

# Demand Estimation 2

## PhD Industrial Organization

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# Plan

1. BLP setup
2. Price elasticity/substitution patterns
3. Estimation: overview and typical data
4. Identification: what if we had micro-data?

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1. **BLP setup**
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## Discrete choice demand models: general setup of 'BLP'

Berry, Levinsohn, and Pakes (1995)

$$u_{ijt} = x_{jt}\beta_{it} + \alpha_{it}p_{jt} + \xi_{jt} + \epsilon_{ijt}$$

- What about income?:
- In the above equation,  $p_{jt}$  should really be  $y_i - p_{jt}$  where  $y_i$  is income.
- Leaving it out has no impact for choices however, and just simplifies exposition. Why?

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  - Leaving it out has no impact for choices however, and just simplifies exposition. Why?
    - Income enters linearly to all options, and only relative differences in utilities matter for choice probabilities
    - If  $y_i - p_{jt}$  entered non-linearly then it would affect the choice probabilities and should be included

# Discrete choice demand models: general setup of 'BLP'

Berry, Levinsohn, and Pakes (1995)

- Even more notation:
  - Recall  $L$  number of demographic vars,  $K$  number of product characteristics
- Define:
  - The **mean utility** of product  $j$  in market  $t$ :  $\delta_{jt} = x_{jt}\beta_0 + \alpha_0 p_{jt} + \xi_{jt}$
  - $\Gamma$ :  $(K + 1) \times L$  matrix with coefficients of demographic variables
  - $\Sigma$ :  $(K + 1) \times (K + 1)$  diagonal matrix with diagonal  $(\alpha_v, \beta_v^{(1)}, \dots, \beta_v^{(K)})$
  - $v_{it} = (v_{it}^{(0)}, \dots, v_{it}^{(K)})^T$
  - $\mu_{ijt} = (x_{jt}, p_{jt}) \cdot (\Gamma D_{it} + \Sigma v_{it})$
- Then we can rewrite our utility equation as:

$$u_{ijt} = \underbrace{\delta_{jt}}_{\text{mean utility}} + \underbrace{\mu_{ijt}}_{\text{interaction between consumer tastes + product characteristics}} + \underbrace{\epsilon_{ijt}}_{\text{idiosyncratic error}}$$

## Discrete choice demand models: general setup of 'BLP'

Berry, Levinsohn, and Pakes (1995)

- Review of where we are: we just characterized a very flexible model of consumer utility.
- Assuming i.i.d. extreme value errors  $\epsilon_{ijt}$  the probability consumer  $i$  chooses product  $j$  is:

$$\frac{\exp(\delta_{jt} + \mu_{ijt})}{1 + \sum_{k=1}^J \exp(\delta_{kt} + \mu_{ikt})}$$

- And **demand** (the share of consumers who purchase good  $j$  in market  $t$ ) is:

$$s_{jt} = \sigma_j(\delta_t, \mathbf{x}_t, \mathbf{p}_t; \Gamma, \Sigma) = \int \frac{\exp(\delta_{jt} + \mu_{ijt})}{1 + \sum_{k=1}^J \exp(\delta_{kt} + \mu_{ikt})} dF(D_{it}, v_{it})$$

- Here:

- $\delta_t, \mathbf{x}_t, \mathbf{p}_t$  are vectors of mean utilities, observed product characteristics, prices, in market  $t$
- $F$  is the joint distribution of observed demographics and unobserved tastes

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## Price elasticity/substitution patterns

- **Question:** is all the complexity in the previous section necessary (in terms of heterogeneous consumers etc)?
  - What would a simpler model (for example, with homogeneous consumers) fail to capture?
- **Answer:** (Typically) it is!
  - Key implication of a demand model: substitution patterns between goods/price elasticity
  - I will now argue that allowing for flexible consumer heterogeneity is **necessary to get the model to generate realistic substitution patterns.**

## Price elasticity/substitution patterns: implications of homogeneous consumer model

- **Thought experiment:** switch off consumer heterogeneity.
  - E.g. accomplish this by setting  $\Gamma = 0$  and  $\Sigma = 0$ . So,  $\mu_{ijt} = 0$ .
- Then, just a Logit model:

$$s_{jt} = \frac{\exp(\delta_{jt})}{1 + \sum_{k=1}^J \exp(\delta_{kt})}$$

- Price elasticities:

$$\eta_{jkt} = \frac{\partial s_{jt}}{\partial p_{kt}} \frac{p_{kt}}{s_{jt}} = \begin{cases} \alpha_0 p_{jt}(1 - s_{jt}) & \text{if } j=k \\ -\alpha_0 p_{kt} s_{kt} & \text{otherwise} \end{cases}$$

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- **Implication 1:**

- Typically,  $\alpha_0(1 - s_{jt}) \approx \alpha_0$  since there are many products and market share of each product is small.
- So, own price-elasticities ( $j=k$ ) are proportional to price.
- What does the model imply about prices vs demand elasticity?
  - This demand model implies that the **lower the price, the more inelastic is demand**
  - Further implication: under typical pricing models → higher markup for these lower priced goods
  - Question: do you think that these implications are reasonable predictions for the model to make?

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- **Implication 2:**

- Consider an increase in the price of product  $k$ . Concretely, think about the market for cars. The price of a BMW goes up. Do you think consumers will substitute towards a Mercedes or a Honda Civic?

## Price elasticity/substitution patterns: implications of homogeneous consumer model

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- **Implication 2:**

- Consider an increase in the price of product  $k$ . Concretely, think about the market for cars. The price of a BMW goes up. Do you think consumers will substitute towards a Mercedes or a Honda Civic?
  - Usually, we'd expect consumers to substitute towards similar products (i.e. the Mercedes)
- But, the homogeneous consumer model predicts the following **diversion ratio**:

$$\frac{\partial s_{jt}}{\partial p_{kt}} / \frac{\partial s_{kt}}{\partial p_{kt}} = s_{jt} / (1 - s_{kt})$$

- Here, substitution is proportional to market share, not how close the products are in terms of their characteristics.
  - Idea: as  $p_k$  increases, consumers who no longer choose  $k$  choose other options at the same frequency as the 'average' consumer (i.e. in proportion to their market share).

## Price elasticity/substitution patterns

- Price elasticities in the full BLP model (which heterogeneous consumers):

$$\eta_{jkt} = \frac{\partial s_{jt}}{\partial p_{kt}} \frac{p_{kt}}{s_{jt}} = \begin{cases} \frac{p_{jt}}{s_{jt}} \int \alpha_{it} s_{ijt} (1 - s_{ijt}) dF(D_{it}, \nu_{it}) & \text{if } j=k \\ -\frac{p_{kt}}{s_{jt}} \int \alpha_{it} s_{ijt} s_{ikt} dF(D_{it}, \nu_{it}) & \text{otherwise} \end{cases}$$

- Notation:  $s_{ijt}$ : probability that  $i$  purchases  $j$  in market  $t$
- **Observation 1:**
- Each consumer has a different price sensitivity, which is averaged to a product-specific mean price sensitivity using the individual probabilities of purchase as weights.
- This relaxes 'implication 1' from before. I.e. model could generate that low-price products have more elastic demand

## Price elasticity/substitution patterns

- Price elasticities in the full BLP model (i.e. including heterogeneous consumers):

$$\eta_{jkt} = \frac{\partial s_{jt}}{\partial p_{kt}} \frac{p_{kt}}{s_{jt}} = \begin{cases} \frac{p_{jt}}{s_{jt}} \int \alpha_{it} s_{ijt} (1 - s_{ijt}) dF(D_{it}, \nu_{it}) & \text{if } j=k \\ -\frac{p_{kt}}{s_{jt}} \int \alpha_{it} s_{ijt} s_{ikt} dF(D_{it}, \nu_{it}) & \text{otherwise} \end{cases}$$

- Notation:  $s_{ijt}$ : probability that  $i$  purchases  $j$  in market  $t$

- **Observation 2:**

- Model generates flexible cross-product substitution patterns.
  - How? Correlation in  $\mu_{ijt}$  and  $\mu_{ikt}$  induces correlation between  $s_{ijt}$  and  $s_{ikt}$ , which then determines substitution patterns.
- Note: alternatively, may be able to generate realistic substitution patterns with a nested logit (e.g. put the luxury cars in the same nest)
  - ...but this requires a-priori decisions about how to segment the market.

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## Estimation: setup of the problem

- What are the parameters we need to estimate?

## Estimation: setup of the problem

- What are the parameters we need to estimate?
- Linear parameters:
  - Parameters from the mean utility equation:  $(\alpha_0, \beta_0)$
- Nonlinear parameters
  - $\Gamma$ : coefficients on (observed) demographics
  - $\Sigma$ : idiosyncratic “taste for characteristics”
- So, full parameter vector to estimate:  $\theta = (\alpha_0, \beta_0, \Gamma, \Sigma)$ .

## Estimation: setup of the problem

- Common assumptions in empirical work:
  - We will also make these assumptions from now on.
  - They simplify the model, but are not necessary, see Chapter 1 of the Handbook for ways to relax these assumptions.
  - (Always good to know common assumptions people make in empirical work, especially when they simplify the model!)
- 1. Distribution of “taste for characteristics”  $\nu_{it} = (\nu_{it}^{(0)}, \dots, \nu_{it}^{(K)})$  is independent of the distribution of demographics  $D_{it}$ .
  - Then,  $F(D_{it}, \nu_{it}) = F_D(D_{it})F_\nu(\nu_{it})$ .
- 2. Each  $\nu_{it}^{(k)}$  is independent across  $k = 0, \dots, K$  and distributed standard normal.

## Data: typically, the data have three types of variables

- 1. Quantities of the  $J$  products purchased in market  $t$ .
  - These are aggregations of individual consumer choices.
  - As we will see, only aggregate data (ie. data on total quantities at the market level) is required for identification. However, information from micro-data (i.e. data on individual choices) can be incorporated.
  - Remember: we are implicitly assuming that definition of a market is narrow enough that consumers in the market face the same prices, characteristics, and demand shocks.
  - Can convert aggregate quantities to market shares if we know the total market size  $I_t$ :  
$$s_{jt} = q_{jt} / I_t.$$

## Data: typically, the data have three types of variables

- - 2. Prices  $p_{jt}$  and “observed” product characteristics  $x_{jt}$  of each of  $J$  products in market  $t$ .
- - 3. Information on consumer demographics.
    - In micro-data, will observe  $D_{ilt}$  i.e. demographics for each  $i$ 
      - e.g. survey data on car purchases
    - In other applications, more aggregated data
      - e.g. *distribution* of demographics  $F_t(D)$
      - could obtain such data from e.g. the Current Population Survey in different cities in the US)
    - In other applications, data at a granularity somewhere between the two above cases.
      - e.g. average age of consumers who purchase product  $j$

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# Identification

- What variation in the data can identify the parameters?
  - Precise econometric definition of identification: see haile.pdf on Canvas, or 'The Identification Zoo' (Lewbel)
- **Thought experiment:** what if we:
  - 1. have micro-data on individual consumers
  - 2. observe a single market
  - 3. switch off  $\Sigma = 0$  (i.e. ignore any idiosyncratic "taste for characteristics", implies heterogeneity is only driven by observed demographics)
- Later, we will build on this intuition to discuss what to do if we had more aggregated market-level data with random taste shocks etc...

## Identification using individual-level data

- Data:
  - $\{y_{ij}, D_i\}_{i=1,\dots,I}$  where  $y_{ij} = 1$  for  $j = 0, 1, \dots, J$  if consumer  $i$  chooses product  $j$  and  $\sum_j y_{ij} = 1$ .
  - All consumers are from the same market (same prices, same product characteristics both observed  $\mathbf{x}$  and unobserved  $\xi$ )
- Comment:
  - Estimating demand might seem hopeless here: we only see one market, so how are we supposed to get how quantities vary with prices if there is no price variation in the data?
  - But, we will now see that it is in fact possible.

## Identification using individual-level data

- (Conditional indirect) utility from product  $j$  (dropping  $t$  subscript and incorporating price  $p_j$  as a 'characteristic' in  $x_j$  to simplify exposition):

$$u_{ij} = \underbrace{x_j \beta_0 + \xi_j}_{\delta_j} + \sum_{k,l} \beta_d^{(l,k)} D_{il} x_{jk} + \epsilon_{ij}$$

- Comment: if we didn't have the (unobserved) demand shock  $\xi_j$  then we could estimate all the parameters of the model at once by maximum likelihood.

## Identification using individual-level data

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- Instead, use a **two-step procedure**:
  - 1. Include a product-specific intercept to capture  $\delta_j = x_j \beta_0 + \xi_j$  (i.e. estimate  $\tilde{\theta} = (\delta_1, \dots, \delta_J, \Gamma)$  using maximum likelihood )
  - 2. Estimate  $\beta_0$  by 'projecting' estimated  $\delta$ 's on the  $x$ 's.
    - If assume  $E(\xi_j | x_j) = 0$  then can use (weighted) least squares
    - If concerned  $x$ 's are correlated with  $\xi$  can use  $E(\xi_j | Z_j) = 0$  where  $Z$  are a vector of exogenous variables (discuss more in a few slides...)

## Identification using individual-level data: step 1

- Estimate the  $\delta$  and  $\Gamma$  parameters by maximum likelihood with utility:

$$u_{ij} = \delta_j + \sum_{k,l} \beta_d^{(l,k)} D_{il} x_{jk} + \epsilon_{ij}$$

- Identifying  $\delta_j$ :
- Take FOC of likelihood  $\rightarrow$  can show that intercepts  $\delta_j$  are found by setting observed market shares equal to the ones predicted by model. That is, if for a fixed  $\Gamma$ , set:

$$\hat{s}_j = \hat{\sigma}(\hat{\delta}_1, \dots, \hat{\delta}_J)$$

- Under some general technical conditions can invert this relationship (the 'Berry inversion'):

$$\hat{\delta}_j = \hat{\sigma}_j^{-1}(\hat{s}_1, \dots, \hat{s}_J)$$

- Asymptotically as  $I \rightarrow \infty$ :

$$\delta_j = \sigma_j^{-1}(s_1, \dots, s_J)$$

## Identification using individual-level data: step 1

- To be continued...