

Quantitative relationship between mutated amino-acid sequence of human copper-transporting ATPases and their related diseases

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Received: 29 January 2008 / Accepted: 19 July 2008 / Published online: 8 August 2008
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Abstract Copper-transporting ATPase 1 and 2 (ATP7A and ATP7B) are two highly homologous P-type copper ATPase exporters. Mutations in ATP7A can lead to Menkes disease which is an X-linked disorder of copper deficiency. Mutations in ATP7B can cause Wilson disease which is an autosomal recessive disorder of copper toxicity. In this study, we attempt to build a quantitative relationship between mutated ATPase and Menkes/Wilson disease. First, we use the amino-acid distribution probability as a measure to quantify the difference in ATPase before and after mutation. Second, we use the cross-impact analysis to define the quantitative relationship between mutant ATPase protein and Menkes/Wilson disease, and compute various probabilities. Finally, we use the Bayesian equation to determine the probability that Menkes/Wilson disease is diagnosed under a mutation. The results show (i) the vast majority of mutations lead to the amino-acid distribution probability increase in mutant ATP7As and decrease in ATP7Bs, and (ii) the probability that a mutation causes Menkes/Wilson disease is about nine tenth. Thus we provide a way to use the descriptively probabilistic method to couple the mutation with its clinical outcome after quantifying mutations in proteins.

Keywords Amino acid · ATP7A · ATP7B · Bayes' law · Computational · Cross-impact analysis · Distribution probability · Menkes disease · Prediction · Wilson disease

Introduction

With worldwide efforts on correlating each disease to genetic level, it is very meaningful to build a quantitative relationship between a mutated protein and its phenotype outcome. For this purpose, we must quantify a protein sequence in order to use a numeric sequence to represent the protein sequence no matter what mechanism we use for quantification. Currently, there are many ways to quantify a protein sequence such as physicochemical property, electrostatic property, steric property, hydrophobic property, hydrogen bond property, etc. [1–10]. These quantifications were developed by physicists and chemists for their own purpose, but not for the study of mutations. When applying these quantifications for mutation studies, it is possible that these quantifications will not change before and after a mutation. For example, pI (isoelectric point) would not change before and after mutation. Furthermore, these borrowed quantification descriptors have their own special dimensions (units), which are not associated with mutations. Although a relationship between structure and mutation is generally modeled using a regression, this system is associated to a phenomenological model rather than to a kinetic/dynamic model. Thus, these borrowed quantifications could be more suitable for studying static state.

In this study, we apply our own quantifications which are subject to mutation and have no problems with dimensions. Since 1999, we have developed three approaches to quantify each amino acid in a protein and a whole protein (for review, see [11–13]). Our three quantifications are not only sensitive to mutations and dimensionless, but also sensitive to the length of protein, composition of amino acids, amino acid positions and neighboring amino acids.

We have developed a kinetic model to predict mutations in proteins from influenza A virus [14–23]. With our approach

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we can look at it from four aspects: (i) prediction of mutation position, (ii) prediction of would-be-mutated amino acids at the predicted positions, (iii) timing of mutation [24,25], and (iv) prediction of new function led by mutation. To accomplish the last prediction, the first step is to develop a quantitative model to couple the changed protein sequence with its outcome, which at this stage would be descriptive. In this study, we attempt to use the amino-acid distribution probability as a measure to build a quantitative relationship between mutated amino-acid sequence of human copper-transporting ATPases (ATP7A and ATP7B) and their related diseases (Menkes and Wilson diseases).

As a critical catalyst, copper is an essential transition metal required for the activity of multiple mammalian enzymes, including cytochrome c oxidase, superoxide dismutase and lysyl oxidase [26,27]. However, copper ions are efficient generators of free radicals, which accounts for the toxicity of copper when homeostatic mechanisms are disrupted under certain conditions [26]. The Menkes and Wilson's disease proteins are two highly homologous P-type copper ATPase exporters [28,29]. They supply copper to the secreted cuproenzyme and ceruloplasmin, which has a role in iron metabolism [30–34].

The Menkes disease gene ATP7A was the first mammalian heavy metal ion transporter cDNA to be cloned, which locates at Xql3.3 and encodes the copper-transporting ATPase 1 consisting of 1,500 amino acids (ATP7A) [35–37]. This gene is expressed in most tissues, except in the liver. Mutations in the gene cause Menkes disease [38–41], which is an X-linked disorder of copper transport leading to Cu^{2+} accumulation predominantly in the intestine and kidney causing damage to certain tissues, neurodegeneration, and death in early childhood [42].

The Wilson disease gene ATP7B was localized to the q14.3 band of chromosome 13 and cloned in 1993 [43–46]. This gene encodes the copper transporting ATPase 2 with 1,465 amino acids (ATP7B) expressed mainly in the liver [47], where it is responsible for biliary excretion of copper [48]. Thus, ATP7B is the copper pump regulating the rate of biliary copper excretion when copper levels start to rise in liver [49–51]. Mutations in the ATP7B gene result in Wilson disease [52–55], which is an autosomal recessive copper toxicosis described firstly by Kinnear Wilson in 1912 [56]. This disease is characterized by a massive accumulation of copper in the liver, with subsequent deposition of copper in other tissues, such as the central nervous system [57–60]. Patients with Wilson disease manifest a spectrum of liver pathologies ranging from hepatitis and cirrhosis to liver failure [61]. About half of patients with Wilson disease develop central nervous system toxicity as the initial clinical manifestation of the disease [62].

Approximately 200 mutations have been identified in unrelated Menkes patients [63], and over 150 in Wilson

patients [64–70]. All types of mutations have been found the copper-transporting ATPases. For their outcome, splice site, nonsense, duplication and deletion mutations would be predicted to truncate the mRNA or the protein and severely affect function [71].

Materials and methods

Data

Both human copper-transporting ATPases with their documented mutations were obtained from UniProtKB/Swiss-Prot entry: ATP7A with its 37 mutations (accession number Q04656; December 4, 2007; Entry version 99, <http://expasy.org/uniprot/Q04656>) and ATP7B with its 165 mutations (accession number P35670; November 13, 2007; Entry version 100, <http://expasy.org/uniprot/P35670>).

Amino-acid distribution probability

Among three approaches developed by us, the amino-acid distribution probability is mainly related to the positions of amino acids along the protein, which is suitable for mutation analysis, and we have used this approach in a number of our previous studies [11–23,25,72–81]. The quantification was developed along such a thought, for example, there are two methionines (M) among 141 amino acids in human hemoglobin α -chain [72]. With regard to their random distribution, our intuition might suggest that there would be one M in the first half of the chain and another M in the second half, which is true in real-life case. In fact, there are only three possible distributions of Ms in human hemoglobin α -chain, i.e. (i) both Ms are in the first half, (ii) one M is in each half and (iii) both Ms are in the second half. Thus each distribution of Ms has the probability of 1/3. If we do not distinguish either the first half or second half but are simply interested in whether both Ms are in both halves or in any half, we have the probability of 1/2 for each distribution.

If we are interested in the distribution probability of three amino acids in a protein, we naturally imagine to grouping the protein into three partitions, and our intuition may suggest that each partition contains an amino acid. If we do not distinguish the first, second and third partition, actually there are totally three types of distributions, i.e. (i) each amino acid is in each partition, (ii) two amino acids are in a partition and an amino acid in another partition, and (iii) three amino acids are in a partition.

At this moment, the distribution probability can be computed according to the statistical mechanics, which classifies the distribution of elementary particles in energy states according to three assumptions of whether distinguishing each particle and energy state, i.e. Maxwell-Boltzmann,

Fermi-Dirac and Bose-Einstein assumptions [82]. We actually used the Maxwell-Boltzmann assumption for computing amino-acid distribution probability, which is [82]

$$\frac{r!}{q_0! \times q_1! \times \dots \times q_n!} \times \frac{r!}{r_1! \times r_2! \times \dots \times r_n!} \times n^{-r},$$

where r is the number of amino acids, n is the number of partitions, r_n is the number of amino acids in the n th partition, q_n is the number of partitions with the same number of amino acids, and $!$ is the factorial function.

Thus, the distribution probabilities are different for these three types of distributions of three amino acids, say, 0.2222 for (i), 0.6667 for (ii) and 0.1111 for (iii). Clearly the protein can only adopt one type of distribution for these three amino acids, which is the actual distribution probability.

For four amino acids, we have five distribution probabilities, i.e. (i) each partition contains an amino acid, (ii) a partition contains two amino acids and two partitions contain an amino acid each, (iii) two partitions contain two amino acids each, (iv) a partition contains an amino acid and a partition contains three amino acids, and (v) a partition contains four amino acids. Their distribution probabilities are 0.0938 for (i), 0.5625 for (ii), 0.1406 for (iii), 0.1875 for (iv), and 0.0156 for (v). Furthermore, there are seven distributions for five amino acids, 11 distributions for 6 amino acids, 15 distributions for 7 amino acids, and so on.

Quantification of normal copper-transporting ATPases

For the human ATP7A and ATP7B before mutation, we used the equation in the above subsection to calculate the amino-acid distribution probability for amino acids listed in Table 1. We showed the simplest example for computation of amino-acid distribution probability in Appendix.

In this way, we quantified each amino acid in normal human ATP7A and 7B although the calculation became more complicated with respect to other types of amino acids in Table 1 because the q , r , n and factorial function increased. Thereafter, we assigned the calculated distribution probability to each amino acid along the sequence in order to have the visualized concept (Fig. 1).

Quantification of mutated copper-transporting ATPases

The amino-acid distribution probability calculated in the above subsection serves as a reference because the mutation does not occur. Obviously any mutation changes the amino-acid composition of copper-transporting ATPases, which certainly changes the distribution pattern of both original and mutated amino acids, thus the amino-acid distribution probability differs for both original and mutated amino acids.

For example, the mutation of the conserved histidine residue is one of the most common mutations found in Wilson

Table 1 Amino acids, their compositions and distribution probability in normal human copper-transporting ATPases

Amino acid	ATP7A		ATP7B	
	Number	Distribution probability	Number	Distribution probability
A	106	0.004930	134	0.000279
R	48	0.000930	53	0.018279
N	67	0.000099	49	0.022866
D	66	0.003838	62	0.003048
C	26	0.040315	29	0.000677
E	92	0.003302	80	0.002652
Q	54	0.000865	69	0.003584
G	92	0.003350	104	0.000924
H	35	0.000143	36	0.013806
I	129	0.000027	106	0.002020
L	135	0.000670	132	0.000038
K	88	0.001071	75	0.004623
M	44	0.000110	44	0.008612
F	43	0.001010	36	0.000407
P	62	0.000964	72	0.004417
S	133	0.000000	126	0.000032
T	100	0.000341	82	0.000460
W	9	0.177028	11	0.040398
Y	30	0.043096	24	0.006342
V	141	0.002731	141	0.003840

A, alanine; R, arginine; N, asparagine; D, aspartic acid; C, cysteine; E, glutamic acid; Q, glutamine; G, glycine; H, histidine; I, isoleucine; L, leucine; K, lysine; M, methionine; F, phenylalanine; P, proline; S, serine; T, threonine; W, tryptophan; Y, tyrosine; V, valine

disease [52,83], which actually changes histidine (H) to glutamine (Q) at position 1069 leading to the Wilson disease. There are 69 Qs and 36 Hs before mutation, but there are 70 Qs and 35 Hs after mutation. For the sake of simplicity, we used Hs as an example in Appendix.

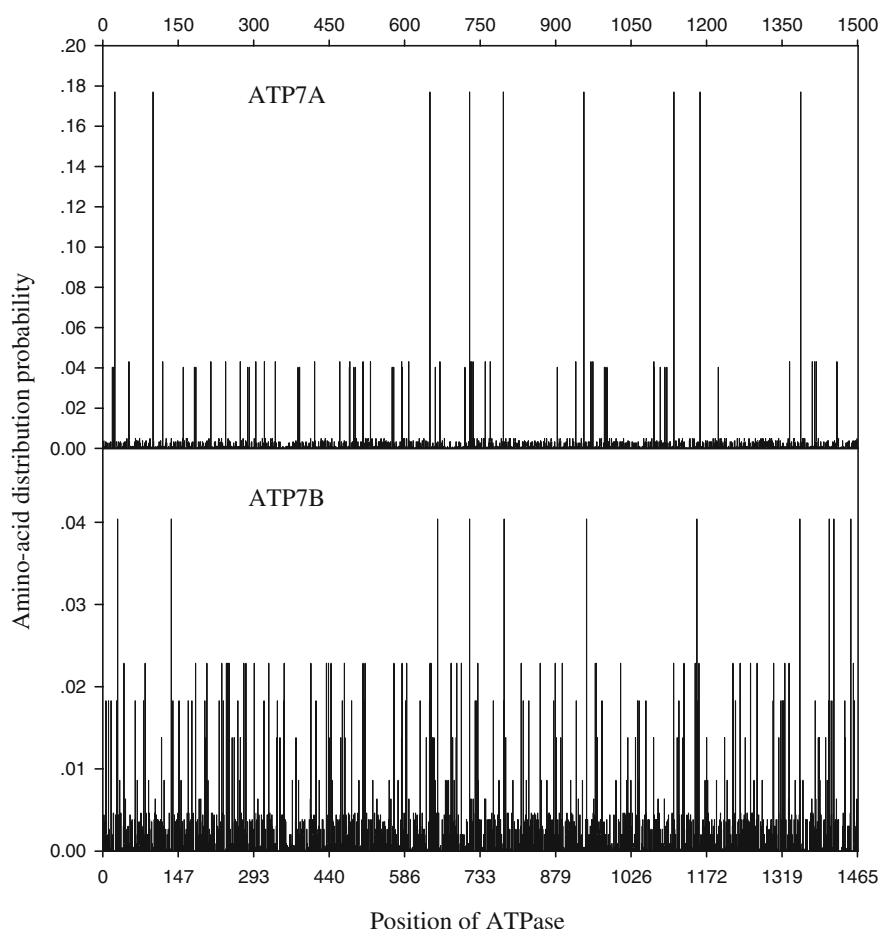
In this manner, we found that the distribution probabilities were 0.0138 and 0.0233 before and after mutation, which means that this mutation led the distribution probability increased in the original amino acid H. On the other hand, the distribution probability was different in the mutated amino acid Q before and after mutation, which were 0.0036 and 0.0111. Therefore the overall effect for this mutation on ATP7B was $(0.0233 - 0.0138) + (0.0111 - 0.0036) = 0.017$, that is, the mutation increased the distribution probability of amino acids.

Hence, we had the quantitative measure for the changed amino-acid distribution probability of mutated copper-transporting ATPases and we also had their related diseases in documentation, thus we could build a quantitative relationship between mutated protein sequence and corresponding clinical outcome.

Probabilistic relationship

For building of quantitative relationship between mutation and clinical outcome, we used the descriptively probabilistic

Fig. 1 Visualization of amino-acid distribution probability along ATP7A and ATP7B sequence



method, because our measure was amino-acid distribution probability and each individual mutation related to its clinical outcome was presented as frequency. Therefore, we used the cross-impact analysis to couple them, because the amino-acid distribution probability either increased or decreased after mutation, which was a 2-possibility event, and the clinical outcome either occurred or did not occur after mutation, which was a yes-and-no event. Thereafter, we used the Bayesian equation to calculate the probability of occurrence of clinical outcome under mutation.

Statistical validation

The descriptive statistics or probability produces the general trends or estimates in a population, such as population weight, in our case the probability of occurrence of clinical outcome under mutation. Therefore, the validation is directly related to the increase in sampling size of population of interest, which could lead the population estimates to move from the current values, such as survey on different years. Therefore we will closely monitor the change in documented mutation and clinical outcome to re-determine

the probability, which is the way to validate the descriptive analysis.

Results and discussion

In principle, the descriptive method was the first step for the development modeling, we therefore firstly dealt with the 37 ATP7A mutations, among which 33 mutations were documented as the Menkes disease, one as the occipital horn syndrome and the rest as polymorphisms. As the cross-impact analysis was particularly suited for two relevant events coupled together [80,84–90], we used this analysis to build a quantitative relationship between the increase/decrease of distribution probability after mutation and the diagnosis of Menkes disease.

Figure 2 showed the cross-impact analysis on the relationship between changed primary structure and diagnosis of Menkes disease. At the level of amino-acid distribution probability, $P(2)$ and $P(\bar{2})$ were the decreased and increased probabilities led by mutations, and 12 point mutations resulted in the probability decreased whereas 25 point mutations led to the probability increased. At the level of diag-

Fig. 2 Cross-impact relationship among mutant human ATP7A, changed amino-acid distribution probability, and diagnosis of Menkes disease

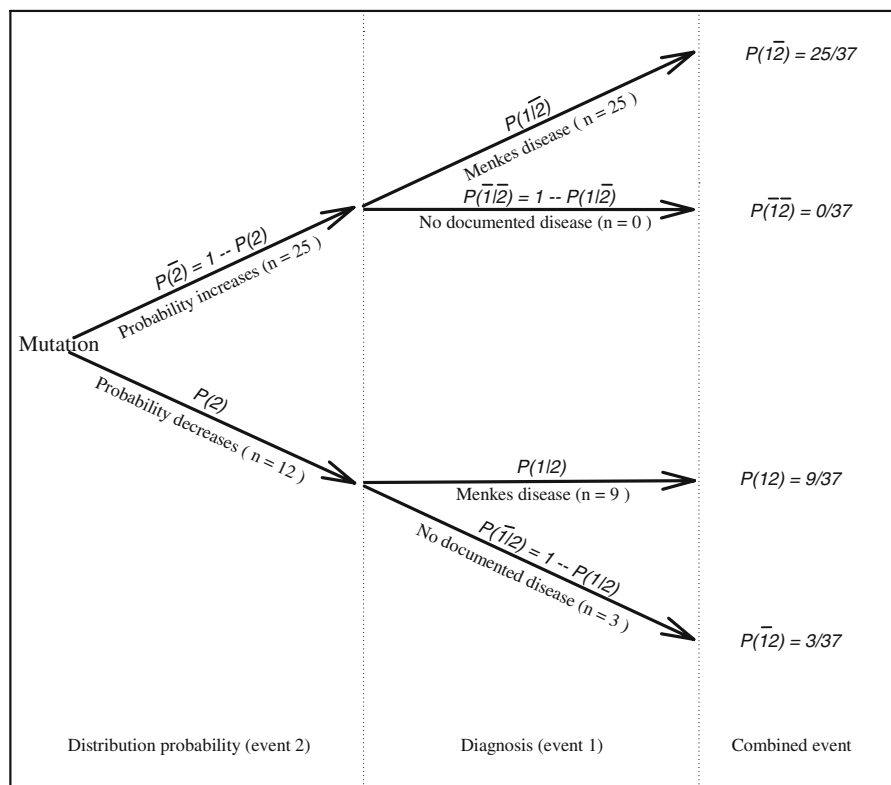


Table 2 Computation on cross-impact analysis in Fig. 2

$P(2) = 12/37 = 0.3243$
$P(\bar{2}) = 1 - P(2) = 1 - 0.3243 = 0.6757 = 25/37$
$P(1 \bar{2}) = 25/25 = 1$
$P(\bar{1} \bar{2}) = 1 - P(1 \bar{2}) = 1 - 1 = 0 = 0/25$
$P(1 2) = 9/12 = 0.75$
$P(\bar{1} 2) = 1 - P(1 2) = 1 - 0.75 = 0.25 = 3/12$
$P(1\bar{2}) = P(1 \bar{2}) \times P(\bar{2}) = 25/25 \times 25/37 = 0.6757 = 25/37$
$P(\bar{1}\bar{2}) = P(\bar{1} \bar{2}) \times P(\bar{2}) = 0/25 \times 25/37 = 0 = 0/37$
$P(12) = P(1 2) \times P(2) = 9/12 \times 12/37 = 0.2432 = 9/37$
$P(\bar{1}2) = P(\bar{1} 2) \times P(2) = 3/12 \times 12/37 = 0.0811 = 3/37$

nosis of Menkes disease under mutation: (i) $P(1|\bar{2})$ was the impact probability (conditional probability) that Menkes disease could be diagnosed under the condition of increased distribution probability, and 25 mutations had such results. (ii) $P(\bar{1}|\bar{2})$ was the impact probability that no disease was documented under the condition of increased distribution probability, and none of mutations worked in such a manner. (iii) $P(1|2)$ was the impact probability that Menkes disease could be diagnosed under the condition of decreased distribution probability, and nine mutations played such a role. (iv) $P(\bar{1}|2)$ was the impact probability that no disease was documented under the condition of decreased distribution probability, and three mutations fell into this category. At the level of combined events, we could see the combined results of changed primary structure and diagnosis of Menkes disease.

Table 2 listed the computed probabilities with respect to Fig. 2, from which we could see several interesting points. (i) As $P(\bar{2})$ was larger than $P(2)$, a mutation had two third chance of increasing the distribution probability in mutant human ATP7A. (ii) As $P(\bar{1}|\bar{2})$ was equal to zero, a mutation that increased the distribution probability had 100% chance of excluding Menkes disease. (iii) As $P(1|2)$ was much larger than $P(\bar{1}|2)$, a mutation that decreased the distribution probability had three fourth chance of bringing out Menkes disease.

Here, it was very meaningful to use the Bayes' law [91],

$$P(1|2) = P(2|1) \frac{P(1)}{P(2)},$$

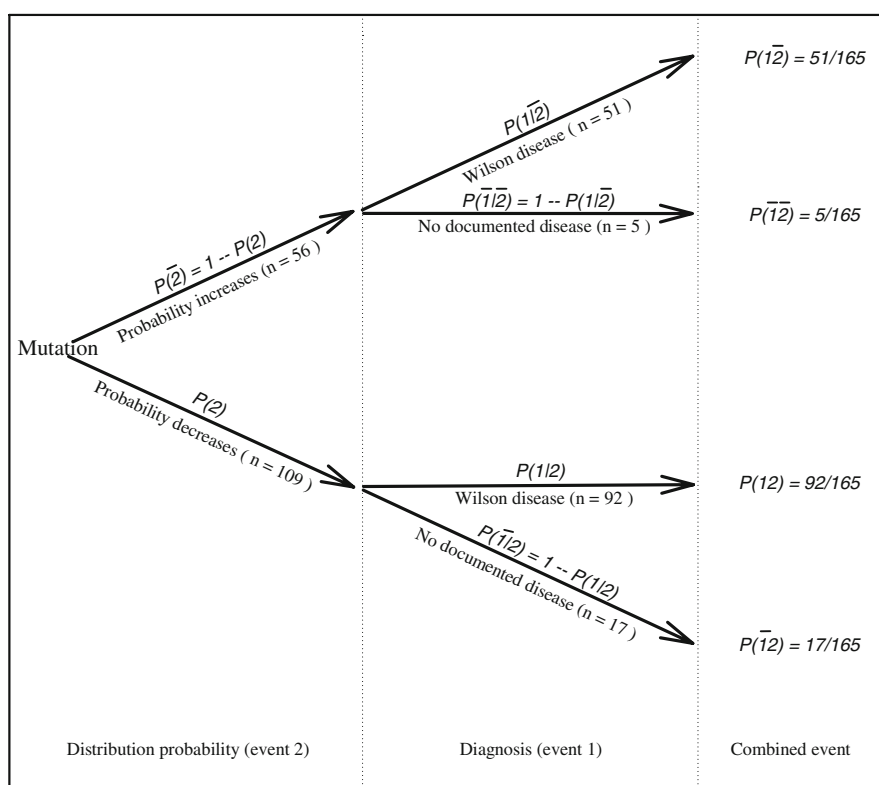
which indicates the probabilities of occurrences of two events, to determine the probability, $P(1)$, the diagnosis of Menkes disease under a mutation, because $P(2)$ and $P(1|2)$ had already been defined in cross-impact analysis, while $P(2|1)$ was the probability that the distribution probability decreased under the condition that the Menkes disease was diagnosed.

Thus, we had $P(1|2) = 9/12 = 0.75$ (Table 2), and $P(2|1) = 9/(9 + 25) = 0.2647$, then we had

$$P(1) = \frac{P(1|2)}{P(2|1)} P(2) = \frac{0.75 \times 0.3243}{0.2647} = 0.9189,$$

namely, the chance that Menkes disease was diagnosed under ATP7A mutation was larger than nine tenth.

Fig. 3 Cross-impact relationship among mutant human ATP7B, changed amino-acid distribution probability, and diagnosis of Wilson disease



Secondly we turned our focus on the human ATP7B with its 165 mutations, of which 143 mutations were documented with Wilson disease. Figure 3 illustrated the cross-impact analysis on the relationship between changed primary structure of ATP7B and diagnosis of Wilson disease.

Similarly, at the level of amino-acid distribution probability, $P(2)$ and $P(\bar{2})$ were the decreased and increased probabilities led by mutations, and 109 and 56 mutations resulted in the probability decreased and increased, respectively. At the level of diagnosis of Wilson disease under a mutation: (i) $P(1|\bar{2})$ was the impact probability that Wilson disease could be diagnosed under the condition of increased distribution probability, and 51 mutations had such results. (ii) $P(\bar{1}|\bar{2})$ was the impact probability that no disease was documented under the condition of increased distribution probability, and five mutations worked in such a manner. (iii) $P(1|2)$ was the impact probability that Wilson disease could be diagnosed under the condition of decreased distribution probability, and 92 mutations played such a role. (iv) $P(\bar{1}|2)$ was the impact probability that no disease was documented under the condition of decreased distribution probability, and 17 mutations fell into this category. At the level of combined events, we could see the combined results of changed primary structure of ATP7B and diagnosis of Wilson disease.

Table 3 displayed the computed probabilities with respect to Fig. 3, from which we could draw some features. (i) As $P(\bar{2})$ was smaller than $P(2)$, a mutation had one third chance of increasing the distribution probability in mutant human

Table 3 Computation on cross-impact analysis in Fig. 3

$P(2) = 109/165 = 0.6606$
$P(\bar{2}) = 1 - P(2) = 1 - 0.6606 = 0.3394 = 56/165$
$P(1 \bar{2}) = 51/56 = 0.9107$
$P(\bar{1} \bar{2}) = 1 - P(1 \bar{2}) = 1 - 0.9107 = 0.0893 = 5/56$
$P(1 2) = 92/109 = 0.8440$
$P(\bar{1} 2) = 1 - P(1 2) = 1 - 0.8440 = 0.1560 = 17/109$
$P(1\bar{2}) = P(1 \bar{2}) \times P(\bar{2}) = 51/56 \times 56/165 = 0.3091 = 51/165$
$P(\bar{1}\bar{2}) = P(\bar{1} \bar{2}) \times P(\bar{2}) = 5/56 \times 56/165 = 0.0303 = 5/165$
$P(12) = P(1 2) \times P(2) = 92/109 \times 109/165 = 0.5576 = 92/165$
$P(\bar{1}2) = P(\bar{1} 2) \times P(2) = 17/109 \times 109/165 = 0.1030 = 17/165$

ATP7B. (ii) As $P(\bar{1}|\bar{2})$ was less than 0.1, a mutation that increased the distribution probability had more than nine tenth chance of leading to Wilson disease. (iii) As $P(1|2)$ was much larger than $P(\bar{1}|\bar{2})$, a mutation that decreased the distribution probability had more than eight tenth chance of resulting in Wilson disease.

Also, we could use the Bayesian equation to determine the probability, $P(1)$, that Wilson disease was diagnosed under ATP7B mutations. As we had $P(1|2) = 92/109 = 0.8440$ (Table 3), and $P(2|1) = 92/(92 + 51) = 0.6434$, so

$$P(1) = \frac{P(1|2)}{P(2|1)} P(2) = \frac{0.8440 \times 0.6606}{0.6434} = 0.8666,$$

say, a ATP7B mutation had about nine tenth chance of causing Wilson disease.

Although both ATP7A and ATP7B genes and their encoding proteins have similar structure [92], the ATP7A mutations in humans cause the Menkes disease characterized by a copper deficiency disorder, while the mutations in ATP7B result in the Wilson disease characterized by a copper toxicity condition. The reason for the very different diseases caused by mutations is in part due to their different pattern of cellular trafficking [93–96] and distinct pattern of tissue expression [97].

Also, different mutation patterns are found in both genes. For example, deletions of varying sizes within the ATP7A gene have been identified in 15–20% of patients with classical Menkes disease [42]. Unlike the situation in Menkes disease, large gene deletions have not been reported for patients with Wilson disease [52]. Thus, less point mutations were documented in ATP7A but much more mutations in ATP7B.

Conclusion

From the probabilistic viewpoint, vast majority of mutations lead to the amino-acid distribution probability increase in

mutant ATP7As and decrease in ATP7Bs, which may throw lights on the difference between Menkes and Wilson disease.

The genotype–phenotype relationship is a crucial issue for diagnosing genetic defective diseases, such as Menkes and Wilson disease [83,97]. There is, in part, a correlation between the severity of the mutation and the disease severity, but other factors are also involved [98]. Using the cross-impact analysis and the Bayesian equation, we can calculate the probability, $P(1)$, that Menkes/Wilson disease can be diagnosed under ATP7A/ATP7B mutations, which will benefit for early clinical diagnosis as treatment can prevent tissue damage.

Appendix

Example of computation of amino-acid distribution probability for normal ATP7A

For the simplest example, there are nine tryptophans (W) in ATP7A, which are the least abundant amino acids in this protein (Table 1). How do these nine Ws distribute along the

Appendix Table 1 All possible distributions of nine tryptophans in human ATP7A

Partition									Distribution probability
I	II	III	IV	V	VI	VII	VIII	IX	
1	1	1	1	1	1	1	1	1	9.3666e-4
	1	1	1	1	1	1	1	2	0.0337
		1	1	1	1	1	1	3	0.0393
			1	1	1	1	1	4	0.0197
				1	1	1	1	5	4.9174e-3
					1	1	1	6	6.5566e-4
						1	1	7	4.6833e-5
							1	8	1.6726e-6
								9	2.3231e-8
		<i>1</i>	<i>1</i>	<i>1</i>	<i>1</i>	<i>1</i>	<i>2</i>	<i>2</i>	<i>0.1770</i>
			1	1	1	1	2	3	0.1967
				1	1	1	2	4	0.0492
					1	1	2	5	5.9009e-3
						1	2	6	3.2783e-4
							2	7	6.6904e-6
			1	1	1	2	2	2	0.1967
				1	1	2	2	3	0.1475
					1	2	2	4	0.0148
						2	2	5	4.9174e-4
				1	2	2	2	2	0.0369
					2	2	2	3	9.8349e-3
				1	1	1	3	3	0.0328
					1	1	3	4	9.8349e-3
						1	3	5	6.5566e-4
							3	6	1.5611e-5
					1	2	3	3	0.0197
						2	3	4	1.6391e-3
						3	3	3	3.6426e-4
						1	4	4	4.0979e-4
							4	5	2.3416e-5

Bold and italic is the real distribution

ATP7A? Accordingly, we can imagine to dividing the ATP7A into nine equal partitions, and each partition has about 167 amino acids (1,500/9) because the human ATP7A is composed of 1,500 amino acids, then there would be 30 configurations for all the possible distributions of nine Ws (Appendix Table 1).

Here, we calculate two distribution probabilities in this Table as example according to the equation. For the nine Ws equally distribute in each partition (the third row in the Table), we have $r_1 = 1, r_2 = 1, \dots, r_9 = 1, q_0 = 0, q_1 = 9, \dots, q_9 = 0$. Thus, we have the distribution probability as

$$\begin{aligned} & \frac{9!}{0! \times 9! \times 0! \times 0! \times 0! \times 0! \times 0! \times 0! \times 0!} \\ & \times \frac{9!}{1! \times 1! \times 1! \times 1! \times 1! \times 1! \times 1! \times 1! \times 1!} \times 9^{-9} \\ & = \frac{362880}{1 \times 362880 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1} \\ & \times \frac{362880}{1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1} \\ & \times \frac{1}{387420489} = 9.3666e^{-4}. \end{aligned}$$

Clearly, the human ATP7A can adopt only one pattern, which is that two partitions contain zero W, five partitions contain one W and two partitions contain two Ws (row 12 in the Table). So we have $r_1 = 0, r_2 = 0, r_3 = 1, r_4 = 1, r_5 = 1, r_6 = 1, r_7 = 1, r_8 = 2, r_9 = 2, q_0 = 2, q_1 = 5, q_2 = 2, q_3 = 0, q_4 = 0, q_5 = 0, q_6 = 0, q_7 = 0, q_8 = 0, q_9 = 0$, that is,

$$\begin{aligned} & \frac{9!}{2! \times 5! \times 2! \times 0! \times 0! \times 0! \times 0! \times 0! \times 0!} \\ & \times \frac{9!}{0! \times 0! \times 1! \times 1! \times 1! \times 1! \times 1! \times 2! \times 2!} \times 9^{-9} \\ & = \frac{362880}{2 \times 120 \times 2 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1} \\ & \times \frac{362880}{1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 1 \times 2 \times 2} \\ & \times \frac{1}{387420489} = 0.1770. \end{aligned}$$

Example of computation of amino-acid distribution probability for ATP7A after mutation

Appendix Table 2 shows the distribution patterns for H before and after mutation at position 1069. Accordingly, we have $r_1 = 0, r_2 = 0, r_3 = 1, r_4 = 0, r_5 = 1, r_6 = 0, r_7 = 3, r_8 = 0, r_9 = 2, r_{10} = 1, r_{11} = 1, r_{12} = 2, r_{13} = 0, r_{14} = 0, r_{15} = 1, r_{16} = 4, r_{17} = 2, r_{18} = 1, r_{19} = 0, r_{20} = 1, r_{21} = 0, r_{22} = 2, r_{23} = 0, r_{24} = 1, r_{25} = 1, r_{26} = 1, r_{27} = 1, r_{28} = 2, r_{29} = 1, r_{30} = 1, r_{31} = 1, r_{32} = 1, r_{33} = 0, r_{34} = 2, r_{35} = 1, r_{36} = 1; q_0 = 11, q_1 = 17, q_2 = 6, q_3 = 1, q_4 = 1, q_5 = 0, q_6 = 0, q_7 = 0, q_8 = 0, q_9 = 0, q_{10} = 0, q_{11} = 0, q_{12} = 0, q_{13} = 0,$

Appendix Table 2 Distribution of histidines before and after mutation at position 1069 in human ATP7B

Partition	Before mutation	After mutation
I	0	0
II	0	0
III	1	1
IV	0	0
V	1	1
VI	0	1
VII	3	2
VIII	0	0
IX	2	3
X	1	0
XI	1	2
XII	2	1
XIII	0	0
XIV	0	1
XV	1	1
XVI	4	3
XVII	2	2
XVIII	1	1
XIX	0	1
XX	1	0
XXI	0	1
XXII	2	1
XXIII	0	1
XXIV	1	0
XXV	1	2
XXVI	1	0
XXVII	1	2
XXVIII	2	1
XXIX	1	1
XXX	1	1
XXXI	1	0
XXXII	1	1
XXXIII	0	0
XXXIV	2	3
XXXV	1	1
XXXVI	1	–

$q_{14} = 0, q_{15} = 0, q_{16} = 0, q_{17} = 0, q_{18} = 0, q_{19} = 0, q_{20} = 0, q_{21} = 0, q_{22} = 0, q_{23} = 0, q_{24} = 0, q_{25} = 0, q_{26} = 0, q_{27} = 0, q_{28} = 0, q_{29} = 0, q_{30} = 0, q_{31} = 0, q_{32} = 0, q_{33} = 0, q_{34} = 0, q_{35} = 0, q_{36} = 0$ and $q_{37} = 0$ before mutation, whose distribution probability is

$$\begin{aligned} & \frac{36!}{11! \times 17! \times 6! \times 1! \times 0! \times 0! \times 0! \times 0! \times 0!} \\ & \times \frac{36!}{0! \times 0! \times 1! \times 0! \times 1! \times 0! \times 0! \times 0! \times 0!} \times 36^{-36} = 0.0138. \end{aligned}$$

We also have $r_1 = 0, r_2 = 0, r_3 = 1, r_4 = 0, r_5 = 1, r_6 = 1, r_7 = 2, r_8 = 0, r_9 = 3, r_{10} = 0, r_{11} = 2, r_{12} = 1, r_{13} = 0, r_{14} = 1, r_{15} = 1, r_{16} = 3, r_{17} = 2, r_{18} = 1, r_{19} = 1, r_{20} = 0, r_{21} = 1, r_{22} = 1, r_{23} = 1, r_{24} = 0, r_{25} = 2, r_{26} = 0, r_{27} = 2, r_{28} = 1, r_{29} = 1, r_{30} = 1, r_{31} = 0, r_{32} = 1, r_{33} = 0, r_{34} = 3, r_{35} = 1; q_0 = 11, q_1 = 16, q_2 = 5, q_3 = 3, q_4 = 0, q_5 = 0, q_6 = 0, q_7 = 0, q_8 = 0, q_9 = 0, q_{10} = 0, q_{11} = 0, q_{12} = 0,$

$q_{13} = 0, q_{14} = 0, q_{15} = 0, q_{16} = 0, q_{17} = 0, q_{18} = 0, q_{19} = 0, q_{20} = 0, q_{21} = 0, q_{22} = 0, q_{23} = 0, q_{24} = 0, q_{25} = 0, q_{26} = 0, q_{27} = 0, q_{28} = 0, q_{29} = 0, q_{30} = 0, q_{31} = 0, q_{32} = 0, q_{33} = 0, q_{34} = 0, q_{35} = 0$, and $q_{36} = 0$ after mutation, whose distribution probability is

$$\frac{35!}{11! \times 16! \times 5! \times 3! \times 0! \dots \times 0!} \times \frac{35!}{0! \times 0! \times 1! \times 0! \times 1! \dots \times 1!} \times 35^{-35} = 0.0233.$$

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