

Study Group Questions # 3

In these questions you consider the relationship between correlation, partial correlations and the multiple correlation coefficient. In Question 4, you consider the partitioned regression model

$$y = X_1\beta_1 + X_2\beta_2 + u \quad (1)$$

where X_1 is a $T \times k_1$ matrix, X_2 is a $T \times k_2$ matrix and y and u are $T \times 1$ vectors. Let $\hat{\beta} = (\hat{\beta}'_1, \hat{\beta}'_2)'$ denote the OLS estimator of $\beta = (\beta'_1, \beta'_2)'$, $\tilde{\beta}_1$ denote the OLS coefficient estimator from the regression of y on X_1 , and \tilde{y} is the prediction of y based on the OLS regression of y on X_1 .

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1. Show that the residual vector can be written as:

$$e = \{I_T - P_1 - \bar{P}_2\}y,$$

where $P_1 = X_1(X'_1 X_1)^{-1} X'_1$ and $\bar{P}_2 = \bar{X}_2(\bar{X}'_2 \bar{X}_2)^{-1} \bar{X}'_2$ for $\bar{X}_2 = M_1 X_2$, and hence that the residual sum of squares, RSS , is given by

$$RSS = y' \{I_T - P_1 - \bar{P}_2\} y = RSS_1 - y' \bar{P}_2 y,$$

where RSS_1 is the residual sum of squares from the OLS regression of y on X_1 .

Consider the partitioned regression model:

$$\begin{aligned} y &= X_1 \beta_1 + X_2 \beta_2 + u \\ \text{Let } X = [X_1 \ X_2] &\Rightarrow X' = \begin{bmatrix} X'_1 \\ X'_2 \end{bmatrix} \\ \hat{\beta} &= \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} \end{aligned}$$

From $\hat{\beta}$ being OLS estimator, we have:

$$\hat{\beta} = (X' X)^{-1} X' y$$

Times both sides with $(X' X)$, we have:

$$(X' X) \hat{\beta} = X' y$$

This equals:

$$\begin{aligned} \begin{bmatrix} X'_1 \\ X'_2 \end{bmatrix} [X_1 \ X_2] \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} &= \begin{bmatrix} X'_1 \\ X'_2 \end{bmatrix} y \\ \Leftrightarrow \begin{bmatrix} X'_1 X_1 & X'_1 X_2 \\ X'_2 X_1 & X'_2 X_2 \end{bmatrix} \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{bmatrix} &= \begin{bmatrix} X'_1 y \\ X'_2 y \end{bmatrix} \\ \Leftrightarrow \begin{bmatrix} X'_1 X_1 \hat{\beta}_1 + X'_1 X_2 \hat{\beta}_2 \\ X'_2 X_1 \hat{\beta}_1 + X'_2 X_2 \hat{\beta}_2 \end{bmatrix} &= \begin{bmatrix} X'_1 y \\ X'_2 y \end{bmatrix} \end{aligned}$$

Then we have the system:

$$\begin{cases} X'_1 X_1 \hat{\beta}_1 + X'_1 X_2 \hat{\beta}_2 = X'_1 y \\ X'_2 X_1 \hat{\beta}_1 + X'_2 X_2 \hat{\beta}_2 = X'_2 y \end{cases} \quad (1)$$

From (1):

$$\hat{\beta}_1 = (X'_1 X_1)^{-1} X'_1 y - (X'_1 X_1)^{-1} X'_1 X_2 \hat{\beta}_2 \quad (1')$$

Substitute $\hat{\beta}_1$ in (2), we have:

$$\begin{aligned} X'_2 X_1 (X'_1 X_1)^{-1} X'_1 y - X'_2 X_1 (X'_1 X_1)^{-1} X'_1 X_2 \hat{\beta}_2 \\ + X'_2 X_2 \hat{\beta}_2 = X'_2 y \end{aligned} \quad (2)$$

$$\begin{cases} P_1 = X_1 (X'_1 X_1)^{-1} X'_1 \\ M_1 = I_T - P_1 \end{cases}$$

We can rewrite (2) as:

$$\begin{aligned} X'_2 P_1 y - X'_2 X_1 X_2 \hat{\beta}_2 + X'_2 I_T X_2 \hat{\beta}_2 &= X'_2 I_T y \\ \Leftrightarrow (X'_2 I_T X_2 - X'_2 P_1 X_2) \hat{\beta}_2 &= X'_2 I_T y - X'_2 P_1 y \\ \Leftrightarrow X'_2 M_1 X_2 \hat{\beta}_2 &= X'_2 M_1 y \\ \Leftrightarrow \hat{\beta}_2 &= (X'_2 M_1 X_2)^{-1} X'_2 M_1 y. \end{aligned}$$

Then, combine with (1'), we have $\hat{\beta}_1$:

$$\hat{\beta}_1 = (X'_1 X_1)^{-1} X'_1 y - (X'_1 X_1)^{-1} X'_1 X_2 (X'_2 M_1 X_2)^{-1} X'_2 M_1 y.$$

From the model, we have:

$$e = y - \hat{\beta}_1 X_1 - \hat{\beta}_2 X_2$$

Plug in $\hat{\beta}_1$ and $\hat{\beta}_2$, we have:

$$e = y - \{X_1 (X'_1 X_1)^{-1} X'_1 y - X_1 (X'_1 X_1)^{-1} X'_1 X_2 (X'_2 M_1 X_2)^{-1} X'_2 M_1 y\}$$

$$- X_2 (X'_2 M_1 X_2)^{-1} X'_2 M_1 y$$

$$\Leftrightarrow e = [I_T - P_1 - (I_T - P_1) X_2 (X'_2 M_1 X_2)^{-1} X'_2 M_1] y \quad (4)$$

$$= [I_T - P_1 - M_1 X_2 (X'_2 M_1 X_2)^{-1} X'_2 M_1] y$$

because M_1 is an orthogonal projection matrix.

$$\begin{cases} \bar{P}_2 = \bar{X}_2 (\bar{X}'_2 \bar{X}_2)^{-1} \bar{X}'_2 \\ \bar{X}_2 = M_1 X_2 \end{cases}$$

We can rewrite (4) as:

$$\begin{aligned} e &= [I_T - P_1 - \bar{X}_2 (\bar{X}'_2 \bar{X}_2)^{-1} \bar{X}'_2] y \\ &= [I_T - P_1 - \bar{P}_2] y. \end{aligned} \quad (\text{Q.E.D.})$$

We have $RSS = \sum_{i=1}^T e_i^2 \quad (i=1, 2, 3, \dots, T)$

$$\begin{aligned} \Rightarrow RSS &= \{[I_T - P_1 - \bar{P}_2] y\}^2 \\ &= y' (I_T - P_1 - \bar{P}_2)^2 (I_T - P_1 - \bar{P}_2) y \\ &= y' (I_T - P_1 - \bar{P}_2) y \end{aligned} \quad (5)$$

because $(I_T - P_1 - \bar{P}_2)$ is an orthogonal projection matrix

Consider the model

$$y = \beta_1 X_1 + u$$

we can derive $\hat{\beta}_1 = (X'_1 X_1)^{-1} X'_1 y$

$$\begin{aligned} \Rightarrow u &= y - X_1 (X'_1 X_1)^{-1} X'_1 y \\ &= I_T y - P_1 y \\ &= M_1 y. \end{aligned}$$

$$\Rightarrow RSS_1 = y' M_1 y.$$

With that, (5) can be rewritten as:

$$\begin{aligned} RSS &= y' (M_1 - \bar{P}_2) y \\ &= y' M_1 y - y' \bar{P}_2 y \\ &= RSS_1 - y' \bar{P}_2 y. \end{aligned} \quad (\text{Q.E.D.})$$

Now consider the version of (1) in which $k_1 = 2$, $k_2 = 1$ and $X_1 = [\iota_T, x_2]$, $X_2 = [x_3]$ that is, the model can be written as

$$y_t = \delta_1 + \delta_2 x_{2,t} + \delta_3 x_{3,t} + u_t, \quad t = 1, 2, \dots, T, \quad (2)$$

where $\beta_1 = (\delta_1, \delta_2)'$ and $\beta_2 = \delta_3$. Re-define R^2 to be the multiple correlation coefficient from the OLS estimation of (2). Define r_2 to be the correlation between y and x_2 , and $r_{y,3|2}$ to be the partial correlation between y and x_3 given (the intercept and) x_2 .

2. Show that:

$$1 - R^2 = (1 - r_2^2)(1 - r_{y,3|2}^2).$$

Hint: Use the result in Question 1.

The model can be rewritten as $y = X_1 \beta_1 + X_2 \beta_2 + u$ with :

$$X_1 = [\iota_T \ x_2] \text{ and } X_2 = [x_3]$$

From Q1, we have $1 - R^2 = RSS_1 - y' \bar{P}_2 y$ with RSS_1 is from

the model $y = \alpha_1 \iota_T + \alpha_2 x_2 + u$,

$$\begin{aligned} RSS_1 &= \sum u_i^2 = \sum (y - \alpha_1 \iota_T - \alpha_2 x_2)^2 = Syy - \alpha_2^2 S_{22} \\ \Rightarrow R_1^2 &= \alpha_2^2 = \frac{S_{22}}{Syy S_{22}} \Rightarrow RSS_1 = Syy (1 - r_2^2) \end{aligned}$$

$$\text{Then we can rewrite : } 1 - R^2 = \frac{RSS}{Syy} = \frac{Syy (1 - r_2^2) - y' \bar{P}_2 y}{Syy}$$

Which means, to prove $1 - R^2 = (1 - r_2^2)(1 - r_{y,3|2}^2)$, we need to prove:

$$\frac{y' \bar{P}_2 y}{Syy} = (1 - r_2^2) r_{y,3|2}^2 \quad \textcircled{2}$$

Start with the RHS. We have $r_{y_3 \bar{S}12}^2 = \frac{(r_{y_3} - r_2 r_{12})^2}{(1-r_2^2)(1-r_{12}^2)}$

$$\begin{aligned} \Rightarrow \text{RHS } \textcircled{1} &= \frac{r_{y_3}^2 - r_2^2 r_{12}^2 + 2 r_{y_3} r_2 r_{12}}{1 - r_{12}^2} \\ &= \frac{\frac{S_{y_3}^2}{S_{yy} S_{22}} - \frac{S_{y_2}^2 S_{12}^2}{S_{yy} S_{22}^2 S_{33}} + 2 \frac{S_{y_3} S_{y_2} S_{12}}{S_{yy} S_{22} S_{33}}}{\frac{S_{22} S_{33} - S_{32}^2}{S_{22} S_{33}}} \\ &= \frac{S_{y_3}^2 S_{22}^2 + S_{y_2}^2 S_{33}^2 - 2 S_{y_3} S_{32} S_{y_2} S_{22}}{S_{yy} S_{22} (S_{33} S_{22} - S_{32}^2)} \end{aligned}$$

Turn to $y^T \bar{P}_2 y$ in LHS $\textcircled{1}$: $y^T \bar{P}_2 y = y^T [I_T - P_3] x_3' [x_2' (I_T - P_1) x_1]^{-1} x_2' [I_T - P_1] y$
with $P_1 = x_1 (x_1' x_1)^{-1} x_1' = \sum_{i=1}^T x_i \begin{bmatrix} i \\ x_i' \end{bmatrix} \begin{bmatrix} i & x_i \end{bmatrix}^{-1} \begin{bmatrix} i \\ x_i' \end{bmatrix}$

$$\text{We have } (x_1' x_1)^{-1} = \begin{bmatrix} I_T I_T & I_T x_2 \\ x_2' I_T & x_2' x_2 \end{bmatrix}^{-1} = \begin{bmatrix} T & \Sigma x_2 \\ \Sigma x_2 & \Sigma x_2^2 \end{bmatrix}^{-1} = \frac{1}{T S_{22}} \begin{bmatrix} \Sigma x_2 & -\Sigma x_2 \\ -\Sigma x_2 & T \end{bmatrix}$$

$$\text{because } \det(x_1' x_1) = T \cdot \Sigma x_2^2 - (\Sigma x_2)^2 = T^2 \left[\frac{1}{T} \Sigma x_2^2 - \left(\frac{1}{T} \Sigma x_2 \right)^2 \right]$$

$$= T (\Sigma x_2^2 - T \bar{x}_2^2) = T \cdot S_{22}.$$

$$\Rightarrow (x_1' x_1)^{-1} \text{ can also be written as } (x_1' x_1)^{-1} = \frac{1}{T S_{22}} \begin{bmatrix} x_1' x_2 & -I_T x_2 \\ -x_2' I_T & I_T I_T \end{bmatrix}$$

$$\Rightarrow P_1 = \frac{1}{T S_{22}} \begin{bmatrix} I_T x_2 & I_T x_2 \\ -x_2' I_T & I_T I_T \end{bmatrix} \begin{bmatrix} I_T \\ x_2' \end{bmatrix}$$

$$= \frac{1}{T S_{22}} \begin{bmatrix} I_T x_2' x_2 I_T - x_2' I_T^2 x_2 I_T^2 - 4 I_T^2 x_2 x_2^2 + T x_2 x_2' \\ I_T x_2' x_2 I_T - x_2' I_T^2 x_2 I_T^2 - 4 I_T^2 x_2 x_2^2 + T x_2 x_2' \end{bmatrix}$$

$$\text{Rewrite } x_3' (I_T - P_1) y = x_3' y - x_3' P_1 y$$

$$\text{we have } x_3' P_1 y = \frac{1}{T S_{22}} x_3' [I_T x_2' x_2 I_T - x_2' I_T^2 x_2 I_T^2 - 4 I_T^2 x_2 x_2^2 + T x_2 x_2'] y$$

$$= \frac{1}{T S_{22}} [\Sigma x_3 \Sigma y x_1' x_2 - \Sigma x_3 \Sigma y x_2' x_2 - \Sigma x_3 \Sigma x_2 x_1' y + T x_2 x_2' x_1' y]$$

$$= \frac{1}{S_{22}} [\Sigma T x_3 \bar{y} x_1' x_2 - \Sigma T x_3 \bar{y} x_2' x_2 - \Sigma T x_3 \bar{x}_2 x_1' y + x_2' x_2 x_1' y]$$

$$= \frac{1}{S_{22}} [(T \bar{x}_3 \bar{y} x_1' x_2 - T^2 \bar{x}_3 \bar{y} \bar{x}_2^2) + (T \bar{x}_3 \bar{y} \bar{x}_2^2 - T \bar{x}_3 \bar{y} x_1' x_2) \\ + (x_2' x_2 x_1' y - T \bar{x}_3 \bar{x}_2 x_1' y)]$$

$$= \frac{1}{S_{22}} (T \bar{x}_3 \bar{y} S_{22} - T \bar{x}_3 \bar{y} S_{22} + S_{22} x_2' x_1' y)$$

$$\Rightarrow x_3' [I_T - P_1] y = x_3' y - T \bar{x}_3 \bar{y} - \frac{S_{22} S_{y2}}{S_{22}} = S_{y3} - \frac{S_{22} S_{y2}}{S_{22}}$$

$$\text{With the same calculation, we can get } \begin{cases} y^T [I_T - P_1] x_2 = S_{y2} - \frac{S_{22} S_{y2}}{S_{22}} \\ x_2' (I_T - P_1) x_2 = S_{33} - \frac{S_{22}^2}{S_{22}} \end{cases}$$

$$\begin{aligned} \Rightarrow \text{We can rewrite } y^T \bar{P}_2 y &= \frac{S_{22} S_{y3} - S_{22} S_{y2}}{S_{22}} \cdot \frac{S_{22}}{S_{yy} S_{22} - S_{22}^2} \cdot \frac{S_{y3} S_{22} - S_{32} S_{y2}}{S_{22}} \\ &= \frac{S_{y3}^2 S_{22}^2 + S_{32}^2 S_{y2}^2 - 2 S_{22} S_{y3} S_{32} S_{y2}}{S_{22} (S_{33} S_{22} - S_{32}^2)} \end{aligned}$$

$$\Rightarrow \text{LHS } \textcircled{1} = \frac{y^T \bar{P}_2 y}{S_{yy}} = \text{RHS } \textcircled{1} \quad (\text{Q.E.D.})$$

For your information: The result in Question 2 holds if we reverse the roles of x_2 and x_3 so that,

$$1 - R^2 = (1 - r_3^2)(1 - r_{y,2|3}^2),$$

where r_3 is the correlation between y and x_3 and $r_{y,2|3}$ is the partial correlation between y and x_2 given (the intercept and) x_3 . This relationship between the correlations also extends to larger models. For example, if we consider the model

$$y_t = \delta_1 + \delta_2 x_{2,t} + \delta_3 x_{3,t} + \delta_4 x_{4,t} + u_t, \quad t = 1, 2, \dots, T, \quad (3)$$

then the R^2 from this regression satisfies:

$$1 - R^2 = (1 - r_2^2)(1 - r_{y,3|2}^2)(1 - r_{y,4|3,2}^2)$$

where $r_{y,4|3,2}$ is the partial correlation between y and x_4 given (ι_T, x_2, x_3) and $r_{y,3|2}$ is the partial correlation between y and x_3 given (ι_T, x_2) . As in the model in Question 2, the subscripts can be interchanged in any order so that, for example, we also have

$$1 - R^2 = (1 - r_3^2)(1 - r_{y,4|3}^2)(1 - r_{y,2|3,4}^2),$$

in the obvious notation for the model in (3).