

The exam is closed book; students are permitted to prepare one 8.5×11 page of formulas, notes, etc. that can be used during the exam. A calculator is permitted but not necessary for the exam.

Do all 4 problems (the problems with the three highest scores are worth 30% each, the problem with the lowest score is worth 10%).

Problem 1. (10 points) Suppose $\begin{pmatrix} y_1 \\ x_1 \end{pmatrix}, \dots, \begin{pmatrix} y_n \\ x_n \end{pmatrix}$ are independent $N_2 \left(\begin{pmatrix} \mu_y \\ \mu_x \end{pmatrix}, \begin{pmatrix} \sigma_{yy} & \sigma_{yx} \\ \sigma_{yx} & \sigma_{xx} \end{pmatrix} \right)$ random variables.

Define $\beta_0 = \mu_y - \sigma_{yx}\sigma_{xx}^{-1}\mu_x$ and $\beta_1 = \sigma_{xx}^{-1}\sigma_{yx}$ so that

$E(y_i|x_i) = \beta_0 + \beta_1 x_i$ and $\sigma^2 = \text{var}(y_i|x_i) = \sigma_{yy} - \sigma_{yx}\sigma_{xx}^{-1}\sigma_{yx}$.

(a - 7 pts) If $n = 10$, $\sum_{i=1}^{10} x_i = 4$ and $\sum_{i=1}^{10} x_i^2 = 2$, find a constant C such that the F -statistic for testing $H_0 : \beta_0 = \beta_1$ versus $H_A : \beta_0 \neq \beta_1$ can be expressed in the form

$$F = C \left(\frac{\hat{\beta}_0 - \hat{\beta}_1}{\hat{\sigma}} \right)^2$$

where $\hat{\beta}_0$, $\hat{\beta}_1$, and $\hat{\sigma}^2$ are the maximum likelihood estimators of β_0 , β_1 , and σ^2 , respectively.

(b - 3 pts) For what values of $\frac{\hat{\beta}_0 - \hat{\beta}_1}{\hat{\sigma}}$ should H_0 be rejected at level .05?

Problem 2. (10 points) Suppose $y_i = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \varepsilon_i$ for $i = 1, \dots, 3$ where $x_i = i$ are known values, the regression parameters β_0 , β_1 , and β_2 are unknown, and $\varepsilon_1, \varepsilon_2, \varepsilon_3$ are independent and identically distributed random variables with mean 0 and unknown variance σ^2 . Now, suppose that the x^2 term is excluded from the model, and least squares estimation is used to model the y 's based on an intercept term and the x 's (that is, suppose that we incorrectly use a linear model and find values $\hat{\beta}_0^*$ and $\hat{\beta}_1^*$, which minimize $\sum_{i=1}^3 (y_i - \beta_0^* - \beta_1^* x_i)^2$).

What is the bias of using $\hat{\beta}_1^*$ to estimate β_1 ? Write your answer as a function of the true (but unknown) parameter values β_0 , β_1 , and β_2 .

Problem 3. (10 points) Suppose $\mathbf{y} \sim N_6(\mathbf{X}\boldsymbol{\beta}, \sigma^2\mathbf{I})$ where \mathbf{X} is an 6×2 matrix and $\boldsymbol{\beta} = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}$ is a vector of fixed but unknown parameters. Also, suppose that

- $\mathbf{X}^\top \mathbf{X} = \begin{pmatrix} 6 & 2 \\ 2 & 1 \end{pmatrix}$,
- the length of the 95% confidence interval for β_0 is 4,
- the 6th row of \mathbf{X} is $(1, 0)$, and
- the residual for the 6th observation is 0.

(a - 3 pts) Compute the maximum likelihood estimate of σ^2 .

(b - 4 pts) Compute the maximum likelihood estimate of σ^2 with the 6th observation removed from the data set.

(c - 3 pts) Compute the length of the 95% confidence interval for β_0 based only on the first 5 observations in the data set (that is, with the 6th observation removed from the data set).

Problem 4. (10 points) Suppose that

$$y_{ijk} = \mu + \alpha_i + \gamma_j + \varepsilon_{ijk} \text{ for } i, j, k = 1, 2$$

where $\alpha_1 + \alpha_2 = \gamma_1 + \gamma_2 = 0$ and ε_{ijk} are independent $\text{Normal}(0, \sigma^2)$ random variables.
 Given the data

		<i>i</i>	
		1	2
<i>j</i>	1	8,2	4,6
	2	0,4	0,0

test the hypothesis $H_0 : \alpha_1 = \alpha_2 = 0$ at level 0.05.

Here are some formulas which might be helpful:

$$\hat{\mathbf{y}} = \mathbf{X}_1\hat{\boldsymbol{\beta}}_1 + \mathbf{X}_2\hat{\boldsymbol{\beta}}_2 \text{ where } \hat{\boldsymbol{\beta}}_1^* = (\mathbf{X}_1^\top \mathbf{X}_1)^{-1} \mathbf{X}_1^\top \mathbf{y}, \mathbf{A} = (\mathbf{X}_1^\top \mathbf{X}_1)^{-1} \mathbf{X}_1^\top \mathbf{X}_2, \text{ and } \hat{\boldsymbol{\beta}}_1 = \hat{\boldsymbol{\beta}}_1^* - \mathbf{A}\hat{\boldsymbol{\beta}}_2$$

$$E(\hat{\boldsymbol{\beta}}_1^*) = \boldsymbol{\beta}_1 + \mathbf{A}\boldsymbol{\beta}_2 \text{ where } \mathbf{A} = (\mathbf{X}_1^\top \mathbf{X}_1)^{-1} \mathbf{X}_1^\top \mathbf{X}_2$$

$$E(s_1^2) = \sigma^2 + \frac{\boldsymbol{\beta}_2^\top \mathbf{X}_2^\top (\mathbf{I} - \mathbf{X}_1(\mathbf{X}_1^\top \mathbf{X}_1)^{-1} \mathbf{X}_1^\top) \mathbf{X}_2 \boldsymbol{\beta}_2}{n - p - 1}$$

$$F = \frac{SSH/q}{SSE/(n - k - 1)} \text{ where}$$

$$SSH = (\mathbf{C}\hat{\boldsymbol{\beta}} - \mathbf{t})^\top (\mathbf{C}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{C}^\top)^{-1} (\mathbf{C}\hat{\boldsymbol{\beta}} - \mathbf{t}) \text{ and } SSE = \mathbf{y}^\top (\mathbf{I} - \mathbf{H}) \mathbf{y} = \mathbf{y}^\top \mathbf{y} - \hat{\boldsymbol{\beta}}^\top \mathbf{X}^\top \mathbf{y}$$

$$\mathbf{a}^\top \hat{\boldsymbol{\beta}} \pm t_{\alpha/2, n-k-1} s \sqrt{\mathbf{a}^\top (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{a}} \text{ is a } 100(1 - \alpha)\% \text{ confidence interval for } \mathbf{a}^\top \boldsymbol{\beta}$$

$$\hat{\boldsymbol{\beta}}_c = \hat{\boldsymbol{\beta}} - (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{C}^\top (\mathbf{C}(\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{C}^\top)^{-1} (\mathbf{C}\hat{\boldsymbol{\beta}} - \mathbf{t}) \text{ and } \hat{\sigma}_c^2 = \frac{Q(\hat{\boldsymbol{\beta}}_c)}{n}$$

$$\hat{\varepsilon}_i = y_i - \mathcal{X}_i^\top \hat{\boldsymbol{\beta}}$$

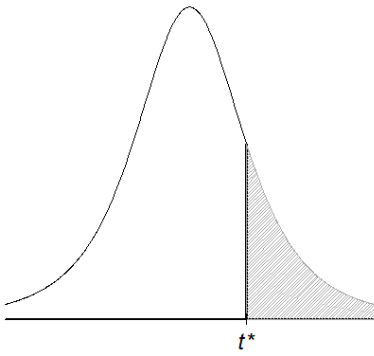
$$\hat{\mathbf{y}}_{(i)} = \mathcal{X}_i^\top \hat{\boldsymbol{\beta}}_{(i)}$$

$$SSE_{(i)} = \mathbf{y}_{(i)}^\top \mathbf{y}_{(i)} - \hat{\boldsymbol{\beta}}_{(i)}^\top \mathbf{X}_{(i)}^\top \mathbf{y}_{(i)}$$

$$r_i = \frac{\hat{\varepsilon}_i}{s\sqrt{1 - h_{ii}}}$$

$$\hat{\boldsymbol{\beta}}_{(i)} = \hat{\boldsymbol{\beta}} - \frac{\hat{\varepsilon}_i}{1 - h_{ii}} (\mathbf{X}^\top \mathbf{X})^{-1} \mathcal{X}_i$$

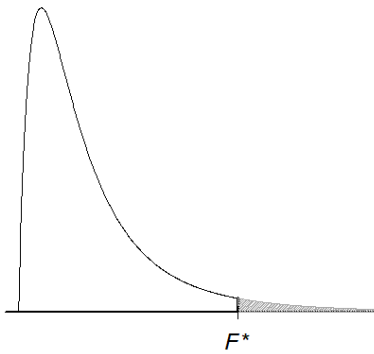
$$s_{(i)}^2 = \left(\frac{n - k - 1 - r_i^2}{n - k - 2} \right) s^2$$



If T is a random variable with a t distribution having df degrees of freedom, then the critical value t^* in the table is the value such that the shaded area is $p = P(T > t^*)$.

t distribution critical values

df	Upper-tail probability p				
	.10	.05	.025	.01	.005
1	3.078	6.314	12.706	31.821	63.657
2	1.886	2.920	4.303	6.965	9.925
3	1.638	2.353	3.182	4.541	5.841
4	1.533	2.132	2.776	3.747	4.604
5	1.476	2.015	2.571	3.365	4.032
6	1.440	1.943	2.447	3.143	3.707
7	1.415	1.895	2.365	2.998	3.499
8	1.397	1.860	2.306	2.896	3.355
9	1.383	1.833	2.262	2.821	3.25
10	1.372	1.812	2.228	2.764	3.169
∞	1.282	1.645	1.960	2.326	2.576



If f is a random variable with an F distribution having $df1$ and $df2$ degrees of freedom, then the critical value F^* in the table is the value such that the shaded area is $P(f > F^*) = p$.

F distribution critical values

$df2$	$p = .05$					$p = .025$				
	$df1$					$df1$				
	1	2	3	4	5	1	2	3	4	5
1	161.45	199.5	215.71	224.58	230.16	1	647.79	799.50	864.16	899.58
2	18.51	19.00	19.16	19.25	19.30	2	38.51	39.00	39.17	39.25
3	10.13	9.55	9.28	9.12	9.01	3	17.44	16.04	15.44	15.10
4	7.71	6.94	6.59	6.39	6.26	4	12.22	10.65	9.98	9.60
5	6.61	5.79	5.41	5.19	5.05	5	10.01	8.43	7.76	7.39
6	5.99	5.14	4.76	4.53	4.39	6	8.81	7.26	6.60	6.23
7	5.59	4.74	4.35	4.12	3.97	7	8.07	6.54	5.89	5.52
8	5.32	4.46	4.07	3.84	3.69	8	7.57	6.06	5.42	5.05
9	5.12	4.26	3.86	3.63	3.48	9	7.21	5.71	5.08	4.72
10	4.96	4.10	3.71	3.48	3.33	10	6.94	5.46	4.83	4.47