

## Exam 1 Solutions

$$1. (a) f(x|\theta) = (\theta-1) x^{-\theta} I_{(1,\infty)}(x)$$

$$= I_{(1,\infty)}(x) (\theta-1) e^{\theta(-\log x)} = h(x) c(\theta) e^{w_1(\theta) \cdot t_1(x)}$$

where  $h(x) = I_{(1,\infty)}(x)$ ,  $c(\theta) = \theta-1$ ,  $w_1(\theta) = \theta$ , and  $t_1(x) = -\log x$

So this family of pdfs is a one-dimensional exponential family.

$$(b) E\left[\frac{dw_1(\theta)}{d\theta} t_1(X)\right] = -\frac{d}{d\theta} \log c(\theta)$$

$$\text{Now, } \frac{dw_1(\theta)}{d\theta} = \frac{d}{d\theta} [\theta] = 1 \quad \text{and} \quad \frac{d}{d\theta} [\log c(\theta)] = \frac{d}{d\theta} [\log(\theta-1)] = \frac{1}{\theta-1}$$

$$\text{So that } E[1 \cdot (-\log X)] = -\frac{1}{\theta-1}$$

$$\Downarrow \\ E[\log X] = \boxed{\frac{1}{\theta-1}}$$

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$$2. (a) P(X=0) = \int_0^{1/2} P(X=0, U=u) du$$

$$= \int_0^{1/2} P(X=0|U=u) \cdot P(U=u) du$$

$$= \int_0^{1/2} (1-u) \cdot 2 du = \left[-(1-u)^2\right]_0^{1/2} = -\frac{1}{4} + 1 = \frac{3}{4}$$

$$P(X=1) = \int_0^{1/2} P(X=1|U=u) \cdot P(U=u) du$$

$$= \int_0^{1/2} u \cdot 2 du = \left[u^2\right]_0^{1/2} = \frac{1}{4} - 0 = \frac{1}{4}$$

So  $X \sim \text{Bernoulli}(\frac{1}{4})$ .

$$(b) EX = \frac{1}{4}, \quad EU = \int_0^{1/2} u \cdot f_u(u) du = \int_0^{1/2} 2u du = \left[u^2\right]_0^{1/2} = \frac{1}{4}$$

$$E[XU] = E[E[XU|U]] = E[U E[X|U]] = E[U \cdot u]$$

$\uparrow$   
since  $X|U=u \sim \text{Bernoulli}(u)$

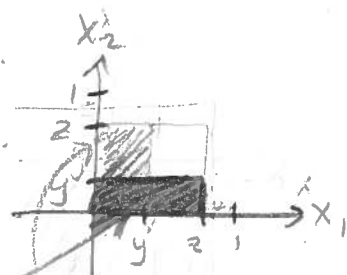
$$\text{So } E[XU] = E[U^2] = \int_0^{1/2} u^2 \cdot 2 du = \left[ \frac{2u^3}{3} \right]_0^{1/2} \\ = \frac{2(\frac{1}{8})}{3} = \frac{1}{12}$$

$$\text{So } \text{Cov}(X, U) = E[XU] - E[X]E[U] \\ = \frac{1}{12} - \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{12} - \frac{1}{16} = \frac{4-3}{48} = \boxed{\frac{1}{48}}$$

$$3. (a) F_{X_{(1)}, X_{(2)}}(y, z) = P(X_{(1)} \leq y \text{ and } X_{(2)} \leq z)$$

If  $\max\{X_1, X_2\} \leq z$ , then  $X_1 \leq z$  and  $X_2 \leq z$ .

If  $\min\{X_1, X_2\} \leq y$ , then  $X_1 \leq y$  or  $X_2 \leq y$ .



So  $\{X_{(1)} \leq y \text{ and } X_{(2)} \leq z\}$  if and only if  $\{X_1 \leq z \text{ and } X_2 \leq y\}$  or  $\{X_1 \leq y \text{ and } y < X_2 \leq z\}$

If  $0 < y < z \leq 1$ , then

$$F_{X_{(1)}, X_{(2)}}(y, z) = P(X_1 \leq z \text{ and } X_2 \leq y) + P(X_1 \leq y \text{ and } y < X_2 \leq z) \\ = P(X_1 \leq z) P(X_2 \leq y) + P(X_1 \leq y) P(y < X_2 \leq z) \\ = z \cdot y + y \cdot (z - y) \\ = y(2z - y)$$

Then  $\frac{\partial^2}{\partial y \partial z} [y(2z - y)] = \frac{\partial}{\partial y} \left\{ \frac{\partial}{\partial z} [y(2z - y)] \right\} \\ = \frac{\partial}{\partial y} [y \cdot 2] = 2$

so  $\boxed{f_{X_{(1)}, X_{(2)}}(y, z) = 2 \cdot I_{(0, z)}(y) I_{(0, 1)}(z)}$

(b)

$$\left. \begin{aligned} v &= \frac{y}{z} \\ w &= z \end{aligned} \right\} \Leftrightarrow \begin{cases} y = vz = vw \\ z = w \end{cases}$$

so this is a one-to-one transformation.

Since  $J = \begin{vmatrix} \frac{\partial x}{\partial v} & \frac{\partial x}{\partial w} \\ \frac{\partial y}{\partial v} & \frac{\partial y}{\partial w} \end{vmatrix} = \begin{vmatrix} w & v \\ 0 & 1 \end{vmatrix} = w,$

the joint density of  $V$  and  $W$  is

$$\begin{aligned} f_{v,w}(v,w) &= f_{x_{(1)}, x_{(2)}}(vw, w) \cdot |J| \\ &= 2 \cdot I_{(0,w)}(vw) \cdot I_{(0,1)}(w) \cdot w \\ &= 2w I_{(0,1)}(w) I_{(0,w)}(vw) \\ &= 2w I_{(0,1)}(w) \cdot I_{(0,1)}(v) \\ &= f_w(w) \cdot f_v(v) \end{aligned}$$

so  $V$  and  $W$  are independent.

4. (a)  $Y_1 + Y_2 \sim \text{Normal}(0, 4)$  since  $E[Y_1 + Y_2] = E[Y_1] + E[Y_2]$

$$\begin{aligned} \text{and } \text{Var}[Y_1 + Y_2] &\stackrel{\text{indep.}}{=} \text{Var}[Y_1] + \text{Var}[Y_2] \\ &\stackrel{\text{indep.}}{=} 2 + 2 = 4 \end{aligned}$$

Normal table  
= .3085

$$P(Y_1 + Y_2 > 1) = P\left(\frac{(Y_1 + Y_2) - 0}{\sqrt{4}} > \frac{1 - 0}{2}\right) = P(Z > \frac{1}{2}) \stackrel{\text{Normal table}}{=} 1 - .6915 = .3085$$

Also,  $\left(\frac{Y_3}{\sqrt{2}}\right), \left(\frac{Y_4}{\sqrt{2}}\right), \left(\frac{Y_5}{\sqrt{2}}\right)$  are iid Normal(0,1)

so  $\frac{Y_3^2}{2} + \frac{Y_4^2}{2} + \frac{Y_5^2}{2} \sim \chi^2_3$

and  $\sqrt{\frac{2}{3}}$

$$P(Y_3^2 + Y_4^2 + Y_5^2 > 5) = P\left(\frac{Y_3^2 + Y_4^2 + Y_5^2}{2} > \frac{5}{2}\right)$$

$$\stackrel{\chi^2 \text{ table}}{=} .48$$

Since  $Y_1 + Y_2$  is independent of  $Y_3^2 + Y_4^2 + Y_5^2$ ,

$$P(Y_1 + Y_2 > 1 \text{ and } Y_3^2 + Y_4^2 + Y_5^2 > 5) = P(Y_1 + Y_2 > 1) P(Y_3^2 + Y_4^2 + Y_5^2 > 5)$$

$$= (.3085)(.48)$$

$$(b) \frac{\bar{Y} - 0}{s/\sqrt{5}} = \frac{\bar{Y}}{\sqrt{\frac{1}{4} \sum (Y_i - \bar{Y})^2} / \sqrt{5}} = \frac{2\sqrt{5} \bar{Y}}{\sum (Y_i - \bar{Y})^2} \sim t_4$$

$$\text{So } P\left(\sqrt{5} \bar{Y} < \sqrt{\frac{5}{2} \sum (Y_i - \bar{Y})^2}\right) = P\left(\frac{2\sqrt{5} \bar{Y}}{\sum (Y_i - \bar{Y})^2} < 2\right)$$

$$= P(T < 2) \stackrel{t\text{-table}}{=} 1 - .06 = .94$$

$$5. (a) E[X(X-1)] = \sum_{x=0}^{\infty} x(x-1) \frac{1}{x!} 4^x e^{-4}$$

$$= \sum_{x=2}^{\infty} x(x-1) \frac{1}{x!} 4^x e^{-4}$$

$$= \sum_{x=2}^{\infty} \frac{1}{(x-2)!} 4^{x-2} 4^2 e^{-4}$$

$$= 4^2 \sum_{x=2}^{\infty} \frac{1}{(x-2)!} 4^{x-2} e^{-4} = 4^2 \cdot 1 = 16$$

$$(b) a = EX = 4, b = \sqrt{\text{Var } X} = \sqrt{EX^2 - (EX)^2} = \sqrt{20 - 4^2} = \sqrt{4} = 2$$

$$\text{since } EX^2 - EX = 16 \Rightarrow EX^2 = 16 + 4 = 20$$

$$(c) P(3.99 < \bar{X} < 4.01) = P\left(\frac{\sqrt{10000}(\bar{X} - 4)}{2} < \frac{\sqrt{10000}(\bar{X} - 4)}{2} < \frac{\sqrt{10000}(4.01 - 4)}{2}\right)$$

$$\stackrel{\text{CLT}}{\approx} P\left(-\frac{1}{2} < Z < \frac{1}{2}\right)$$

$$= P(Z < \frac{1}{2}) - P(Z < -\frac{1}{2})$$

$$= .6915 - (1 - .6915) = .3830$$