The Bootstrap

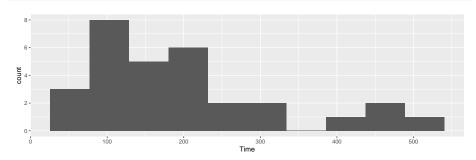
Packages for this section

```
library(tidyverse)
library(bootstrap)
```

ref

Is my sampling distribution normal enough?

Recall the IRS data that we used as a motivation for the sign test:



• We said that a t procedure for the mean would not be a good idea because the distribution is skewed

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What actually matters

- It's not the distribution of the data that has to be approx normal (for a t procedure).
- What matters is the sampling distribution of the sample mean.
- If the sample size is large enough, the sampling distribution will be normal enough even if the data distribution is not.
 - This is why we had to consider the sample size as well as the shape.
- But how do we know whether this is the case or not? We only have one sample.

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The (nonparametric) bootstrap

- Typically, our sample will be reasonably representative of the population.
- Idea: pretend the sample *is* the population, and sample from it *with* replacement.
- Calculate test statistic, and repeat many times.
- This gives an idea of how our statistic might vary in repeated samples: that is, its sampling distribution.
- Called the bootstrap distribution of the test statistic.
- If the bootstrap distribution is approx normal, infer that the true sampling distribution also approx normal, therefore inference about the mean such as t is good enough.
- If not, we should be more careful.

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Why it works

- We typically estimate population parameters by using the corresponding sample thing: eg. estimate population mean using sample mean.
- This called plug-in principle.
- The fraction of sample values less than a value x called the **empirical** distribution function (as a function of x).
- By plug-in principle, the empirical distribution function is an estimate of the population CDF.
- In this sense, the sample *is* an estimate of the population, and so sampling from it is an estimate of sampling from the population.

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Bootstrapping the IRS data

 Sampling with replacement is done like this (the default sample size is as long as the original data):

```
boot=sample(irs$Time, replace=T)
mean(boot)
```

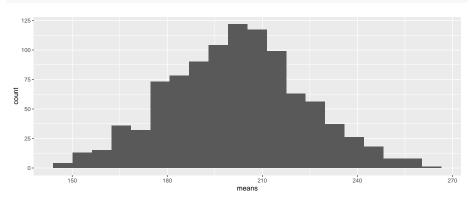
```
## [1] 217.3667
```

- That's one bootstrapped mean. We need a whole bunch.
- Use the same idea as for simulating power:

```
rerun(1000, sample(irs$Time, replace=T)) %>%
  map_dbl(~mean(.)) -> means
```

Sampling distribution of sample mean

ggplot(tibble(means), aes(x=means))+geom_histogram(bins=20)

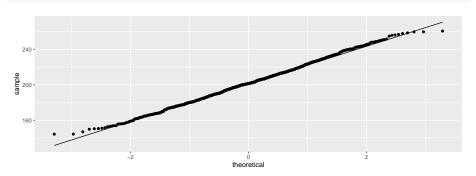


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Comments

This is not so bad: a long right tail, maybe:

```
ggplot(tibble(means), aes(sample=means))+
stat_qq()+stat_qq_line()
```



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Confidence interval from the bootstrap distribution

There are two ways (at least):

• percentile bootstrap interval: take the 2.5 and 97.5 percentiles (to get the middle 95%). This is easy, but not always the best:

```
(b_p=quantile(means, c(0.025, 0.975)))
```

```
## 2.5% 97.5%
## 159.5292 244.1075
```

 bootstrap t: use the SD of the bootstrapped sampling distribution as the SE of the estimator of the mean and make a t interval:

```
n=length(irs$Time)
t_star=qt(0.975, n-1)
(b_t=mean(means)+c(-1, 1)*t_star*sd(means))
```

```
## [1] 157.9020 245.4334
```

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Comparing

get ordinary t interval:

```
my_names=c("LCL", "UCL")
o_t=t.test(irs$Time)$conf.int
```

Compare the 2 bootstrap intervals with the ordinary t-interval:

limit	o_t	b_t	b_p
LCL	155.0081	157.9020	159.5292
UCL	247.4585	245.4334	244.1075

- ullet The bootstrap t and the ordinary t are very close
- The percentile bootstrap interval is noticeably shorter (common) and higher (skewness).

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Which to prefer?

- If the intervals agree, then they are all good.
- If they disagree, they are all bad!
- In that case, use BCA interval (over).

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Bias correction and acceleration

- this from "An introduction to the bootstrap", by Brad Efron and Robert J. Tibshirani.
- there is way of correcting the CI for skewness in the bootstrap distribution, called the BCa method
- complicated (see the Efron and Tibshirani book), but implemented in bootstrap package.

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Run this on the IRS data:

```
bca=bcanon(irs$Time, 1000, mean)
bca$confpoints
```

```
## alpha bca point
## [1,] 0.025 162.8000
## [2,] 0.050 166.8000
## [3,] 0.100 174.4667
## [4,] 0.160 179.5667
## [5,] 0.840 224.6000
## [6,] 0.900 234.3000
## [7,] 0.950 247.4333
## [8,] 0.975 256.3000
```

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use 2.5% and 97.5% points for CI

[1] 162.8 256.3

```
bca$confpoints %>% as_tibble() %>%
  filter(alpha %in% c(0.025, 0.975)) %>%
  pull(`bca point`) -> b_bca
b_bca
```

The Bootstrap

Comparing

tibble(limit=my_names, o_t, b_t, b_p, b_bca)

limit	o_t	b_t	b_p	b_bca
LCL	155.0081	157.9020	159.5292	162.8
UCL	247.4585	245.4334	244.1075	256.3

- The BCA interval says that the mean should be estimated even higher than the bootstrap percentile interval does.
- The BCA interval is the one to trust.

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Bootstrapping the correlation

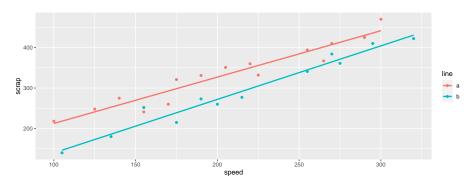
Recall the soap data:

```
url="http://www.utsc.utoronto.ca/~butler/c32/soap.txt"
soap=read delim(url," ")
##
   -- Column specification
## cols(
##
     case = col_double(),
##
     scrap = col_double(),
##
     speed = col_double(),
     line = col_character()
##
## )
```

The data

```
ggplot(soap, aes(x=speed, y=scrap, colour=line))+
  geom_point()+geom_smooth(method="lm", se=F)
```

`geom_smooth()` using formula 'y ~ x'



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Comments

- Line B produces less scrap for any given speed.
- For line B, estimate the correlation between speed and scrap (with a confidence interval.)

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Extract the line B data; standard correlation test

```
soap %>% filter(line=="b") -> line b
with(line_b, cor.test(speed, scrap))
##
##
   Pearson's product-moment correlation
##
## data: speed and scrap
## t = 15.829, df = 10, p-value = 2.083e-08
## alternative hypothesis: true correlation is not equal to 0
## 95 percent confidence interval:
## 0.9302445 0.9947166
## sample estimates:
##
         cor
```

0.9806224

Bootstrapping a correlation 1/2

- ullet This illustrates a different technique: we need to keep the x and y values together.
- Sample rows of the data frame rather than individual values of speed and scrap:

line_b %>% sample_frac(replace=T)

case	scrap	speed	line
25	422	320	b
25	422	320	b
25	422	320	b
20	215	175	b
24	252	155	b
20	215	175	b
24	252	155	b
22	260	200	b
25	422	320	b
26	273	190	b
19	341	255	b
27	410	295	b

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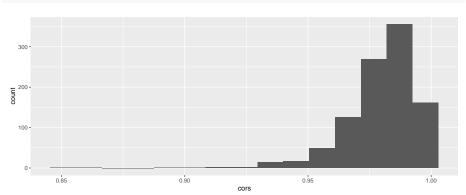
Bootstrapping a correlation 2/2

1000 times:

```
rerun(1000, sample_frac(line_b, replace=T)) %>%
  map_dbl(~with(.,cor(speed, scrap))) -> cors
```

A picture of this

ggplot(tibble(cors), aes(x=cors))+geom_histogram(bins=15)



Comments and next steps

- This is very left-skewed.
- Bootstrap percentile interval is:

```
(b_p=quantile(cors, c(0.025, 0.975)))
```

```
## 2.5% 97.5%
## 0.9432836 0.9963609
```

• We probably need the BCA interval instead.

Getting the BCA interval 1/2

 To use bcanon, write a function that takes a vector of row numbers and returns the correlation between speed and scrap for those rows:

```
theta=function(rows, d) {
  d %>% slice(rows) %>% with(., cor(speed, scrap))
}
theta(1:3, line_b)
```

```
## [1] 0.9928971
```

case	scrap	speed	line
16	140	105	b
17	277	215	b
18	384	270	b

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Getting the BCA interval 2/2

- Inputs to bcanon are now:
 - row numbers (1 through 12 in our case: 12 rows in line_b)
 - number of bootstrap samples
 - the function we just wrote
 - the data frame:

```
points=bcanon(1:12, 1000, theta, line_b)$confpoints
points %>% as_tibble() %>%
  filter(alpha %in% c(0.025, 0.975)) %>%
  pull(`bca point`) -> b_bca
b_bca
```

```
## [1] 0.9281762 0.9945797
```

Comparing the results

tibble(limit=my_names, o_c, b_p, b_bca)

limit	o_c	b_p	b_bca
LCL	0.9302445	0.9432836	0.9281762
UCL	0.9947166	0.9963609	0.9945797

- The bootstrap percentile interval doesn't go down far enough.
- The BCA interval seems to do a better job than the ordinary cor.test interval in capturing the skewness of the distribution.

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