

# Lesson 3: Measurement of Association for Contingency Tables

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# Poll Everywhere Question 1

Make sure to remember your answer!! We'll use this on Wednesday!

# Learning Objectives

1. Understand the difference between testing for association and measuring association
2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.
3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.
4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

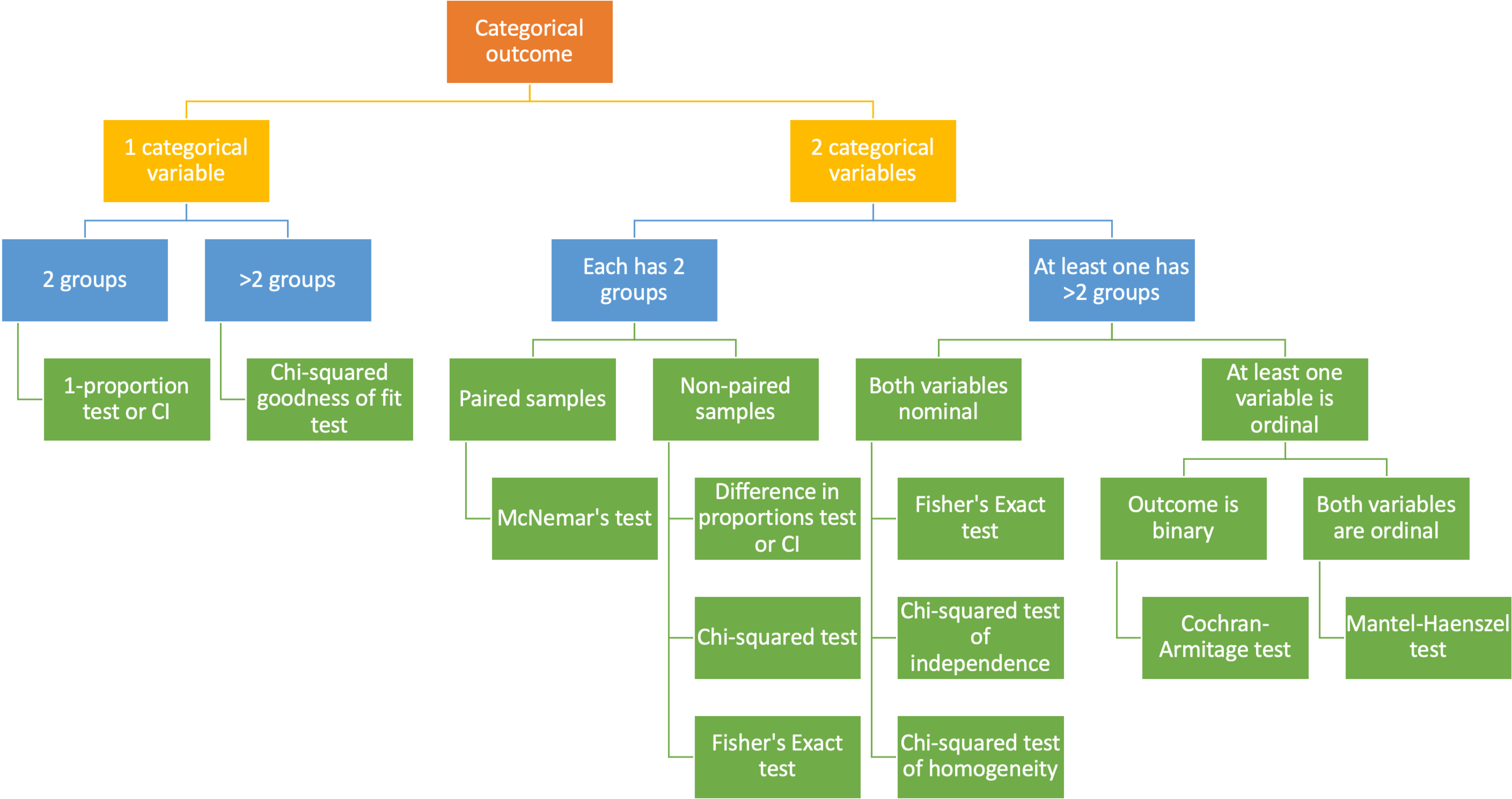
# Learning Objectives

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# Review of Test of Association (1/2)

- Last class: learned some tests of association for contingency tables
- For studies with two independent samples
  - General association
    - Chi-squared test
    - Fisher's Exact test
  - Test of trends
    - Cochran-Armitage test
    - Mantel-Haenszel test

# Review of Test of Association (2/2)



# Test of association *does not* measure association

- Test of association **does not** provide an effective measure of association.
- The p-value alone is not enough
  - $p\text{-value} < 0.05$  suggests there is a statistically significant association between the group and outcome
  - $p\text{-value} = 0.00001$  vs.  $p\text{-value} = 0.01$  does not mean the **magnitude** of association is different
- But it does not tell **how different** the risks are between the two groups
- We want a measurement to quantify different risks across the groups

# Measures of Association

- When we have a **2x2 contingency table** (binary outcome and explanatory variable) and **independent samples**, we have three main options to measure of association:

1. Risk difference (RD)

2. Relative risk (RR)

3. Odds ratio (OR)

Each measures association by comparing the proportion of successes/failures from each categorical group of our explanatory variable.



# Before we discuss each further...

Let's define the cells within a 2x2 contingency table:

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	$n_{11}$	$n_{12}$	$n_1$
2	$n_{21}$	$n_{22}$	$n_2$
Total	$n_+$ (or $n_S$ )	$n_-$ (or $n_F$ )	$n$

- Then we can define risk: the proportion of “successes”
  - Risk of successful response for explanatory group 1:  $\text{Risk}_1 = \frac{n_{11}}{n_1}$

# Learning Objectives

1. Understand the difference between testing for association and measuring association

2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.

3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.

4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

# Risk Difference (RD)

- Risk difference computes the absolute difference in risk for the two groups (from the explanatory variable)

- Point estimate:

$$\widehat{RD} = \hat{p}_1 - \hat{p}_2 = \frac{n_{11}}{n_1} - \frac{n_{21}}{n_2}$$

- With range of point estimate from  $[-1, 1]$

- Approximate standard error:

$$SE_{\widehat{RD}} = \sqrt{\frac{\hat{p}_1 \cdot (1 - \hat{p}_1)}{n_1} + \frac{\hat{p}_2 \cdot (1 - \hat{p}_2)}{n_2}}$$

- 95% Wald confidence interval for  $\widehat{RD}$ :

$$\widehat{RD} \pm 1.96 \cdot SE_{\widehat{RD}}$$

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	$n_{11}$	$n_{12}$	$n_1$
2	$n_{21}$	$n_{22}$	$n_2$
Total	$n_+$ (or $n_S$ )	$n_-$ (or $n_F$ )	$n$

# Recall the Strong Heart Study

The [Strong Heart Study](#) is an ongoing study of American Indians residing in 13 tribal communities in three geographic areas (AZ, OK, and SD/ND). We will look at data from this study examining the **incidence of diabetes** at a follow-up visit and **impaired glucose tolerance (ITG)** at baseline (4 years apart).

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664



# SHS Example: Risk Difference

## Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## Needed steps:

- 1. Compute the risk difference
- 2. Compute 95% confidence interval
- 3. Interpret the estimate

# SHS Example: Risk Difference (1/4)

## Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1. Compute the risk difference

$$\widehat{RD} = \hat{p}_1 - \hat{p}_2 = \frac{n_{11}}{n_1} - \frac{n_{21}}{n_2} = \frac{198}{532} - \frac{128}{1132} = 0.3722 - 0.1131 = 0.2591$$

# SHS Example: Risk Difference (2/4)

## Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

2. Compute 95% confidence interval

$$\begin{aligned} & \widehat{RD} \pm z_{(1-\frac{\alpha}{2})}^* \times SE_{\widehat{RD}} \\ &= \widehat{RD} \pm z_{(1-\frac{\alpha}{2})}^* \times \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}} \\ &= 0.2591 \pm 1.96 \times \sqrt{\frac{0.3722(1-0.3722)}{532} + \frac{0.1131(1-0.1131)}{1132}} \\ &= (0.2141, 0.3041) \end{aligned}$$

# SHS Example: Risk Difference (3/4)

Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1/2. Compute risk difference and 95% confidence interval

```
1 fmsb::riskdifference(198, 128, 532, 1132)
```

	Cases	People at risk	Risk
Exposed	198.0000000	532.0000000	0.3721805
Unexposed	128.0000000	1132.0000000	0.1130742
Total	326.0000000	1664.0000000	0.1959135

Risk difference and its significance probability (H0: The difference equals to zero)

```
data: 198 128 532 1132
p-value < 2.2e-16
95 percent confidence interval:
 0.2140779 0.3041346
sample estimates:
[1] 0.2591062
```



# SHS Example: Risk Difference (4/4)

## Risk difference

Compute the point estimate and 95% confidence interval for the diabetes risk difference between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## 3. Interpret the estimate

The diabetes diagnosis risk difference between impaired and normal glucose tolerance is 0.2591 (95% CI: 0.2141, 0.3041). Since the 95% confidence interval does not contain 0, we have sufficient evidence that the risk of diabetes diagnosis within 4 year follow-up for people with impaired versus normal glucose tolerance is different.

# When is the risk difference misleading?

- The same risk differences can have very different clinical meanings depending on the risk for each group
- **Example:** for two treatments A and B, we know the risk difference (RD) is 0.009. Is it a meaningful difference?
  - If the risk is 0.01 for Trt A and 0.001 for Trt B?
  - If the risk is 0.41 for Trt A and 0.401 for Trt B?
- Using the RD alone to summarize the difference in risks for comparing the two groups can be **misleading**
  - The **ratio of risk** can provide an informative descriptive measure of the “relative risk”

# Learning Objectives

1. Understand the difference between testing for association and measuring association
2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.
3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.
4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

# Relative Risk (RR)

- Relative risk computes the **ratio** of each group's proportions of “success”
  - Also called **risk ratio**
- Point estimate:

$$\widehat{RR} = \frac{\hat{p}_1}{\hat{p}_2} = \frac{n_{11}/n_1}{n_{21}/n_2}$$

- Range:  $[0, \infty]$

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	$n_{11}$	$n_{12}$	$n_1$
2	$n_{21}$	$n_{22}$	$n_2$
Total	$n_+$ (or $n_S$ )	$n_-$ (or $n_F$ )	$n$

# Poll Everywhere Question 2

# Log-transformation of RR

- Sampling distribution of the relative risk is **highly skewed** unless sample sizes are quite large
  - **Log transformation** results in *approximately normal distribution*
  - Thus, **compute confidence interval using normally distributed, log-transformed RR**
  - Then we convert back to the RR
- We take the log (natural log) of RR:  $\ln(\widehat{RR})$  or  $\log(\widehat{RR})$ 
  - Whenever I say “log” I mean natural log (base  $e$ , very common in statistics)
- Then we need to find approximate standard error for  $\ln(\widehat{RR})$

$$SE_{\ln(\widehat{RR})} = \sqrt{\frac{1}{n_{11}} - \frac{1}{n_1} + \frac{1}{n_{21}} - \frac{1}{n_2}}$$

- 95% confidence interval for  $\ln(\widehat{RR})$ :

$$\ln(\widehat{RR}) \pm 1.96 \times SE_{\ln(\widehat{RR})}$$

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	$n_{11}$	$n_{12}$	$n_1$
2	$n_{21}$	$n_{22}$	$n_2$
Total	$n_+$ (or $n_S$ )	$n_-$ (or $n_F$ )	$n$

# How do we get back to the RR scale?

- We computed confidence interval using normally distributed, log-transformed RR ( $\ln(\widehat{RR})$ ):

$$\left( \ln(\widehat{RR}) - 1.96 \times SE_{\ln(\widehat{RR})}, \ln(\widehat{RR}) + 1.96 \times SE_{\ln(\widehat{RR})} \right)$$

- Now we need to exponentiate the CI to get back to interpretable values
  - Take exponential of lower and upper bounds
- 95% confidence interval for RR: two ways to display equation

$$\left( e^{\ln(\widehat{RR}) - 1.96 \times SE_{\ln(\widehat{RR})}}, e^{\ln(\widehat{RR}) + 1.96 \times SE_{\ln(\widehat{RR})}} \right)$$

$$\left( \exp \left[ \ln(\widehat{RR}) - 1.96 \times SE_{\ln(\widehat{RR})} \right], \exp \left[ \ln(\widehat{RR}) + 1.96 \times SE_{\ln(\widehat{RR})} \right] \right)$$

# Relative Risk (RR)

- Can you compute the estimated RRs for the previous example?
  - If the risk for Trt A is 0.01 and Trt B is 0.001?  $\widehat{RR} = 10$
  - If the risk for Trt A is 0.41 and Trt B is 0.401?  $\widehat{RR} = 1.02$
- When  $\widehat{RR} = 1$  ...
  - Risk is the same for the two groups
  - In other words, the **group and the outcome are independent**
- When computing  $\widehat{RR}$  it is important to **identify which variable is the response variable and which is explanatory variable**
  - We may say “risk for Trt A” but this translates to the risk (or probability) of outcome success for those receiving Trt A



# SHS Example: Relative Risk (1/6)

## Relative risk

Compute the point estimate and 95% confidence interval for the diabetes  
Relative risk between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## Needed steps:

1. Compute the relative risk
2. Find confidence interval of log RR
3. Convert back to RR
4. Interpret the estimate

# SHS Example: Relative Risk (2/6)

## Relative risk

Compute the point estimate and 95% confidence interval for the diabetes  
Relative risk between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1. Compute the relative risk

$$\widehat{RR} = \frac{\hat{p}_1}{\hat{p}_2} = \frac{n_{11}/n_1}{n_{21}/n_2} = \frac{198/532}{128/1132} = \frac{0.3722}{0.1131} = 3.2915$$

# SHS Example: Relative Risk (3/6)

## Relative risk

Compute the point estimate and 95% confidence interval for the diabetes Relative risk between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

2. Find confidence interval of log RR

$$\begin{aligned} & \ln(\widehat{RR}) \pm 1.96 \times SE_{\ln(\widehat{RR})} \\ & = \ln(\widehat{RR}) \pm z_{(1-\frac{\alpha}{2})}^* \times \sqrt{\frac{1}{n_{11}} - \frac{1}{n_1} + \frac{1}{n_{21}} - \frac{1}{n_2}} \\ & = 1.1913 \pm 1.96 \times \sqrt{\frac{1}{198} - \frac{1}{532} + \frac{1}{128} - \frac{1}{1132}} \\ & = (0.9944, 1.3883) \end{aligned}$$

# SHS Example: Relative Risk (4/6)

## Relative risk

Compute the point estimate and 95% confidence interval for the diabetes  
Relative risk between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

### 3. Convert back to RR

$$\begin{aligned} &(\exp(0.9944), \exp(1.3883)) \\ &=(2.703, 4.0081) \end{aligned}$$

# SHS Example: Relative Risk (5/6)

## Relative risk

Compute the point estimate and 95% confidence interval for the diabetes  
Relative risk between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1/2/3. Compute risk ratio and 95% confidence interval

```
1 library(epitools)
2 SHS_ct = table(SHS$glucimp, SHS$case)
3 riskratio(x = SHS_ct, rev = "rows")$measure
```

	risk ratio with 95% C.I.		
	estimate	lower	upper
Normal	1.000000	NA	NA
Impaired	3.291471	2.702998	4.008061

# Pause: other option in pubh package

```
1 SHS = SHS %>% mutate(glucimp = as.factor(glucimp) %>% relevel(ref = "Normal"))
2 contingency(case ~ glucimp, data = SHS)
```

		Outcome			
Predictor		1	0		
Impaired		198	334		
Normal		128	1004		

	Outcome +	Outcome -	Total	Inc risk *
Exposed +	198	334	532	37.22 (33.10 to 41.48)
Exposed -	128	1004	1132	11.31 (9.52 to 13.30)
Total	326	1338	1664	19.59 (17.71 to 21.58)

Point estimates and 95% CIs:

Inc risk ratio	3.29 (2.70, 4.01)
Inc odds ratio	4.65 (3.61, 6.00)
Attrib risk in the exposed *	25.91 (21.41, 30.41)
Attrib fraction in the exposed (%)	69.62 (63.00, 75.05)
Attrib risk in the population *	8.28 (5.63, 10.94)
Attrib fraction in the population (%)	42.28 (34.71, 48.98)

Uncorrected chi2 test that OR = 1:  $\chi^2(1) = 154.239$   $\text{Pr}>\chi^2 = <0.001$   
Fisher exact test that OR = 1:  $\text{Pr}>\chi^2 = <0.001$   
Wald confidence limits  
CI: confidence interval  
\* Outcomes per 100 population units

Pearson's Chi-squared test with Yates' continuity correction

data: dat  
X-squared = 152.6, df = 1, p-value < 2.2e-16

# SHS Example: Relative Risk (6/6)

Relative risk

Compute the point estimate and 95% confidence interval for the diabetes Relative risk between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

### 3. Interpret the estimate

The estimated risk of diabetes is 3.29 times greater for American Indians who had impaired glucose tolerance at baseline compared to those who had normal glucose tolerance (95% CI: 2.70, 4.01).

Additional interpretation of 95% CI (not needed): We are 95% confident that the (population) relative risk is between 2.70 and 4.01.

Since the 95% confidence interval does not include 1, there is sufficient evidence that the risk of diabetes differs significantly between impaired and normal glucose tolerance at baseline.

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# Odds (building up to Odds Ratio)

- For a probability of success  $p$  (or sometimes referred to as  $\pi$ ), the odds of success is:

$$\widehat{\text{odds}} = \frac{\hat{p}}{1 - \hat{p}} = \frac{\hat{\pi}}{1 - \hat{\pi}}$$

- **Example:** if  $\hat{\pi} = 0.75$ , then odds of success =  $\frac{0.75}{0.25} = 3$
- If odds > 1, it implies a success is more likely than a failure
  - Example: for  $\widehat{\text{odds}} = 3$ , we expect to observe three times as many successes as failures
- If odds is known, the probability of success can be computed

$$\hat{\pi} = \frac{\widehat{\text{odds}}}{\widehat{\text{odds}} + 1}$$

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	$n_{11}$	$n_{12}$	$n_1$
2	$n_{21}$	$n_{22}$	$n_2$
Total	$n_+$ (or $n_S$ )	$n_-$ (or $n_F$ )	$n$

# Odds Ratio (OR)

- Odds ratio is the ratio of two odds:

$$\widehat{OR} = \frac{\widehat{\text{odds}}_1}{\widehat{\text{odds}}_2} = \frac{\hat{p}_1 / (1 - \hat{p}_1)}{\hat{p}_2 / (1 - \hat{p}_2)}$$

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	n <sub>11</sub>	n <sub>12</sub>	n <sub>1</sub>
2	n <sub>21</sub>	n <sub>22</sub>	n <sub>2</sub>
Total	n <sub>+</sub> (or n <sub>S</sub> )	n <sub>-</sub> (or n <sub>F</sub> )	n

- **Range:** [0, ∞]
  - **Interpretation:** The odds of success for “group 1” is “ $\widehat{OR}$ ” times the odds of success for “group 2”
- 
- What do values of odds ratios mean?

Odds Ratio	Clinical Meaning
$\widehat{OR} < 1$	Odds of success is smaller in group 1 than in group 2
$\widehat{OR} = 1$	Explanatory and response variables are independent
$\widehat{OR} > 1$	Odds of success is greater in group 1 than in group 2

# Poll Everywhere Question 3

# Odds Ratio (OR)

- Values of OR farther from 1.0 in a given direction represent stronger association
  - An OR = 4 is farther from independence than an OR = 2
  - An OR = 0.25 is farther from independence than an OR = 0.5
  - For OR = 4 and OR = 0.25, they are equally away from independence (because  $\frac{1}{4} = 0.25$ )
- We take the inverse of the OR for **success** of **group 1** compared to **group 2** to get...
  - OR for **failure** of **group 1** compared to **group 2**
  - OR for **success** of **group 2** compared to **group 1**

# Log-transformation of $OR$

- Like RR, sampling distribution of the odds ratio is highly skewed
  - Log transformation results in approximately normal distribution
  - Thus, compute confidence interval using normally distributed, log-transformed OR
- Approximate standard error for  $\ln(\widehat{OR})$ :

$$SE_{\ln(\widehat{OR})} = \sqrt{\frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}}}$$

- 95% confidence interval for  $\ln(\widehat{OR})$ :

$$\ln(\widehat{OR}) \pm 1.96 \times SE_{\ln(\widehat{OR})}$$

Explanatory Variable	Response Variable		Total
	Success	Failure	
1	$n_{11}$	$n_{12}$	$n_1$
2	$n_{21}$	$n_{22}$	$n_2$
Total	$n_+$ (or $n_S$ )	$n_-$ (or $n_F$ )	$n$

# How do we get back to the OR scale?

- We computed confidence interval using normally distributed, log-transformed OR ( $\ln(\widehat{OR})$ ):

$$\left( \ln(\widehat{OR}) - 1.96 \times SE_{\ln(\widehat{OR})}, \ln(\widehat{OR}) + 1.96 \times SE_{\ln(\widehat{OR})} \right)$$

- Now we need to **exponentiate the CI to get back to interpretable values**
  - Take exponential of lower and upper bounds
- 95% confidence interval for RR: two ways to display equation

$$\left( e^{\ln(\widehat{OR}) - 1.96 \times SE_{\ln(\widehat{OR})}}, e^{\ln(\widehat{OR}) + 1.96 \times SE_{\ln(\widehat{OR})}} \right)$$

$$\left( \exp \left[ \ln(\widehat{OR}) - 1.96 \times SE_{\ln(\widehat{OR})} \right], \exp \left[ \ln(\widehat{OR}) + 1.96 \times SE_{\ln(\widehat{OR})} \right] \right)$$

# SHS Example: Odds Ratio (1/6)

## Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## Needed steps:

1. Compute the odds ratio
2. Find confidence interval of log OR
3. Convert back to OR
4. Interpret the estimate

# SHS Example: Odds Ratio (2/6)

## Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1. Compute the odds ratio

$$\hat{p}_1 = 198/532 = 0.3722, \hat{p}_2 = 128/1132 = 0.1131$$

$$\widehat{OR} = \frac{\hat{p}_1 / (1 - \hat{p}_1)}{\hat{p}_2 / (1 - \hat{p}_2)} = \frac{0.3722 / (1 - 0.3722)}{0.1131 / (1 - 0.1131)} = 4.6499$$



# SHS Example: Odds Ratio (3/6)

## Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

2. Find confidence interval of log OR

$$\begin{aligned} & \ln(\widehat{OR}) \pm 1.96 \times SE_{\ln(\widehat{OR})} \\ & = \ln(\widehat{OR}) \pm z_{(1-\frac{\alpha}{2})}^* \times \sqrt{\frac{1}{n_{11}} + \frac{1}{n_{12}} + \frac{1}{n_{21}} + \frac{1}{n_{22}}} \\ & = 1.5368 \pm 1.96 \times \sqrt{\frac{1}{198} + \frac{1}{334} + \frac{1}{128} + \frac{1}{1004}} \\ & = (1.2824, 1.7913) \end{aligned}$$

# SHS Example: Odds Ratio (4/6)

## Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## 3. Convert back to OR

$$\begin{aligned} &(\exp(1.2824), \exp(1.7913)) \\ &=(3.6053, 5.9971) \end{aligned}$$

# SHS Example: Odds Ratio (5/6)

## Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

1/2/3. Compute OR and 95% confidence interval

```
1 library(epitools)
2 SHS_ct = table(SHS$glucimp, SHS$case)
3 # no `rev` needed below bc we set the reference level in slide 32
4 oddsratio(x = SHS_ct, method = "wald")$measure
```

	odds ratio with 95% C.I.		
	estimate	lower	upper
Normal	1.000000	NA	NA
Impaired	4.649888	3.605289	5.997148

# Pause: other option in pubh package

```
1 contingency(case ~ glucimp, data = SHS, digits = 3)
```

		Outcome			
Predictor		1	0		
Impaired		198	334		
Normal		128	1004		

	Outcome +	Outcome -	Total	Inc risk *
Exposed +	198	334	532	37.218 (33.097 to 41.482)
Exposed -	128	1004	1132	11.307 (9.521 to 13.298)
Total	326	1338	1664	19.591 (17.709 to 21.581)

Point estimates and 95% CIs:

Inc risk ratio	3.291 (2.703, 4.008)
Inc odds ratio	4.650 (3.605, 5.997)
Attrib risk in the exposed *	25.911 (21.408, 30.413)
Attrib fraction in the exposed (%)	69.618 (63.004, 75.050)
Attrib risk in the population *	8.284 (5.631, 10.937)
Attrib fraction in the population (%)	42.284 (34.713, 48.976)

Uncorrected chi2 test that OR = 1: chi2(1) = 154.239 Pr>chi2 = <0.001

Fisher exact test that OR = 1: Pr>chi2 = <0.001

Wald confidence limits  
CI: confidence interval  
\* Outcomes per 100 population units

Pearson's Chi-squared test with Yates' continuity correction

data: dat  
X-squared = 152.6, df = 1, p-value < 2.2e-16

# SHS Example: Odds Ratio (6/6)

Odds ratio

Compute the point estimate and 95% confidence interval for the diabetes Odds ratio between impaired and normal glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

### 3. Interpret the estimate

The estimated odds of diabetes for American Indians with impaired glucose tolerance at baseline is 4.65 times the odds for American Indians with normal glucose tolerance at baseline.

Additional interpretation of 95% CI (not needed): We are 95% confident that the odds ratio is between 3.61 and 6.00.

Since the 95% confidence interval does not include 1, there is sufficient evidence that the odds of diabetes differs significantly between impaired and normal glucose tolerance at baseline.

# Inversing an Odds Ratio

- Some people prefer interpretations of  $OR > 1$  instead of an  $OR < 1$
- The transformation can easily be done by inverse
  - Remember we discussed that  $OR = 4$  is an equivalent a strong association as  $OR = 0.25$  (1/4)
- OR comparing **group 1** to **group 2** = inverse of OR comparing **group 2** to **group 1**

$$OR_{1v2} = \frac{\hat{p}_1 / (1 - \hat{p}_1)}{\hat{p}_2 / (1 - \hat{p}_2)} = \frac{1}{\frac{\hat{p}_2 / (1 - \hat{p}_2)}{\hat{p}_1 / (1 - \hat{p}_1)}} = \frac{1}{OR_{2v1}}$$

# Poll Everywhere Question 4

# SHS Example: Inversing Odds Ratio

## Inversing odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between **normal** and **impaired** glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## Needed steps:

1. Inverse point estimate and confidence interval

$$\widehat{OR} = \frac{1}{4.6499} = 0.2151$$

The 95% Confidence interval is then

$$\left( \frac{1}{5.9971}, \frac{1}{3.6053} \right) = (0.1667, 0.2774)$$



# SHS Example: Inversing Odds Ratio

## Inversing odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between **normal** and **impaired** glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

## Needed steps:

### 1. Inverse point estimate and confidence interval

```
1 library(epitools)
2 oddsratio(x = SHS_ct, method = "wald", rev = "rows")$measure
```

```
      odds ratio with 95% C.I.
      estimate      lower      upper
Impaired 1.000000      NA      NA
Normal   0.215059 0.1667459 0.2773702
```

# SHS Example: Inversing Odds Ratio

## Inversing odds ratio

Compute the point estimate and 95% confidence interval for the diabetes odds ratio between **normal** and **impaired** glucose tolerance.

Glucose tolerance	Diabetes		Total
	No	Yes	
Impaired	334	198	532
Normal	1004	128	1132
Total	1338	326	1664

### Needed steps:

2. Interpret the estimate

The estimated odds of diabetes for American Indians with normal glucose tolerance at baseline is 0.22 times the odds for American Indians with impaired glucose tolerance at baseline.

Additional interpretation of 95% CI (not needed): We are 95% confident that the odds ratio is between 0.17 and 0.28.

Since the 95% confidence interval does not include 1, there is sufficient evidence that the odds of diabetes differs significantly between impaired and normal glucose tolerance at baseline.

# Learning Objectives

1. Understand the difference between testing for association and measuring association
2. Estimate the risk difference (and its confidence interval) from a contingency table and interpret the estimate.
3. Estimate the risk ratio (and its confidence interval) from a contingency table and interpret the estimate.
4. Estimate the odds ratio (and its confidence interval) from a contingency table and interpret the estimate.

# pubh vs. epitools

- In `pubh` with `contingency()`
  - Get all the info at once
  - Really nice to double check how the code is interpreting your input
- In `epitools` with `riskratio()` or `oddsratio()`
  - Much easier to grab the numbers!
  - In Quarto you can take R code and directly put it in your text

```
1 g = oddsratio(x = SHS_ct, method = "wald", rev = "rows")  
2 g$measure[2,1]
```

```
[1] 0.215059
```

- I can write `{r eval="false" echo="true"} round(g$measure[2,1], 3)` to print the number 0.215