# Problem Set 6

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## Problem 1

(a) From the moment equations,

$$E(1 - w\theta) = 0$$

$$E(w)\theta = 1$$

$$\theta = 1$$

and

$$E(\theta^3 - x) = 0$$

$$E(x) = 1$$

Hence, the optimal weighting matrix is

$$E(g'g)^{-1} = E \begin{bmatrix} (\theta^3 - x)^2 & (\theta^3 - x)(1 - w\theta) \\ (\theta^3 - x)(1 - w\theta) & (1 - w\theta)^2 \end{bmatrix}^{-1}$$

$$= E \begin{bmatrix} (1 - x)^2 & (1 - x)(1 - w) \\ (1 - x)(1 - w) & (1 - w)^2 \end{bmatrix}^{-1}$$

$$= E \begin{bmatrix} x^2 - 2x + 1 & 1 - x - w + wx \\ 1 - x - w + wx & w^2 - 2w + 1 \end{bmatrix}^{-1}$$

$$= \begin{bmatrix} 3 - 2 + 1 & 1 - 1 - 1 + 1/2 \\ 1 - 1 - 1 + 1/2 & 2 - 2 + 1 \end{bmatrix}^{-1}$$

$$= \begin{bmatrix} 2 & -1/2 \\ -1/2 & 1 \end{bmatrix}^{-1}$$

$$= \begin{bmatrix} 4/7 & 2/7 \\ 2/7 & 8/7 \end{bmatrix}$$

#### (b) (i) From lecture, we know

$$\sqrt{n}(\hat{\theta} - \theta_0) \to_d N(0, ((G'WG)^{-1})'G'W\Omega W'G(G'WG)^{-1})$$

When  $W = \Omega^{-1}$ , we get

$$((G'\Omega^{-1}G)^{-1})'G'(\Omega^{-1})'G(G'\Omega^{-1}G)^{-1}$$

$$= (G'\Omega^{-1}G)^{-1}$$

$$= \left(\left(\frac{\partial}{\partial \theta}Eg\right)'\Omega^{-1}\left(\frac{\partial}{\partial \theta}Eg\right)\right)^{-1}$$

$$= \left(\begin{bmatrix}3 & -1\end{bmatrix}\begin{bmatrix}4/7 & 2/7\\2/7 & 8/7\end{bmatrix}\begin{bmatrix}3\\-1\end{bmatrix}\right)^{-1}$$

$$= \left(\frac{4}{7}9 + \frac{2}{7}(-3) + \frac{2}{7}(-3) + \frac{8}{7}\right)^{-1}$$

$$= \frac{7}{32}$$

So

$$\sqrt{n}(\hat{\theta} - \theta_0) \rightarrow_d N(0, 7/32)$$

#### (ii) If W = I, then

$$((G'WG)^{-1})'G'W\Omega W'G(G'WG)^{-1} = ((G'G)^{-1})'G'\Omega G(G'G)^{-1}$$

$$= \frac{1}{100} \begin{bmatrix} 3 & -1 \end{bmatrix} \begin{bmatrix} 2 & -1/2 \\ -1/2 & 1 \end{bmatrix} \begin{bmatrix} 3 \\ -1 \end{bmatrix}$$

$$= \frac{1}{100} \left( 2(9) - \frac{1}{2}(-3) - \frac{1}{2}(-3) + 1 \right)$$

$$= \frac{11}{50}$$

So

$$\sqrt{n}(\hat{\theta} - \theta_0) \rightarrow_d N(0, 11/50)$$

### Problem 2

(a) The Poisson pdf of y|x is

$$e^{-e^{x'\theta}} \frac{(e^{x'\theta})^y}{y!}$$

So the likelihood is

$$\prod_{i=1}^{n} e^{-e^{x_i'\theta}} \frac{(e^{x_i'\theta})^{y_i}}{y_i!}$$

The log-likelihood is then

$$\sum_{i=1}^{n} -e^{x_i'\theta} + y_i x_i'\theta - \log(y_i!)$$

(b) (i) We have  $\hat{\theta}_{ML}$  satisfies

$$0 = \sum_{i=1}^{n} -x_{i,j} e^{x_i' \hat{\theta}_{ML}} + y_i x_{i,j}$$

$$0 = \sum_{i=1}^{n} x_i (y_i - e^{x_i' \hat{\theta}_{ML}})$$

Define  $g(y, x, \theta) = x(y - e^{x'\theta})$ . Then  $\hat{\theta}_{ML}$  is the GMM estimator associated with this Eg = 0 moment condition. We just need to show that  $\theta_0$  satisfies the moment condition and the SOC holds. First, clearly

$$Eg(y, x, \theta_0) = E(x(y - e^{x'\theta_0})) = E(E(x(y - e^{x'\theta_0})|x))$$

$$= E(xE(y|x) - xe^{x'\theta_0}) = E(xe^{x'\theta_0} - xe^{x'\theta_0}) = 0$$

So  $\theta_0$  satisfies the moment condition, and we know the GMM estimator is consistent so  $\hat{\theta}_{ML}$  is consistent as long as the SOC holds:

$$G = \frac{\partial}{\partial \theta} Eg = -E(xx'e^{x'\theta_0})$$

Then

$$-v' E(xx' e^{x'\theta}) v = -E(vxx' v' e^{x'\theta}) = -E(||vx||^2 e^{x'\theta}) < 0$$

as long as  $\exists x$  such that P(vx) > 0 for all  $v \neq 0$ .

(ii) Note that G is symmetric from the previous part. The GMM variance under identity weight matrix is

$$((G'G)^{-1})'G'\Omega G(G'G)^{-1}=G^{-1}\Omega G^{-1}$$

Assuming this is nondegenerate (G invertible), then since  $H(\hat{\theta})^{-1}$  is a consistent estimator of  $G^{-1}$ , we only need G invertible for  $H(\hat{\theta})^{-1}\Omega(\hat{\theta})H(\hat{\theta})^{-1}$  to be a consistent estimator of the asymptotic variance. However, for  $-H(\hat{\theta})^{-1}$  to be a consistent estimator of the asymptotic variance, we need

$$\Omega = -G^{-1}$$

which happens iff

$$E(x(y - e^{x'\theta})(x(y - e^{x'\theta}))'|x) = E((y - e^{x'\theta})^2 x x'|x) = E(xx'e^{x'\theta}|x)$$
$$E((y - e^{x'\theta})^2|x) = e^{x'\theta} \implies V(y|x) = e^{x'\theta}$$

Hence we also need  $V(y|x) = e^{x'\theta}$  for  $-H(\hat{\theta})^{-1}$  to be a consistent estimator of the variance.