PROVIDENCE: an Efficient and Secure Ballot Polling Risk-Limiting Audit

Oliver Broadrick

Thesis Committee:

Arkady Yerukhimovich, Assistant Professor, GW, Committee Chair Poorvi Vora, Professor, GW, Thesis Advisor, Committee Member Filip Zagórski, Assistant Professor, University of Wrocław, Thesis Co-Adviser, Committee Member

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Outline

- ► Risk-Limiting Audits (RLAs)
- Problem statement and contributions
- ▶ Background: From BRAVO, to MINERVA...
- Lambda in the Providence
 - Theoretical properties
 - Experimental verification
- Workload models
- ► Pilot use
- Conclusion

Election security

What do we want?

► The right winner *and* strong evidence that they are the right winner

An approach:

- Cast votes on paper ballots
- ► Tally the votes with electronic scanners
- Perform compliance and tabulation audits

Risk-Limiting Audit (RLA)

Risk-Limiting Audit (RLA) with risk limit α : A tabulation audit that, if the reported outcome is wrong, will detect and correct it with probability $1-\alpha$.

 α is an error measure. Low values of α are good.

Ballot Polling RLAs

- Assume a voter-verified paper trail stored in numbered boxes, successfully completed compliance audits, and a ballot manifest, which describes how many ballots in each box. Then publicly:
- 1. Choose a round size, n
- 2. Select *n* ballots uniformly at random, with replacement, using a pseudorandom number generator

Ballot Polling RLAs

Rolling the dice!



Ballot Polling RLAs

- Assume a voter-verified paper trail, successfully completed compliance audits, and a ballot manifest. Then publicly:
- 1. Choose a round size, n
- 2. Select *n* ballots uniformly at random, with replacement, using a pseudorandom number generator
- 3. Find and manually interpret the selected ballots, recording the interpretations
- 4. Compute the stopping condition $\mathcal A$ which outputs:
 - (a) Correct: stop, confirming the reported outcome, or
 - (b) Undetermined: sample more ballots, or
 - (c) Hand count: stop and perform a full recount
- 5. If more ballots are to be drawn, return to step 1 (or, when election officials choose, proceed to a full, manual recount)

Existing ballot polling RLAs

Bravo

- ▶ In the two candidate case is an instance of Wald's classic Sequential Probability Ratio Test (SPRT)
- Most efficient RLA when the stopping condition is checked after each ballot is drawn (ballot-by-ballot)
- ▶ In real audits, decisions are taken after many ballots are drawn (round-by-round), and BRAVO is implemented as:
 - Selection-Ordered (SO) BRAVO, where ballot selection order is retained, and the decisions are taken as though the audit were ballot-by-ballot
 - ► End-of-Round (EoR) BRAVO, where the decision using the BRAVO stopping rule is taken once, after the entire round of ballots is drawn

Existing ballot polling RLAs

Minerva

- Recent RLA designed for round-by-round use
- Known to be risk-limiting if all round sizes are predetermined, before the audit begins
- ▶ In a first round chosen to give a 0.90 probability of stopping, MINERVA requires
 - ▶ 50% as many ballots as EoR BRAVO
 - ▶ 70-80% as many ballots as SO BRAVO
- ► For smaller stopping probability rounds, the benefit of MINERVA decreases (i.e. as the audit approaches the ballot-by-ballot case)

Problems We Will Address

- 1. Predetermined round sizes are limiting
 - e.g. After nearly meeting the stopping condition in a round, MINERVA requires escalation to the predetermined next round size
 - May be more efficient to choose future round sizes as a function of previous samples; would allow greater flexibility
- 2. Existing workload measures don't capture the cost of a round
 - We are unaware of any RLAs that have ever actually drawn a single ballot at a time
 - More efficient to draw many at once

Contributions

- 1. RLA Providence
 - An audit with the efficiency of MINERVA and flexibility of BRAVO
 - ▶ Proofs of theoretical properties, plus experiments
- 2. Evaluation of BRAVO, MINERVA, and PROVIDENCE under simple, new workload models that account for the cost of a round

Background

RLA definition

An audit A is a function from the space of possible samples, X, to the set {Correct, Undetermined} where:

- (a) Correct: stop the audit
- (b) Undetermined: sample more ballots

Definition (α -Risk-Limiting Audit (α -RLA))

An audit A is an α -RLA if for all samples $X \in \mathcal{X}$

$$\Pr[\mathcal{A}(X) = Correct \mid H_0] \leq \alpha,$$

where H_0 corresponds to the incorrectly reported outcome *closest* to the reported outcome (i.e. a tie)

Binary hypothesis test:

- H_a, the outcome was correctly reported
 - $ightharpoonup p = p_a$ the reported proportion
- \triangleright H_0 , the outcome was incorrectly reported
 - ▶ $p = \frac{1}{2}$ the reported proportion

Risk :=
$$Pr[A(X) = Correct \mid H_0]$$

Stopping Probability := $Pr[A(X) = Correct \mid H_a]$

Notation:

 n_j is the cumulative round size in round j k_j is the cumulative tally of winner ballots in round j

Definitions

Bravo:

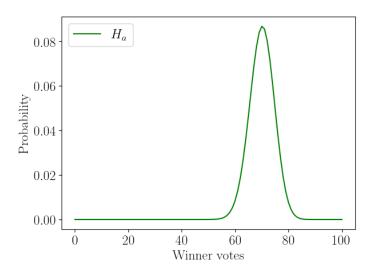
$$\sigma_{j}(k_{j}) \triangleq \frac{p_{a}^{k_{j}}(1-p_{a})^{n_{j}-k_{j}}}{p_{0}^{k}(1-p_{0})^{n_{j}-k_{j}}} = \frac{\Pr[K_{j}=k_{j}|H_{a}]}{\Pr[K_{j}=k_{j}|H_{0}]} \geq \frac{1}{\alpha}$$

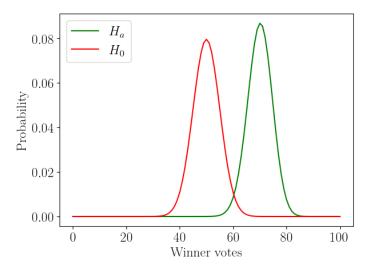
MINERVA:

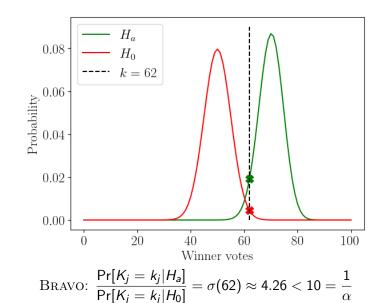
$$\tau_{j}(k_{j}) \triangleq \frac{Pr[K_{j} \geq k_{j} \land \forall_{i < j}(\mathcal{A}(X_{i}) \neq \mathsf{Correct}) \mid H_{a}, \mathbf{n_{j}}]}{Pr[K_{j} \geq k_{j} \land \forall_{i < j}(\mathcal{A}(X_{i}) \neq \mathsf{Correct}) \mid H_{0}, \mathbf{n_{j}}]} \geq \frac{1}{\alpha}$$

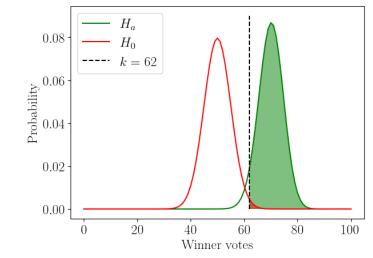
A toy example:

- ▶ Two candidates
- ▶ Reported proportion of votes for the winner $p_a = 0.7$
- Risk limit $\alpha = 0.1$
- First round size $n_1 = 100$

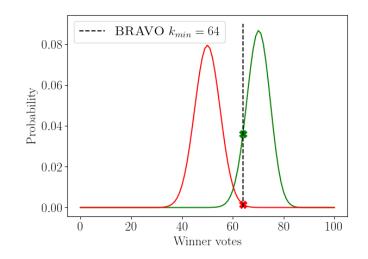




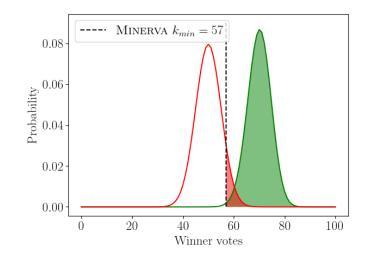




MINERVA:
$$\frac{Pr[K_j \ge k_j \mid H_a, \mathbf{n_j}]}{Pr[K_j \ge k_j \mid H_0, \mathbf{n_j}]} = \tau(62) \approx 92.10 \ge 10 = \frac{1}{\alpha}$$



$$\sigma(63) \approx 9.95 < 10 \le 23.21 \approx \sigma(64)$$



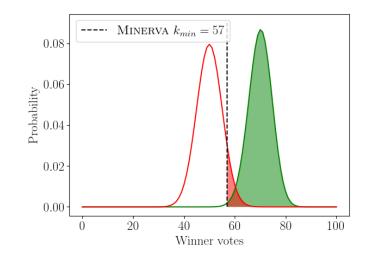
$$\tau(56)\approx7.37<10\leq10.32\approx\tau(57)$$

Why does this tail ratio give an RLA?

Let $R_j = \Pr[\text{stop in round j and no earlier} | H_0]$ and $S_j = \Pr[\text{stop in round j and no earlier} | H_a]$. Then,

$$\frac{S_j}{R_j} \ge \frac{1}{\alpha} \implies R_j \le \alpha S_j \implies \sum_j R_j \le \alpha \sum_j S_j \le \alpha$$

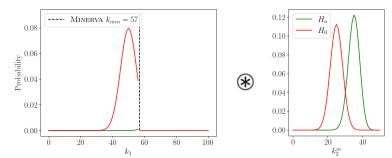
So, audits that enforce $\left(\frac{S_j}{R_j} \geq \frac{1}{\alpha}\right) \forall j$ are risk-limiting!



$$\tau(56)\approx7.37<10\leq10.32\approx\tau(57)$$

$$\tau_{j}(k_{j}) \triangleq \frac{Pr[K_{j} \geq k_{j} \land \forall_{i < j}(\mathcal{A}(X_{i}) \neq \mathsf{Correct}) \mid H_{a}, \mathbf{n_{j}}]}{Pr[K_{j} \geq k_{j} \land \forall_{i < j}(\mathcal{A}(X_{i}) \neq \mathsf{Correct}) \mid H_{0}, \mathbf{n_{j}}]} \geq \frac{1}{\alpha}$$

If $n_2 = 150$, what's pdf for $K_2 = k_2 \wedge \mathcal{A}(X_1) \neq \textit{Correct}$? What's $\Pr[K_2 = 100 \wedge \mathcal{A}(X_1) \neq \textit{Correct}]$?



Implicitly assumes that n_2 is the same for all k_1

Our work

Adversarial round sizes

Goal of adversary: increase the risk above $\boldsymbol{\alpha}$

Definition (Weakly round-choosing adversary)

One who, before the audit begins, selects round sizes n_j for all j as a function of audit and contest parameters:

$$n_j(\alpha, p_a, p_0, ballot_manifest)$$

EoR $\rm Bravo,\ SO\ Bravo,\ Minerva$ are all resistant to a weakly round-choosing adversary

Adversarial round sizes

Definition (Strongly Round-Choosing Adversary)

One who selects n_1 :

$$n_1(\alpha, p_a, p_0, ballot_manifest),$$

and n_j for $j \geq 2$:

$$n_j(\alpha, p_a, p_0, ballot_manifest, \mathbf{k_{j-1}}, \mathbf{n_{j-1}})$$

EoR Bravo, SO Bravo are resistant to a strongly round-choosing adversary; $\operatorname{Minerva}$ is not

PROVIDENCE

Providence:

$$\omega_{j} \triangleq \frac{Pr[K_{j} \geq k_{j} \wedge K_{j-1} = k_{j-1} \mid H_{a}, n_{j-1}, n_{j}]}{Pr[K_{j} \geq k_{j} \wedge K_{j-1} = k_{j-1} \mid H_{0}, n_{j-1}, n_{j}]} \geq \frac{1}{\alpha}$$

RLA proof idea:

- ▶ Sum over all k_{j-1}
- ▶ Again, $\left(\frac{S_j}{R_i} \ge \frac{1}{\alpha}\right) \forall j \implies \mathsf{RLA}$

$$\omega_{j} = \frac{\Pr[K_{j-1} = k_{j-1} \mid H_{a}, n_{j-1}] \Pr[K_{j}^{m} \geq k_{j}^{m} \mid k_{j-1}, H_{a}, n_{j-1}, n_{j}]}{\Pr[K_{j-1} = k_{j-1} \mid H_{0}, n_{j-1}] \Pr[K_{j}^{m} \geq k_{j}^{m} \mid k_{j-1}, H_{0}, n_{j-1}, n_{j}]}$$

 $\omega_i(k_i, k_{i-1}) = \sigma(k_{i-1})\tau_1(k_i^m)$

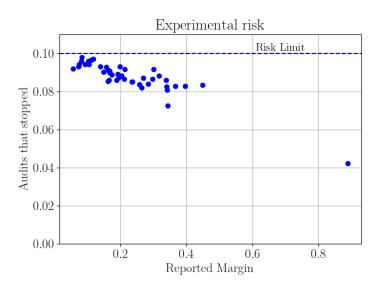
Simulations

What if Oliver put some errors in the proof?

Simulations:

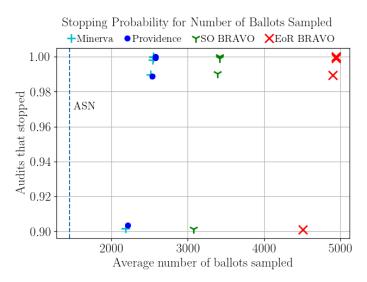
- ➤ 2020 Presidential Contest (states with a pairwise margin at least 0.05)
- ightharpoonup Risk limit $\alpha = 0.1$
- ▶ 10^4 trials per state assuming H_a (i.e. p as reported)
- ▶ 10^4 trials per state assuming H_0 (i.e. p = 0.5)
- 5 rounds
- ▶ Round sizes chosen to each give stopping probability of 0.9, conditioned on the current sample (except MINERVA which uses first round size to achieve 0.9 stopping probability and then a multiplier of 1.5)

PROVIDENCE risk



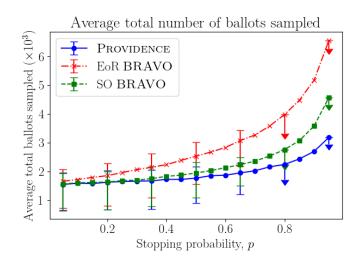
Number of ballots

Texas, with margin 0.057



Number of ballots

Now parameterize round sizes by stopping probability 2016 Presidential contest in VA with margin $\approx 0.053\,$



Workload

With a round cost:

$$W(E_b, E_r) = E_b c_b + E_r c_r$$

 E_b : expected number of ballots

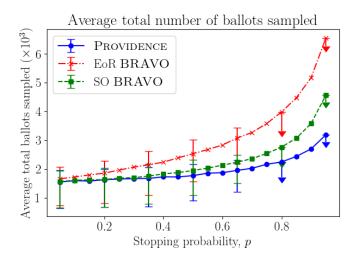
 E_r : expected number of rounds

 c_b : per ballot cost

 c_r : per round cost

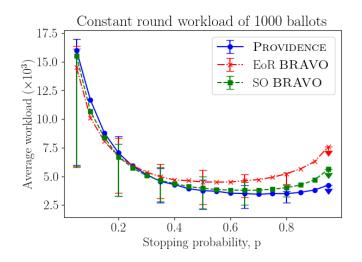
Number of ballots

$$W(E_b, E_r) = E_b c_b + E_r c_r$$
 with $c_b = 1$ and $c_r = 0$

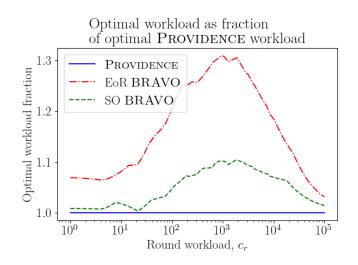


Workload

$$W(E_b, E_r) = E_b c_b + E_r c_r$$
 with $c_b = 1$ and $c_r = 1000$



Optimal workload



Pilot audit

February 2022 in Providence, Rhode Island

 $\mathsf{Margin} \approx 0.256$

ballots	Providence	Minerva	SO Bravo	EoR Bravo
140	0.0418	0.0418	0.0541	0.366

Conclusion

- ► PROVIDENCE: efficient and secure
- Piloted in RI, and implemented in Arlo, most popular RLA software
- Introduction of workload models accounting for round and precinct costs; misleading samples
- Future work: Optimal audits
 - ▶ Bravo is SPRT which is optimal for the round schedule of $[1,1,1,\ldots]$
 - ► MINERVA and PROVIDENCE simplify to BRAVO for this case
 - Nothing is known about the existence of an optimal test for arbitrary round schedules $[n_1, n_2, n_3, ...]$
 - Application of round-by-round tests to other domains (e.g. medical trials)

Thank you for listening