

# Last Class!

Revisit interactions + Intro to Poisson Regression

# Announcements

- Do you all get requests for the formal course evaluations?
- Remaining grading
  - Homework 4, 5, and 6 redos
  - Project Analysis, Report, and Presentation
- Anything I'm missing?

# Class 17 Learning Objectives

Not as formal today:

1. Calculate odds ratios for variables (categorical x continuous) involved in interactions
  - If we want, we can look at categorical x categorical and continuous x continuous
2. Run and interpret a simple Poisson Regression

# Odds Ratio in the Presence of Interaction (I)

- When interaction exists between a risk factor (F) and another variable (X), the estimate of the odds ratio for F depends on the value of X
- When an interaction term (F\*X) exists in the model
  - $OR_F \neq \exp(\beta_F)$  in general
- Assume we want to compute the odds ratio for ( $F = f_1$  and  $F = f_0$ ), the correct model-based estimate is

$$\widehat{OR}_F = \exp(\hat{g}(F = f_1, x) - \hat{g}(F = f_0, x))$$

# Odds Ratio in the Presence of Interaction (II)

- We may write the two logits for given  $x$  as below:

$$\hat{g}(f_1, x) = \hat{\beta}_0 + \hat{\beta}_1 f_1 + \hat{\beta}_2 x + \hat{\beta}_3 f_1 \times x$$

$$\hat{g}(f_0, x) = \hat{\beta}_0 + \hat{\beta}_1 f_0 + \hat{\beta}_2 x + \hat{\beta}_3 f_0 \times x$$

- The difference in two logits is:

$$\hat{g}(f_1, x) - \hat{g}(f_0, x) = \hat{\beta}_1 (f_1 - f_0) + \hat{\beta}_3 x (f_1 - f_0)$$

- Therefore,

$$\begin{aligned} \widehat{OR}(F = f_1, F = f_0, X = x) \\ = \exp[\hat{\beta}_1 (f_1 - f_0) + \hat{\beta}_3 x (f_1 - f_0)] \end{aligned}$$

# Odds Ratio in the Presence of Interaction (III)

To find the confidence interval for the  $\widehat{OR}$ , we need to find confidence interval for  $\hat{g}(f_1, x) - \hat{g}(f_0, x)$  first

$$\begin{aligned} & Var(\hat{g}(f_1, x) - \hat{g}(f_0, x)) \\ &= (f_1 - f_0)^2 \times Var(\hat{\beta}_1) + [x(f_1 - f_0)]^2 \times var(\hat{\beta}_3) + 2x(f_1 - f_0)^2 \times cov(\hat{\beta}_1, \hat{\beta}_3) \end{aligned}$$

- We can follow the same procedure as we did previously to construct 95% CI for  $\hat{g}(f_1, x) - \hat{g}(f_0, x)$ , denoted by  $[L_{g01}, U_{g01}]$ .

$$\widehat{OR} = \exp(\hat{g}(f_1, x) - \hat{g}(f_0, x))$$

The 95% CI for  $\widehat{OR}$  is:  $[\exp(L_{g01}), \exp(U_{g01})]$

# Example: GLOW – Computing OR under interaction

**Step 1:** The two sets of values of the covariates are  
(priorfrac = 1, Age = a) compared to  
(priorfrac = 0, Age = a)

**Step 2:** Substituting these values into the general expression

**Step 3:** Taking the difference in the two functions

**Step 4:** Exponentiating the result

# Example: GLOW – Step 1

**Step 1:** The two sets of values of the covariates are ( $\text{priorfrac} = 1$ ,  $\text{Age} = a$ ) compared to ( $\text{priorfrac} = 0$ ,  $\text{Age} = a$ )



# Example: GLOW – Step 2

**Step 2:** Substituting these values into the general expression

$$\text{logit}(\pi(\mathbf{x})) = \hat{\beta}_0 + \hat{\beta}_1 \text{PF} + \hat{\beta}_2 \text{Age} + \hat{\beta}_3 \text{PF} \times \text{Age}$$

$$\begin{aligned} \text{logit}(\pi(\text{PF} = 1, \text{Age} = a)) &= \hat{\beta}_0 + \hat{\beta}_1 \times 1 + \hat{\beta}_2 a + \hat{\beta}_3 \times 1 \times a \\ &= \hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 a + \hat{\beta}_3 a \end{aligned}$$

$$\begin{aligned} \text{logit}(\pi(\text{PF} = 0, \text{Age} = a)) &= \hat{\beta}_0 + \hat{\beta}_1 \times 0 + \hat{\beta}_2 a + \hat{\beta}_3 \times 0 \times a \\ &= \hat{\beta}_0 + \hat{\beta}_2 a \end{aligned}$$

# Example: GLOW – Step 3

**Step 3:** Taking the difference in the two functions

$$\begin{aligned} & [\text{logit}(\pi(\text{PF} = 1, \text{Age} = a))] - [\text{logit}(\pi(\text{PF} = 0, \text{Age} = a))] \\ &= [\hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 a + \hat{\beta}_3 a] - [\hat{\beta}_0 + \hat{\beta}_2 a] \\ &= \hat{\beta}_1 + \hat{\beta}_3 a \end{aligned}$$

# Example: GLOW – Step 4

**Step 4:** Exponentiating the result

| Model | Variable                               | Coeff. | Std. Err. | z     | p      | 95% CI  |        |
|-------|--|--------|-----------|-------|--------|---------|--------|
| 3     | $\hat{\beta}_1$ PRIORFRAC              | 4.961  | 1.8102    | 2.74  | 0.006  | 1.413,  | 8.509  |
|       | $\hat{\beta}_2$ AGE                    | 0.063  | 0.0155    | 4.04  | <0.001 | 0.032,  | 0.093  |
|       | $\hat{\beta}_3$ PRIORFRAC $\times$ AGE | -0.057 | 0.0250    | -2.29 | 0.022  | -0.106, | -0.008 |
|       | $\hat{\beta}_0$ Constant               | -5.689 | 1.0841    | -5.25 | <0.001 | -7.814, | -3.565 |

$$\widehat{OR}[(PF = 1, Age = a), (PF = 0, Age = a)] = \exp(\hat{\beta}_1 + \hat{\beta}_3 a)$$
$$= \exp(4.961 - 0.057a)$$

If we let  $a = 60$ , i.e., compute OR for age = 60, then

- $\widehat{OR}_{a=60} = \exp(4.961 - 0.057 * 60) = 4.669$

If we let  $a = 70$ , i.e., compute OR for age = 70, then

- $\widehat{OR}_{a=70} = \exp(4.961 - 0.057 * 70) = 2.64$

# Example: GLOW – Table of ORs

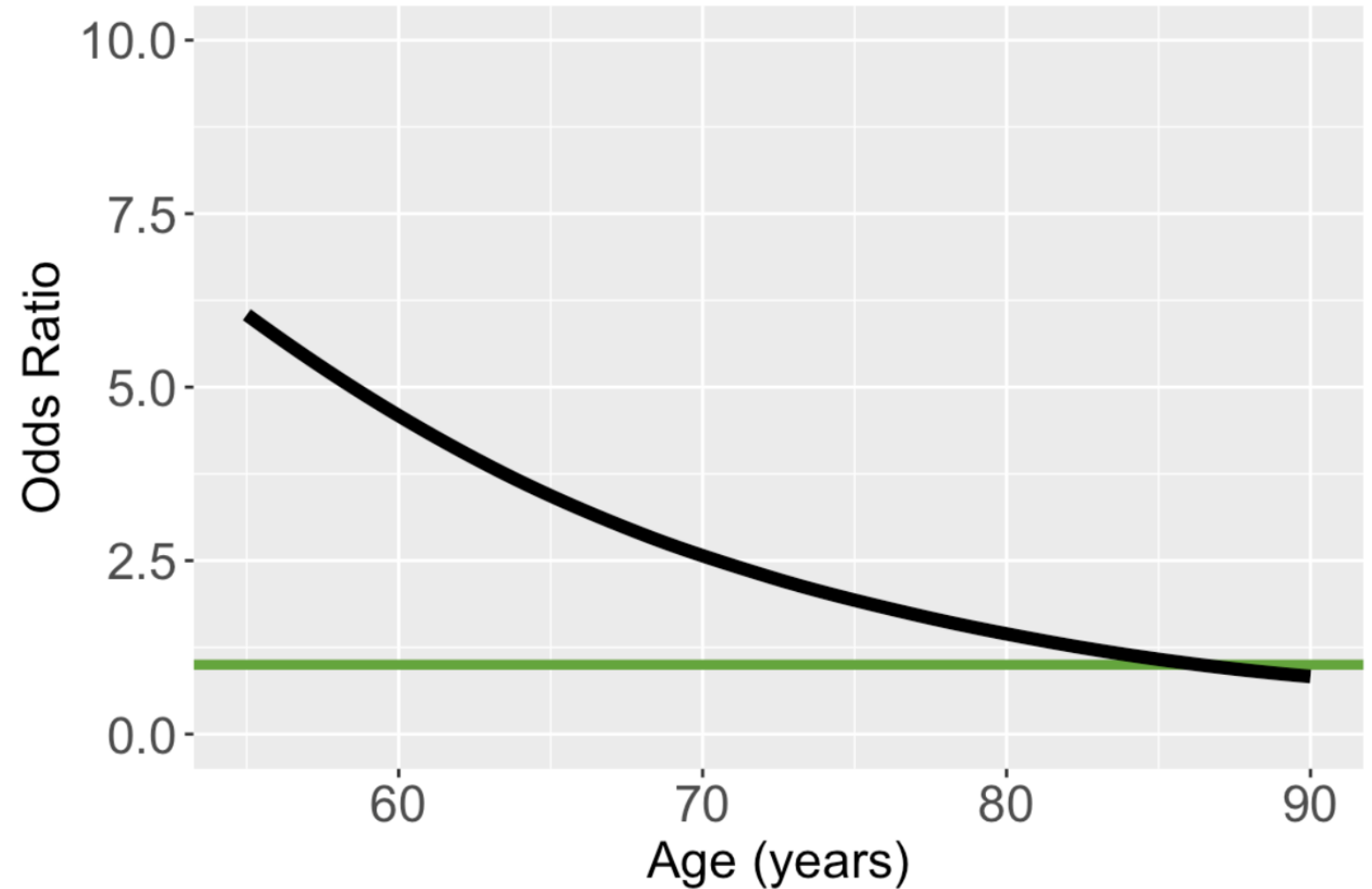
- Sometimes a Table listing the estimated odds ratio and 95% CI at different value of X will also be helpful and informative

**Table 3.13   Estimated Odds Ratios for Prior Fracture  
as a Function of Age from Model 3 in Table 3.12**

| Age | Odds Ratio | 95% CI      |
|-----|------------|-------------|
| 55  | 6.1        | 2.38, 15.53 |
| 60  | 4.6        | 2.20, 9.49  |
| 65  | 3.4        | 1.96, 5.99  |
| 70  | 2.6        | 1.63, 4.06  |
| 75  | 1.9        | 1.20, 3.11  |
| 80  | 1.4        | 0.79, 2.65  |

# Plot of Odds Ratio Involving Interaction Term

- Odds ratio comparing prior fracture to no prior fracture

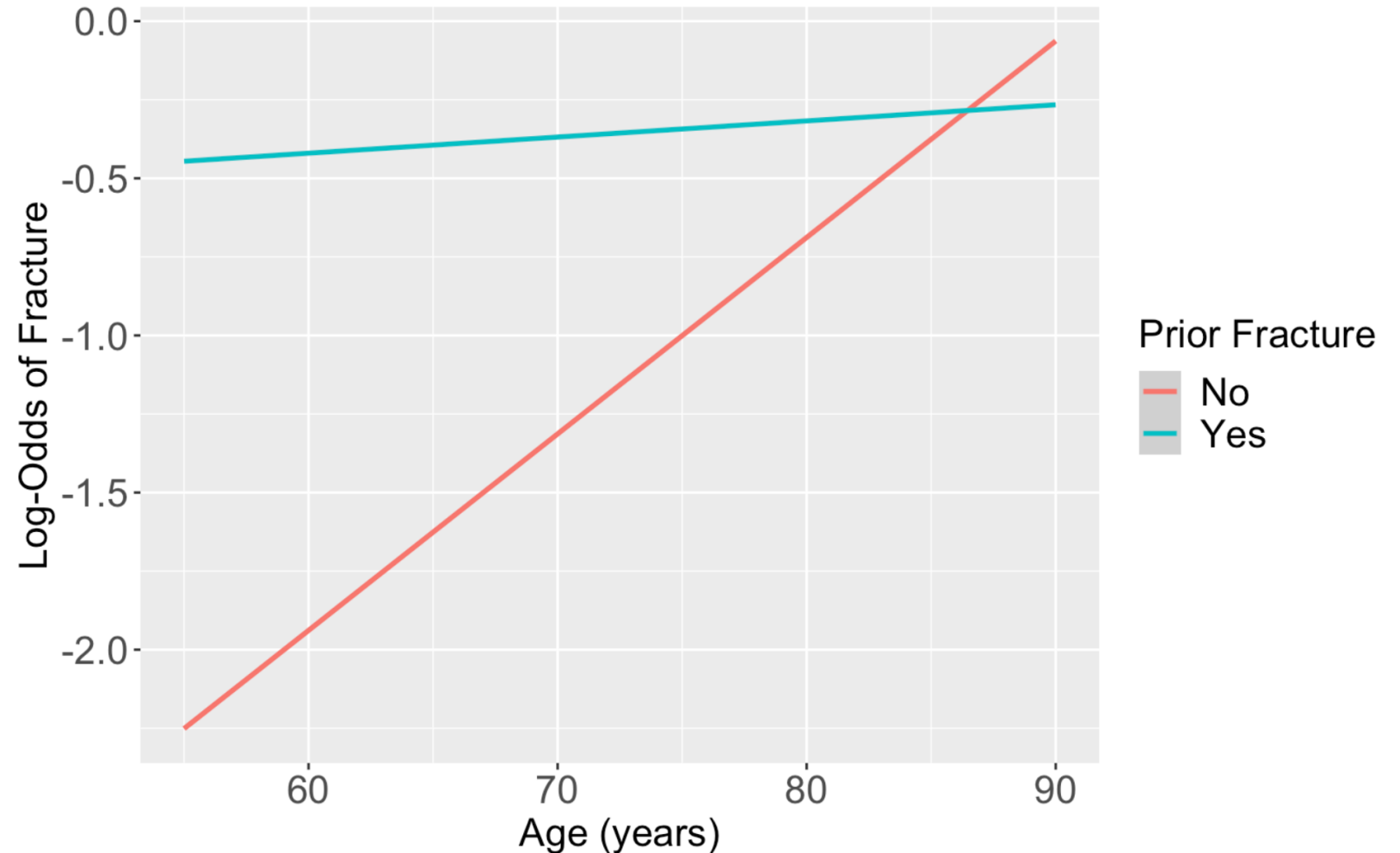


# Graphically Showing Interaction Revisited

- One easy way to see the nature of the interaction between F and X is to plot the two logit functions
- Using the GLOW fracture example:

$$\begin{aligned}\text{logit}(\pi(\text{PF} = 1, \text{Age} = a)) \\ &= \hat{\beta}_0 + \hat{\beta}_1 + \hat{\beta}_2 a + \hat{\beta}_3 a \\ &= (\hat{\beta}_0 + \hat{\beta}_1) + (\hat{\beta}_2 + \hat{\beta}_3) a\end{aligned}$$

$$\begin{aligned}\text{logit}(\pi(\text{PF} = 0, \text{Age} = a)) \\ &= \hat{\beta}_0 + \hat{\beta}_2 a\end{aligned}$$



# Simple Poisson Regression

# Poisson Distribution (I)

- This distribution is often used to model count data
- Examples:
  - Distribution of number of deaths due to lung cancer
  - Distribution of number of individuals diagnosed with leukemia
  - Distribution of number of hospitalizations



# Poisson Distribution (II)

- The probability function of Poisson distribution:

$$P(Y = y|\mu) = \frac{e^{-\mu} \mu^y}{y!}$$

- Where  $y$  are non-negative integers  $y = 0, 1, 2, \dots$
  - Where  $\mu$  is the mean of  $Y$ , that is  $E(Y) = \mu$
  - And also,  $\text{Var}(Y) = \mu$
- 
- For a Poisson distribution,  $Y \sim \text{Poisson}(\mu)$ 
    - Range:  $[0, \infty)$

# Poisson Distribution (III)

- If we look at the probability of  $y$  events in a time period  $t$  for a Poisson random variable, we could write:

$$P(Y = y|\mu) = \frac{e^{-\mu} \mu^y}{y!}$$

- Where  $y$  are non-negative integers  $y = 0, 1, 2, \dots$
- And  $\mu = \lambda t$ , where  $\lambda$  is the expected number of events per unit time
- Then  $\mu$  is the expected number of events over time  $t$

# Poisson Distribution (IV)

- What does  $\lambda$  represent here?
  - A **rate**, the expected number of events in a given population over a given period time
    - **Example:** Number of patient arrivals into the Emergency Room *per hour*
- The Poisson distribution is the prototype for assigning probabilities of observing any number of events

# **Poll Everywhere**

## **Question 1**

# Why Person-Years? (I)

- In the example of number of patient arrivals, an event does not conclude the study
  - If someone arrives within the first minute of the study, then we keep counting
  - We may be able to study the association of arrivals with qualities of the hospital, but we can't measure qualities of the individuals arriving
- What happens if we want to measure qualities of the individual?
  - We can measure a hospitalization rate

# Why Person-Years? (II)

- If we are measuring at the individual level and counting something that is “terminal” then our count will always be 0 or 1
  - Example: Number of individuals diagnosed with leukemia
    - This only happens once, so how do we measure the rate here?
- Since rate involves the counts and time – we can use the time to diagnosis to estimate the rate
  - Often expressed in units such as events per thousand person-years

# Calculating Person-Year

- One **person-year** is a unit of time defined as **one person being followed for one year**
- Person-years for a sample of  $n$  subjects is calculated as the total years followed for the  $n$  subjects, where each subject could have different follow-up time
- **Example:** suppose we have 5 subjects, two of the subjects were followed for 2 years, and two of them are followed for 3 years and the fifth subject was followed for 3.8 years

$$\begin{aligned} &\text{person} - \text{years} \\ &= 2 \text{ people} \times 2 \text{ years} + 2 \text{ people} \times 3 \text{ years} \\ &+ 1 \text{ person} \times 3.8 \text{ years} = 13.8 \text{ person} - \text{years} \end{aligned}$$

# Calculating Rate (II)

- Suppose that we observe one event during the follow-up period, then

$$\begin{aligned}\text{Rate of event} &= \frac{\# \text{ events}}{\text{person} - \text{years}} = \frac{1 \text{ event}}{13.8 \text{ person} - \text{years}} \\ &= 0.072 \text{ events per person} - \text{year} \\ &= 72 \text{ events per } 1000 \text{ person} - \text{years}\end{aligned}$$



# Stop here?

If we do:

- Good tutorial in R: <https://www.dataquest.io/blog/tutorial-poisson-regression-in-r/>

# Review: Simple Logistic Regression

- Let  $Y$  is the dependent variable of interest and  $x$  is a predictor variable,
- In simple logistic regression, we have

$$\log \left( \frac{\pi(x_i)}{1-\pi(x_i)} \right) = \beta_0 + \beta_1 x_i$$

where  $\pi(x_i) = P(Y_i = 1|x_i)$

# Simple Poisson Regression Model

- What do we model in a Poisson regression?
- Log of conditional mean of Y given x
  - Let **Y** be a **Poisson count** for **a given unit of time**, then
$$\mu(x) = \lambda(x)$$
  - In a simple Poisson regression, we have

$$\ln(\mu(x_i)) = \ln(\lambda(x_i)) = \beta_0 + \beta_1 x_i$$

- So this is also called a **log-linear** model

# Parameter Interpretation (I)

- In simple Poisson regression:

$$\ln(\mu(x_i)) = \ln(\lambda(x_i)) = \beta_0 + \beta_1 x_i$$

- When  $x$  is a **binary variable**: How do we interpret  $\beta_0$  and  $\beta_1$ ?

- When  $x_i = 0$ :

$$\ln(\mu(x_i = 0)) = \ln(\lambda(x_i = 0)) = \beta_0$$

- $\mu(x_i = 0) = \exp(\beta_0)$ : the mean (rate) of  $Y_i$  when  $x_i = 0$

- When  $x_i = 1$ :

$$\ln(\mu(x_i = 1)) = \ln(\lambda(x_i = 1)) = \beta_0 + \beta_1$$

- $\mu(x_i = 1) = \exp(\beta_0 + \beta_1)$ : the mean (rate) of  $Y_i$  when  $x_i = 1$

# Parameter Interpretation (II)

- When  $x$  is a **binary variable**: How do we interpret  $\beta_0$  and  $\beta_1$ ?
- By subtraction, we have

$$\beta_1 = \ln \left( \frac{E(Y_i | x_i = 1)}{E(Y_i | x_i = 0)} \right) = \ln \left( \frac{\lambda(x_i = 1)}{\lambda(x_i = 0)} \right)$$

- $\beta_1$ : Log-rate ratio
- **And  $\exp(\beta_1)$  is the rate ratio**

# Further reading / tutorials on Poisson regression

- Good tutorial in R: <https://www.dataquest.io/blog/tutorial-poisson-regression-in-r/>
- When people are followed for different amounts of time, we should include an offset
  - Poisson Regression Modeling Using Rate Data: section from above site that discusses offsets
- We can use Wald test and LRT in the same way as logistic regression to test our coefficients and variables

# Wrap-up

- 4-minute exit ticket
- Thanks for a great quarter!

## Class 17 Exit Ticket



<https://forms.office.com/r/5q7YxDi58s>