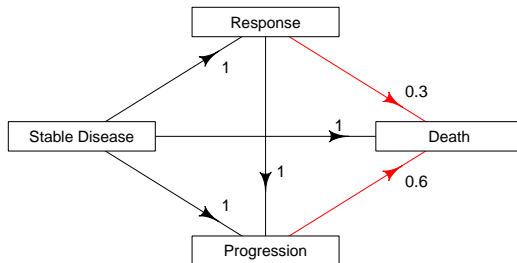


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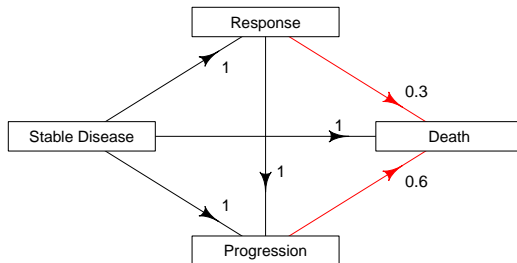
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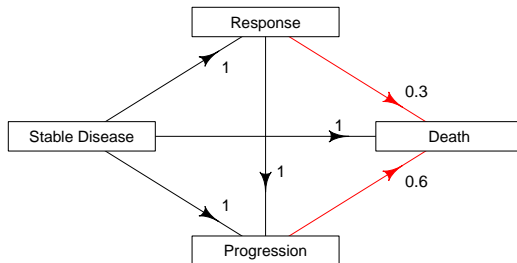
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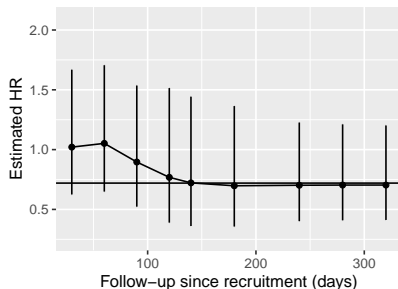
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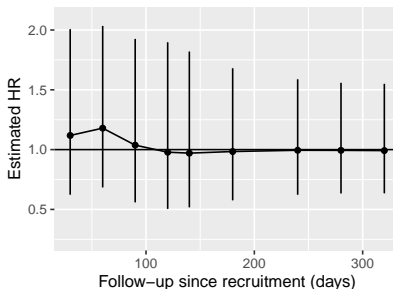
- Resulting **OS HR = 0.73**. Close to Oak OS HR.
- Follow-up post-PD for experimental arm truncated at 30, 60, 90, 120, 150, 180 and 240 days after recruitment.
- Repeat 1000 times.

# Stability of hazard ratio estimate

**A** Treatment Effect



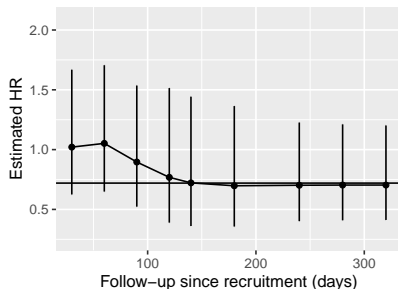
**B** No Treatment Effect



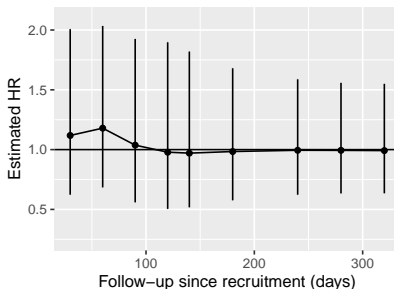
**180-240 days sufficient** to obtain stable point estimate over time.

# Stability of hazard ratio estimate

**A** Treatment Effect



**B** No Treatment Effect



**180-240 days sufficient** to obtain stable point estimate over time.

Typical early phase follow-up: Post-PD deaths censored **180 days after recruitment** in experimental arm.

## Example 1: Cleopatra

# Cleopatra

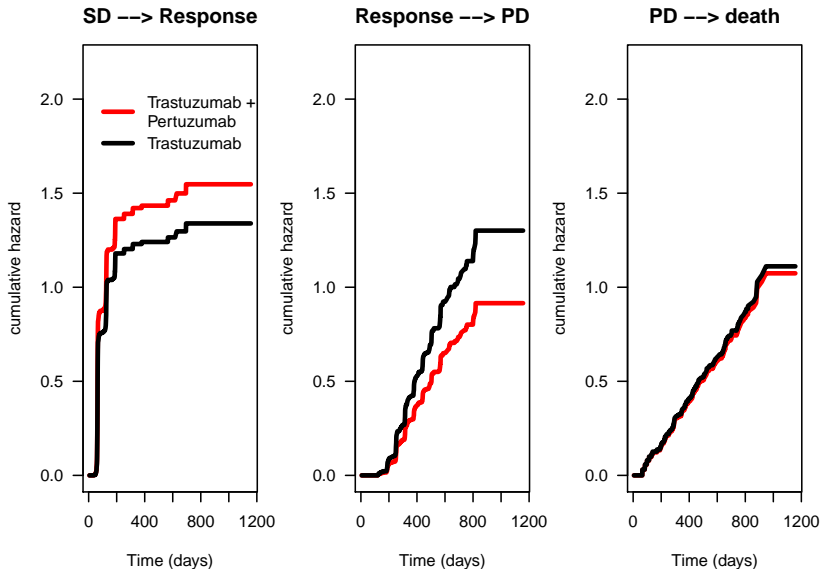
Baselga and Cortes (2012), Swain and Baselga (2015).

Previously untreated HER2-positive metastatic breast cancer patients.

	Pertuzumab+Trastuzumab	Trastuzumab	HR (95% CI)
<b>Survival</b>	N=402	N=406	
Overall Survival			<b>0.64</b> <b>(0.47,0.88)</b>
Progression-free Survival			0.62 (0.51,0.75)
<b>Response</b>	N=343	N=336	
Objective Response	275 (80.2%)	233 (69.3%)	
Stable Disease	50 (14.6%)	70 (20.8%)	
Progressive Disease	13 (3.8%)	28 (8.3%)	
<b>Duration of Response</b>	N=275	N=233	
Median (months, 95% CI)	<b>20.2 (16.0,24.0)</b>	<b>12.5 (10.0-15.0)</b>	

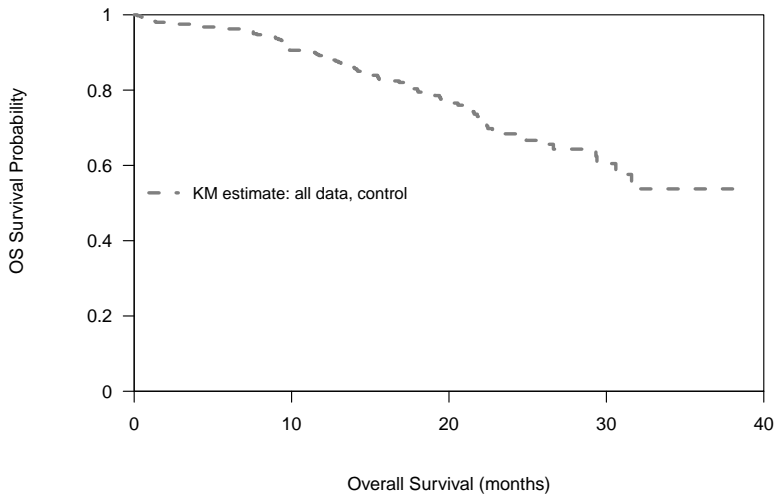
- Moderate difference in response.
- Prolonged **duration of response** in experimental arm.
- Clear OS benefit.
- Experimental treatment induces antibody-dependent cellular cytotoxicity  $\Rightarrow$  no benefit beyond PD expected  $\Rightarrow$   $\lambda_{34}$  **same in both arms.**

# Cleopatra: raw cumulative hazard estimates (of interest)





## Cleopatra: estimates / predictions of $S_{\text{exp}}$



406

347

150

28

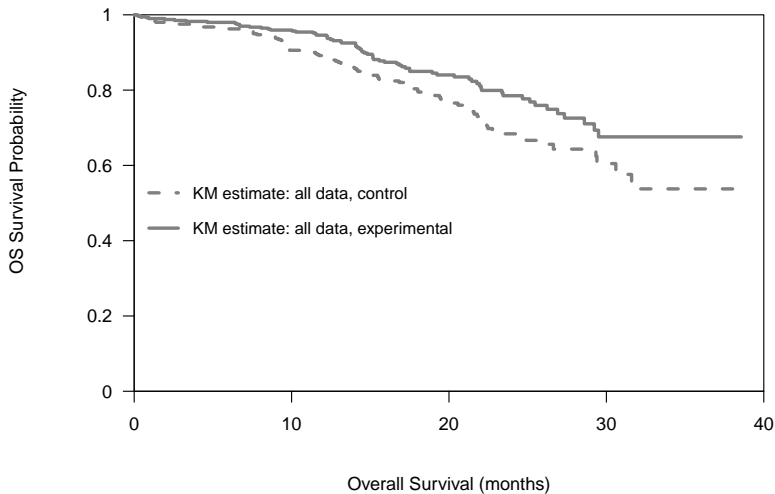
402

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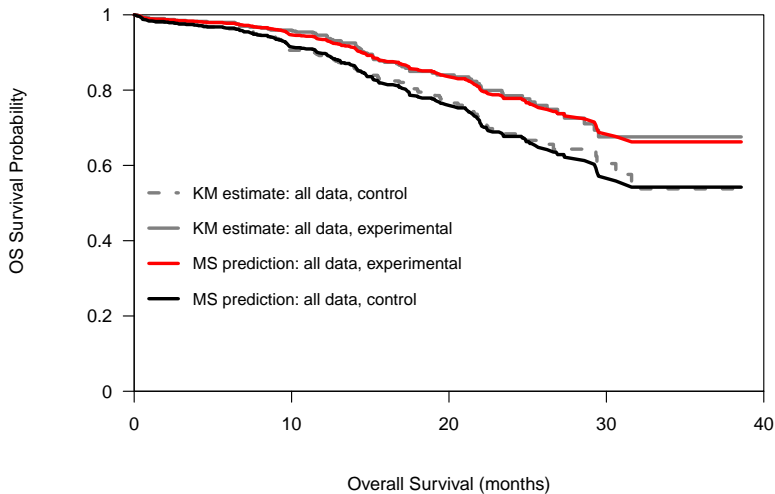
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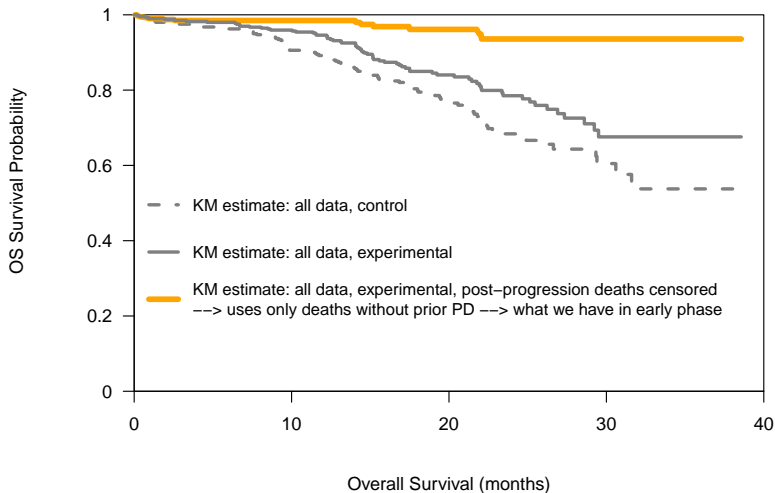
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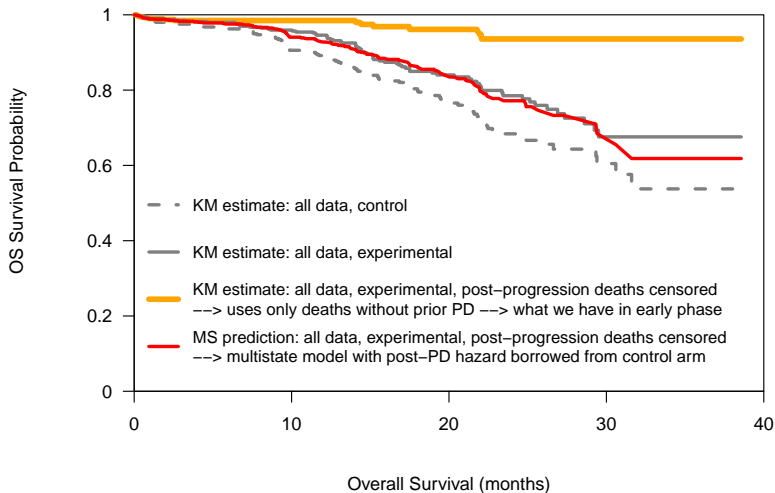
35

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406	347	150	28
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# Conclusions for Cleopatra

For estimated / predicted survival function in experimental arm, based on **all data**:

- Majority of patients dies after observed PD.
- KM estimate of simply censoring post-PD deaths does not work  $\Rightarrow$  very **few deaths observed**.
- Multistate model prediction assuming post-PD hazards as in control provides good prediction.

## **Early phase decision based on multistate prediction:**

**Early phase decision based on  
multistate prediction:**

**Operating characteristics?**



# OS prediction from mimicked early phase data

Sample early phase trial from **Cleopatra experimental arm**:

- 40 patients,
- 6 months uniform recruitment,
- analysis 15 months after first patient entered,
- censor post-PD follow-up **one day after PD**,
- estimate  $\lambda_{12}, \lambda_{13}, \lambda_{14}, \lambda_{23}, \lambda_{24}$  from this data,
- **borrow  $\hat{\lambda}_{34}$  from historical data** = Cleopatra control arm in idealized scenario,
- compute prediction of  $S_{\text{exp}}$  as described above.

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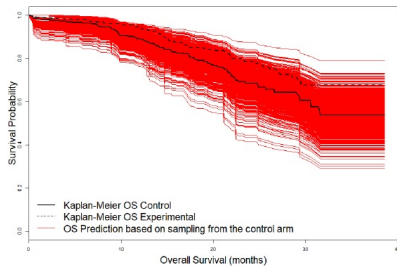
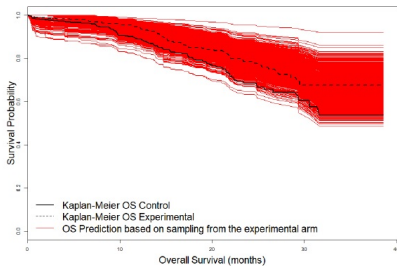
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Resampling easily allows for **quantification of uncertainty**.

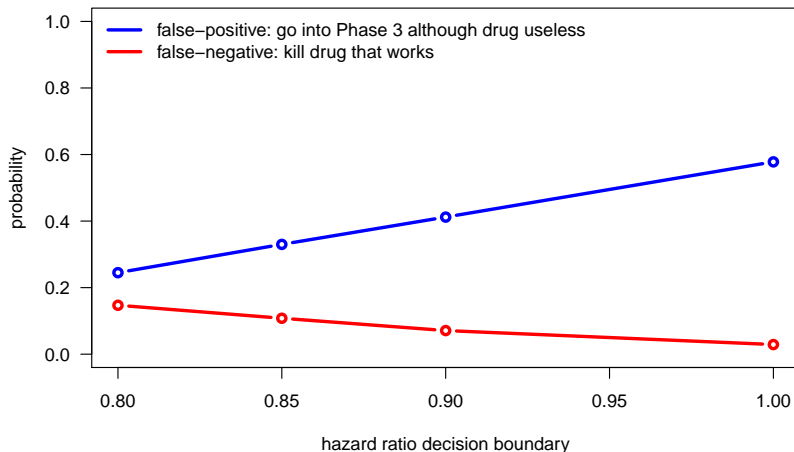
# Cleopatra: operating characteristics

Sampled from **experimental** and **control** arm.



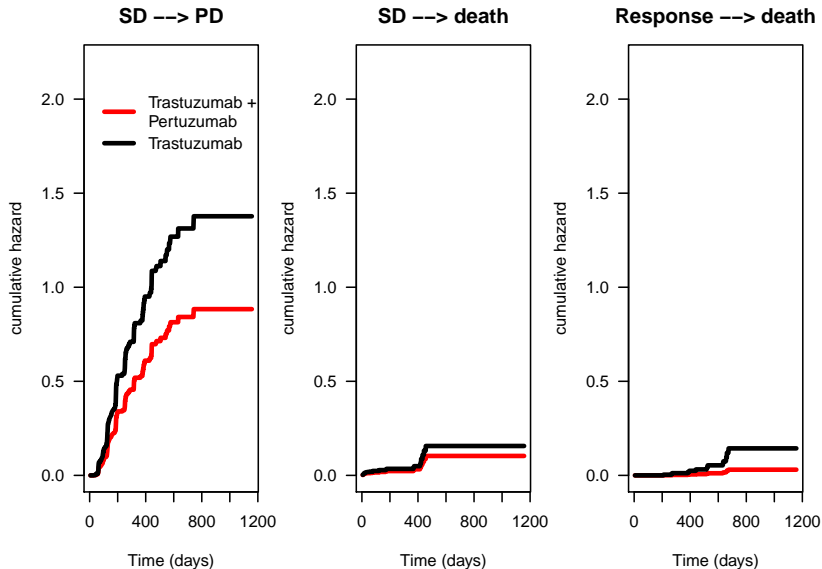
# Cleopatra: operating characteristics

probability to go into Phase 3:  $P(\text{approximated HR} \leq \text{boundary})$



Decision based on response:  $\approx 10\%$  difference, some prolongation of DOR  $\Rightarrow$  moved to Phase 3.

# Cleopatra: cumulative hazards of secondary interest



Previously treated non-small-cell lung cancer. [Rittmeyer et al. \(2017\)](#).

- Control: no benefit post-PD expected.
- Experimental: CIT  $\Rightarrow$  benefit post-PD expected.

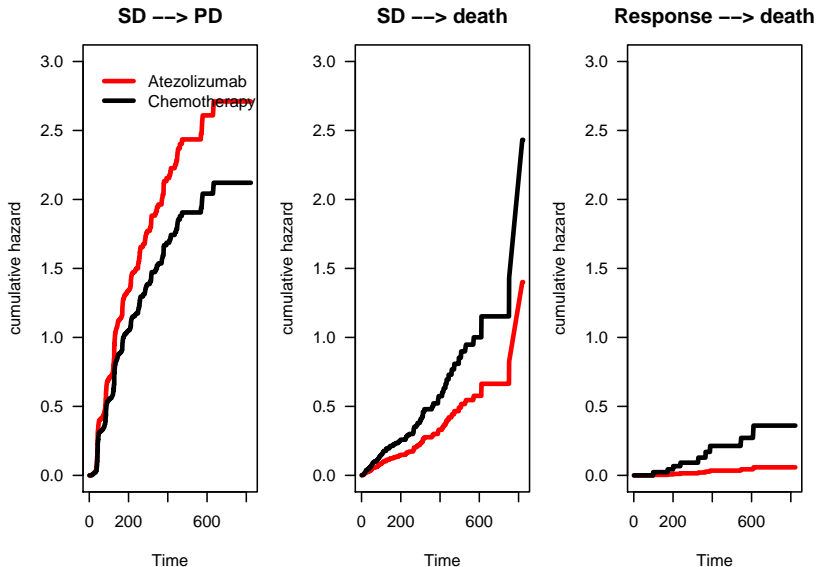
	Atezolizumab	Chemotherapy	HR (95% CI)
<b>Survival</b>	N=425	N=425	
Overall Survival			0.73 (0.62,0.87)
Progression-free Survival			0.95 (0.82,1.10)
<b>Response</b>	N=425	N=425	
<b>Objective Response</b>	<b>58 (13.6%)</b>	<b>57 (13.4%)</b>	
Stable Disease	150 (35%)	177 (42%)	
Progressive Disease	187 (44%)	117 (28%)	
Duration of Response	N=58	N=57	
Median (months, 95% CI)	26.3 (10,NE)	6.2 (4.9-7.6)	

No observed difference in response.

Prolonged duration of response in experimental arm.

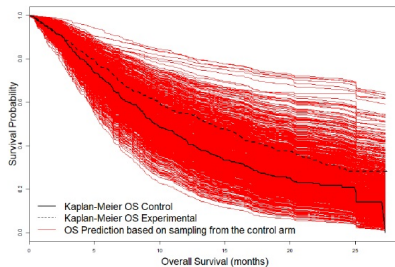
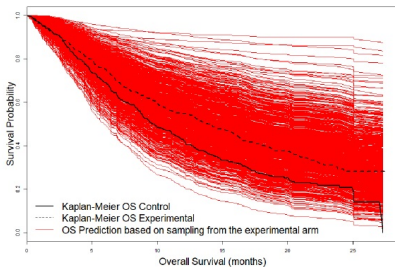
Clear survival benefit.

# Oak: cumulative hazards of secondary interest



# Oak: operating characteristics

Sampled from **experimental** and **control** arm.



# Non-proportional hazards via multistate model



**Immunotherapy:**  
**1) no difference in PFS,**  
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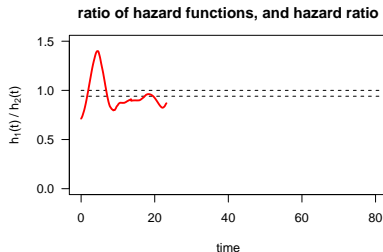
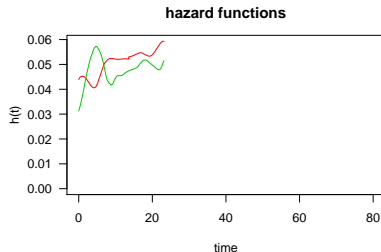
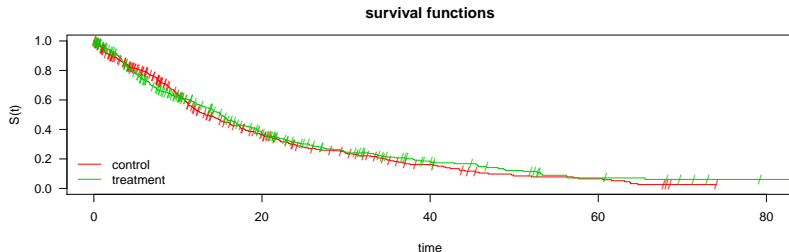
**How to quantify effect?**

# A fictional clinical trial

Simulated clinical trial:

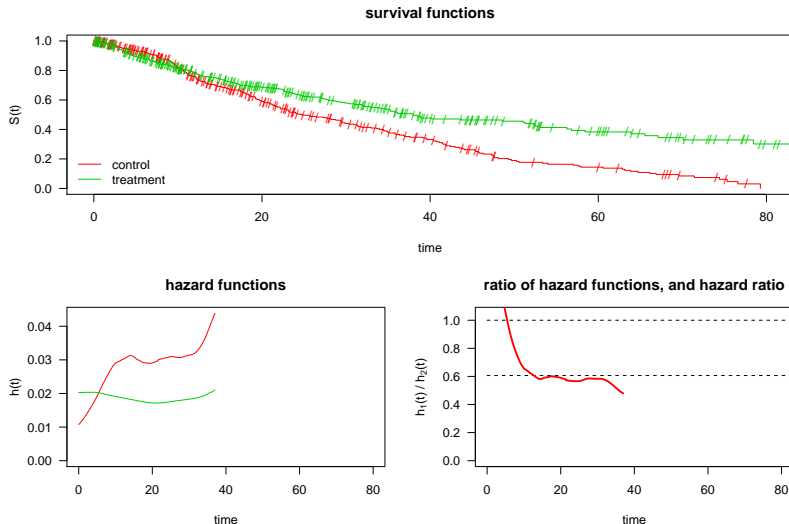
- 1:1 randomized, 400 and 400 patients per arm.
- No administrative censoring, but drop-out.

# PFS for simulated clinical trial



- Estimated hazard ratio: 0.94, 95% confidence interval [0.80, 1.11].
- Hypothesis test for PH:  $p = 0.24$ .

# OS for simulated clinical trial



- Estimated hazard ratio: 0.61, 95% confidence interval [0.50, 0.74].
- Hypothesis test for PH:  $p < 0.0001$ .

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**Non-proportional hazards** for OS. How to summarize effect of treatment?

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Transition	Control arm	Treatment arm
$0 \rightarrow 1$	$\lambda_{01}^c = \log(2)/25$	$\lambda_{01}^t = \lambda_{01}^c \cdot \mathbf{1}$
$0 \rightarrow 2$	$\lambda_{02}^c = \log(2)/30$	$\lambda_{02}^t = \lambda_{02}^c \cdot \mathbf{0.8}$
$1 \rightarrow 2$	$\lambda_{12}^c = \log(2)/15$	$\lambda_{12}^t = \lambda_{12}^c \cdot \mathbf{0.4}$

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	coef	HR = exp(coef)	95% CI	p-value
transition event-free $\rightarrow$ PD	-0.04	0.96	[0.77, 1.19]	0.72
transition event-free $\rightarrow$ death	-0.09	0.91	[0.70, 1.18]	0.49
transition PD $\rightarrow$ death	-1.09	0.34	[0.24, 0.46]	< 0.0001

Gaschler-Markefski *et al.* (2014).



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Biostatisticians ideally placed to contribute to this!

# *Doing now what patients need next*

## **R version and packages used to generate these slides:**

R version: R version 3.6.0 (2019-04-26)

Base packages: stats / graphics / grDevices / utils / datasets / methods / base

Other packages: nls2 / proto / diagram / shape / ggplot2 / rocheBCE / muhaz / flexsurv / reporttools / xtable / mstate / etm / dplyr / mvna / prodlm / biostatKR / survival

This document was generated on 2019-11-13 at 11:53:01.