

# Multivariate Analysis of Variance

## Multivariate analysis of variance

- Standard ANOVA has just one response variable.
- What if you have more than one response?
- Try an ANOVA on each response separately.
- But might miss some kinds of interesting dependence between the responses that distinguish the groups.

# Packages

```
library(car) # may need to install first  
library(tidyverse)  
library(MVTests) # also may need to install
```

## Small example

- Measure yield and seed weight of plants grown under 2 conditions: low and high amounts of fertilizer.
- Data (fertilizer, yield, seed weight):

```
url <- "http://ritsokiguess.site/datafiles/manova1.txt"
hilo <- read_delim(url, " ")
```

- 2 responses, yield and seed weight.

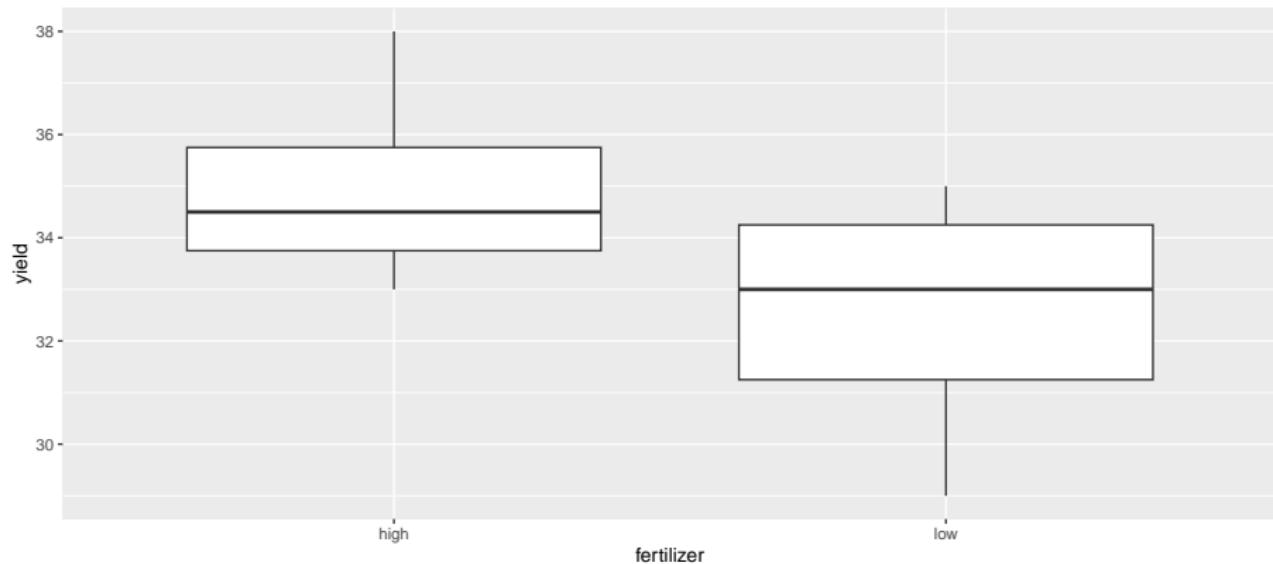
# The data

```
hilo
```

```
# A tibble: 8 x 3
  fertilizer yield weight
  <chr>      <dbl>   <dbl>
1 low         34     10
2 low         29     14
3 low         35     11
4 low         32     13
5 high        33     14
6 high        38     12
7 high        34     13
8 high        35     14
```

## Boxplot for yield for each fertilizer group

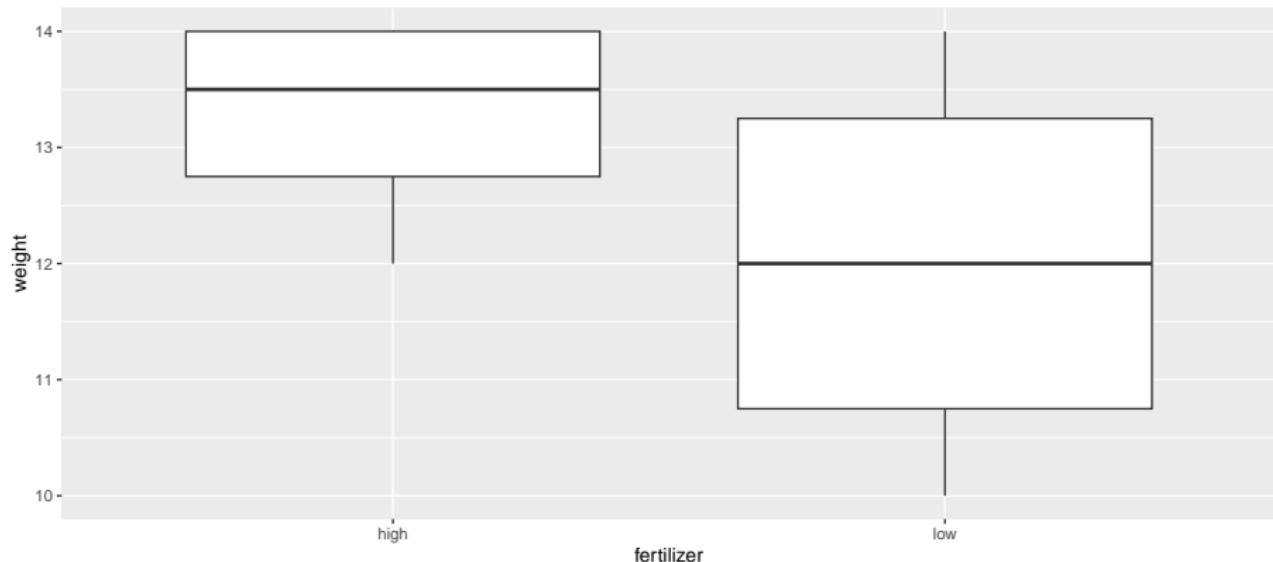
```
ggplot(hilo, aes(x = fertilizer, y = yield)) + geom_boxplot()
```



Yields overlap for fertilizer groups.

## Boxplot for weight for each fertilizer group

```
ggplot(hilo, aes(x = fertilizer, y = weight)) + geom_boxplot()
```



Weights overlap for fertilizer groups.

## ANOVAs for yield and weight

```
hilo.y <- aov(yield ~ fertilizer, data = hilo)
summary(hilo.y)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
fertilizer	1	12.5	12.500	2.143	0.194
Residuals	6	35.0	5.833		

```
hilo.w <- aov(weight ~ fertilizer, data = hilo)
summary(hilo.w)
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
fertilizer	1	3.125	3.125	1.471	0.271
Residuals	6	12.750	2.125		

Neither response depends significantly on fertilizer. But...

## Plotting both responses at once

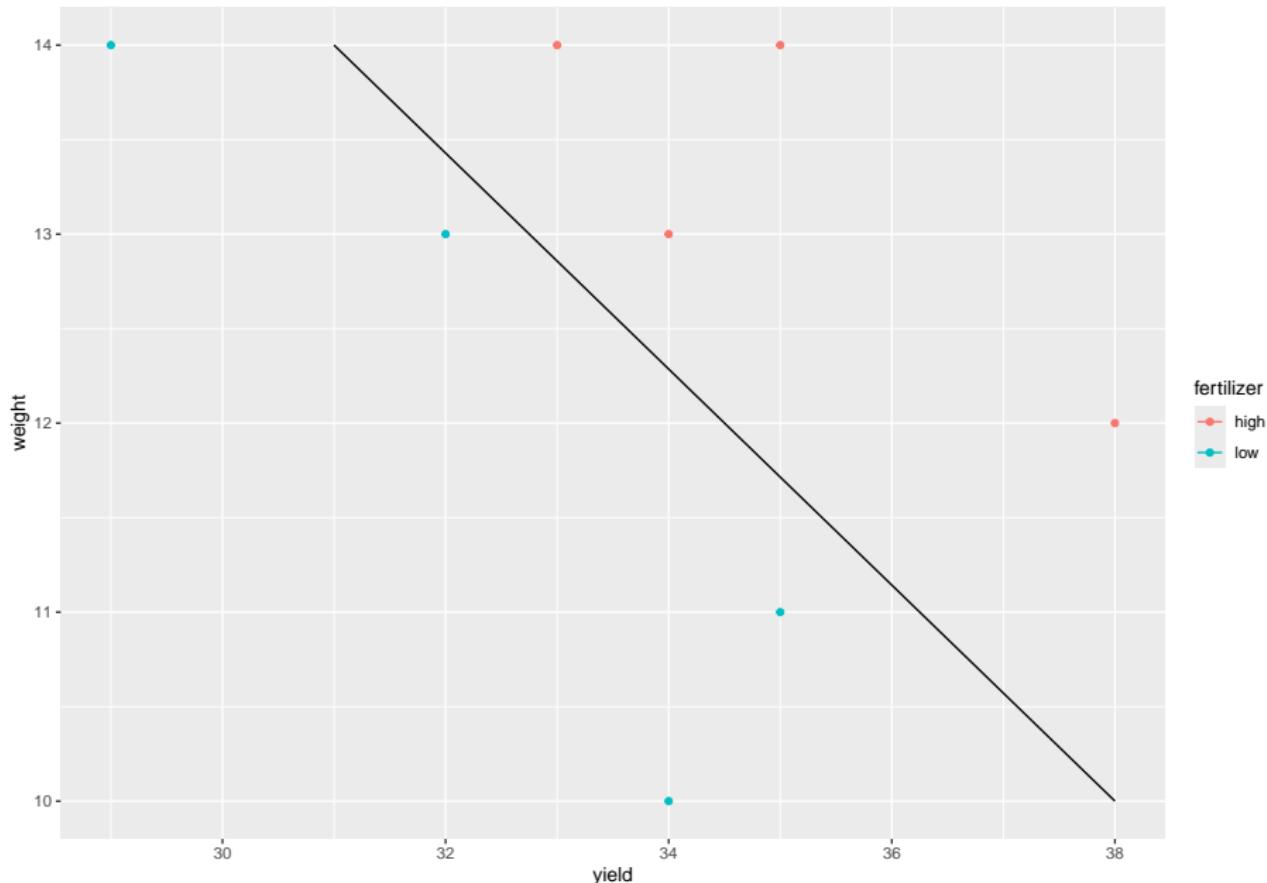
- Have two response variables (not more), so can plot the response variables against *each other*, labelling points by which fertilizer group they're from.
- First, create data frame with points (31, 14) and (38, 10) (why? Later):

```
d <- tribble(  
  ~line_x, ~line_y,  
  31, 14,  
  38, 10  
)
```

- Then plot data as points, and add line through points in d:

```
ggplot(hilo, aes(x = yield, y = weight,  
                  colour = fertilizer)) + geom_point() +  
  geom_line(data = d,  
            aes(x = line_x, y = line_y, colour = NULL)) -> g
```

# The plot



## Comments

- Graph construction:
  - ▶ Joining points in d by line.
  - ▶ geom\_line inherits colour from aes in ggplot.
  - ▶ Data frame d has no fertilizer (previous colour), so have to unset.
- Results:
  - ▶ High-fertilizer plants have both yield and weight high.
  - ▶ True even though no sig difference in yield or weight individually.
  - ▶ Drew line separating highs from lows on plot.

# MANOVA finds multivariate differences

- Is difference found by diagonal line significant? MANOVA finds out.

```
response <- with(hilo, cbind(yield, weight))
hilo.1 <- manova(response ~ fertilizer, data = hilo)
summary(hilo.1)
```

```
Df Pillai approx F num Df den Df Pr(>F)
fertilizer 1 0.80154   10.097      2       5 0.01755 *
Residuals  6
---
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

- Yes! Difference between groups is *diagonally*, not just up/down (weight) or left-right (yield). The *yield-weight combination* matters.

# Strategy

- Create new response variable by gluing together columns of responses, using `cbind`.
- Use `manova` with new response, looks like `lm` otherwise.
- With more than 2 responses, cannot draw graph. What then?
- If MANOVA test significant, cannot use Tukey. What then?
- Use *discriminant analysis* (of which more later).

# Another way to do MANOVA

using Manova from package car:

```
hilo.2.lm <- lm(response ~ fertilizer, data = hilo)
hilo.2 <- Manova(hilo.2.lm)
summary(hilo.2)
```

Type II MANOVA Tests:

Sum of squares and products for error:

	yield	weight
yield	35	-18.00
weight	-18	12.75

---

Term: fertilizer

Sum of squares and products for the hypothesis:

	yield	weight
yield	12.50	6.250
weight	6.25	3.125

Multivariate Tests: fertilizer

	Df	test	stat	approx F	num Df	den Df	Pr(>F)
Pillai	1	0.801542	10.09714	2	5	0.017546	*
Wilks	1	0.198458	10.09714	2	5	0.017546	*
Hotelling-Lawley	1	4.038855	10.09714	2	5	0.017546	*
Roy	1	4.038855	10.09714	2	5	0.017546	*

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

## Comments

- Same result as small-m manova.
- Manova will also do *repeated measures*, coming up later.

# Assumptions

- normality of each response variable within each treatment group
  - ▶ this is actually *multivariate* normality, with correlations
- equal spreads: each response variable has same variances and correlations (with other response variables) within each treatment group. Here:
  - ▶ yield has same spread for low and high fertilizer
  - ▶ weight has same spread for low and high fertilizer
  - ▶ correlation between yield and weight is same for low and high fertilizer
- test equal spread using Box's  $M$  test
  - ▶ a certain amount of unequalness is OK, so only a concern if P-value from  $M$ -test is very small (eg. less than 0.001).

## Assumptions for yield-weight data

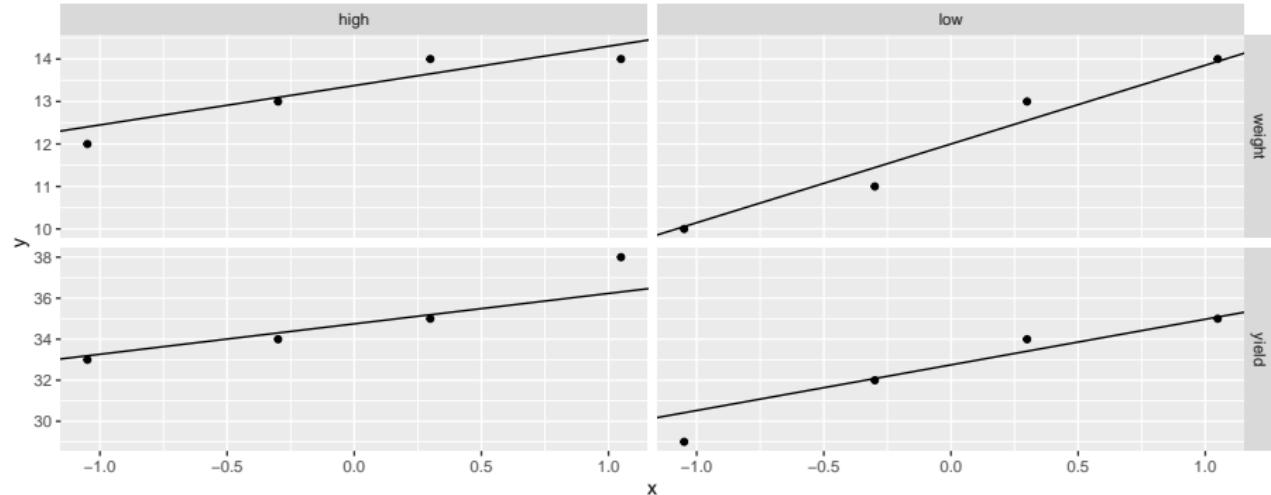
For normal quantile plots, need “extra-long” with all the data values in one column:

```
hilo %>%
  pivot_longer(-fertilizer, names_to = "xname",
              values_to = "xvalue") %>%
  ggplot(aes(sample = xvalue)) + stat_qq() +
  stat_qq_line() +
  facet_grid(xname ~ fertilizer, scales = "free") -> g
```

There are only four observations per response variable - treatment group combination, so graphs are not very informative (over):

# The plots

g



## Box M test

- Make sure package MVTests loaded first.
- inputs:
  - ▶ the response matrix (or, equivalently, the response-variable columns from your dataframe)
  - ▶ the column with the grouping variable in it (most easily gotten with \$).

```
library(MVTests)
# hilo %>% select(yield, weight) -> numeric_values
summary(BoxM(response, hilo$fertilizer))
```

### Box's M Test

Chi-Squared Value = 1.002964 , df = 3 and p-value: 0.801

- No problem at all with unequal spreads.

## Another example: peanuts

- Three different varieties of peanuts (mysteriously, 5, 6 and 8) planted in two different locations.
- Three response variables:  $y$ ,  $smk$  and  $w$ .

```
u <- "http://ritsokiguess.site/datafiles/peanuts.txt"
peanuts.orig <- read_delim(u, " ")
```

# The data

peanuts.orig

# A tibble: 12 x 6					
	obs	location	variety	y	smk
	<dbl>	<dbl>	<dbl>	<dbl>	<dbl>
1	1	1	5	195.	153.
2	2	1	5	194.	168.
3	3	2	5	190.	140.
4	4	2	5	180.	121.
5	5	1	6	203	157.
6	6	1	6	196.	166
7	7	2	6	203.	166.
8	8	2	6	198.	162.
9	9	1	8	194.	164.
10	10	1	8	187	165.
11	11	2	8	202.	167.
12	12	2	8	200	174.

## Setup for analysis

```
peanuts.orig %>%
  mutate(
    location = factor(location),
    variety = factor(variety)
  ) -> peanuts
peanuts
```

	obs	location	variety	y	smk	w
	<dbl>	<fct>	<fct>	<dbl>	<dbl>	<dbl>
1	1	1	5	195.	153.	51.4
2	2	1	5	194.	168.	53.7
3	3	2	5	190.	140.	55.5
4	4	2	5	180.	121.	44.4
5	5	1	6	203	157.	49.8
6	6	1	6	196.	166	45.8
7	7	2	6	203.	166.	60.4

## Analysis (using manova)

```
peanuts.1 <- manova(response ~ location * variety, data = peanuts)
summary(peanuts.1)
```

	Df	Pillai	approx F	num Df	den Df	Pr(>F)
location	1	0.89348	11.1843	3	4	0.020502
variety	2	1.70911	9.7924	6	10	0.001056
location:variety	2	1.29086	3.0339	6	10	0.058708
Residuals	6					

location \*

variety \*\*

location:variety .

Residuals

---

Signif. codes:

0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

## Comments

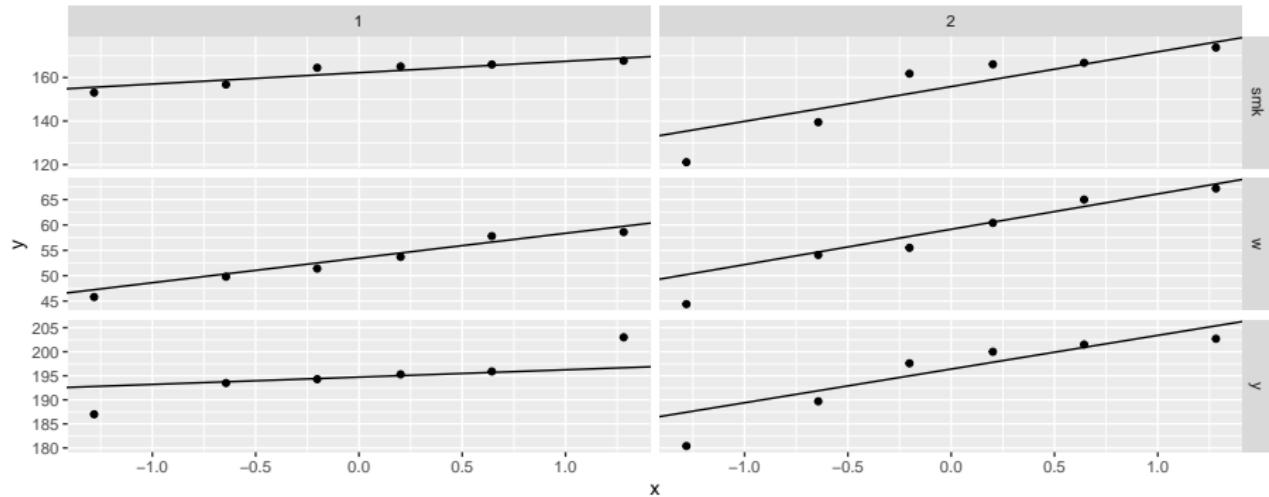
- Interaction not quite significant, but main effects are.
- Combined response variable ( $y$ ,  $smk$ ,  $w$ ) definitely depends on location and on variety
- Weak dependence of ( $y$ ,  $smk$ ,  $w$ ) on the location-variety *combination*.
- Understanding that dependence beyond our scope right now.

## Comments

- this time there are only six observations per location and four per variety, so normality is still difficult to be confident about
- $y$  at location 1 seems to be the worst for normality (long tails / outliers), and maybe  $y$  at location 2 is skewed left, but the others are not bad
- there is some evidence of unequal spread (slopes of lines), but is it bad enough to worry about? (Box M-test, over).

# Assessing normality

```
peanuts %>% pivot_longer(y:w, names_to = "yname",
                           values_to = "y") %>%
  ggplot(aes(sample = y)) + stat_qq() + stat_qq_line() +
  facet_grid(yname ~ location, scales = "free_y")
```



## Box's M tests

- One for location, one for variety:

```
summary(BoxM(response, peanuts$location))
```

Box's M Test

Chi-Squared Value = 12.47797 , df = 6 and p-value: 0.0521

```
summary(BoxM(response, peanuts$variety))
```

Box's M Test

Chi-Squared Value = 10.56304 , df = 12 and p-value: 0.567

- Neither of these P-values is low enough to worry about. (Remember, the P-value here has to be *really* small to indicate a problem.)
- Box's M test does not work well (and can fail to work at all) if the sample sizes are too small.

## Addendum: Box's M for interaction

- Create a combo column that is the combination of location and variety:

```
peanuts %>% mutate(combo =  
                      str_c(location, "-", variety)) -> d  
d
```

# A tibble: 12 x 7

	obs	location	variety	y	smk	w	combo
	<dbl>	<fct>	<fct>	<dbl>	<dbl>	<dbl>	<chr>
1	1	1	5	195.	153.	51.4	1-5
2	2	1	5	194.	168.	53.7	1-5
3	3	2	5	190.	140.	55.5	2-5
4	4	2	5	180.	121.	44.4	2-5
5	5	1	6	203	157.	49.8	1-6
6	6	1	6	196.	166	45.8	1-6
7	7	2	6	203.	166.	60.4	2-6
8	8	2	6	198.	162.	54.1	2-6

Then run Box's M test as usual:

```
summary(BoxM(response, d$combo))
```

Box's M Test

Chi-Squared Value = -Inf , df = 30 and p-value: 1

except that the result makes no sense. This is because there are only two observations per location-variety combination, which is not enough to estimate anything, and so the test no longer works.