

# ECON550: Problem Set 4

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## HMC Exercises (7th edition)

### 2.1.1

$$P(0 < X_1 < 1/2, 1/4 < X_2 < 1) = \int_0^{1/2} \int_{1/4}^1 4x_1x_2 = \frac{1}{4} \left(1 - \frac{1}{16}\right) = \frac{15}{64}$$

$$P(X_1 = X_2) = 0$$

$$P(X_1 < X_2) = \int_0^1 \int_0^{x_2} 4x_1x_2 \, dx_1 \, dx_2 = \int_0^1 2x_2(x_2^2) = \frac{2}{4} = 1/2$$

$$P(X_1 \leq X_2) = 1/2$$

### 2.1.6

$$P(Z \leq 0) = 0$$

$$P(Z \leq 6) = \int_0^6 \int_0^{6-x} e^{-x-y} \, dy \, dx = \int_0^6 (e^{-x} - e^{-6}) \, dy = 1 - e^{-6} - 6e^{-6} \approx 0.9826$$

$$F_Z(t) = \int_0^t \int_0^{t-x} e^{-x-y} \, dy \, dx = \int_0^t e^{-x} - e^{-t} \, dx = 1 - e^{-t} - te^{-t}$$

$$f_Z(t) = te^{-t}$$

where  $t \geq 0$ .

### 2.1.10 Marginal distribution:

$$f_{X_1}(x_1) = \int_{x_1}^1 15x_1^2x_2 \, dx_2 = \frac{15}{2}x_1^2(1 - x_1^2)$$

$$f_{X_2}(x_2) = \int_0^{x_2} 15x_1^2x_2 \, dx_1 = 5x_2^4$$

$$P(X_1 + X_2 \leq 1) = \int_0^{1/2} \int_{x_1}^{1-x_1} 15x_1^2x_2 \, dx_2 \, dx_1 = \int_0^{1/2} (15/2)x_1^2(1 - 2x_1) \, dx_1 = (5/16) - (15/64) = 5/64$$

**2.1.16**

$$\begin{aligned} P(2X + 3Y < 1) &= \int_0^{1/2} \int_0^{(1-2x)/3} 6(1-x-y) dy dx = \int_0^{1/2} 6 \left( \frac{(1-x)(1-2x)}{3} - \frac{(1-2x)^2}{18} \right) dx \\ &= \frac{1}{3} \int_0^{1/2} (5 - 14x + 8x^2) dx = \frac{1}{3} ((5/2) - (7/4) + (1/3)) = 13/36 \end{aligned}$$

**2.3.1** Conditional mean:

$$E[X_2|X_1 = x_1] = \int_0^1 x_2 \frac{x_1 + x_2}{\frac{1}{2} + x_1} dx_2 = \frac{x_1/2 + 1/3}{\frac{1}{2} + x_1} = \frac{3x_1 + 2}{6x_1 + 3}$$

Conditional variance:

$$\begin{aligned} E[X_2^2|X_1 = x_1] &= \int_0^1 x_2^2 \frac{x_1 + x_2}{\frac{1}{2} + x_1} dx_2 = \frac{x_1/3 + 1/4}{x_1 + 1/2} = \frac{4x_1 + 3}{12x_1 + 6} \\ E[X_2^2|X_1 = x_1] - (E[X_2|X_1 = x_1])^2 &= \frac{4x_1 + 3}{12x_1 + 6} - \left( \frac{3x_1 + 2}{6x_1 + 3} \right)^2 \\ &= \frac{(4x_1 + 3)(2x_1 + 1)}{6(2x_1 + 1)^2} - \frac{(3x_1 + 2)(3x_1 + 2)}{9(2x_1 + 1)^2} = \frac{6x_1^2 + 6x_1 + 1}{18(2x_1 + 1)^2} \end{aligned}$$

**2.3.2a, b** To determine  $c_1$  and  $c_2$ , we need normalization to 1. For  $c_1$ , we get

$$\int_0^{x_2} c_1 x_1 / x_2^2 dx_1 = 1$$

So we need  $c_1 = 2$ . Likewise, we get that  $c_2 = 5$ . Then the joint pdf is

$$f(x_1, x_2) = f_{1|2}(x_1|x_2)f_2(x_2) = 10x_1x_2^2$$

$$0 < x_1 < x_2 < 1$$

**2.3.10** The marginal probability function of  $x_1$ :

$$p_1(0) = 4/18$$

$$p_1(1) = 7/18$$

$$p_1(2) = 7/18$$

The marginal probability function of  $x_2$ :

$$p_2(0) = 11/18$$

$$p_2(1) = 7/18$$

The conditional means are:

$$E[X_1|X_2 = 0] = 16/11$$

$$E[X_1|X_2 = 1] = 5/7$$

$$E[X_2|X_1 = 0] = 3/4$$

$$E[X_2|X_1 = 1] = 3/7$$

$$E[X_2|X_1 = 2] = 1/7$$

**2.4.1a**

$$\frac{Cov(X, Y)}{\sqrt{V(X)V(Y)}} = \frac{(2/3)}{(2/3)} = 1$$

**2.4.6**

$$E[Y|x] = \int_{-x}^x \frac{y}{2x} = \frac{(1/2)x^2 - (1/2)(-x)^2}{2x} = 0$$

and hence this is a straight line.

**2.4.11** Consider  $Y = X_1 + X_2$ . Then by linearity of expectation,  $E[Y] = \mu_1 + \mu_2$ . Also

$$E[Y^2] = E[X_1^2 + 2X_1X_2 + X_2^2] = E[X_1^2] + E[X_2^2] + 2E[X_1X_2]$$

$$V[Y] = E[Y^2] - E[Y]^2 = E[X_1^2] + E[X_2^2] + 2E[X_1X_2] - (\mu_1 + \mu_2)^2$$

$$= E[X_1^2] + E[X_2^2] + 2E[X_1X_2] - \mu_1^2 - \mu_2^2 - 2\mu_1\mu_2$$

$$= \sigma^2 + \sigma^2 + 2Cov(X_1, X_2)$$

$$= 2\sigma^2 + 2Cov(X_1, X_2)$$

$$= 2\sigma^2(1 + \rho)$$

Then by Chebyshev, we have

$$P(|Y - \mu_1 - \mu_2| \geq k\sigma) \leq \frac{2\sigma^2(1 + \rho)}{(k\sigma)^2} = \frac{2(1 + \rho)}{k^2}$$

**2.5.2** Can't be independent because the support is non-rectangular, as we argued in class. Since  $x_1 < x_2$ , we have dependence.

**2.5.5** The desired probability is 5

$$\frac{2}{3} + \frac{5}{8} - \frac{5}{12} = \frac{7}{8}$$

**2.5.8** These are not independent, the support isn't a rectangle.

$$E(X|y) = \int_y^1 \frac{3x^2}{(3/2)(1-y^2)} \frac{2}{3}$$

**2.5.11** Let  $X_1$  be the first midpoint,  $X_2$  the second. Then

$$\begin{aligned}
 P(X_1 > X_2 + 1) &= P(X_1 - X_2 > 1) = \int_7^1 4 \int_6^{x-1} \frac{1}{196} dy dx = \frac{1}{8} \\
 P(X_2 > X_1 + 1) &= P(X_2 - X_1 > 1) = \int_6^2 0 \int_0^{y-1} \frac{1}{196} dy dx = \frac{311}{392} \\
 1 - (P(X_1 > X_2 + 1) + P(X_2 > X_1 + 1)) &= \frac{4}{49}
 \end{aligned}$$

**2.8.1**

$$\begin{aligned}
 (n-1)^{-1} \sum_{i=1}^n (X_i - \bar{X})^2 &= (n-1)^{-1} \sum_{i=1}^n X_i^2 - 2\bar{X}X_i + \bar{X}^2 \\
 &= (n-1)^{-1} \left( \sum_{i=1}^n X_i^2 - \sum_{i=1}^n 2\bar{X}X_i + \sum_{i=1}^n \bar{X}^2 \right) \\
 &= (n-1)^{-1} \left( \sum_{i=1}^n X_i^2 - 2\bar{X} \sum_{i=1}^n X_i + n\bar{X}^2 \right) \\
 &= (n-1)^{-1} \left( \sum_{i=1}^n X_i^2 - 2n\bar{X}^2 + n\bar{X}^2 \right) \\
 &= (n-1)^{-1} \left( \sum_{i=1}^n X_i^2 - n\bar{X}^2 \right)
 \end{aligned}$$

**2.8.4**

$$\begin{aligned}
 E[X_1 X_2] &= \mu_1 \mu_2 + Cov(X_1, X_2) = \mu_1 \mu_2 \\
 V(X_1 X_2) &= E[X_1^2 X_2^2] - \mu_1^2 \mu_2^2 = E[X_1]^2 E[X_2]^2 - \mu_1^2 \mu_2^2 \\
 &= (\sigma_1^2 + \mu_1^2)(\sigma_2^2 + \mu_2^2) - \mu_1^2 \mu_2^2 \\
 &= \sigma_1^2 \sigma_2^2 + \mu_1^2 \sigma_2^2 + \mu_2^2 \sigma_1^2
 \end{aligned}$$

**2.8.10**

$$\begin{aligned}
 V(X + 2Y) &= Cov(X + 2Y, X + 2Y) = Var(X) + 4Var(Y) + 4Cov(X, Y) = 15 \\
 Cov(X, Y) &= 3/4 \\
 \rho &= Cov(X, Y) / \sqrt{V(X)V(Y)} = \frac{3}{8\sqrt{2}}
 \end{aligned}$$

**2.8.16**

$$V\left(\sum X_i\right) = \sum V(X_i) + \sum_i \sum_{j \neq i} Cov(X_i, X_j) = 10 \cdot (5) + 90 \cdot (0.5) = 50 + 45 = 95$$

**2.8.18** Consider  $f(x) = \sqrt{x}$ . By concavity and Jensen's inequality,

$$E[f(S^2)] < f(E[S^2])$$

$$E[S] < f(\sigma^2) = \sigma$$

as desired.

**2.6.1a**

$$f_X(x) = \frac{2x+2}{3}$$

$$f_Y(y) = \frac{2y+2}{3}$$

$$f_Z(z) = \frac{2z+2}{3}$$

**2.6.1b**

$$P(0 < X, Y, Z < 1/2) = \int_0^{1/2} \int_0^{1/2} \int_0^{1/2} \frac{2x+2y+2z}{3} = \int_0^{1/2} \int_0^{1/2} \frac{1}{12} + \frac{y+z}{3} = \int_0^{1/2} \frac{1}{24} + \frac{1}{24} + \frac{z}{6} = \frac{1}{24} + \frac{1}{48} = \frac{1}{16}$$

$$P(0 < X < 1/2) = P(0 < Y < 1/2) = P(0 < Z < 1/2) = \int_0^{1/2} \frac{2x+2}{3} = \frac{1}{12} + \frac{1}{3} = \frac{5}{12}$$

**2.6.1c** No. We can verify using part b that  $P(A \cap B \cap C) \neq P(A)P(B)P(C)$

**2.6.1e** The joint cdf is

$$\int \int \int f(x, y, z) = \frac{x^2 + y^2 + z^2}{3}$$

The marginal cdfs are given by

$$F_X(x) = \frac{x^2 + 2x}{3}$$

$$F_Y(y) = \frac{y^2 + 2y}{3}$$

$$F_Z(z) = \frac{z^2 + 2z}{3}$$

**2.6.1f** Conditional distribution:

$$f_{X,Y|Z} = \frac{f(x, y, z)}{f_Z(z)} = \frac{2x+2y+2z}{2z+2} = \frac{x+y+z}{z+1}$$

$$\begin{aligned} E[X+Y|Z=z] &= \int \int (x+y) \frac{x+y+z}{z+1} \\ &= \frac{1}{z+1} \int \int x^2 + x(y+z) + xy + y^2 + yz = \frac{1}{z+1} \int \frac{1}{3} + y + (z/2) + y^2 + yz \end{aligned}$$

$$= \frac{1}{z+1} \left( \frac{1}{3} + \frac{1}{2} + z + \frac{1}{3} \right) \\ = \frac{6z+7}{6z+6}$$

**3.3.2** The PDF is given by:

$$f(x) = \frac{x^{3/2} e^{-x/2}}{3\sqrt{2\pi}}$$

By numerically bashing or consulting a table, we get that  $c = 0.831$ ,  $d = 12.833$

**3.4.2**

$$P(X < 60) = P((X - 75)/10 < 1.5) = 0.0668$$

$$P(70 < X < 100) = P(-0.5 < (X - 75)/10 < 2.5) = 0.6853$$

**3.4.3** We can take  $b \approx 1.645$ .

**3.4.18** Skewness:

$$\frac{E[(X - \mu)^3]}{\sigma^3} = E \left[ \left( \frac{X - \mu}{\sigma} \right)^3 \right] = \int \frac{x^3 e^{-x^2/2}}{\sqrt{2\pi}} = 0$$

Kurtosis

$$\frac{E[(X - \mu)^4]}{\sigma^4} = \int \frac{x^4 e^{-x^2/2}}{\sqrt{2\pi}} = 3$$

**3.5.8** We rewrite

$$f(x, y) = \frac{1}{2\pi} \left( \exp \left( -\frac{1}{2}(x^2 + y^2) \right) + xy \exp \left( -(x^2 + y^2 - 1) \right) \right)$$

Note that when we integrate through the domain of only  $x$  or only  $y$ , the second term disappears due to symmetry about 0, and we will be left with

$$f_X(x) = \frac{1}{\sqrt{2\pi}} \exp \left( -\frac{x^2}{2} \right)$$

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \exp \left( -\frac{y^2}{2} \right)$$

Hence the marginal distributions are normal. However, the pdf is still joint, as the distribution depends on both  $x, y$  and characterizes the realization on  $X, Y$ .

**3.5.10** By HMC Theorem 3.5.1, we have that

$$\mu_Z = a\mu_X + b\mu_Y = 0$$

$$\sigma_Z^2 = a^2\sigma_X^2 + 2ab\rho\sigma_X\sigma_Y + b^2\sigma_Y^2 = a^2 + b^2 + 2ab\rho$$

Then  $Z$  is distributed as  $N(\mu_Z, \sigma_Z^2)$ .

**3.5.15** Take  $a = [1/n, 1/n, 1/n, \dots]$ . Then by Theorem 3.5.1, we have the distribution of  $\bar{X}$  is given by  $N(a\mu, a\Sigma a')$ . If all of its components have the same mean  $\mu$ , then the distribution is then  $N(\mu, a\Sigma a')$ .

## Problem 1

We know that

$$\begin{aligned}
 V(AX + b) &= E[(AX + b) - E[AX + b])(AX + b) - E[AX + b]]' \\
 &= E[(AX - AE[X])(AX - AE[X])'] \\
 &= E[A(X - E[X])(X - E[X])'A'] \\
 &= AE[(X - E[X])(X - E[X])']A' \\
 &= AV(X)A'
 \end{aligned}$$

## Problem 2

$$\begin{aligned}
 Cov(AX + BY, CZ + DW) &= E[(AX + BY - E[AX + BY])(CZ + DW - E[CZ + DW])'] \\
 &= E[(AX + BY - AE[X] - BE[Y])(CZ + DW - CE[Z] - DE[W])'] \\
 &= E[(A(X - E[X]) + B(Y - E[Y]))(C(Z - E[Z]) + D(W - E[W]))'] \\
 &= E[A(X - E[X])(Z - E[Z])'C' + B(Y - E[Y])(Z - E[Z])'C' + A(X - E[X])(W - E[W])'D' + B(Y - E[Y])(W - E[W])'D'] \\
 &= AE[(X - E[X])(Z - E[Z])']C' + BE[(Y - E[Y])(Z - E[Z])']C' \\
 &\quad + AE[(X - E[X])(W - E[W])']D' + BE[(Y - E[Y])(W - E[W])']D' \\
 &= ACov(X, Z)C' + BCov(Y, Z)C' + ACov(X, W)D' + BCov(Y, W)D'
 \end{aligned}$$