## Clustering and Latent Variable Models

Mengye Ren

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K-means Clustering

## Unsupervised learning

Goal Discover interesting structure in the data.

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Formulation Density estimation:  $p(x;\theta)$  (often with *latent* variables).

### Unsupervised learning

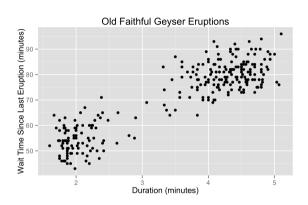
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Examples

- Discover *clusters*: cluster data into groups.
- Discover *factors*: project high-dimensional data to a small number of "meaningful" dimensions, i.e. dimensionality reduction.
- Discover *graph structures*: learn joint distribution of correlated variables, i.e. graphical models.

## Example: Old Faithful Geyser

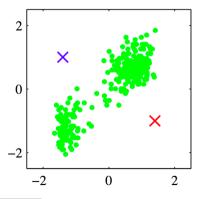


- Looks like two clusters.
- How to find these clusters algorithmically?

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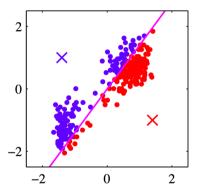
### k-Means: By Example

- Standardize the data.
- Choose two cluster centers.



From Bishop's Pattern recognition and machine learning, Figure 9.1(a).

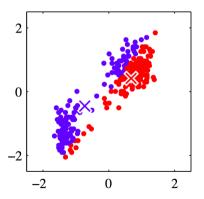
• Assign each point to closest center.



From Bishop's Pattern recognition and machine learning, Figure 9.1(b).

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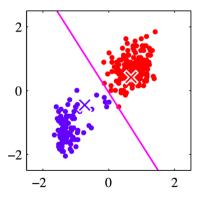
• Compute new cluster centers.



From Bishop's Pattern recognition and machine learning, Figure 9.1(c).

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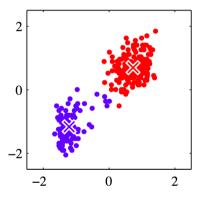
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From Bishop's Pattern recognition and machine learning, Figure 9.1(d).

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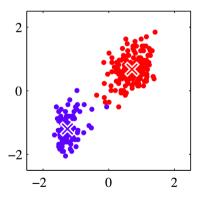
Compute cluster centers.



From Bishop's Pattern recognition and machine learning, Figure 9.1(e).

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• Iterate until convergence.



From Bishop's Pattern recognition and machine learning, Figure 9.1(i).

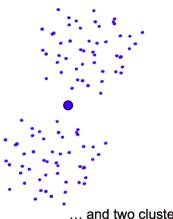
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## Suboptimal Local Minimum

• The clustering for k = 3 below is a local minimum, but suboptimal:



Would be better to have one cluster here



... and two clusters here

From Sontag's DS-GA 1003, 2014, Lecture 8.

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• The *k*-means objective is to minimize the distance between each example and its cluster centroid:

$$J(c, \mu) = \sum_{i=1}^{n} \|x_i - \mu_{c_i}\|^2.$$
 (2)

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k-means converges to a local minimum.

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Re-run with random initial centroids.

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    - Randomly choose next centroid with probability proportional to the computed distance squared.

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### Summary

#### We've seen

• Clustering—an unsupervised learning problem that aims to discover group assignments.

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Next, probabilistic model of clustering.

- A generative model of x.
- Maximum likelihood estimation.

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Gaussian Mixture Models

# Probabilistic Model for Clustering

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  - We have a probability distribution for each cluster.

# Probabilistic Model for Clustering

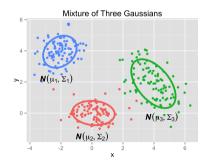
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#### Example:

- Choose  $z \in \{1, 2, 3\}$  with  $p(1) = p(2) = p(3) = \frac{1}{3}$ .
- 2 Choose  $x \mid z \sim \mathcal{N}(X \mid \mu_z, \Sigma_z)$ .



# Gaussian mixture model (GMM)

Generative story of GMM with k mixture components:

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#### Probability density of x:

• Sum over (marginalize) the latent variable z.

$$p(x) = \sum_{z} p(x, z) \tag{5}$$

$$=\sum_{z}p(x\mid z)p(z)\tag{6}$$

$$= \sum_{k} \pi_k \mathcal{N}(\mu_k, \Sigma_k) \tag{7}$$

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Suppose we have found parameters

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

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- Assuming all clusters are distinct, there are *k*! equivalent solutions.
- Not a problem per se, but something to be aware of.

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- MLE (also called maximize marginal likelihood).
- Log likelihood of data:

$$L(\theta) = \sum_{i=1}^{n} \log p(x_i; \theta)$$
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- Cannot push log into the sum... z and x are coupled.
- No closed-form solution for GMM—try to compute the gradient yourself!

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• What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

## Gradient Descent / SGD for GMM

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$$\hat{\Sigma}_{z} = \frac{1}{n_{z}} \sum_{i:z_{i}=z} (x_{i} - \hat{\mu}_{z}) (x_{i} - \hat{\mu}_{z})^{T}.$$
 empirical cluster covariance (13)

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### Learning GMMs: inference

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(16)

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- $p(z \mid x)$  is a soft assignment.
- If we know the parameters  $\mu$ ,  $\Sigma$ ,  $\pi$ , this would be easy to compute.

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  - E-step: fill in latent variables by inference.
    - compute soft assignments  $p(z | x_i)$  for all i.
  - **2** M-step: standard MLE for  $\mu$ ,  $\Sigma$ ,  $\pi$  given "observed" variables.
    - Equivalent to MLE in the observable case on data weighted by  $p(z \mid x_i)$ .

### M-step for GMM

• Let  $p(z \mid x)$  be the soft assignments:

$$\gamma_i^j = \frac{\pi_j^{\text{old}} \mathcal{N}\left(x_i \mid \mu_j^{\text{old}}, \Sigma_j^{\text{old}}\right)}{\sum_{c=1}^k \pi_c^{\text{old}} \mathcal{N}\left(x_i \mid \mu_c^{\text{old}}, \Sigma_c^{\text{old}}\right)}.$$

Exercise: show that

$$n_{z} = \sum_{i=1}^{n} \gamma_{i}^{z}$$

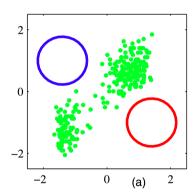
$$\mu_{z}^{\text{new}} = \frac{1}{n_{z}} \sum_{i=1}^{n} \gamma_{i}^{z} x_{i}$$

$$\Sigma_{z}^{\text{new}} = \frac{1}{n_{z}} \sum_{i=1}^{n} \gamma_{i}^{z} (x_{i} - \mu_{z}^{\text{new}}) (x_{i} - \mu_{z}^{\text{new}})^{T}$$

$$\pi_{z}^{\text{new}} = \frac{n_{z}}{n}.$$

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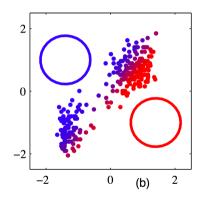
#### Initialization



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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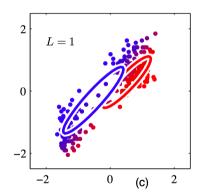
#### • First soft assignment:



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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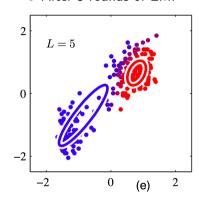
#### • First soft assignment:



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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#### • After 5 rounds of EM:

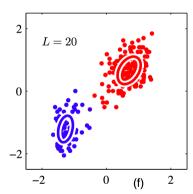


From Bishop's Pattern recognition and machine learning, Figure 9.8.

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## EM for GMM

#### • After 20 rounds of EM:



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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- EM is a general algorithm for learning latent variable models.
- Key idea: if data was fully observed, then MLE is easy.
  - E-step: fill in latent variables by computing  $p(z \mid x, \theta)$ .
  - M-step: standard MLE given fully observed data.
- Simpler and more efficient than gradient methods.
- Can prove that EM monotonically improves the likelihood and converges to a local minimum.
- k-means is a special case of EM for GMM with hard assignments, also called hard-EM.

Latent Variable Models

### General Latent Variable Model

- Two sets of random variables: z and x.
- z consists of unobserved hidden variables.
- x consists of **observed variables**.

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#### Definition

A latent variable model is a probability model for which certain variables are never observed.

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$$p(x, z \mid \theta)$$

#### **Definition**

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e.g. The Gaussian mixture model is a latent variable model.

## Complete and Incomplete Data

• Suppose we observe some data  $(x_1, ..., x_n)$ .

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- To simplify notation, take x to represent the entire dataset

$$x = (x_1, \ldots, x_n),$$

and z to represent the corresponding unobserved variables

$$z = (z_1, \ldots, z_n)$$
.

- An observation of x is called an **incomplete data set**.
- An observation (x, z) is called a **complete data set**.

## Our Objectives

• Learning problem: Given incomplete dataset x, find MLE

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# Our Objectives

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$$\hat{\theta} = \arg\max_{\theta} p(x \mid \theta).$$

• Inference problem: Given x, find conditional distribution over z:

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.

- For Gaussian mixture model, learning is hard, inference is easy.
- For more complicated models, inference can also be hard. (See DSGA-1005)

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Note that

$$\underset{\theta}{\arg\max} p(x \mid \theta) = \underset{\theta}{\arg\max} [\log p(x \mid \theta)].$$

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- We often call p(x,z) the **joint**. (for "joint distribution")
- Similarly,  $\log p(x)$  is the marginal log-likelihood.

# EM Algorithm

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Maximize the expected complete data log-likelihood:

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Maximize the **expected complete data log-likelihood**:

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EM assumption: the expected complete data log-likelihood is easy to optimize

Why should this work?

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Math Prerequisites

Jensen's Inequality

### Theorem (Jensen's Inequality)

If  $f : R \to R$  is a **convex** function, and x is a random variable, then

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Moreover, if f is **strictly convex**, then equality implies that  $x = \mathbb{E}x$  with probability 1 (i.e. x is a constant).

• e.g.  $f(x) = x^2$  is convex. So  $\mathbb{E}x^2 \geqslant (\mathbb{E}x)^2$ . Thus

$$\operatorname{Var}(x) = \mathbb{E}x^2 - (\mathbb{E}x)^2 \geqslant 0.$$

# Kullback-Leibler Divergence

- Let p(x) and q(x) be probability mass functions (PMFs) on  $\mathcal{X}$ .
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(Assumes q(x) = 0 implies p(x) = 0.)

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(Assumes 
$$q(x) = 0$$
 implies  $p(x) = 0$ .)

Can also write this as

$$\mathrm{KL}(p\|q) = \mathbb{E}_{x \sim p} \log \frac{p(x)}{q(x)}.$$

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Gibbs Inequality 
$$(KL(p||q) \ge 0 \text{ and } KL(p||p) = 0)$$

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Let p(x) and q(x) be PMFs on  $\mathfrak{X}$ . Then

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with equality iff p(x) = q(x) for all  $x \in \mathcal{X}$ .

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• KL divergence measures the "distance" between distributions.

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,

with equality iff p(x) = q(x) for all  $x \in \mathcal{X}$ .

- KL divergence measures the "distance" between distributions.
- Note:
  - KL divergence not a metric.
  - KL divergence is not symmetric.

$$\mathrm{KL}(p\|q) = \mathbb{E}_p\left[-\log\left(\frac{q(x)}{p(x)}\right)\right]$$

$$\begin{split} \mathrm{KL}(p\|q) &= \mathbb{E}_{p}\left[-\log\left(\frac{q(x)}{p(x)}\right)\right] \\ &\geqslant -\log\left[\mathbb{E}_{p}\left(\frac{q(x)}{p(x)}\right)\right] \end{aligned} \qquad \text{(Jensen's)}$$

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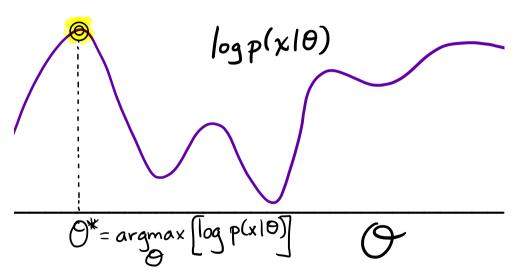
$$\begin{aligned} \mathrm{KL}(p\|q) &=& \mathbb{E}_{p}\left[-\log\left(\frac{q(x)}{p(x)}\right)\right] \\ &\geqslant &-\log\left[\mathbb{E}_{p}\left(\frac{q(x)}{p(x)}\right)\right] \quad \text{(Jensen's)} \\ &=& -\log\left[\sum_{\{x\mid p(x)>0\}}p(x)\frac{q(x)}{p(x)}\right] \\ &=& -\log\left[\sum_{x\in\mathcal{X}}q(x)\right] \\ &=& -\log 1=0. \end{aligned}$$

$$\begin{split} \mathrm{KL}(\rho \| q) &= \mathbb{E}_{\rho} \left[ -\log \left( \frac{q(x)}{\rho(x)} \right) \right] \\ &\geqslant -\log \left[ \mathbb{E}_{\rho} \left( \frac{q(x)}{\rho(x)} \right) \right] \quad \text{(Jensen's)} \\ &= -\log \left[ \sum_{\{x \mid \rho(x) > 0\}} \rho(x) \frac{q(x)}{\rho(x)} \right] \\ &= -\log \left[ \sum_{x \in \mathcal{X}} q(x) \right] \\ &= -\log 1 = 0. \end{split}$$

• Since  $-\log$  is strictly convex, we have strict equality iff q(x)/p(x) is a constant, which implies a=p .

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The ELBO: Family of Lower Bounds on  $\log p(x \mid \theta)$ 



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# Lower bound of the marginal log-likelihood

$$\log p(x;\theta) = \log \sum_{z \in \mathcal{Z}} p(x,z;\theta)$$

$$\log p(x;\theta) = \log \sum_{z \in \mathcal{Z}} p(x,z;\theta)$$
$$= \log \sum_{z \in \mathcal{Z}} q(z) \frac{p(x,z;\theta)}{q(z)}$$

$$\begin{aligned} \log p(x;\theta) &=& \log \sum_{z \in \mathcal{Z}} p(x,z;\theta) \\ &=& \log \sum_{z \in \mathcal{Z}} q(z) \frac{p(x,z;\theta)}{q(z)} \\ &\geqslant & \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(x,z;\theta)}{q(z)} \end{aligned}$$

# Lower bound of the marginal log-likelihood

$$\begin{array}{lcl} \log p(x;\theta) & = & \log \sum_{z \in \mathcal{Z}} p(x,z;\theta) \\ \\ & = & \log \sum_{z \in \mathcal{Z}} q(z) \frac{p(x,z;\theta)}{q(z)} \\ \\ & \geqslant & \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(x,z;\theta)}{q(z)} \\ \\ & \stackrel{\mathrm{def}}{=} & \mathcal{L}(q,\theta) \end{array}$$

- Evidence:  $\log p(x; \theta)$
- Evidence lower bound (ELBO):  $\mathcal{L}(q, \theta)$
- q: chosen to be a family of tractable distributions
- Idea: maximize the ELBO instead of  $log p(x; \theta)$

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• The MLE is defined as a maximum over  $\theta$ :

$$\hat{\theta}_{\mathsf{MLE}} = \operatorname*{arg\,max}_{\theta} \left[ \log p(x \mid \theta) \right].$$

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$$\log p(x \mid \theta) \geqslant \mathcal{L}(q, \theta).$$

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$$\log p(x \mid \theta) \geqslant \mathcal{L}(q, \theta).$$

• In EM algorithm, we maximize the lower bound (ELBO) over  $\theta$  and q:

$$\hat{\theta}_{\mathsf{EM}} \approx \operatorname*{arg\,max}_{\theta} \left[ \operatorname*{max}_{q} \mathcal{L}(q,\theta) \right]$$

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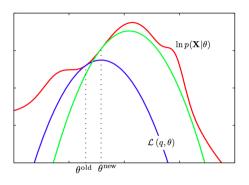
• In EM algorithm, q ranges over all distributions on z.

• Choose sequence of q's and  $\theta$ 's by "coordinate ascent" on  $\mathcal{L}(q,\theta)$ .

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- EM Algorithm (high level):
  - Choose initial  $\theta^{\text{old}}$ .

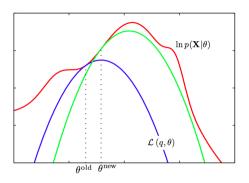
- Choose sequence of q's and  $\theta$ 's by "coordinate ascent" on  $\mathcal{L}(q,\theta)$ .
- EM Algorithm (high level):
  - Choose initial  $\theta^{\text{old}}$ .
  - 2 Let  $q^* = \arg\max_{q} \mathcal{L}(q, \theta^{\text{old}})$
  - **3** Let  $\theta^{\text{new}} = \arg\max_{\theta} \mathcal{L}(q^*, \theta)$ .

- Choose sequence of q's and  $\theta$ 's by "coordinate ascent" on  $\mathcal{L}(q,\theta)$ .
- EM Algorithm (high level):
  - Choose initial  $\theta^{\text{old}}$ .
  - 2 Let  $q^* = \arg\max_{q} \mathcal{L}(q, \theta^{\text{old}})$
  - 3 Let  $\theta^{\text{new}} = \arg\max_{\theta} \mathcal{L}(q^*, \theta)$ .
  - Go to step 2, until converged.
- Will show:  $p(x \mid \theta^{new}) \geqslant p(x \mid \theta^{old})$
- ullet Get sequence of  $\theta$ 's with monotonically increasing likelihood.



• Start at  $\theta^{\text{old}}$ .

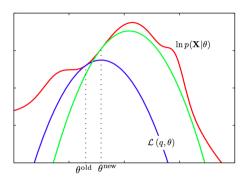
From Bishop's Pattern recognition and machine learning, Figure 9.14.



- Start at  $\theta^{\text{old}}$ .
- ② Find q giving best lower bound at  $\theta^{\text{old}} \Longrightarrow \mathcal{L}(q,\theta)$ .

From Bishop's Pattern recognition and machine learning, Figure 9.14.

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From Bishop's Pattern recognition and machine learning, Figure 9.14.

Is ELBO a "good" lowerbound?

$$\begin{split} \mathcal{L}(q,\theta) &= \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(x,z \mid \theta)}{q(z)} \\ &= \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(z \mid x,\theta) p(x \mid \theta)}{q(z)} \\ &= -\sum_{z \in \mathcal{Z}} q(z) \log \frac{q(z)}{p(z \mid x,\theta)} + \sum_{z \in \mathcal{Z}} q(z) \log p(x \mid \theta) \\ &= -\mathrm{KL}(q(z) \| p(z \mid x,\theta)) + \underbrace{\log p(x \mid \theta)}_{\mathbf{z} \in \mathcal{Z}} \end{split}$$

- KL divergence: measures "distance" between two distributions (not symmetric!)
- $KL(a||p) \ge 0$  with equality iff a(z) = p(z|x).
- ELBO = evidence KL ≤ evidence

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• Find q maximizing

$$\mathcal{L}(q,\theta) = -KL[q(z), p(z \mid x, \theta)] + \log p(x \mid \theta)$$

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$$\mathcal{L}(q^*, \theta) = -\underbrace{\mathrm{KL}[p(z \mid x, \theta), p(z \mid x, \theta)]}_{=0} + \log p(x \mid \theta)$$

• Summary:

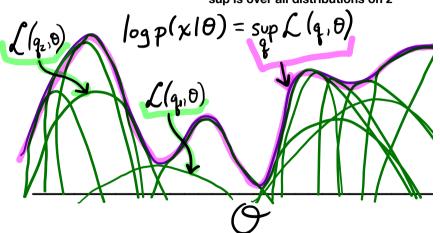
$$\log p(x \mid \theta) = \sup_{q} \mathcal{L}(q, \theta) \qquad \forall \theta$$

• For any  $\theta$ , sup is attained at  $q(z) = p(z \mid x, \theta)$ .

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#### sup is over all distributions on z



## Summary

Latent variable models: clustering, latent structure, missing lables etc.

Parameter estimation: maximum marginal log-likelihood

Challenge: directly maximize the evidence  $\log p(x; \theta)$  is hard

Solution: maximize the evidence lower bound:

$$\mathsf{ELBO} = \mathcal{L}(q, \theta) = -\mathsf{KL}(q(z) || p(z \mid x; \theta)) + \log p(x; \theta)$$

Why does it work?

$$q^*(z) = p(z \mid x; \theta) \quad \forall \theta \in \Theta$$
$$\mathcal{L}(q^*, \theta^*) = \max_{\theta} \log p(x; \theta)$$

Coordinate ascent on  $\mathcal{L}(q,\theta)$ 

- **1** Random initialization:  $\theta^{\text{old}} \leftarrow \theta_0$
- Repeat until convergence

Expectation (the E-step): 
$$q^*(z) = p(z \mid x; \theta^{\text{old}})$$
  
 $J(\theta) = \mathcal{L}(q^*, \theta)$ 

**Maximization** (the M-step):  $\theta^{\text{new}} \leftarrow \underset{\theta}{\text{arg max}} J(\theta)$ 

## EM Algorithm

#### Expectation Step

• Let  $q^*(z) = p(z \mid x, \theta^{\text{old}})$ . [ $q^*$  gives best lower bound at  $\theta^{\text{old}}$ ]

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- Let

$$J(\theta) := \mathcal{L}(q^*, \theta) = \underbrace{\sum_{z} q^*(z) \log \left( \frac{p(x, z \mid \theta)}{q^*(z)} \right)}_{\text{expectation w.r.t. } z \sim q^*(z)}$$

- Expectation Step
  - Let  $q^*(z) = p(z \mid x, \theta^{\text{old}})$ .  $[q^*]$  gives best lower bound at  $\theta^{\text{old}}$
  - Let

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$$\theta^{\mathsf{new}} = \argmax_{\theta} J(\theta).$$

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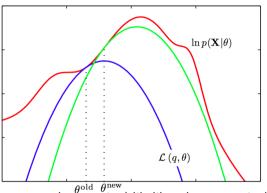
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[Equivalent to maximizing expected complete log-likelihood.]

EM puts no constraint on q in the E-step and assumes the M-step is easy. In general, both steps can be hard.

## Monotonically increasing likelihood



Exercise: prove that EM increases the marginal likelihood monotonically

$$\log p(x; \theta^{\mathsf{new}}) \geqslant \log p(x; \theta^{\mathsf{old}}) .$$

Does EM converge to a global maximum?

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Variations on EM

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#### EM and More General Variational Methods

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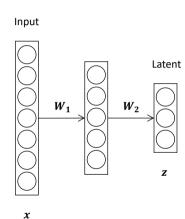
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- Lower bound now looser:

$$q^* = \underset{q \in \Omega}{\operatorname{arg\,min}\, \mathrm{KL}}[q(z), p(z \mid x, \theta^{\mathrm{old}})]$$

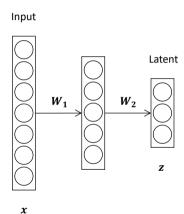
## Deep Latent Variable Models

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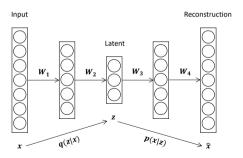
- Neural network is a flexible function class to represent transformation between random variables e.g., q(z).
- In neural networks, the hidden activations do not have probabilistic interpretation as they are not random variables.
- What if we let the hidden represent some learned latent code?

# Input Latent $W_1$ $W_2$ $\boldsymbol{z}$

 $\boldsymbol{x}$ 

# Variational Autoencoders (VAE) <sup>1</sup>

• An autoencoder (AE) is a neural network that reconstructs the same input.



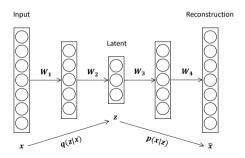
<sup>&</sup>lt;sup>1</sup>Diederik P Kingma, Max Welling. Auto-Encoding Variational Bayes. ICLR 2014.

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# Variational Autoencoders (VAE) 1

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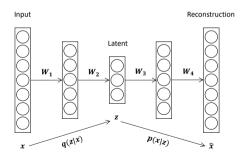
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- How to make q a probability distribution?



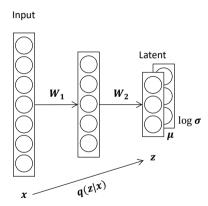
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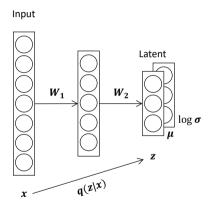
## Reparameterization Trick

• Let's assume that q(z|x) is a Gaussian distribution.



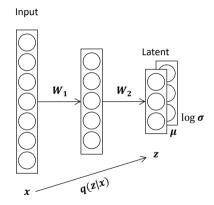
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- A stochastic z can be sampled from  $\mathcal{N}(\mu, \sigma^2)$ :  $z = \mu + \sigma \cdot \epsilon$ , where  $\epsilon \sim \mathcal{N}(0, 1)$ .



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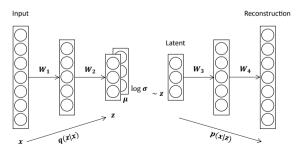
Divergence between q and the prior distribution Reconstruction based on z

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#### Stochastic Gradient

• The loss function needs to take expectation over q:

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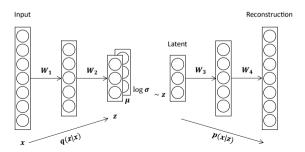
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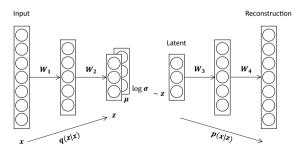
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- VAE: Introducing variational inference to neural networks. A classic starting example for deep generative modeling.

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Conclusion and Outlook

### Acknowledgement

- Most content developed by David Rosenberg (now at Bloomberg).
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- This is a very challenging grad-level course.
- Congrats, you are almost done.

### Next Lecture: Project Presentation

- Dec 12, in-person presentations.
- 24 groups, 120mins.
- Aim for 3 mins per group, hard stop at 4 mins, and 1 min max for Q&A.
- Send me your slides in PDF with your group number by Dec 11 11:59pm.

Linear Perceptron, conditional probability models, SVMs

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- Start from the task requirements, e.g. amount of data, computation resource
- The best lesson is to practice!

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  - Regularization: balance estimation error and generalization error.

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  - Empirical risk minimization, i.e. average loss on the training data.
  - Regularization: balance estimation error and generalization error.
- Bayesian approach: expectation over parameters.
  - Posterior: prior belief updated by observed data.
  - Bayes action minimizes the posterior risk.

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### Algorithms

Learning Find model parameters—often an optimization problem.

- (Stocahstic) (sub)gradient descent
- Functional gradient descent (gradient boosting)
- Convex vs non-convex objectives

Inference Answer questions given a learned model.

- Bayesian inference: compute various quantities given the posterior.
- Dynamic programming: compute arg max in structured prediction.

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- Classic ML sheds new insight into understand DL.
- Classic ML lays down foundation when we innovate DL algorithms.

- Computer Vision (Prof. Rob Fergus)
- Deep Learning (Prof. Yann LeCun)
- Deep Reinforcement Learning (Prof. Lerrel Pinto)
- Foundations of Deep Learning Theory (Prof. Matus Telgarsky)
- Inference and Representation (Prof. Joan Bruna)
- Learning with Large Language and Vision Models (Prof. Saining Xie)
- Mathematics of Deep Learning (Prof. Joan Bruna)
- Natural Language Processing (Prof. He He)