Clustering and Latent Variable Models

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(Slides credit to David Rosenberg, He He, et al.)

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K-means Clustering

Unsupervised learning

Goal Discover interesting structure in the data.

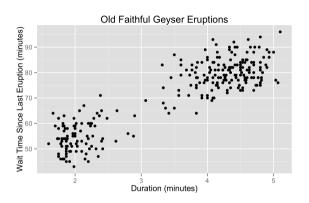
Formulation Density estimation: $p(x;\theta)$ (often with *latent* variables).

Examples

- Discover *clusters*: cluster data into groups.
- Discover *factors*: project high-dimensional data to a small number of "meaningful" dimensions, i.e. dimensionality reduction.
- Discover *graph structures*: learn joint distribution of correlated variables, i.e. graphical models.

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Example: Old Faithful Geyser

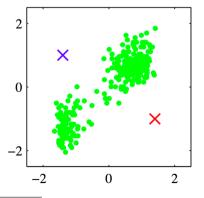


- Looks like two clusters.
- How to find these clusters algorithmically?

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k-Means: By Example

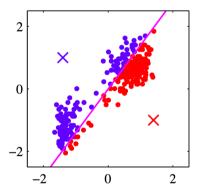
- Standardize the data.
- Choose two cluster centers.



From Bishop's Pattern recognition and machine learning, Figure 9.1(a).

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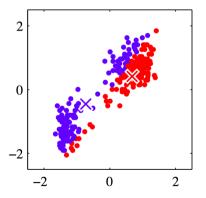
• Assign each point to closest center.



From Bishop's Pattern recognition and machine learning, Figure 9.1(b).

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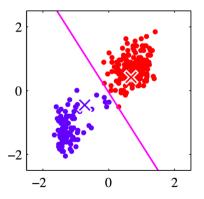
• Compute new cluster centers.



From Bishop's Pattern recognition and machine learning, Figure 9.1(c).

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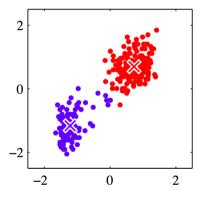
• Assign points to closest center.



From Bishop's Pattern recognition and machine learning, Figure 9.1(d).

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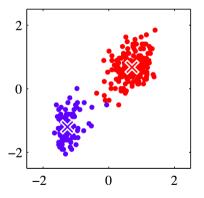
Compute cluster centers.



From Bishop's Pattern recognition and machine learning, Figure 9.1(e).

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• Iterate until convergence.

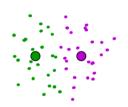


From Bishop's Pattern recognition and machine learning, Figure 9.1(i).

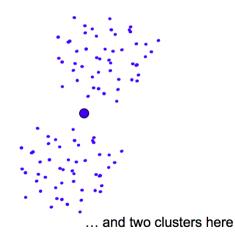
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Suboptimal Local Minimum

• The clustering for k = 3 below is a local minimum, but suboptimal:



Would be better to have one cluster here



From Sontag's DS-GA 1003, 2014, Lecture 8.

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Formalize k-Means

- Dataset $\mathcal{D} = \{x_1, \dots, x_n\} \subset \mathcal{X}$ where $\mathcal{X} = \mathbb{R}^d$.
- Goal: Partition data \mathcal{D} into k disjoint sets C_1, \ldots, C_k .
- Let $c_i \in \{1, ..., k\}$ be the cluster assignment of x_i .
- The **centroid** of C_i is defined to be

$$\mu_i = \underset{\mu \in \mathcal{X}}{\operatorname{arg\,min}} \sum_{x \in C_i} \|x - \mu\|^2.$$
 mean of C_i (1)

• The *k*-means objective is to minimize the distance between each example and its cluster centroid:

$$J(c, \mu) = \sum_{i=1}^{n} \|x_i - \mu_{c_i}\|^2.$$
 (2)

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k-Means: Algorithm

- **1** Initialize: Randomly choose initial centroids $\mu_1, \ldots, \mu_k \in \mathbb{R}^d$.
- 2 Repeat until convergence (i.e. c_i doesn't change anymore):
 - For all *i*, set

$$c_i \leftarrow \underset{j}{\operatorname{arg\,min}} \|x_i - \mu_j\|^2$$
. Minimize J w.r.t. c while fixing μ (3)

$$\mu_j \leftarrow \frac{1}{|C_j|} \sum_{x \in C} x.$$
 Minimze J w.r.t. μ while fixing c . (4)

• Recall the objective: $J(c, \mu) = \sum_{i=1}^{n} ||x_i - \mu_{c_i}||^2$.

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Avoid bad local minima

k-means converges to a local minimum.

• *J* is non-convex, thus no guarantee to converging to the global minimum.

Avoid getting stuck with bad local minima:

- Re-run with random initial centroids.
- k-means++: choose initial centroids that spread over all data points.
 - Randomly choose the first centroid from the data points \mathfrak{D} .
 - Sequentially choose subsequent centroids from points that are farther away from current centroids:
 - Compute distance between each x_i and the closest already chosen centroids.
 - Randomly choose next centroid with probability proportional to the computed distance squared.

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Summary

We've seen

- Clustering—an unsupervised learning problem that aims to discover group assignments.
- k-means:
 - Algorithm: alternating between assigning points to clusters and computing cluster centroids.
 - Objective: minmizing some loss function by coordinate descent.
 - Converge to a local minimum.

Next, probabilistic model of clustering.

- A generative model of x.
- Maximum likelihood estimation.

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Gaussian Mixture Models

Gaussian mixture model (GMM)

Generative story of GMM with k mixture components:

- Choose cluster $z \sim \text{Categorical}(\pi_1, \ldots, \pi_k)$.
- **2** Choose $x \mid z \sim \mathcal{N}(\mu_z, \Sigma_z)$.

Probability density of x:

• Sum over (marginalize) the latent variable z.

$$p(x) = \sum_{z} p(x, z) \tag{5}$$

$$=\sum_{z}p(x\mid z)p(z)\tag{6}$$

$$=\sum_{k}\pi_{k}\mathcal{N}(\mu_{k},\Sigma_{k})\tag{7}$$

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Identifiability Issues for GMM

• Suppose we have found parameters

Cluster probabilities: $\pi = (\pi_1, \dots, \pi_k)$

Cluster means : $\mu = (\mu_1, ..., \mu_k)$

Cluster covariance matrices: $\Sigma = (\Sigma_1, \dots \Sigma_k)$

that are at a local minimum.

- What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.
- We'll get the same likelihood. How many such equivalent settings are there?
- Assuming all clusters are distinct, there are k! equivalent solutions.

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Learning GMMs

How to learn the parameters π_k, μ_k, Σ_k ?

- MLE (also called maximize marginal likelihood).
- Log likelihood of data:

$$L(\theta) = \sum_{i=1}^{n} \log p(x_i; \theta)$$
 (8)

$$=\sum_{i=1}^{n}\log\sum_{z}p(x,z;\theta)$$
(9)

- Cannot push log into the sum... z and x are coupled.
- No closed-form solution for GMM—try to compute the gradient yourself!

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Gradient Descent / SGD for GMM

What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done, in principle but need to be clever about it.
- For example, each covariance matrix $\Sigma_1, \ldots, \Sigma_k$ has to be positive semidefinite.
- How to maintain that constraint?
 - Rewrite $\Sigma_i = M_i M_i^T$, where M_i is an unconstrained matrix.
 - Then Σ_i is positive semidefinite.

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Learning GMMs: observable case

Suppose we observe cluster assignments z. Then MLE is easy:

$$n_z = \sum_{i=1}^n \mathbb{1}[z_i = z]$$
 # examples in each cluster (10)

$$\hat{\pi}(z) = \frac{n_z}{n}$$
 fraction of examples in each cluster (11)

$$\hat{\mu}_z = \frac{1}{n_z} \sum_{i: z_i = z} x_i$$
 empirical cluster mean (12)

$$\hat{\Sigma}_{z} = \frac{1}{n_{z}} \sum_{i:z_{i}=z} (x_{i} - \hat{\mu}_{z}) (x_{i} - \hat{\mu}_{z})^{T}. \qquad \text{empirical cluster covariance}$$
 (13)

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Learning GMMs: inference

The inference problem: observe x, want to know z.

$$p(z = j \mid x_i) = p(x, z = j)/p(x)$$
 (14)

$$= \frac{p(x \mid z = j)p(z = j)}{\sum_{k} p(x \mid z = k)p(z = k)}$$
(15)

$$= \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_k \pi_k \mathcal{N}(x_i \mid \mu_k, \Sigma_k)}$$
(16)

- $p(z \mid x)$ is a soft assignment.
- If we know the parameters μ, Σ, π , this would be easy to compute.

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Let's compute the cluster assignments and the parameters iteratively.

The expectation-minimization (EM) algorithm:

- **1** Initialize parameters μ , Σ , π randomly.
- 2 Run until convergence:
 - E-step: fill in latent variables by inference.
 - compute soft assignments $p(z | x_i)$ for all i.
 - **2** M-step: standard MLE for μ , Σ , π given "observed" variables.
 - Equivalent to MLE in the observable case on data weighted by $p(z \mid x_i)$.

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M-step for GMM

• Let $p(z \mid x)$ be the soft assignments:

$$\gamma_i^j = \frac{\pi_j^{\text{old}} \mathcal{N}\left(x_i \mid \mu_j^{\text{old}}, \Sigma_j^{\text{old}}\right)}{\sum_{c=1}^k \pi_c^{\text{old}} \mathcal{N}\left(x_i \mid \mu_c^{\text{old}}, \Sigma_c^{\text{old}}\right)}.$$

Exercise: show that

$$n_{z} = \sum_{i=1}^{n} \gamma_{i}^{z}$$

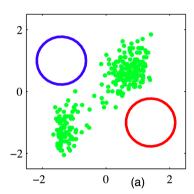
$$\mu_{z}^{\text{new}} = \frac{1}{n_{z}} \sum_{i=1}^{n} \gamma_{i}^{z} x_{i}$$

$$\Sigma_{z}^{\text{new}} = \frac{1}{n_{z}} \sum_{i=1}^{n} \gamma_{i}^{z} (x_{i} - \mu_{z}^{\text{new}}) (x_{i} - \mu_{z}^{\text{new}})^{T}$$

$$\pi_{z}^{\text{new}} = \frac{n_{z}}{n}.$$

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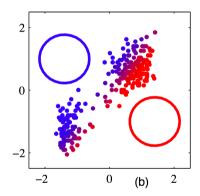
Initialization



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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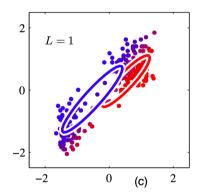
• First soft assignment:



From Bishop's $Pattern\ recognition\ and\ machine\ learning,\ Figure\ 9.8.$

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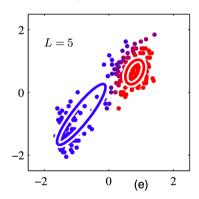
• First soft assignment:



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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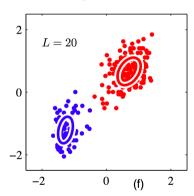
• After 5 rounds of EM:



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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• After 20 rounds of EM:



From Bishop's Pattern recognition and machine learning, Figure 9.8.

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EM for GMM: Summary

- EM is a general algorithm for learning latent variable models.
- Key idea: if data was fully observed, then MLE is easy.
 - E-step: fill in latent variables by computing $p(z \mid x, \theta)$.
 - M-step: standard MLE given fully observed data.
- Simpler and more efficient than gradient methods.
- Can prove that EM monotonically improves the likelihood and converges to a local minimum.
- k-means is a special case of EM for GMM with hard assignments, also called hard-EM.

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Latent Variable Models

General Latent Variable Model

- Two sets of random variables: z and x
- z consists of unobserved hidden variables.
- x consists of observed variables.
- Joint probability model parameterized by $\theta \in \Theta$:

$$p(x,z \mid \theta)$$

Definition

A latent variable model is a probability model for which certain variables are never observed.

e.g. The Gaussian mixture model is a latent variable model.

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Complete and Incomplete Data

- Suppose we observe some data (x_1, \ldots, x_n) .
- To simplify notation, take x to represent the entire dataset

$$x = (x_1, \ldots, x_n),$$

and z to represent the corresponding unobserved variables

$$z = (z_1, \ldots, z_n)$$
.

- An observation of x is called an **incomplete data set**.
- An observation (x, z) is called a **complete data set**.

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Our Objectives

• Learning problem: Given incomplete dataset x, find MLE

$$\hat{\theta} = \underset{\theta}{\operatorname{arg\,max}} p(x \mid \theta).$$

• Inference problem: Given x, find conditional distribution over z:

$$p(z | x, \theta)$$
.

- For Gaussian mixture model, learning is hard, inference is easy.
- For more complicated models, inference can also be hard.

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Log-Likelihood and Terminology

Note that

$$\underset{\theta}{\arg\max} p(x \mid \theta) = \underset{\theta}{\arg\max} [\log p(x \mid \theta)].$$

- Often easier to work with this "log-likelihood".
- We often call p(x) the marginal likelihood.
 - because it is p(x,z) with z "marginalized out":

$$p(x) = \sum_{z} p(x, z)$$

- We often call p(x,z) the **joint**. (for "joint distribution")
- Similarly, $\log p(x)$ is the marginal log-likelihood.

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EM Algorithm

Intuition

Problem: marginal log-likelihood $\log p(x;\theta)$ is hard to optimize (observing only x)

Observation: complete data log-likelihood $\log p(x,z;\theta)$ is easy to optimize (observing both x and z)

Idea: guess a distribution of the latent variables q(z) (soft assignments)

Maximize the **expected complete data log-likelihood**:

$$\max_{\theta} \sum_{z \in \mathcal{Z}} q(z) \log p(x, z; \theta)$$

EM assumption: the expected complete data log-likelihood is easy to optimize

Why should this work?

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Math Prerequisites

Jensen's Inequality

Theorem (Jensen's Inequality)

If $f: R \to R$ is a **convex** function, and x is a random variable, then

$$\mathbb{E}f(x) \geqslant f(\mathbb{E}x).$$

Moreover, if f is strictly convex, then equality implies that $x = \mathbb{E}x$ with probability 1 (i.e. x is a constant).

• e.g. $f(x) = x^2$ is convex. So $\mathbb{E}x^2 \ge (\mathbb{E}x)^2$. Thus

$$\operatorname{Var}(x) = \mathbb{E}x^2 - (\mathbb{E}x)^2 \geqslant 0.$$

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Kullback-Leibler Divergence

- Let p(x) and q(x) be probability mass functions (PMFs) on X.
- How can we measure how "different" p and q are?
- The Kullback-Leibler or "KL" Divergence is defined by

$$\mathrm{KL}(p\|q) = \sum_{x \in \mathcal{X}} p(x) \log \frac{p(x)}{q(x)}.$$

(Assumes
$$q(x) = 0$$
 implies $p(x) = 0$.)

Can also write this as

$$\mathrm{KL}(p\|q) = \mathbb{E}_{x\sim p}\log\frac{p(x)}{q(x)}.$$

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Gibbs Inequality
$$(KL(p||q) \ge 0 \text{ and } KL(p||p) = 0)$$

Theorem (Gibbs Inequality)

Let p(x) and q(x) be PMFs on \mathfrak{X} . Then

$$KL(p||q)\geqslant 0$$
,

with equality iff p(x) = q(x) for all $x \in \mathcal{X}$.

- KL divergence measures the "distance" between distributions.
- Note:
 - KL divergence not a metric.
 - KL divergence is **not symmetric**.

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Gibbs Inequality: Proof

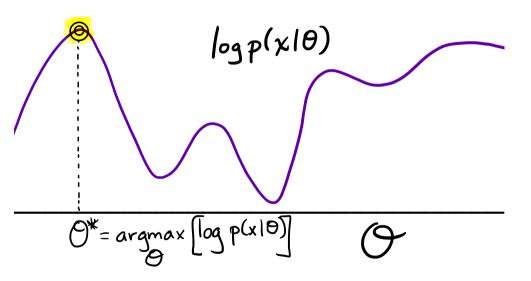
$$\begin{aligned} \mathrm{KL}(\rho \| q) &= \mathbb{E}_{p} \left[-\log \left(\frac{q(x)}{p(x)} \right) \right] \\ &\geqslant -\log \left[\mathbb{E}_{p} \left(\frac{q(x)}{p(x)} \right) \right] \quad \text{(Jensen's)} \\ &= -\log \left[\sum_{\{x \mid p(x) > 0\}} p(x) \frac{q(x)}{p(x)} \right] \\ &= -\log \left[\sum_{x \in \mathcal{X}} q(x) \right] \\ &= -\log 1 = 0. \end{aligned}$$

• Since $-\log$ is strictly convex, we have strict equality iff q(x)/p(x) is a constant, which implies q=p.

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The ELBO: Family of Lower Bounds on $\log p(x \mid \theta)$

The Maximum Likelihood Estimator



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Lower bound of the marginal log-likelihood

$$\begin{array}{lcl} \log p(x;\theta) & = & \log \sum_{z \in \mathcal{Z}} p(x,z;\theta) \\ \\ & = & \log \sum_{z \in \mathcal{Z}} q(z) \frac{p(x,z;\theta)}{q(z)} \\ \\ & \geqslant & \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(x,z;\theta)}{q(z)} \\ \\ & \stackrel{\mathrm{def}}{=} & \mathcal{L}(q,\theta) \end{array}$$

- Evidence: $\log p(x;\theta)$
- Evidence lower bound (ELBO): $\mathcal{L}(q, \theta)$
- q: chosen to be a family of tractable distributions
- Idea: maximize the ELBO instead of $log p(x; \theta)$

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MLE, EM, and the ELBO

• The MLE is defined as a maximum over θ :

$$\hat{\theta}_{\mathsf{MLE}} = \operatorname*{arg\,max}_{\theta} [\log p(x \mid \theta)] \, .$$

• For any PMF q(z), we have a lower bound on the marginal log-likelihood

$$\log p(x \mid \theta) \geqslant \mathcal{L}(q, \theta).$$

• In EM algorithm, we maximize the lower bound (ELBO) over θ and q:

$$\hat{\theta}_{\mathsf{EM}} pprox rg \max_{\theta} \left[\max_{q} \mathcal{L}(q, \theta)
ight]$$

• In EM algorithm, q ranges over all distributions on z.

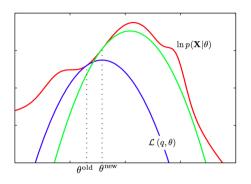
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EM: Coordinate Ascent on Lower Bound

- Choose sequence of q's and θ 's by "coordinate ascent" on $\mathcal{L}(q,\theta)$.
- EM Algorithm (high level):
 - Choose initial θ^{old} .
 - 2 Let $q^* = \arg\max_{q} \mathcal{L}(q, \theta^{\text{old}})$
 - 3 Let $\theta^{\text{new}} = \arg\max_{\theta} \mathcal{L}(q^*, \theta)$.
 - Go to step 2, until converged.
- Will show: $p(x \mid \theta^{new}) \ge p(x \mid \theta^{old})$
- ullet Get sequence of θ 's with monotonically increasing likelihood.

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EM: Coordinate Ascent on Lower Bound



- Start at θ^{old} .
- ② Find q giving best lower bound at $\theta^{\text{old}} \Longrightarrow \mathcal{L}(q,\theta)$.

From Bishop's Pattern recognition and machine learning, Figure 9.14.

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Is ELBO a "good" lowerbound?

$$\mathcal{L}(q,\theta) = \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(x,z \mid \theta)}{q(z)}$$

$$= \sum_{z \in \mathcal{Z}} q(z) \log \frac{p(z \mid x,\theta)p(x \mid \theta)}{q(z)}$$

$$= -\sum_{z \in \mathcal{Z}} q(z) \log \frac{q(z)}{p(z \mid x,\theta)} + \sum_{z \in \mathcal{Z}} q(z) \log p(x \mid \theta)$$

$$= -KL(q(z) || p(z \mid x,\theta)) + \underbrace{\log p(x \mid \theta)}_{\mathbf{z} \in \mathcal{Z}}$$

- KL divergence: measures "distance" between two distributions (not symmetric!)
- $KL(a||p) \ge 0$ with equality iff a(z) = p(z|x).
- ELBO = evidence KL ≤ evidence

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Maximizing over q for fixed θ .

Find q maximizing

$$\mathcal{L}(q,\theta) = -\text{KL}[q(z), p(z \mid x, \theta)] + \underbrace{\log p(x \mid \theta)}_{\text{no } q \text{ here}}$$

- Recall $KL(p||q) \ge 0$, and KL(p||p) = 0.
- Best q is $q^*(z) = p(z \mid x, \theta)$ and

$$\mathcal{L}(q^*, \theta) = -\underbrace{\mathrm{KL}[p(z \mid x, \theta), p(z \mid x, \theta)]}_{=0} + \log p(x \mid \theta)$$

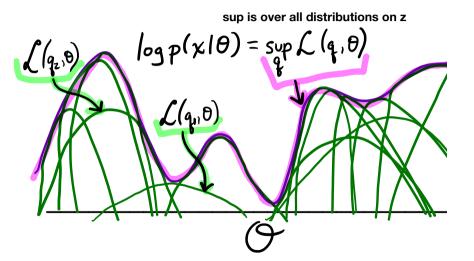
Summary:

$$\log p(x \mid \theta) = \sup_{q} \mathcal{L}(q, \theta) \qquad \forall \theta$$

• For any θ , sup is attained at $q(z) = p(z \mid x, \theta)$.

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Marginal Log-Likelihood **IS** the Supremum over Lower Bounds



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Summary

Latent variable models: clustering, latent structure, missing lables etc.

Parameter estimation: maximum marginal log-likelihood

Challenge: directly maximize the evidence $\log p(x; \theta)$ is hard

Solution: maximize the evidence lower bound:

$$\mathsf{ELBO} = \mathcal{L}(q, \theta) = -\mathsf{KL}(q(z) || p(z \mid x; \theta)) + \log p(x; \theta)$$

Why does it work?

$$q^*(z) = p(z \mid x; \theta) \quad \forall \theta \in \Theta$$
$$\mathcal{L}(q^*, \theta^*) = \max_{\theta} \log p(x; \theta)$$

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EM algorithm

Coordinate ascent on $\mathcal{L}(q, \theta)$

- Random initialization: $\theta^{\text{old}} \leftarrow \theta_0$
- Repeat until convergence

Expectation (the E-step):
$$q^*(z) = p(z \mid x; \theta^{\text{old}})$$

 $J(\theta) = \mathcal{L}(q^*, \theta)$

Maximization (the M-step): $\theta^{\text{new}} \leftarrow \underset{\theta}{\text{arg max}} J(\theta)$

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EM Algorithm

- Expectation Step
 - Let $q^*(z) = p(z \mid x, \theta^{\text{old}})$. $[q^*]$ gives best lower bound at θ^{old}
 - let

$$J(\theta) := \mathcal{L}(q^*, \theta) = \underbrace{\sum_{z} q^*(z) \log \left(\frac{p(x, z \mid \theta)}{q^*(z)} \right)}_{\text{expectation w.r.t. } z \sim q^*(z)}$$

Maximization Step

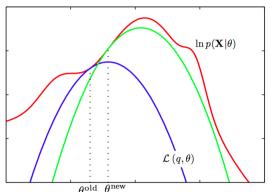
$$\theta^{\mathsf{new}} = \underset{\theta}{\mathsf{arg}} \max_{\theta} J(\theta).$$

[Equivalent to maximizing expected complete log-likelihood.]

EM puts no constraint on q in the E-step and assumes the M-step is easy. In general, both steps can be hard.

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Monotonically increasing likelihood



Exercise: prove that EM increases the marginal likelihood monotonically

$$\log p(x; \theta^{\mathsf{new}}) \geqslant \log p(x; \theta^{\mathsf{old}}) .$$

Does EM converge to a global maximum?

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Variations on EM

FM Gives Us Two New Problems

• The "E" Step: Computing

$$J(\theta) := \mathcal{L}(q^*, \theta) = \sum_{z} q^*(z) \log \left(\frac{p(x, z \mid \theta)}{q^*(z)} \right)$$

• The "M" Step: Computing

$$\theta^{\mathsf{new}} = \underset{\theta}{\mathsf{arg}} \max_{\boldsymbol{\theta}} J(\boldsymbol{\theta}).$$

• Either of these can be too hard to do in practice.

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Generalized EM (GEM)

- Addresses the problem of a difficult "M" step.
- Rather than finding

$$\theta^{\text{new}} = \underset{\theta}{\text{arg max}} J(\theta),$$

find any θ^{new} for which

$$J(\theta^{\text{new}}) > J(\theta^{\text{old}}).$$

- Can use a standard nonlinear optimization strategy
 - \bullet e.g. take a gradient step on J.
- We still get monotonically increasing likelihood.

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EM and More General Variational Methods

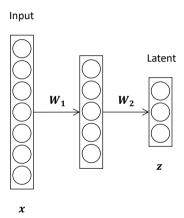
- Suppose "E" step is difficult:
 - Hard to take expectation w.r.t. $q^*(z) = p(z \mid x, \theta^{\text{old}})$.
- Solution: Restrict to distributions Q that are easy to work with.
- Lower bound now looser:

$$q^* = \underset{q \in \Omega}{\operatorname{arg\,min}\, \mathrm{KL}[q(z), p(z \mid x, \theta^{\mathsf{old}})]}$$

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Deep Latent Variable Models

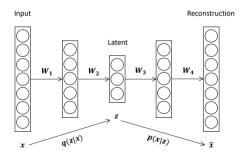
- Neural network is a flexible function class to represent transformation between random variables e.g., q(z).
- In neural networks, the hidden activations do not have probabilistic interpretation as they are not random variables.
- What if we let the hidden represent some learned latent code?



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Variational Autoencoders (VAE) 1

- An autoencoder (AE) is a neural network that reconstructs the same input.
- The first half is an encoder, from input to latent. The second half is a decoder.
- How to make q a probability distribution?

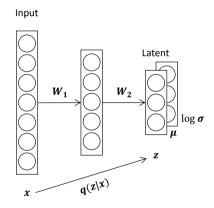


¹Diederik P Kingma, Max Welling. Auto-Encoding Variational Bayes. ICLR 2014.

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Reparameterization Trick

- Let's assume that q(z|x) is a Gaussian distribution.
- Instead of letting the neural network to output a stochastic variable, we can let it predict deterministically the distribution parameters μ and σ .
- A stochastic z can be sampled from $\mathcal{N}(\mu, \sigma^2)$: $z = \mu + \sigma \cdot \epsilon$, where $\epsilon \sim \mathcal{N}(0, 1)$.



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Variational Lower Bound

- Encoder q weights: φ ; Decoder p weights: θ .
- Now maximize ELBO:

$$L(q; \phi, \theta) = \sum_{z} q(z) \log \frac{p_{\theta}(x, z)}{q_{\phi}(z|x)}$$

$$= \mathbb{E}_{z \sim q} [-\log q_{\phi}(z|x) + \log p_{\theta}(x, z)]$$

$$= \mathbb{E}_{z \sim q} [-\log q_{\phi}(z|x) + \log p_{\theta}(x|z) + \log p_{\theta}(z)]$$

$$(17)$$

$$= \mathbb{E}_{z \sim q} [-\log q_{\phi}(z|x) + \log p_{\theta}(x|z) + \log p_{\theta}(z)]$$

$$(19)$$

 $= \underbrace{-KL(q_{\Phi}(z|x)||p_{\theta}(z))}_{+ \underbrace{\mathbb{E}_{z \sim q}(\log p_{\theta}(x|z))}_{}}$ (20)

Divergence between q and the prior distribution Reconstruction based on z

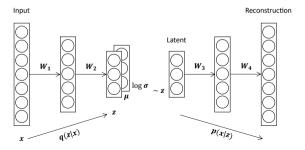
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Stochastic Gradient

• The loss function needs to take expectation over q:

$$L(q; \phi, \theta) = -KL(q_{\phi}(z|x)||p_{\theta}(z)) + \mathbb{E}_{z \sim q}(\log p_{\theta}(x|z))$$

- Turns out we just need to have a Monte Carlo sample size of 1:
 - For each x, sample one z from g(z|x).
- Backprop through reparameterization.



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Learned Manifold

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Today's Summary

- Motivation: Unsupervised learning
- K-means: A simple algorithm for discovering clusters
- Making k-means probabilistic: Gaussian mixture models
- More generally: Latent variable models
- Learning of latent variable models: EM
- Underlying principle: Maximizing ELBO
- VAE: Introducing variational inference to neural networks. A classic starting example for deep generative modeling.

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Conclusion and Outlook

Acknowledgement

- Most content developed by David Rosenberg (now at Bloomberg).
- Later adapted by He He, Tal Linzen, and others.
- This is a very challenging grad-level course.
- Congrats, you are almost done.

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Next Lecture: Project Presentation

- Dec 10, in-person presentations.
- 22 groups, 120mins.
- Aim for 3 mins per group, hard stop at 4 mins, and 1 min max for Q&A.
- Send your slides in PDF with your group number by Dec 9 11:59pm (via Google form).

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Models

Linear Perceptron, conditional probability models, SVMs Non-linear Kernelized models, trees, basis function models, neural nets

How to choose the model family?

- Trade-offs:
 - approximation error and estimation error (bias and variance),
 - accuracy and efficiency (during both training and inference).
- Start from the task requirements, e.g. amount of data, computation resource
- The best lesson is to practice!

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Objectives

Loss functions How far off a prediction is from the target, e.g. 0-1 loss, margin-based loss, squared loss.

Risk Expected loss - but expectation over what?

- Frequentist approach: expectation over data.
 - Empirical risk minimization, i.e. average loss on the training data.
 - Regularization: balance estimation error and generalization error.
- Bayesian approach: expectation over parameters.
 - Posterior: prior belief updated by observed data.
 - Bayes action minimizes the posterior risk.

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Algorithms

Learning Find model parameters—often an optimization problem.

- (Stocahstic) (sub)gradient descent
- Functional gradient descent (gradient boosting)
- Convex vs non-convex objectives

Inference Answer questions given a learned model.

- Bayesian inference: compute various quantities given the posterior.
- Dynamic programming: compute arg max in structured prediction.

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Do We Still Need ML?

- Deep Learning (DL) has been overwhelmingly popular in the past few years.
- Many ML methods are considered out-dated.
- However, DL is not necessarily good for all types of data (data availability, data quality, data modality etc.). Classic methods may also have their sweet spots.
- Classic ML sheds new insight into understand DL.
- Classic ML lays down foundation when we innovate in DL algorithms.

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Other ML Related Advanced Courses in CS/DS

- Bayesian Machine Learning(Andrew Wilson)
- Computer Vision (Saining Xie)
- Deep Learning (Yann LeCun)
- Deep Reinforcement Learning (Lerrel Pinto)
- Enbodied Learning and Vision (Mengye Ren)
- Foundations of Deep Learning Theory (Matus Telgarsky)
- Inference and Representation (Joan Bruna)
- Learning with Large Language and Vision Models (Saining Xie)
- Mathematics of Deep Learning (Joan Bruna)
- Natural Language Processing (He He)

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